| 1 | Identifying conservation priorities in a defaunated tropical | | | | | | |
|----------|--|--|--|--|--|--|--|
| 2 | biodiversity hotspot | | | | | | |
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| 21 | Aim: Unsustainable hunting is leading to widespread defaunation across the tropics. To | | | | | | |
| 22 | mitigate against this threat with limited conservation resources, stakeholders must | | | | | | |
| 23 | make decisions on where to focus anti-poaching activities. Identifying priority areas in a | | | | | | |
| 24 | robust way allows decision-makers to target areas of conservation importance, | | | | | | |
| 25 | therefore maximizing the impact of conservation interventions. | | | | | | |
| 26 | Location: Annamite mountains, Vietnam and Laos. | | | | | | |

Methods: We conducted systematic landscape-scale surveys across five study sites (four protected areas, one unprotected area) using camera-trapping and leech-derived environmental DNA. We analyzed detections within a Bayesian multi-species occupancy framework to evaluate species responses to environmental and anthropogenic influences. Species responses were then used to predict occurrence to unsampled regions. We used predicted species richness maps and occurrence of endemic species to identify areas of conservation importance for targeted conservation interventions.

Results: Analyses showed that habitat-based covariates were uninformative. Our final model therefore incorporated three anthropogenic covariates as well as elevation, which reflects both ecological and anthropogenic factors. Conservation-priority species tended to found in areas that are more remote now or have been less accessible in the past, and at higher elevations. Predicted species richness was low and broadly similar across the sites, but slightly higher in the more remote site. Occupancy of the three endemic species showed a similar trend.

Main conclusion: Identifying spatial patterns of biodiversity in heavily-defaunated 41 landscapes may require novel methodological and analytical approaches. Our results 42 indicate to build robust prediction maps it is beneficial to sample over large spatial 43 scales, use multiple detection methods to increase detections for rare species, include 44 anthropogenic covariates that capture different aspects of hunting pressure, and analyze 45 data within a Bayesian multi-species framework. Our models further suggest that more 46 remote areas should be prioritized for anti-poaching efforts to prevent the loss of rare 47 and endemic species. 48

Keywords: Annamites, camera-trapping, defaunation, environmental DNA, multispecies occupancy, species richness, tropical rainforest, unsustainable hunting

51 Introduction

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Tropical biodiversity is declining at an alarming rate as a result of intense anthropogenic 54 pressures (Bradshaw et al, 2009). Although habitat loss and degradation are major 55 drivers of these declines (Rosa et al, 2016), unsustainable hunting is increasingly 56 emerging as the primary threat to wildlife in tropical biodiversity hotspots (Benítez-57 López et al, 2017). Large and medium-sized mammals tend to be particularly vulnerable 58 to hunting because they often occur at lower average population densities, have lower 59 intrinsic rates of increase, and longer generation times (Bodmer et al, 1997; Cardillo et 60 al, 2005; Davidson et al, 2009). Indeed, the "empty forest syndrome" that Redford 61 (1992) warned about almost three decades ago is now a commonplace phenomenon 62 and, given the ever-increasing demand for wildlife products in the world's tropical 63 regions (Rosen & Smith, 2010; Ripple et al. 2016), this trend is unlikely to slow in the 64 coming years. Without urgent and effective measures to address overexploitation, 65 tropical wildlife populations will continue to decline, and species extinctions will follow. 66 67 Confronting the pantropical defaunation crisis has become one of the most important challenges facing conservation today (Bradshaw et al, 2009). 68

69 Defaunation has been particularly severe in Southeast Asia, where high human densities, a thriving illegal wildlife trade, weak protected area governance, and rapid 70 71 infrastructure development have synergistically contributed to unsustainable, industrial-scale hunting (Duckworth et al, 2012; Wilcove et al, 2013; Harrison et al, 72 73 2016). Within Southeast Asia, the Annamites ecoregion on the border of Vietnam and Laos has undergone severe defaunation as a result of widespread illegal hunting 74 75 (Harrison et al, 2016; Timmins et al, 2016). Poaching in the Annamites is primarily accomplished by the setting of indiscriminate wire snares (Gray et al, 2018). Numerous 76

mammals are regionally extinct (Walston et al, 2010; Brook et al, 2014), and even once 77 common species now survive at low densities (Duckworth et al, 2016). High levels of 78 79 unsustainable hunting pressure are particularly worrisome from a conservation perspective, because the region is home to several endemic mammal species. Mammals 80 restricted to this ecoregion include the saola *Pseudoryx nghetinhensis*, large-antlered 81 muntiac Muntiacus vuquangensis, Annamites dark muntiac species complex Muntiacus 82 rooseveltorum / truongsonensis, Owston's civet Chrotogale owstoni, and Annamite 83 striped rabbit Nesolagus timminsi (Tordoff et al, 2003; Hurley et al, 2005; Long et al, 84 2005). Taken together, the high poaching pressure and unique biodiversity in the 85 Annamites make it one of the highest priority tropical regions in the world for the 86 87 prevention of imminent hunting-driven extinctions.

To maximize the effectiveness of conservation interventions to prevent unsustainable 88 hunting in tropical biodiversity hotspots, it is imperative to make optimal use of limited 89 90 conservation resources. In the Annamites, the magnitude of the snaring crisis (Gray et al, 2018), coupled with nascent protected area enforcement capacities and lack of sufficient 91 92 resources, has overwhelmed efforts to adequately reduce this threat at the landscape level. Given these limitations, targeting snare removal efforts to specific areas within a 93 landscape may be critical to reduce snaring to levels that would allow population 94 95 recovery. To implement this approach, it is first necessary to identify priority areas. In the Annamites, areas that harbor threatened and endemic species are top priorities for 96 targeted *in situ* protection measures. These species often occur at low densities, and are 97 therefore particularly susceptible to local extirpation. To identify priority areas, it is 98 important to apply appropriate analytical techniques. Species distribution modeling 99 100 provides an ideal framework for mapping spatial patterns of biodiversity, and thus identifying conservation-priority areas (Rodríguez et al, 2007; Guisan et al, 2013). 101

There are, however, two fundamental challenges to the modeling of species distributions in tropical rainforest environments. First, tropical mammal species are often difficult to detect because they are rare, elusive, and occur at low densities. Second, even when these species can be detected, it may be difficult to obtain enough data to construct robust species distribution models (Cayuela et al, 2009), particularly in defaunated areas, where mammal populations are depleted.

