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Convergent evolution of polyploid genomes from across the eukaryotic tree of life

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Abstract:

By modeling the homoeologous gene losses that occurred in fifty genomes deriving from ten distinct polyploidy events, we show that the evolutionary forces acting on polyploids are remarkably similar, regardless of whether they occur in flowering plants, ciliates, fishes or yeasts. The models suggest these events were nearly all allopolyploidies, with two distinct progenitors contributing to the modern species. We show that many of the events show a relative rate of duplicate gene loss prior to the first post-polyploidy speciation that is significantly higher than in later phases of their evolution. The relatively low selective constraint seen for the single-copy genes these losses produced lead us to suggest that most of the purely selectively neutral duplicate gene losses occur in the immediate post-polyploid period. We also find ongoing and extensive reciprocal gene losses (RGL; alternative losses of duplicated ancestral genes) between these genomes. With the exception of a handful of closely related taxa, all of these polyploid organisms are separated from each other by tens to thousands of reciprocal gene losses. As a result, it is very unlikely that viable diploid hybrid species could form between these taxa, since matings between such hybrids would tend to produce offspring lacking essential genes. It is therefore possible that the relatively high frequency of recurrent polyploidies in some lineages may be due to the ability of new polyploidies to bypass RGL barriers.

53 **Introduction**

54 That organisms with doubled genomes existed was evident early in the history of genetics
55 (Kuwada 1911; Clausen and Goodspeed 1925), and a lively debate was entered as to the
56 implications of this fact. Wagner (1970) declared polyploidy to be “evolutionary noise” the same
57 year that Susumu Ohno (1970) was giving it pride of place among the forces generating
58 evolutionary innovations. The advent of genome sequencing changed the ground of this debate,
59 opening new horizons of time for studies of the prevalence and influence of polyploidy. We
60 know now that great branches of the eukaryotic evolutionary tree, including the vertebrates, all
61 flowering plants and many yeasts, descend from ancient polyploids (Van de Peer, et al. 2017),
62 events that were difficult or impossible to detect with older data. For reasons that are not yet
63 fully understood, many of these groups also show recurrent polyploidies, especially flowering
64 plants (Soltis, et al. 2009) and teleost fishes (Braasch and Postlethwait 2012).

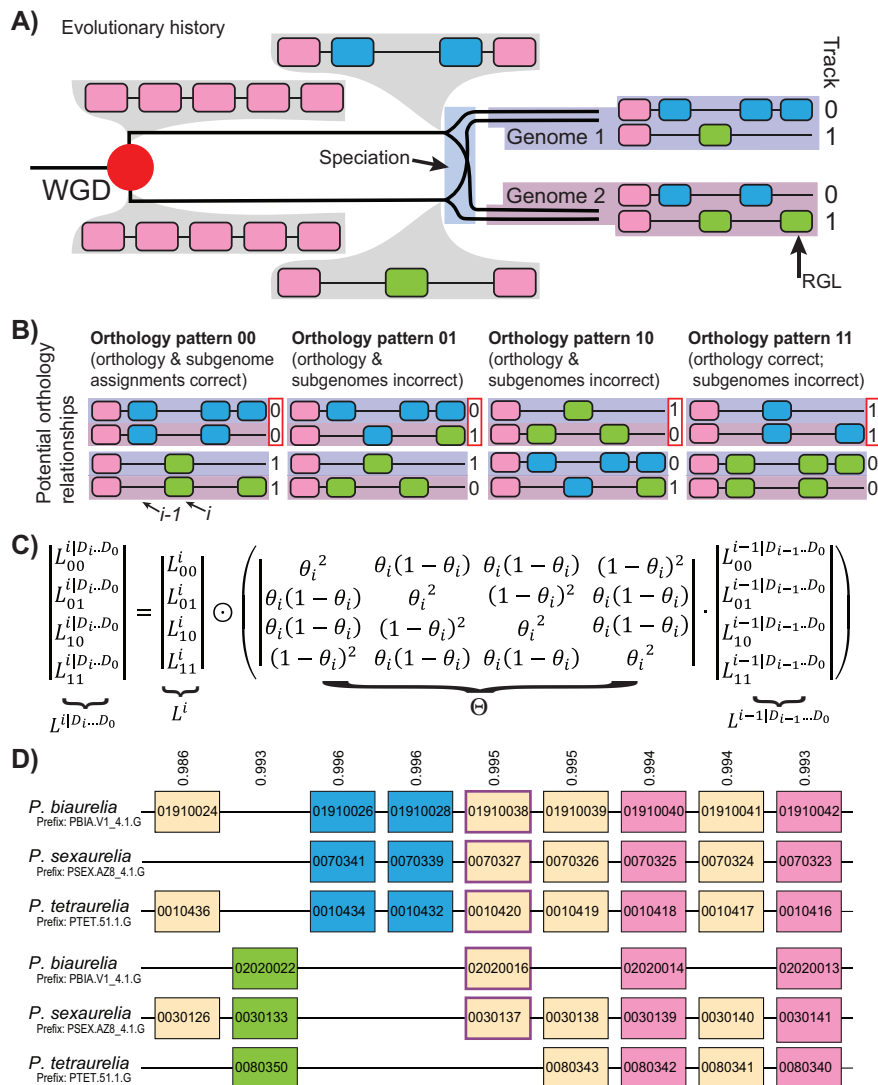
65 With this extensive new set of polyploidies as a resource, other old questions can also be
66 revisited, such as the relative prevalence of auto- and allopolyploids (Stebbins Jr 1947).
67 Allopolyploidy refers to hybridizations between distinct species that result in doubled (or more)
68 genomes, while autopolyploids are derived from a single progenitor species (Kuwada 1911;
69 Clausen and Goodspeed 1925; Stebbins Jr 1947). Analyses of several paleopolyploid genomes
70 have shown that while gene losses are common after polyploidy, in many cases the losses are not
71 experienced equally by the two parental subgenomes (Thomas, et al. 2006; Emery, et al. 2018), a
72 pattern known as biased fractionation. These biases are plausible but not definitive indicators of
73 allopolyploidy.