Advances in noninvasive survey methods and statistical modeling techniques provide 108 ways to address these challenges. Two noninvasive methods have revolutionized 109 surveys for tropical mammals: camera-traps (Tobler et al, 2008) and high-throughput 110 sequencing of environmental DNA (eDNA) (Bohmann et al, 2013). Camera-traps are a 111 well-established method, and have been used to gather data on even the rarest of 112 tropical mammal species (Whitfield, 1998; Raloff, 1999; Ganas & Lindsell, 2010). The 113 use of eDNA is relatively new but shows considerable promise. Invertebrate-derived 114 115 DNA (iDNA) approaches using terrestrial hematophagous leeches, in particular, have proven adept at detecting tropical mammals (Schnell et al, 2012; Schnell et al, 2018; 116 117 Weiskopf et al, 2018). Recently, Abrams et al (2019) showed that combining cameratrapping and iDNA leech data has the potential to improve detection probabilities for 118 tropical mammal species beyond what would be provided by each method 119 120 independently. The joint camera-trap and iDNA approach thus opens new possibilities for obtaining detections of elusive tropical rainforest mammals, which in turn can be 121 used to build robust species distribution models. 122

Even with improved detection methods and combined datasets, however, it may not be possible to obtain sufficient records for rare species. This shortfall represents a major issue because the rarest species are often the species of highest conservation concern. Multi-species occupancy models offer an analytical framework to address this challenge,

as species with few detections borrow information from more abundant species, which
allows parameter estimation for rare species (Tobler et al, 2015; Drouilly et al, 2018; Li
et al, 2018). Because species-specific responses to covariates can be projected to
unsampled areas, this approach can be used to generate maps of species potential
occurrence (MacKenzie et al, 2017; Sollmann et al. 2017).

132 Here, we collected a landscape-scale systematic camera-trapping and iDNA dataset across a protected area complex in the central Annamites landscape to identify priority 133 areas for targeted conservation interventions. We used a multi-species occupancy 134 framework and environmental and anthropogenic covariates to estimate species 135 occurrence and predict species richness across the surveyed landscape. Our prediction 136 maps provide insight into where to focus conservation efforts among individual study 137 sites at the landscape scale, and more specifically can inform deployment of snare-138 removal teams within protected areas. We discuss our results within the context of 139 informing targeted conservation interventions to prevent further defaunation, and 140 species extinctions, within tropical biodiversity hotspots. 141

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143 Methods
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146 Study area
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We conducted landscape-scale surveys in a large contiguous forest in the central Annamites landscape of Vietnam and Laos. The study area spans both countries and is divided into five administrative units. In Vietnam, we surveyed three sites: Bach Ma National Park (NP), the Hue Saola Nature Reserve (NR), and the Quang Nam Saola NR. In Laos, we surveyed the eastern section of Xe Sap National Protected Area (NPA) and an

adjacent ungazetted forest near the village of Ban Palé (Fig. 1). Together these areas 152 comprise approximately 900 km² of mountainous terrain with elevations ranging 153 between 100 - 2000 m asl. The dominant habitat type is wet evergreen tropical 154 rainforest. Although the wider central Annamites region has experienced extensive past 155 disturbance from defoliation and logging, habitat loss, degradation, and fragmentation 156 within the past 20 years has been minimal within our study sites (Meyfroidt et al, 2008; 157 Matusch, 2014). At the landscape scale, forest structure and habitat type are consistent 158 across the study sites, characterized by mature secondary forest with a multi-tiered 159 160 closed canopy (Fig. S1). The Vietnam sites are surrounded by a densely-populated matrix consisting of human settlements, agricultural fields, and timber plantations. 161 162 Human population density in the Lao sites is low and, aside from small-scale shifting cultivation, the landscape surrounding the survey areas has not been heavily modified. 163 164 However, Vietnamese incursion into these areas for poaching and illegal gold mining is widespread (Tilker, 2014), and has been facilitated by the recent construction of a road 165 166 connecting Vietnam and Laos that bisects the Palé area.

167 Poaching pressure is high across the landscape (Wilkinson 2016; WWF, 2017). Measures to mitigate illegal hunting differ in intensity and effectiveness among the five sites. 168 Patrolling in Bach Ma NP is not intensive and has received less technical and financial 169 170 support than the adjacent sites. The Hue and Quang Nam Saola NRs have benefited from WWF investment in enforcement since 2011 under the Carbon and Biodiversity (CarBi) 171 project, maintaining active Forest Guard patrol teams to strengthen enforcement 172 capacities in the field and provision of capacity development in patrol strategy, data 173 collection and adaptive management for park staff. The Forest Guard teams are 174 comprised of local community members and their primary role is to remove wire snares 175 and destroy poacher camps (Wilkinson, 2017). Between 2011 and 2017, the patrols 176

removed > 110,000 snares from the Hue and Quang Nam Saola NRs (WWF, 2017). The
eastern section of Xe Sap NPA has also benefited from WWF-supported snare removal
operations, although these efforts have not been as regular or intensive as in the Saola
NRs. There are no active patrols in Palé, as it is outside of the Xe Sap NPA.

181 Data collection and preparation

We conducted systematic camera-trapping and leech surveys from November 2014 -182 December 2016. We set up a total of 140 camera-trap stations: 53 stations in Bach Ma 183 NP, 21 in the Hue Saola NR, 25 in the Quang Nam Saola NR, 15 in eastern Xe Sap NPA, 184 and 26 in the Palé area (Fig. 1; Table S2). Stations were spaced approximately 2.5 km 185 apart (mean = 2.47 ± 0.233), aiming at spatial independence of sampling locations, and 186 left in the forest for a minimum of 60 days (mean = 71.60 ± 16.39). Cameras were set 20 187 - 40 cm off the ground, operational 24 hour/day, and programed to take a three-photo 188 burst with no delay between photographic events. To maximize detection probabilities, 189 190 we set two camera-traps (Hyperfire Professional PC850, Reconyx®, Holmen, USA) at each station facing different directions. We treated the two cameras as a single station in 191 192 our analyses. Camera-trap data was managed using the package *camtrapR* (Niedballa et al, 2015). We excluded arboreal species from our final species list, as these species are 193 unlikely to be reliably detected by camera-traps placed at ground-level (Abrams et al, 194 2018). We also removed rodents and squirrels, given the difficulty of identifying these 195 mammals to species-level using camera-trap images alone, and all domestic animals. 196

We complemented camera-trapping with the collection of terrestrial haematophagous leeches around the camera-trap stations. Leeches were collected once during cameratrap setup and again during retrieval. In Vietnam, leeches were collected in 20 x 20m sampling plots set up to assess microhabitat characteristics (see below). In Laos, we 201 collected leeches in a grid around each camera trap station, with one camera-trap station per grid. We altered the leech collection strategy in Laos because sampling 202 203 occurred during the dry season; increasing spatial coverage around the stations allowed us to collect leech numbers similar to the Vietnam sites. We separated the two types of 204 leeches, brown and tiger, because the leeches potentially differ in their feeding behavior 205 (Schnell et al 2015). All leeches of the same type from the same station and occasion 206 were combined and processed as one leech bulk sample. Leeches were immediately 207 placed in RNAlater and stored long-term at -20° C. 208

Leeches were processed using the laboratory procedures and bioinformatics pipeline 209 described in Axtner et al (2019). The workflow is designed to minimize the risk of false 210 positives that could arise from laboratory artifacts or misidentification during 211 taxonomic assignment. To address these risks, it employs different levels of replication 212 (i.e. extraction, PCRs), a curated reference database, and the probabilistic taxonomic 213 214 assignment method PROTAX (Somervuo et al, 2017) that has been shown to be robust even when reference databases are incomplete (Rodgers et al., 2017, Richardson et al., 215 216 2017). Leech samples were digested, DNA was extracted, and then mitochondrial target DNA of host species was amplified with PCR and sequenced using Illumina high-217 218 throughput sequencing. We trained PROTAX models and weighted them toward 127 219 mammal and bird species expected to occur in the study area by assigning a prior probability of 90% to these species and a 10% probability to all others (Somervuo et al, 220 2017; see Table S2 for full weighted species list). Our protocol was slightly modified 221 from Axtner et al (2019) in that we amplified the mitochondrial marker 16S in six PCR 222 replicates for all samples and used genetic markers 12S and CytB only for samples 223 where taxonomic assignment was still uncertain due to interspecific invariance or 224 missing references (e.g. porcupines, viverrids, muntjacs). We accepted a species 225

assignment when it was present in at least two independent PCR replicates (Axtner et al.

227 2019, Abrams et al. 2019). As with the camera-trapping data, we excluded arboreal

228 species, rodents, squirrels, and domestic animals from the final species list.

229 Covariates

We hypothesized that mammal occurrence may be influenced by both environmental 230 and anthropogenic factors. We measured three environmental features that characterize 231 232 different aspects of microhabitat structure: canopy closure, vegetation density, and leaf litter. We used canopy closure as an indication of forest degradation, with lower values 233 representing more disturbed habitat (Chazdon, 2003). Previous studies have shown that 234 vegetation density may be an important microhabitat feature for some tropical 235 mammals (Goulart et al, 2009; Martin et al, 2015; Mathai et al, 2017). Leaf litter impacts 236 multiple aspects of vegetation community composition (Facelli & Pickett, 1991). It is 237 also an important microhabitat for invertebrates and small vertebrates (Burghouts et al, 238 1992; Vitt & Caldwell, 1994), which are important food resources for insectivores and 239 small carnivores. 240

To assess microhabitat features, we set up a 20 x 20 m plot around the camera-trap 241 242 stations, with the centerpoint halfway between the two cameras, and oriented along the cardinal axes. To measure canopy closure we took vertical photographs at the 243 centerpoint and at the corners of the grid. Canopy photographs were manually 244 converted to black and white images using the GNU Image Manipulation Program (GIMP, 245 246 2017). We calculated percentage canopy closure (white pixels) for each image using R 3.4.0 (R Development Core Team, 2016). Values for each image were averaged to give a 247 248 single canopy closure value for each station. To measure vegetation density we took photographs in each cardinal direction of a 1 x 1.5 m orange sheet positioned 10 m from 249

the centerpoint. Photographs were processed using the canopy closure protocol, giving a single average vegetation density value for each station. We measured leaf litter percent cover in nine 1 x 1 m subplots located at the centerpoint, 10 m from the centerpoint in each cardinal direction, and at the plot corners. Each subplot was visually assigned a value from 0 to 4 based on the amount of leaf litter versus bare ground visible in each plot. Leaf litter values were averaged to give a single value for each station. For a detailed explanation of the microhabitat assessment see Abrams et al (2018).

257 In addition to the environmental covariates, we measured anthropogenic features that approximate hunting pressure. We use proxies for hunting pressure, rather than direct 258 measures, for two reasons. First, we are not aware of any existing datasets that directly 259 measure hunting pressure within our study sites. Second, robustly assessing poaching 260 represents a difficult undertaking because illegal hunting is such a cryptic phenomenon. 261 Although some studies have used presence or absence of people from camera-trapping 262 263 data to represent direct measures of hunting pressure (Dias et al, 2019), such measures are not applicable in our landscape, because some local communities are allowed to 264 265 legally enter the study sites to collect non-timber forest products. Further complicating the situation is the fact that these local people may engage in both legal non-timber 266 forest product collection and illegal hunting in order to maximize potential profit. Given 267 268 the difficulties in assessing hunting directly, we used measures of accessibility as proxies for hunting pressure in our study sites. Previous studies have shown accessibility and 269 hunting to be correlated (Rao et al, 2005; Espinosa et al, 2014; Koerner et al, 2017). We 270 used three covariates that capture different aspect of accessibility: distance from major 271 cities, village density, and least cost path from major roads. We used city distance as a 272 proxy for hunting pressure captured at the landscape scale. Although we measure 273 distance to the nearest major city (Hue or Da Nang, both with population > 350,000), we 274

also interpret this covariate as an approximation of accessibility to the densely-275 populated coastal areas of Vietnam. We chose to measure distance to the cities, rather 276 277 than other points along the urbanized coastal areas, because Hue and Da Nang are known to be major hubs for the illegal wildlife trade (VanSong, 2003; Sandalj et al, 278 2016). Given the volume of bushmeat that passes through these markets (Sandalj et al, 279 2016), it is likely that these urban population centers create substantial natural resource 280 demand shadows across the landscape, as has been shown in other tropical regions (Ape 281 Alliance, 1998). We derived the city distance covariate by calculating the Euclidean 282 distance from the camera-trap stations to the nearest major city using the package 283 *gDistance* (Van Etten, 2017), then taking the lower of the two values. The city distance 284 covariate is measured in meters, with increasing values indicating more remote areas. 285 We then took the log of the covariate to approximate the non-linear effect that 286 287 increasing distance likely has on accessibility. Village density serves as a proxy for hunting at the local scale. Local villagers often supplement their income by providing 288 289 bushmeat to the bushmeat markets in regional towns and cities, and are therefore a primary driver of poaching in the central Annamites (MacMillan & Nguyen, 2014). 290 291 Studies in other tropical regions have demonstrated mammal depletions surrounding local villages (Rao et al, 2005; Koerner et al, 2017; Abrahams et al, 2017). To calculate 292 village density we first created a ground-truthed point shapefile layer documenting local 293 villages around our study sites. We then created a heatmap in QGIS 2.18.9 (QGIS 294 Development Team, 2016) using the village shapefile as the input point layer. To create 295 the heatmap, we used the default quartic kernel decay function and set the radius to 15 296 297 km. The village density radius was chosen so that all individual sampling stations in our study landscape were covered in the final heatmap. Observations in the field indicate 298 299 that all stations, even those in the most in the most remote areas, were subject to some