74 There has also been controversy as to whether and how polyploidy affects the rate of
75 speciation. Werth and Windham (1991) proposed that reciprocal gene losses (RGLs), the
76 alternative loss of one of the two duplicated genes from different populations, could create
77 Bateson–Dobzhansky–Muller incompatibilities between populations, because matings between
78 them would give rise to offspring with no copies of the genes. Were those genes essential, the
79 offspring lacking them would be inviable (Werth and Windham 1991) Such incompatibilities
80 have been observed both in the wild and the laboratory (Mizuta, et al. 2010; Maclean and Greig
81 2011). Muir and Hahn (2015) emphasize that RGL requires a period of reproductive isolation to
82 form.

83 In the case of the yeast polyploidy, RGLs are commonly found between the descendant
84 genomes, suggesting the potential for polyploidy to create new species by purely neutral means
85 (Scannell, et al. 2006; Scannell, et al. 2007). However, direct analyses of the speciation and
86 extinction rates of polyploid and nonpolyploid lineages has yielded inconclusive results, with
87 some studies claiming reduced net diversification rates among polyploids and others disagreeing
88 (Mayrose, et al. 2011; Soltis, Segovia-Salcedo, et al. 2014). More generally, the immediate and
89 long-term adaptive value of polyploidy remains unclear: for instance, allopolyploids combine
90 hybridizations with genome doubling and may derive immediate advantages from the
91 hybridization effects rather than the doubling itself (Soltis, Visger, et al. 2014). Increased stress
92 tolerance in polyploid organisms has also been invoked to argue for a radiation of polyploidy
93 coincident with global catastrophes such as the KT mass extinction (Fawcett, et al. 2009).

94 Using our tool for modeling the evolution of polyploid genomes, POInT (the Polyploidy
95 Orthology Inference Tool; Conant and Wolfe 2008), we explored the resolution of ten
96 independent polyploidies. We adopt the term “homoeolog” below to refer to homologous genes
97 produced by any type of polyploidy rather than “duplicate” or “ohnolog” because the events
98 considered comprise several distinct types of polyploidy. The hallmark of polyploidy in a
99 genome is a pattern of interleaved synteny, comprising not just the surviving homoeologs but
100 also single-copy genes that are now found in interleaved positions on pairs (or more) of
101 chromosomal segments homologous to the ancestral single-copy regions. In Figure 1A, we show
102 an example of this evolutionary process, which yields conserved synteny blocks in the extant
103 genomes. Those synteny blocks differ between genomes, meaning it is necessary to “phase”
104 them into orthologous regions. As shown in Figure 1B, for a set of n tetraploid genomes, there
105 are 2^n possible orthology relationships at each ancestral locus. We use the term “pillar” to denote
106 all of the genes or lost homoeologs at such a locus. POInT computes the likelihood of the
107 observed homoeolog presence/absence data at each pillar for each possible orthology
108 relationship. Via a hidden Markov model (HMM) that combines the possible orthology
109 relationships for each pillar with the syntenic organization among pillars (Figure 1C), POInT
110 employs posterior decoding to infer orthology estimates for each pillar with associated posterior
111 probabilities (top of Figure 1D) as well as estimates of the model parameters describing the
112 process of homoeolog loss (Figure 2B and C).

113 Our analyses here encompass a total of 50 polyploid genomes and more than 460,000
 114 individual genes (Figure 2A). We find that the patterns of gene loss after these different events
 115 show strikingly similar patterns, with strong evidence for biased fractionation and homoeolog
 116 fixation. Using synonymous substitutions as an evolutionary clock, we show that the rate of gene
 117 loss immediately after a polyploidy is generally higher than in later periods. RGL is also
 118 prevalent after all of these polyploidies, and we suggest it might introduce barriers to
 119 hybridization that could be overcome through subsequent allopolyploidy events.
 120



121 **Figure 1: Inferring orthologous chromosome regions between polyploid genomes with POInT (the**
 122 **Polyploidy Orthology Inference Tool). A) Cartoon model of gene losses and a speciation event**
 123 **after a whole-genome duplication. Immediately after the WGD, all five genes are present in two**
 124 **homoeologous copies. Three homoeologous gene losses occur prior to the split of the two**
 125 **species, one in the less fractionated subgenome (Track “0;” yielding the green gene in the lower**
 126 **window) and two from the more fractionated subgenome (Track “1;” yielding the two blue genes**
 127 **in the upper window). After the speciation event, Genome 1 loses a homoeolog from the more**
 128

129 fractionated subgenome and Genome 2 loses one from the less fractionated subgenome, a case
130 of reciprocal gene loss (RGL). **B**) There are $2^n=4$ potential ways of phasing the two chromosomal
131 regions from Genome 1 relative to Genome 2 (e.g., of assigning orthology between the two
132 regions). We identify these 4 states with the subgenome assignment for the top track for each of
133 the two genomes (00→11; red boxes at the right of each diagram). POInT uses a model of
134 homoeolog loss to compute the likelihood of the observed gene presence/absence data at each
135 locus (or “pillar”) for each of these 2^n relationships. These relationships each constitute a hidden
136 state of the HMM implemented by POInT whereas a likelihood of observed gene
137 presence/absence data for a relationship represents an emission probability for the HMM. **C**)
138 Recurrence equation for computing the likelihood of each orthology assignment at pillar i
139 conditional on the data at pillars 0 through $i-1$ (see **B**). For pillar i , we define a vector L^i to be the
140 likelihood of the orthology states, with elements L_{00}^i , L_{01}^i , L_{10}^i and L_{11}^i being POInT’s estimates of
141 the likelihood of each such state based on the gene presence/absence data at that pillar. We then
142 use a transition probability matrix Θ , with each entry representing the probability that pillar i has a
143 particular orthology state conditional upon another orthology state at $i-1$. The probability that the
144 orthology state is maintained between pillars $i-1$ and i is $1-\theta_i$ for each genome (and $(1-\theta_i)^2$ in total);
145 the chance that one genome changes orthology state is $\theta_i(1-\theta_i)$ and the chance that both change
146 is θ_i^2 . Here, $\theta_i=\theta$, a global constant estimated from the data by maximum likelihood, except when
147 synteny is not maintained between pillars, in which case $\theta_i=0.5$ (adjacent pillars do not inform on
148 each other’s orthology state; *Methods*). To compute a likelihood for the entire data set, POInT
149 implements an HMM forward algorithm that expresses $L^{i|D_i\dots D_0}$, the probabilities of orthology
150 relationships for pillar i and the observed data at pillars 0 through i (denoted $D_i \dots D_0$), in terms of
151 the emission probabilities L^i , the transition probabilities Θ and the probabilities $L^{i-1|D_{i-1}\dots D_0}$ that
152 were already computed for pillar $i-1$. The vector of $L^{i|D_i\dots D_0}$ is then the element-wise vector product
153 (indicated with the “ \odot ”) of $\Theta \cdot L^{i-1|D_{i-1}\dots D_0}$ and L^i . This formula can be applied sequentially starting
154 at pillar 0, with the base case $L^{0|D_0} = L^0$. For m pillars, the overall likelihood of the dataset is then
155 the sum of the elements of $L^{m|D_m\dots D_0}$. **D**) Given an inferred ancestral gene order prior to the
156 polyploidy (*Methods*), POInT employs posterior decoding to infer the orthology relationships at
157 each pillar. Here we illustrate a small region of such an ancestral order from the more recent
158 *Paramecium* WGD (after phasing from the earlier duplication, see *Methods*), showing the set of
159 orthology relationships inferred by posterior decoding. For reference, genes in adjacent pillars
160 that are also neighbors in an extant genome are shown connected by lines. The number above
161 each pillar is the posterior probability of the inferred orthology relationship. The upper set of three
162 tracks correspond to the less-fractionated parental subgenome, the lower three to the more
163 fractionated one, illustrating the possibility for local variation in biased fractionation. Gene
164 retained from *only* the less-fractionated genome are colored blue, from *only* the more fractionated
165 one green, and fully retained duplicates are shown in pink. All other patterns of duplicate retention
166 are shown in beige for clarity. See also Figure 2B.