level of hunting pressure. We then used the extract function in the raster package 300 (Hijmans, 2019) to obtain heatmap values for each station. The village density covariate 301 302 is unitless, with lower values indicating areas that are more remote. Finally, the least cost path covariate also serves as a proxy for hunting pressure at the local scale. 303 However, it differs from the village density measure in two fundamental ways. First, the 304 least cost path covariate explicitly incorporates accessibility based on terrain 305 ruggedness characteristics, therefore providing a more accurate representation of 306 remoteness than linear measures. Second, we calculated the least cost path covariate 307 308 over three time periods (1994, 2004, and 2014) to better capture the amount of time 309 that an area has been subjected to poaching pressure. The least cost path covariate 310 therefore captures both spatial and temporal dimensions. To create the least cost path covariate we first used the timelapse function in Google Earth Engine (Gorelick et al, 311 312 2017) to generate a GIS layer of major roads in and around our study site for three time periods: 1994, 2004, and 2014. We converted the roads layer to points in QGIS. Next, we 313 314 used the R package *movecost* (Alberti, 2018) to calculate travel time along least cost path routes from the stations to the nearest 50 points along the road layer, using a shuttle 315 radar topography mission (SRTM) 30 m digital elevation model as the cost surface 316 raster, and then selected the lowest value as the final least cost path value. We averaged 317 the three values to give a single least cost path value for each station, which we use as an 318 approximation of the time that an area has been accessible over the past 20 years. The 319 320 roads least cost path covariate is measured in hours, with higher values indicating areas that take longer to access, and are therefore more remote. 321

We also included elevation as a covariate in our models. We consider elevation as both an anthropogenic and ecological covariate. Because higher elevation areas are more difficult to access, elevation serves as a measure of remoteness within our landscape.

Elevation is also linked to a complex range of ecological attributes in the central Annamites, including subtle variations in forest structure and microclimate (Tordoff et al, 2003; Long, 2005).

We standardized all covariates. We tested for correlations between all possible pairs of covariates using Pearson's correlation plots. None of our covariates were highly correlated (r < 0.6; Fig. S2).

331 Modeling framework

We adopted a hierarchical multi-species occupancy model to estimate species occupancy 332 and richness (Dorazio & Royle, 2005; Dorazio et al., 2006). Occupancy models estimate 333 the probability of species occupancy, ψ , while accounting for species detection, p, using 334 repeated species detection/non-detection data collected across multiple sampling 335 locations (MacKenzie et al., 2003). To convert camera-trapping data to an occupancy 336 format, we divided the active camera-trapping time for each station into 10-day 337 sampling periods, yielding a minimum of six occasions for each station. We chose to use 338 339 a 10-day sampling period to minimize zero-inflation in the detection history matrix. We treated each leech collection event as a separate occasion for the stations. We defined z_{ii} 340 341 as the true occupancy state (0 or 1) of species *i* at sampling station *j*. Occupancy state can be modeled as a Bernoulli random variable with the success probability ψ_{ii} , the 342 343 occupancy probability of species *i* at site *j*. We defined *p*_{iik} as detection probability for species *i* at station *j* during the *k*th sampling occasion, and y_{ijk} the observation (i.e., y_{ijk} = 344 345 1 if species *i* is observed at site *j*, occasion *k*, and 0 otherwise). Observing a species is conditional on its occurrence, so that *y*_{ijk} can be modeled as a Bernoulli random variable 346 347 with success probability $z_{ii} \cdot p_{iik}$.

Covariate effects on both parameters can be modeled on the logit scale. We included 348 habitat and anthropogenic covariates on ψ_{ii} to investigate their potential effects on 349 species occurrence. To avoid overparametizing the model, we first ran single-covariate 350 models using each of the seven covariates that we selected a priori, and assessed 351 covariate importance by evaluating effect sizes for each species in the community. 352 Because the environmental covariates did not show strong effects on occupancy, with all 353 species having 95% BCIs overlapping zero and most species showing overlapping 75% 354 Bayesian Confidence Intervals (BCIs) (Fig. S3), these covariates were not included in the 355 final model. Our final community model included four covariates on ψ_{ii} : city distance, 356 village density, roads least cost path, and elevation. Following Abrams et al (2019), we 357 used survey method (camera-trap, brown leech, or tiger leech) as a covariate on p. We 358 accounted for varying survey effort by including number of days each camera-trap 359 360 station was operational during each 10-day occasion or number of leeches per sample 361 on *p*.

We implemented the models in a Bayesian framework using JAGS (Plummer, 2003) 362 363 accessed through the package rjags (Plummer, 2018). We used vague priors (e.g. normal distributions with mean zero and variance 100 for community-level occupancy and 364 detection coefficients). We ran three parallel Markov chains with 250,000 iterations, of 365 366 which we discarded 50,000 as burn-in. We assessed chain convergence using the Gelman-Rubin statistic, with values close to 1 indicating convergence (Gelman et al, 367 2004). We report results as posterior mean and standard deviation. We consider a 368 coefficient to have strong support if the 95% Bayesian confidence interval (95% BCI, the 369 2.5% and 97.5% percentiles of the posterior distribution) does not overlap zero, and 370 moderate support if the posterior 75% BCI does not overlap zero. The full model 371 description is provided in Appendix 1. 372

To test for spatial autocorrelation in response variables not accounted for by predictor 373 variables, we followed an approach put forth by Moore and Swihart (2003). We 374 calculated Moran's I using the moransI function from the R package lctools for residuals 375 from occupancy models using a neighborhood distance of 2.5 km (average spacing of our 376 sampling stations). We only found evidence of low to moderate spatial autocorrelation 377 in occupancy model residuals in only 2 of the 23 species analyzed. We acknowledge that 378 for these species we may underestimate occupancy. However, our analysis is concerned 379 380 with comparisons of patterns across study sites, not among species, and for a given 381 species, any bias should be similar across the sites. Estimates of Moran's I and associated p-values for species are shown in Appendix 2. 382