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169 **Results**

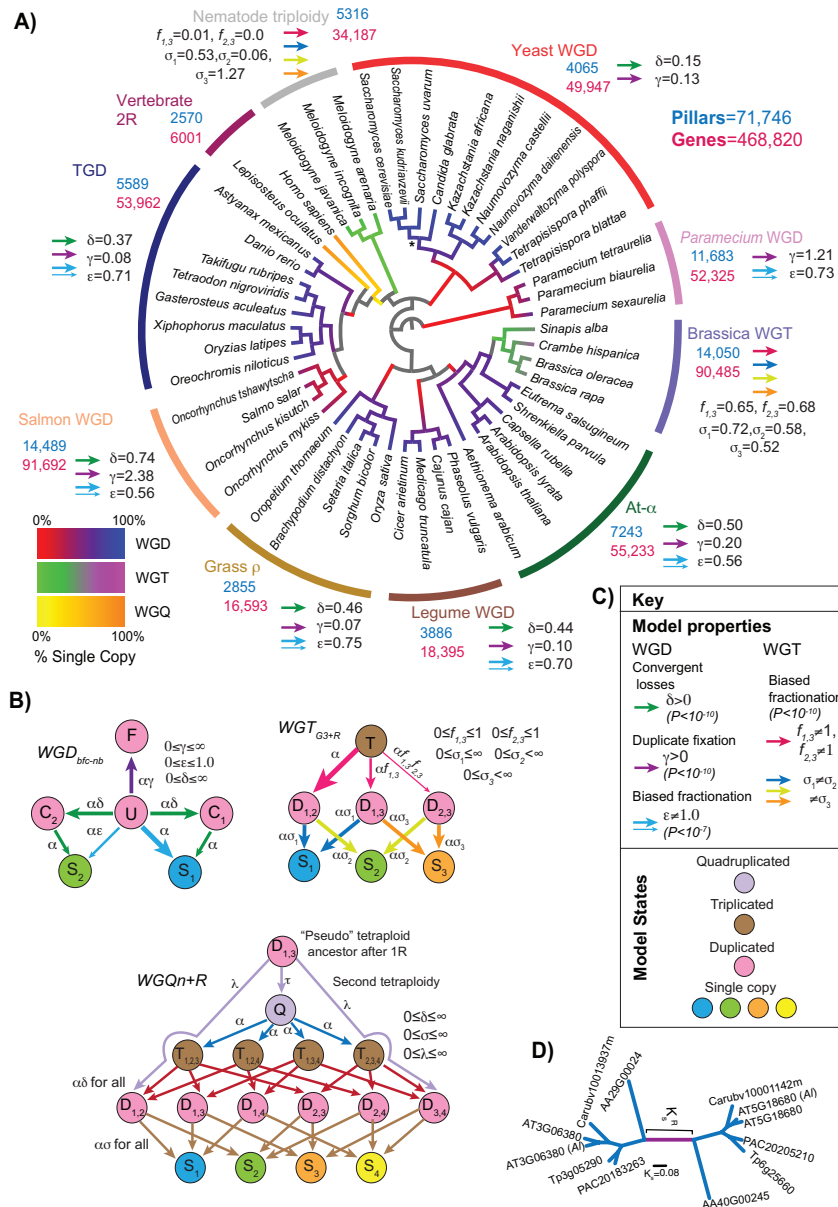
170 *Modeling evolution after ten independent polyploidies.*

171 Using POInT, we assembled a set of ~70,000 homoeologous loci produced by ten
172 different polyploidies. For each polyploidy, we inferred a set of pillars that it created and ordered
173 them so as to maximize the retained synteny among the extant genes, approximating the
174 ancestral order of the single-copy genes just prior to polyploidy (*Methods*). Six of the events are
175 whole genome duplications (WGDs or tetraploidies): At- α in *Arabidopsis thaliana* and its

176 relatives, a WGD found in legumes, the ρ event from grasses, the teleost-specific genome
177 duplication (TGD), and WGDs from salmonids and yeasts. We further analyzed an asexual
178 triploidy in nematodes, a hexaploidy (whole genome triplication; WGT) in cabbages and their
179 relatives (*Brassica* WGT) and two octoploidies: the vertebrate 2R polyploidy and another in the
180 paramecia (Figure 2A). Analyzing octoploidies in POInT is computationally expensive. As a
181 result, we modeled the octoploidy among the paramecia as occurring via two sequential genome
182 duplications and then extracted and analyzed only the more recent of these two events for the
183 remainder of our work (*Methods*). This approach failed with the vertebrate 2R event, presumably
184 because the two events are very ancient and closely spaced in time. A visual interface to these
185 data is available from the POInT browser (<http://wgd.statgen.ncsu.edu>).

186 For the WGD events, we compared nested models of evolution (Figure 2B and
187 Supplemental Table 1) that describe the process of homoeolog loss after polyploidy: these
188 models differ as to whether they include biased fractionation, duplicate fixation and convergent
189 homoeolog losses. For all seven tetraploidies, models that allow for the fixation of homoeologs
190 after polyploidy fit the observed loss data better than models without such an effect ($\gamma \neq 0$; $P < 10^{-10}$;
191 likelihood ratio test or LRT; Figure 2). In addition, every event save that in yeast shows
192 strong evidence for biased fractionation ($\epsilon \neq 1$; $P < 10^{-7}$; LRT; Figure 2), while all but the
193 *Paramecium* event show a pattern of independent yet convergent losses to the same homoeolog
194 in independent lineages ($\delta \neq 0$; $P < 10^{-10}$; LRT; Figure 2). The nematode triploidy and the *Brassica*
195 WGT also share similar patterns of biased fractionation (Figure 2 and Supplemental Table 2).

196 The fact that these events are of widely differing ages is evident from the different
197 degrees of loss/resolution seen in the extant genomes. The branches of Figure 2A are color-
198 coded by POInT's inferences of the proportion of single-copy genes (e.g., loci where all but one
199 of the homoeologous genes have been lost) present at their beginning and ending. While the
200 yeast WGD is inferred to be nearly "fully" resolved (nearly all homeologous loci reduced to
201 single-copy), the tetraploidy in salmonid fishes and the nematode triploidy show proportionally
202 few single-copy genes. The nematode triploidy differs from the remaining events in that these
203 animals are asexual triploids and are likely under a different selective regime in their gene losses,
204 (Schoonmaker, et al. 2020). The continued occurrence of meiotic chromosome pairings of
205 homoeologous chromosomes created by the salmonid event may have reduced the rate of
206 homoeolog loss in those genomes (Allendorf, et al. 2015).



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Figure 2: Modeling the resolution of ten polyploidy events with POInT. **A)** The assumed phylogenetic relationships of the ten polyploidies studied. Grey branches indicate where no polyploidy event was studied. The relationships of the taxa were inferred from the homoeolog loss data for the legume WGD, grass ρ , nematode triploidy, salmonid WGD and the Paramecium WGD. For the yeast WGD, At- α , the TGD and the Brassica WGT, the relationships were taken from published sources (the vertebrate 2R tree is trivial). Because the temporal divergences of various groups are not well established, the tree is illustrated in an ultrametric format with nonmeaningful branch lengths (Scaled topologies for each event are shown in Supplemental Figure 4). However, each polyploid branch is colored using POInT's estimates of the proportion of loci that were single-copy at its beginning and ending. Corresponding color keys for tetra, hexa and octoploidies are shown. The number of "pillars" (homoeologous loci) and the total number of gene models studied across each event are noted, as are the total number of loci and genes considered. The "*" on the yeast WGD branch indicates the branch where the proportion of genes returned to single-copy that are presently essential was tested (Supplemental Table 5). Next to each event, we show arrows and parameter estimates indicating post-polyploidy evolutionary processes such as biased

224 fractionation for which we found significant evidence in that event (see key in panel **C**). **B**
225 Nested models of post-polyploidy evolution for the three types of events (WGD: whole-genome
226 duplication/tetraploidy, WGT: whole-genome triplication/hexaploidy and WGQ: whole-genome
227 quadruplication/octoploidy). Using POInT, we fit nested models of gene loss after polyploidy
228 with likelihood ratio tests (*Methods*). **WGD**: all pillars start in state **U** (**U**ndifferentiated), from
229 which they can transition to either the three other duplicated states, **C**₁ (**C**onverging state 1), **C**₂
230 (**C**onverging state 2) and **F** (**F**ixed) or to the two single-copy states **S**₁ (**S**ingle-copy 1) and **S**₂
231 (**S**ingle-copy 2). **C**₁ and **S**₁ are states where the gene from the less-fractionated parental
232 subgenome will be or are preserved, and **C**₂ and **S**₂ the corresponding states for the more-
233 fractionated parental subgenome. The null model has parameters $\gamma=\delta=0$ and $\varepsilon=1.0$. Duplicate
234 fixation is inferred when $\gamma\neq 0$, convergent losses when $\delta\neq 0$ and biased fractionation when
235 $\varepsilon<1.0$. **WGT**: in the base model all pillars start in state **T** (**T**riplicated) and transition first to
236 duplicated states (**D**_{x,y}) and hence to the single-copy states (**S**_x). Genome 1 is assumed to be
237 favored (fewer losses) and the identity of that genome inferred in the POInT computation.
238 Losses from the triplicated state are then increasingly disfavored first to **D**_{1,3} (parameter $f_{1,3}$)
239 and **D**_{2,3} (parameter $f_{2,3}$). There are also individual rates of loss from the duplicated to single-
240 copy states (δ_x). In the null model, $f_{1,3}=f_{2,3}=1.0$ and $\delta_1=\delta_2=\delta_3$. We also fit a separate model that
241 allow this set of parameters to take on separate values on the root branch and on the remaining
242 branches (Supplemental Table 1). **WGQ**: Models of octoploid formation. The null model simply
243 treats the four subgenomes as equivalent and as starting in the quadruplicated state (**Q**). This
244 model has different loss rates from triplicated to duplicated loci (**T**_{x,y,z} to **D**_{x,y}, parameter δ) and
245 duplicated to single-copy loci (**D**_{x,y} to **S**_x, parameter σ). A formation model for the octoploidy can
246 then be added: all pillars start in state **D**_{1,3} and can symmetrically experience a gene loss from
247 genome 1 or 3 (parameter λ) and transition to state **D**_{1,2} or **D**_{3,4} or become quadruplicated (null
248 transition). The three models illustrated here are the most complex model fit to the various
249 events, including the parameters associated and their numerical range. **C**) Description of the
250 various modeled features from panels **A** and **B** (top) and the model states from **B** (bottom). **D**)
251 An example mirrored gene tree for a completely retained set of homoeologs from At- α ,
252 illustrating the trees from which synonymous divergences were estimated. The branch lengths
253 are given in number of synonymous substitution per synonymous site (e.g., K_s), with the shared
254 internal (e.g., “root”) branch shown in purple (K_s^R). For analysis purposes, the length of this
255 branch was always divided by two to be comparable to the remaining branches (e.g., split at its
256 midpoint).