To predict species richness across the landscape we first divided the study area into 200 383 x 200 m grid cells. We included proposed extensions for the Hue and Quang Nam Saola 384 Nature Reserves in the prediction area. Next, we derived covariate values for each cell. 385 386 For the city distance, village density, and roads least cost path covariates we followed the same protocols described above. Elevation values were extracted from an SRTM 30 387 388 m digital elevation model (see Fig. S4 for covariate rasters). We used estimates of the coefficients from the multi-species model linking covariates to occupancy probability to 389 390 predict occupancy values for each species and grid cell and then summed the occupancy 391 probabilities for all species per cell to produce species richness maps. To highlight areas of high richness for conservation-priority species, we produced a separate species 392 richness map for the endemic species and those listed as Near Threatened or higher on 393 The IUCN Red List of Threatened Species. To provide a further level of detail for the 394 endemic species, we also produced single-species occupancy maps for Annamite striped 395 396 rabbit, Annamite dark muntjac, and Owston's civet. We note that, although our sampling 397 stations only covered part of the study sites, covariate values at the stations were largely representative of values across the sites. When we filtered the raster cells to remove cells that fell outside the range of our covariates at the sampling stations, we found that a small number raster cells were excluded. However, we present the full prediction maps here, both because the differences between the complete and filtered rasters were minor, and for visualization purposes (see Fig S5 for modified prediction maps).

403 We further used a modified Bray-Curtis index to assess compositional dissimilarity among the five study sites. The Bray-Curtis index calculates dissimilarity values by 404 405 comparing composition in a reference assemblage with one or more target assemblages (Bray & Curtis, 1957). We adapted the index to compare predicted species occupancy 406 probabilities between all possible site combinations, following the general framework 407 proposed by Giocomini & Galetti (2013). To do this we sampled random values from the 408 posterior distributions of species-specific occupancy probabilities for both the focal and 409 target study sites. We repeated this procedure 30,000 times using Monte Carlo sampling 410 411 to generate a distribution of values and took the mean of the posterior distribution. The final value gives an indication of how dissimilar the predicted community-level 412 413 occupancies are among the sites. Dissimilarity values can range between -1 and 1. A value of 0 indicates no differences in occupancy between the focal and reference sites, a 414 415 value of 1 indicates complete dissimilarity with the reference site having higher occupancies than the focal site, and a value of -1 indicates complete dissimilarity with 416 the focal site having higher occupancies than the reference site. We calculated Brav-417 Curtis dissimilarities first for the entire community, and then for endemic and 418 threatened species. Further details on the Bray-Curtis dissimilarity index are provided 419 in Appendix 3. 420

422 **Results**

We obtained data from 139 camera-trap stations totaling 17,393 trap-nights (Table S1). 423 The camera-trapping yielded 5,261 independent detections ($\Delta = 60$ min between 424 425 subsequent pictures of the same species at the same camera trap) of 27 terrestrial mammals. We identified all mammals to species, with the exception of the ferret badgers 426 427 (Melogale personata and M. moschata) and pangolins (Manis pentadactyl and M. *javanica*), which we identified to the genus level due to the difficulty of identifying to 428 species using camera-trap photographs, and the Annamite dark muntiac species 429 complex *Muntiacus rooseveltorum / truongsonensis*, due to its unresolved taxonomic 430 status. Our final species list resulted in 22 mammals. We obtained 193 leech samples 431 totaling 2,043 leeches (1,888 brown, 155 tiger) from 98 stations (mean leeches / station 432 = 21, standard deviation = 22; Table S1). We were able to amplify and sequence DNA 433 from 104 samples. PROTAX identified 25 mammals to the species level and 7 to the 434 genus level. The final species list from the leeches included 19 terrestrial mammals. 435 Overall, the two survey methods provided similar species lists. The exceptions were 436 437 pangolin, pig-tailed macaque Macaca leonina, spotted linsang Prionodon pardicolor, and yellow-bellied weasel Mustela kathiah, which were detected only in the camera-traps, 438 and marbled cat *Pardofelis marmorata*, which was detected only in the leeches. The final 439 440 species list used for the community occupancy analysis included 23 mammals. Four of these were threatened, three were Annamite endemics, and one species fit both 441 categories. The full species list and classifications can be found in Table S3. 442

Detection probabilities (*p*) within the mammal community varied among species and with respect to survey method. Estimates of occupancy (ψ) showed extreme heterogeneity among individual species. Estimates of *p* and ψ can be found in Fig. S6, and full model results are provided in Table S4.

Species-specific responses to the covariates were highly variable within the community 447 (Fig. 2). Six species showed a moderate positive relationship with city distance, three 448 species showed a moderate negative relationship, and four species showed a strong 449 negative relationship. There was a moderate positive effect of elevation on occupancy 450 for nine species and a strong positive effect for four species. One species showed a 451 moderate negative response to elevation and three species showed a strong negative 452 response. There was a moderate positive relationship for elevation at the community 453 level. Five species showed a moderate positive response to the road least cost path 454 covariate. Three species showed a moderate negative relationship with this covariate 455 456 and one species had a strong negative response. Most species showed a negative 457 relationship with village density, with 16 species having a moderate negative relationship and four species showing a strong negative relationship. The community 458 459 response for the village density covariate showed a strong negative response.

460 Average predicted species richness for the full community was substantially lower than the total number of species detected in the study, and was similar among the study sites 461 462 (Table 1, Fig. 3A). The Palé area had slightly higher predicted species richness than the other four areas. Richness of threatened and endemic species followed a similar pattern, 463 with all sites showing low richness relative to the total number of conservation-priority 464 465 species detected, and the highest richness in the Palé area (Table 1, Fig. 3B). Predicted occupancies for the three Annamite endemic mammals showed heterogeneity among 466 species and sites (Table 1, Fig. 4). Annamite dark muntjac had the highest predicted 467 occupancy, followed by Annamite striped rabbit, followed by Owston's civet. All three 468 endemics had highest predicted occupancies in the Palé area, followed by Xe Sap NPA, 469 470 Quang Nam Saola NR, Hue Saola NR, and finally Bach Ma NP. Single-species occupancy predictions for all species can be found in S7. 471

The Bray-Curtis dissimilarity index values showed similar values for full community 472 occupancies among Bach Ma NP, the Hue and Quang Nam Saola NRs, and Xe Sap NPA 473 (Table 2). However, the Palé area had negative and high defaunation index values when 474 compared to every other site, indicating that occupancies for the full suite of species are 475 higher for this area. Dissimilarity values for endemic and threatened species showed a 476 similar pattern. The fact that Palé area showed negative and higher dissimilarity values 477 for both the full community and for conservation-priority species suggests that, within 478 the context of species occurrence, the site has undergone less severe defaunation than 479 the other sites. 480