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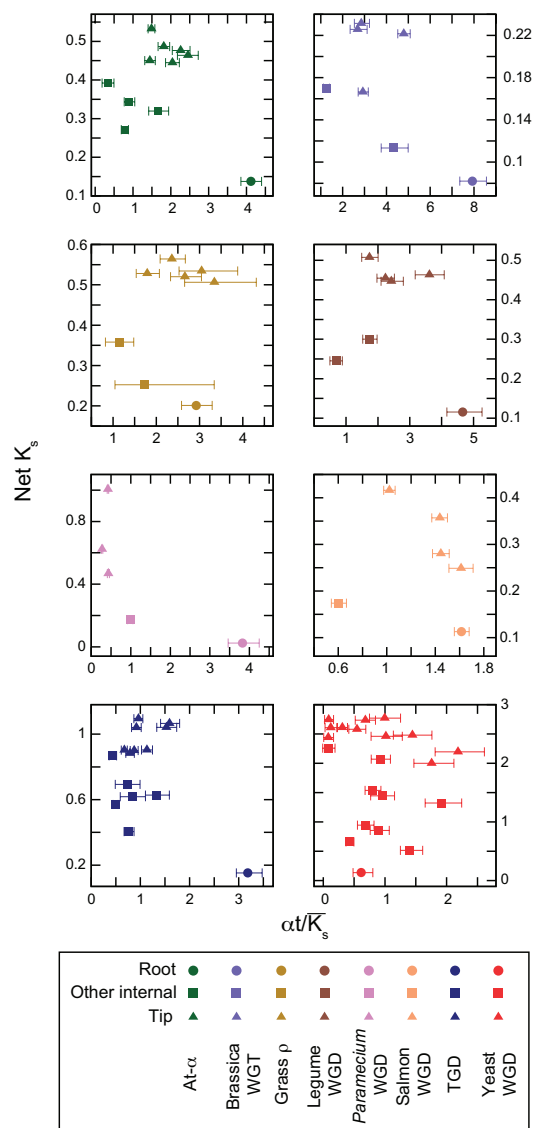
259 *Many events show rapid homoeolog loss immediately after polyploidy.*

260 Loss of duplicate genes immediately after polyploidy can be rapid (Scannell, et al. 2006;
261 Scannell, et al. 2007), and at least two non-exclusive hypotheses exist as to why. The first is that
262 genetic drift should eliminate truly redundant gene copies quickly (Li 1980; Lynch and Conery
263 2000). The second is the potential for “selected” duplicate losses. These losses might occur if the
264 increases in gene copy number after polyploidy induce disadvantageous dosage conflicts, such
265 that natural selection acts to remove the homoeologous copies in question (Edger and Pires 2009;
266 De Smet, et al. 2013).

267 To study the pattern of early losses, we examined the divergence that occurred
268 immediately after the polyploidy and prior to any speciation events. In the context of a gene tree
269 for a pair of homoeologous genes produced by a WGD, this period corresponds to the internal

270 branch of the gene tree separating that pair of homoeologs. For a WGT, the situation is
271 analogous except that there are three such branches separating the three homoeologous copies.
272 For simplicity, we refer to these branch(es) as the “root” (purple in Figure 2D). For all branches
273 in each polyploidy, we obtained a rough estimate of the time encompassed by that branch by
274 using the mean number of synonymous substitutions per synonymous site (\overline{K}_s) across many
275 homoeologous genes as a neutral clock (*Methods*). The rate of homoeolog loss for each branch is
276 given by POInT’s branch length estimate (α), computed with an irreversible exponential loss
277 model proportional to the number of homoeologous copies at the beginning of that branch
278 (meaning that they are not biased by the fact that later branches have fewer total homoeologs
279 available for loss, *Methods*). The ratio of $\alpha \cdot t / \overline{K}_s$ gives a sense of whether homoeolog losses per
280 time are unusually high or low for a given branch relative to other branches in the same
281 polyploidy. For the majority of the polyploidies, we found that the $\alpha \cdot t / \overline{K}_s$ ratio was higher for
282 the root branch than any other branch, consistent with a more rapid loss of homoeologs along
283 this branch (Figure 3). This result is the more striking because the inferred mean K_s value for the
284 root branch (\overline{K}_s^R) should, in the case of an allopolyploidy, also include the pre-polyploidy
285 progenitor divergences. Hence, the \overline{K}_s^R values for these events should be over-estimates, making
286 the $\alpha \cdot t / \overline{K}_s^R$ ratio an underestimate of the relative homoeolog loss rate along the root branch.

287 If natural selection were actively favoring the loss of some homoeologous copies
288 immediately after polyploidy, we might expect that the genes involved in those early losses
289 would display a stronger selective constraint than do homoeologous copies lost later in the
290 polyploidy’s history. We hence compared the average selective constraint, measured as the ratio
291 of nonsynonymous to synonymous substitutions, or K_a/K_s , of fully single-copy genes whose
292 homoeologs were lost along the root branch to that of other fully single-copy genes where the
293 preservation of homoeologous copies from alternative subgenomes means that the losses must
294 have occurred after the first speciation event. For most events we observe little difference
295 between these two groups, while for the Legume WGD the single-copy genes lost later are
296 actually *more* constrained, the opposite of the prediction for selected losses (Supplemental
297 Figure 1).



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Figure 3: Rapid loss of homoeologs immediately after polyploidy. On the x-axis is the ratio of rate of homoeolog loss (the αt branch length estimate from POInT's models, see Figure 2) and the estimated mean synonymous divergence for that branch (\overline{K}_s ; see *Methods*). Hence, larger values of this ratio indicate more homoeolog losses per unit K_s . For the At- α , Brassica WGT, Legume WGD, Paramecium WGD and the TGD, the $\alpha \cdot t / \overline{K}_s$ ratio for the root branch is significantly larger than seen on any other branch (c.f., the 95% confidence intervals shown, computed as described in the *Methods* section). For these panels, we used a model excluding duplicate fixation here because including fixation in the model occasionally results in very long estimates of tip branch lengths (*Methods*). However, our conclusions are similar under the fully WGD_{bfc-nb} model (see Supplemental Figure 7). On the y-axis is the net synonymous divergence to the end of the branch in question: in other words, the sum of the synonymous divergence of that branch and all its ancestors back to the root branch. This net divergence value is a rough indicator of the time since the polyploidy for each branch. The root branch is indicated with a circle, other internal branches with squares and tip branches with triangles.

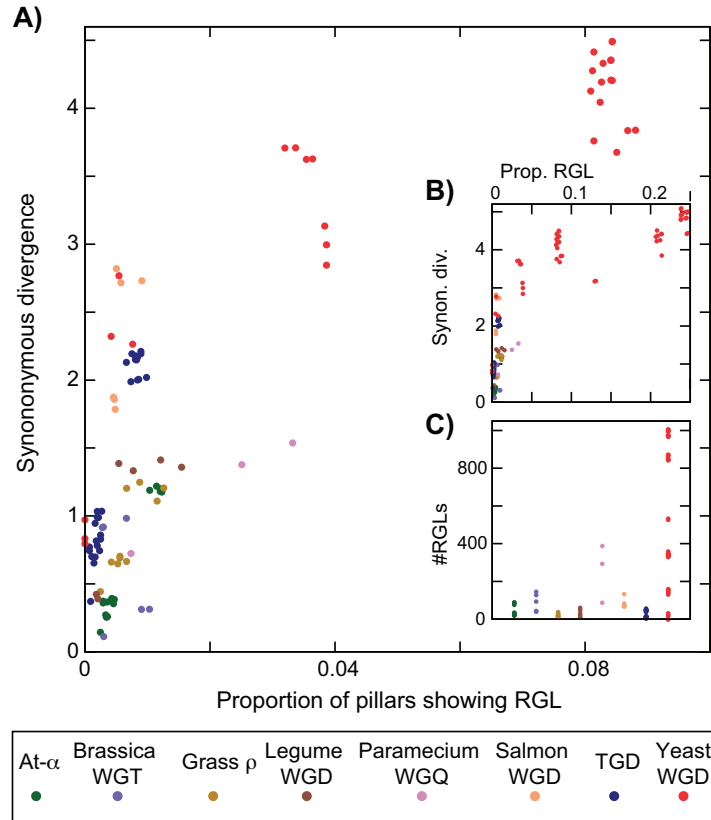
315 *Extensive reciprocal gene loss between pairs of polyploid taxa.*

316 Following Scannell and colleagues (2006; 2007), we searched for post-polyploidy
317 reciprocal gene losses (RGL). We omitted the vertebrate 2R and nematode triploidy from this
318 analysis due to the fragmented nature of the genomes used. With the exception of three closely
319 related yeast species in the *Saccharomyces* genus, every pair of genomes in our remaining eight
320 polyploidies were separated by at least 4 RGLs (this minimal number was seen in the platyfish,
321 tilapia and medaka clade of the TGD; Figure 4C), with the number rising to over a thousand for a
322 few of the yeast taxa pairs. These conclusions are also robust to the confidence cutoffs used to
323 infer the RGLs (Supplemental Figure 2). Our results are in accord with previous work in yeasts
324 and grasses (Scannell, et al. 2006; Scannell, et al. 2007; Schnable, et al. 2012), and there appears
325 to be a relatively direct relationship between the synonymous divergence of a pair of taxa (a
326 proxy for divergence time) and the number of RGLs separating them (Figure 4A and B). Such a
327 relationship would be expected if both RGLs and synonymous substitutions were accumulating
328 through neutral evolutionary processes (Figure 4A). However, the proportionality between
329 synonymous substitutions and RGLs differs between polyploidy events, with the yeast WGD
330 showing more RGLs per unit K_s than the other events. When we compared the genes involved in
331 reciprocal losses in zebrafish, *A. thaliana* and bakers' yeast to other single-copy genes, there
332 were no significant functional differences between these two sets, again as one would expect
333 were RGL a neutral process (*Methods*).

334 The evolutionary importance of RGLs can be assessed by the biological role of the genes
335 that experienced it. For instance, were only “non-essential” genes to experience RGL, then it
336 might not present significant barriers to hybridization. We can use experimental data on gene
337 essentiality from bakers' yeast, *A. thaliana* and zebrafish (*Methods*) to ask whether the
338 proportion of RGLs that include an essential gene differs from the overall proportion of essential
339 single-copy genes. For the At- α and TGD events, the proportion of RGLs where the surviving
340 gene in *A. thaliana* or zebrafish is essential does not differ from the proportion of other single-
341 copy genes that are essential (Supplemental Table 3). Curiously, the RGLs found when
342 comparing bakers' yeast to some of its nearer relatives are actually *more* likely to be essential
343 than other single-copy genes (Supplemental Table 3). This overrepresentation is likely
344 attributable to the shared duplicate losses that occurred prior to the first speciation event being
345 underrepresented in essential genes (Supplemental Table 4). As a result, RGLs, which must have

346 occurred *after* the first speciation event (see the yeast clade of Figure 2A), would be enriched in
 347 essential genes simply because more essential genes survived in duplicate past that first
 348 speciation.