481

482 **Discussion**

483

Our study highlights the landscape-scale effects of unsustainable hunting on the 484 485 occurrence and distribution of terrestrial mammals within a tropical biodiversity hotspot. For a structurally-intact tropical rainforest habitat, average predicted species 486 487 richness (7.02) was low (Table 1; Fig. 3A & B). For comparison, Deere et al (2018) found a predicted species richness of 14.12 in logged forests in Malaysian Borneo, and a 488 489 richness of 4.54 in adjacent oil palm plantations. Given the largely homogenous landscape-scale forest structure and habitat in our study sites, the low predicted species 490 491 richness is likely indicative of a community that has undergone severe hunting-driven defaunation. The extent of faunal impoverishment is further supported by the fact that 492 493 we failed to record almost half of the mammal community would be expected to occur in these sites based on historical distribution maps (Tilker & Abrams, in press). 494

At the landscape level predicted richness was broadly similar among the five study sites,
although the more remote Palé area showed the highest richness values, especially for

the endemic and threatened species (Fig. 3). The Bray-Curtis dissimilarity values show a 497 similar geographic pattern, with the Palé area having higher dissimilarity values 498 compared to the other four areas, indicating less defaunation (Table 2). For the three 499 endemic species, Bach Ma NP showing the lowest predicted occupancy, followed by the 500 Hue Saola NR, Quang Nam Saola NR, Xe Sap NPA, and Palé. These findings indicate a 501 strong landscape-scale defaunation gradient for the three endemic species (Fig. 4). Our 502 covariate responses suggest that this gradient reflects an increasing level of remoteness 503 (Fig. 2). Bach Ma NP lies near densely-populated coastal areas of Vietnam, has lower 504 average elevations, and has been accessible for decades by a well-established road 505 network. In the westernmost section is the Palé area, which is far from major cities, has 506 few villages, higher elevations, and has only recently been accessible by road. The Saola 507 508 NRs and Xe Sap NPA fall between these two extremes. Furthermore, these areas have 509 had some level of active enforcement in the last few years, which may have slowed the 510 decline of mammal populations.

511 Our results provide information that is directly applicable to conservation planning in this landscape. From a biogeographic perspective, protecting the Palé area is a top 512 513 priority for the conservation of threatened and endemic species. Indeed, it may be the only place in our survey sites to harbor Owston's civet, Asian black bear Ursus 514 thibetanus, and marbled cat (Fig. S7). Our predictive maps offer a robust scientific 515 framework to support ongoing initiatives to grant Palé formal protected area status as a 516 first step to implementing active protection measures. Our maps also provide 517 information to guide targeted snare-removal efforts within protected areas. This 518 519 information is especially useful for the Hue and Quang Nam Saola NRs, where WWF and local partners are operating snare-removal teams, but have not yet been able to 520 521 significantly reduce snaring pressure across the wider protected area complex

(Wilkinson, 2016). We suggest that, to maximize the impact of snare-removal efforts on 522 conservation-priority species, the teams should focus on the more remote areas along 523 the border of the two reserves, and in the border area of the Quang Nam Saola NR and 524 Bach Ma NP. It is possible that these areas have maintained higher occupancies of 525 conservation-priority species because they are more difficult to access and, as a result, 526 have cumulatively experienced less snaring pressure. However, remoteness will not 527 protect these areas for long. An increase in road development in recent years has 528 created a situation where even the most remote locations in the Saola NRs can now be 529 reached within a single day from the nearest access point, meaning that no area is 530 inaccessible for a motivated hunter. Given the likely relationship between accessibility 531 and increased hunting pressure, it seems inevitable that, in the absence of scaled-up 532 enforcement efforts, snaring pressure will continue to increase in the more remote 533 534 areas, especially as other parts of the protected areas become increasingly empty and poachers are forced to travel further distances to maintain comparable levels of offtake 535 536 (Kümpel et al, 2010). Although threatened and endemic species appear to be absent from much of Bach Ma NP, there are isolated high-elevation areas that should be 537 considered for intensive anti-poaching efforts. Our models indicate that the border areas 538 of eastern Xe Sap NPA Area have undergone moderate to severe defaunation. Given that 539 this area appears to be heavily hunted by Vietnamese poachers (Tilker, 2014), such a 540 finding is not unexpected. The eastern section of the protected area is nonetheless a top 541 priority for continued protection efforts, both because it may be a stronghold for the 542 endemic Annamite striped rabbit (Fig. S7), and more generally because effective 543 enforcement can serve as a buffer from further cross-border incursions into the Palé 544 area. In a best-case scenario, reducing snaring pressure in core areas within this 545 546 landscape could not only prevent the local extirpation of conservation-priority species,

but also allow their populations to rebound (see Steinmetz et al, 2010 for a case study
on large mammal population recovery following mitigation of unsustainable hunting
pressure).

In many ways, our study landscape exemplifies classical "empty forest syndrome" 550 (Redford, 1992). All large and medium-sized predators (with the exception of Asiatic 551 black bear), as well as all megaherbivores, appear to be locally extirpated (Tilker & 552 Abrams et al, in press). Large ungulates have been hunted out from most of the 553 landscape (Fig. S7). Yet our findings show that even in this empty forest, conservation-554 priority species still persist, albeit at extremely low occupancies. Based on these results, 555 we suggest that the conservation potential of defaunated landscapes should not 556 automatically be dismissed in the absence of comprehensive surveys. It is important that 557 558 such surveys use sufficient sampling effort and be conducted over a large spatial extent for two reasons. First, species often show extreme spatial heterogeneity in defaunated 559 560 landscapes because local extinctions necessarily result in reduced, often patchy, 561 distributions. Surveys over wider areas are more likely to detect remnant populations. Second, working over larger spatial scales may better capture the underlying factors 562 563 influencing species distribution, which can be especially important in landscapes characterized by complex anthropogenic pressures operating at multiple spatial scales. 564 In our study, it was only by sampling the wider forest complex that we were able to 565 adequately characterize the full spectrum of anthropogenic factors that appear to impact 566 species occurrence patterns. Large-scale surveys require substantial resources. We 567 acknowledge that, with multiple competing conservation objectives and finite resources, 568 569 landscape-scale surveys may not always be possible. However, we note that because this approach can enhance the efficiency of targeted interventions, it is possible that limited 570 571 conservation resources may be saved in the long term.

To overcome the challenge of detecting rare and elusive tropical mammal species, we 572 used two complimentary survey methods: camera-trapping and leech collection. 573 Although camera-trapping detected more species overall, leeches provided our sole 574 detection for marbled cat, and doubled the number of records for two rare species, 575 Owston's civet and Asian black bear. Moreover, while camera-trapping detection 576 probabilities were higher for most species in our analysis, leeches had higher average 577 detection rates for both Asian black bear and the endemic dark muntjac (Fig. S7). Our 578 results are consistent with the findings of Abrams et al (2019) that demonstrate the 579 advantages of using both camera-trapping and eDNA to increase detection probabilities 580 for tropical mammals. We further suggest that because utilizing multiple methods may 581 increase detections of rare species, this approach could be especially important when 582 surveying faunally impoverished systems. Future surveys using joint detection methods 583 584 need not rely only on leeches but could use other sources of eDNA, such as water (Ushio et al, 2017) or ticks (Gariepy et al, 2012), or incorporate other noninvasive sampling 585 586 techniques, such as acoustic monitoring devices (Kalan et al, 2015). The Bayesian modeling approach that we used, adapted from Abrams et al (2019), is flexible with 587 regard to the underlying detection method used to generate spatial or temporal 588 replicates. 589

We found that species occurrence in our study area appears to be primarily driven by anthropogenic factors, with no strong influence from the habitat covariates that we assessed in our models (Fig. S3). The lack of a strong signal with the habitat covariates was unexpected, given the importance of vegetation structure in explaining mammal occurrence patterns in other tropical rainforests (Goulart et al, 2009; Mathai et al, 2017; Sollmann et al, 2017). One possible explanation is that, as anthropogenic pressures in a landscape increase, ecological relationships weaken. Several hypothetical scenarios

597 could give rise to this situation. Spatially nonrandom hunting pressure could, for example, differentially impact areas of preferred habitat, leaving higher occupancies in 598 less suitable areas. Alternatively, intensive hunting across a landscape could drive 599 stochastic local extinctions, leaving remnant populations that are distributed randomly 600 with respect to habitat. Regardless of the underlying process, the failure of habitat-based 601 indices to reflect faunal biodiversity, thus the "environmental decoupling" of species-602 habitat relationships, has broad implications. Biodiversity assessments that rely solely 603 on remote-sensed habitat-based measures may provide information that is inaccurate 604 because they do not accurately capture species occurrence patterns. Recently, a growing 605 606 number of scientists have called for the development of standardized remote-sensing 607 parameters, often referred to as Satellite Remote-Sensing Essential Biodiversity Variables (SRS-EBVs), to monitor biodiversity at the global scale (O'Connor et al, 2015; 608 Skidmore et al, 2015; Pettorelli et al, 2016). While we acknowledge the value of earth 609 observation data to provide insight into biodiversity patterns and processes at large 610 611 scales (Bush et al, 2017), our results indicate that remote-sensed habitat-based measures may provide little information on the status or distribution of wildlife in 612 defaunated landscapes. In tropical rainforests subject to hunting pressure, there is likely 613 no substitute for large-scale in situ surveys to collect primary biodiversity data. 614

Our results underscore the importance of incorporating anthropogenic factors in studies that seek to explain or predict species occurrence in landscapes characterized by high human pressure. Furthermore, our findings suggest that to build robust distribution models it may be beneficial to incorporate a diverse suite of anthropogenic covariates that capture different aspects of this pressure. We used measures of village density and city distance as proxies for accessibility at the local and landscape scales, respectively. Previous studies have shown the impact of similar accessibility measures on wildlife

communities at different spatial scales (Schuette et al, 2013; Koerner et al, 2017; Torres 622 et al, 2018). Our least cost path covariate adds an additional dimension to these 623 accessibility measures, both because it takes into account the ruggedness of the terrain 624 in our landscape, and because it is calculated over a 20-year window. We see potential 625 for further development of anthropogenic covariates that include both spatial and 626 temporal components. Finally, we use elevation as a proxy for both local accessibility 627 and a complex set of ecological attributes. The relative contribution of anthropogenic 628 and environmental traits to species occurrence along elevational gradients in the 629 Annamites represents an intriguing question. Future studies in the region that measure 630 631 a wider range of microhabitat characteristics, and ideally are conducted in areas under 632 less severe hunting pressure, may provide insight into this issue.

Given the current magnitude of hunting across the world's tropical rainforests 633 (Harrison, 2011; Ripple et al, 2016; Benítez-López et al, 2017), and future projections for 634 population growth (Gerland et al, 2014) and road expansion in developing countries 635 (Lawrence et al, 2014), it is likely that defaunation will become increasingly prevalent in 636 637 tropical regions. Confronting the pantropical defaunation crisis will require a wellresourced, multi-faceted approach from conservation stakeholders worldwide. Because 638 specific threats and potential solutions necessarily depend on local context, effective 639 640 strategies to prevent unsustainable hunting must be site-specific. One constant that is applicable to conservation initiatives in all tropical hotspots, however, is that resources 641 are limited. We show that, within this context, understanding spatial patterns of 642 defaunation can help stakeholders prioritize areas for conservation activities, and 643 therefore more effectively use finite conservation resources. 644

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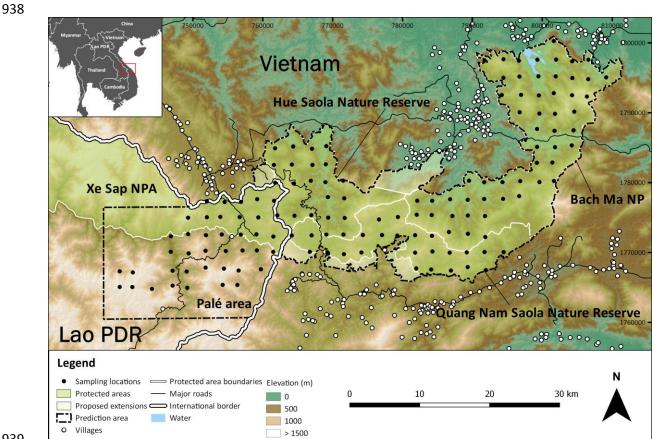
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Figures 937



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Figure 1: Map of camera-traps and leech collection stations across five areas in the 940

central Annamites landscape. 941

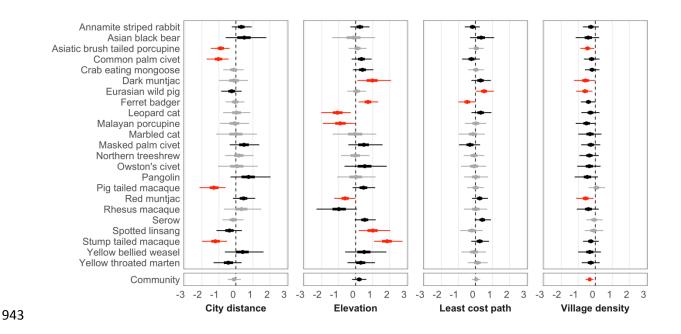


Figure 2: Standardized beta coefficients (mean, 95% BCI, 75% BCI, on the logit scale) showing covariate effects on species occupancy Gray bars show relationships in which the 75% Bayesian Credible Interval (BCI) does not overlap zero, black bars indicate that the 75% interval does not overlap zero but the 95% interval does overlap zero, and red bars indicate that the 95% interval does not overlap zero. The community response is shown in the lower panel.



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Figure 3: Predicted species richness across five study sites in the central Annamites landscape with histogram showing proportion of cells in each study area for predicted species numbers (left panel) and prediction map (right panel) based on community occupancy model fit to camera-trapping and iDNA data for 23 mammal species. (A) Predicted richness for all species. (B) Predicted species richness for threatened and endemic species.

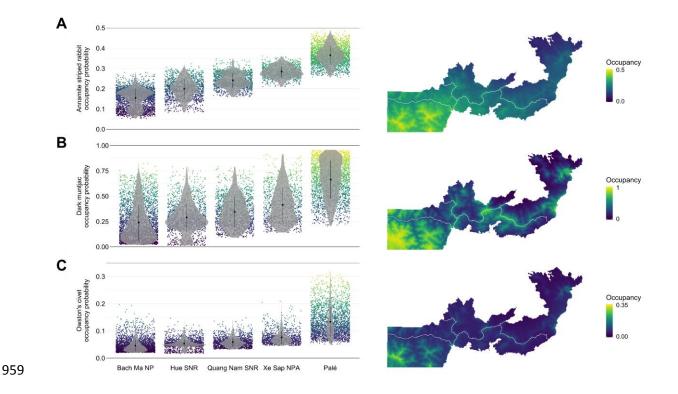


Figure 4: Predicted occupancies for three Annamite endemic species with violin plot 960 showing predicted occupancy values for each of the five study sites (left panel) and 961 prediction map (right panel) based on community occupancy model fit to camera-962 trapping and iDNA data for 23 mammal species. (A) Predicted occupancy for Annamite 963 striped rabbit. (B) Predicted occupancy for Annamite dark muntjac. (C) Predicted 964 occupancy for Owston's civet. Note that, for visualization purposes, occupancy values 965 are scaled independently for each species. Single-species prediction maps with 966 standardized scaling for occupancy values can be found in Fig. S7. 967

| | | Study site | | | | | All sites |
|------------------|--------------------------------|-------------|--------------------|--------------|-------------|-------------|-------------|
| | | Bach Ma NP | Quang Nam Saola NR | Hue Saola NR | Xe Sap NPA | Palé | |
| ichness | Full community | 6.52 ± 1.39 | 6.32 ± 1.00 | 7.00 ± 1.24 | 6.89 ± 0.98 | 8.25 ± 1.13 | 7.02 ± 1.41 |
| Species richness | Threatened and endemic species | 1.79 ± 0.77 | 1.80 ± 0.65 | 1.85 ± 0.66 | 2.10 ± 0.70 | 2.97 ± 0.85 | 2.12 ± 0.89 |
| occupancy | Annamite striped rabbit | 0.15 ± 0.04 | 0.24 ± 0.04 | 0.20 ± 0.05 | 0.28 ± 0.03 | 0.36 ± 0.05 | 0.24 ± 0.09 |
| | Annamite dark muntjac | 0.24 ± 0.18 | 0.35 ± 0.15 | 0.30 ± 0.18 | 0.41 ± 0.17 | 0.66 ± 0.19 | 0.39 ± 0.24 |
| Species | Owston's civet | 0.05 ± 0.02 | 0.06 ± 0.02 | 0.05 ± 0.02 | 0.08 ± 0.02 | 0.13 ± 0.05 | 0.07 ± 0.05 |

Table 1: Predicted species richness and species occupancies (mean ± SD) for five study sites in the central Annamites landscape, from
 multi-species community occupancy model fit to 23 mammal species. Full community indicates richness for all 23 species. Threatened and
 endemic species indicates richness for 10 species that are endemic and / or listed as Near Threatened or higher on *The IUCN Red List of Threatened Species.* Species occupancy shows predicted occupancies for each of the three Annamite endemic mammals. Occupancy values
 range from 0 to 1.

| | Bach Ma NP | Quang Nam Saola NR | Hue Saola NR | Xe Sap NPA | Palé |
|--------------------|----------------------|----------------------|----------------------|---------------------|---------------------|
| Bach Ma NP | 0 | -0.0006 ± 0.0205 | 0.0406 ± 0.0119 | 0.0350 ± 0.0233 | 0.1338 ± 0.0315 |
| | | 0.0431 ± 0.0205 | 0.0049 ± 0.0119 | 0.0841 ± 0.0233 | 0.2547 ± 0.0315 |
| Quang Nam Saola NR | 0.0006 ± 0.0205 | 0 | 0.0412 ± 0.0206 | 0.0357 ± 0.0061 | 0.1346 ± 0.0193 |
| | -0.0431 ± 0.0205 | | -0.0382 ± 0.0205 | 0.04127 ± 0.006 | 0.2150 ± 0.0193 |
| Hue Saola NR | -0.0406 ± 0.0119 | -0.0412 ± 0.0205 | 0 | -0.0055 ± 0.0206 | 0.0938 ± 0.0297 |
| | -0.0049 ± 0.0119 | 0.0382 ± 0.0206 | | 0.0793 ± 0.0206 | 0.2503 ± 0.0297 |
| Xe Sap NPA | -0.0350 ± 0.0233 | -0.0357 ± 0.0061 | 0.0055 ± 0.0206 | 0 | 0.0995 ± 0.0174 |
| | -0.0840 ± 0.0233 | -0.0413 ± 0.006 | -0.0792 ± 0.0206 | | 0.1754 ± 0.0174 |
| Palé | -0.1338 ± 0.0315 | -0.1346 ± 0.0193 | -0.0938 ± 0.0297 | -0.0995 ± 0.0174 | 0 |
| | -0.2547 ± 0.0315 | -0.2149 ± 0.0192 | -0.2502 ± 0.0297 | -0.1754 ± 0.0174 | |

975

976 **Table 2:** Bray-Curtis dissimilarity values (mean ± SE) calculated using predicted occupancy values per site for all mammal species (top

value) and for threatened and endemic species (bottom value). Values range from 0 to 1, with 0 indicating complete similarity in species

⁹⁷⁸ occupancies between the sites, and 1 indicating complete dissimilarity.