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Figure 4: Reciprocal gene loss (RGL) after polyploidy. **A)** Reciprocal gene losses (RGLs) between pairs of polyploid taxa (x-axis, normalized by the total number of loci/pillars analyzed for that event) as a function of the inferred synonymous divergence of those taxa (y-axis). Panel **A** gives a cropped view that focuses on RGLs in the non-yeast taxa, while panel **B** shows how the RGL frequencies in yeast dramatically exceed those for the remaining events. For each pair of taxa from a given event, we identified all single-copy loci in the two genomes where POInT infers a 95% or greater confidence that those genes are paralogs created by the ancient polyploidy and not more recent orthologs produced by the post-polyploidy speciation events. There are roughly linear relationships between RGL frequency and synonymous divergence. Because the data points shown are phylogenetically dependent (different species pairs share considerable common evolutionary history), we have not attempted to fit regression lines to these data. Standard approaches to phylogenetically independent contrasts (Felsenstein 1985) do not apply here as the inferred RGLs are pairwise species traits and not independent measures on each taxon. It is however notable that the asexually reproducing yeasts appear to accumulate more RGLs per unit K_s than other taxa. **B)** As for **A** but including the full range of RGL prevalence in the taxa sharing the yeast WGD. **C)** Total numbers of RGLs inferred for each pair of taxa for each event.

371 The importance of RGL in driving speciation events among polyploid taxa has been
372 questioned on theoretical grounds, as the appearance of RGLs is subject to the same requirement
373 of reproductive isolation as are the appearances of other genetic incompatibilities among
374 populations (Muir and Hahn 2015). This objection has more force for obligately sexual
375 organisms than it does for organisms such as bakers' yeast, where it is estimated that there are
376 1000 mitotic cell divisions for every meiosis and that only about 1% of meioses are out-crosses
377 (Tsai, et al. 2008). Indeed, Figure 4 suggests that RGL may occur more frequently in yeasts (and
378 potentially in some plants, which may also reproduce asexually) than in the teleost fishes and
379 particularly the salmonids.

380 Even if RGL does not drive speciation, it still represents a barrier to diploid hybrids: most
381 of the taxa pairs for which essentiality data are available are separated from each other by at least
382 one RGL for an essential gene, the exceptions being some of the closest relatives of *A. thaliana*,
383 zebrafish and bakers' yeast studied (Supplemental Table 3). This observation is consistent with
384 studies of the relative frequency of diploid and polyploid hybridizations in flowering plants. In
385 these lineages, it is rare to find successful diploid hybrids involving distantly related parental
386 species (where RGLs could be common). However, allopolyploid hybrids appear to form at
387 similar rates across a much larger range of divergence times (Buggs, et al. 2009). A potential
388 explanation for the frequency of recurrent polyploidy is therefore simply that a new
389 allopolyploidy can allow paleopolyploids to again enjoy the benefits of hybridization (such as
390 hybrid vigor and heterosis; Birchler, et al. 2006; Chen 2010) in the face of their isolation due to
391 RGL.

392

393 **Discussion**

394 There are a surprising number of similarities seen in the manner of polyploidy resolution
395 across these independent polyploidies. Biased fractionation and other patterns in the homoeolog
396 losses are similar across many events: reciprocal gene losses are also present for most pairs of
397 polyploid taxa. The rate of homoeolog loss immediately after polyploidy is very high for many,
398 but not all, events (Figure 3).

399 Moreover, the differences in evolutionary patterns we do see are often in keeping with
400 what we know about the history of the events themselves. For instance, the salmonid WGD is
401 marked by continuing pairing of homoeologous chromosomes in meiosis (Allendorf, et al. 2015).

402 These pairings appear to limit the number of homoeolog losses and, for this event, loss rates at
403 the phylogeny tips and root are similar (per unit K_s). The grass ρ and yeast events have loss rates
404 that are roughly similar (again per unit K_s) across time, a fact for which we currently do not have
405 an operating hypothesis.

406 For the events that do show rapid losses along the root branch, which of the two
407 hypotheses mentioned, drift or selected losses, seems to best explain our data? The homoeologs
408 lost along the root are not more selectively constrained than other purely single-copy genes
409 known to have been lost later (Supplemental Figure 1). This fact probably speaks against any
410 very large number of selected losses. The single-copy genes as a whole are also generally
411 somewhat less selectively constrained than are genes with surviving homoeologs (Supplemental
412 Figure 1). Moreover, there is a clear pattern in most events whereby most of the fully single-copy
413 genes that exist are predicted to have been lost on the root branch (Supplemental Figure 3). The
414 yeast, nematode, and Paramecium events may violate this pattern because the nematode event is
415 an asexual triploidy while the other two involve lineages that have significant rates of asexual
416 reproduction. In such cases, restoring proper meiotic pairing is less necessary than in taxa with
417 primarily sexual reproduction. As a result, we expect that asexually reproducing lineages could
418 more easily form viable new species immediately after polyploidy, meaning that the post-
419 polyploid “lag” in speciation might be less evident (Schranz, et al. 2012). As a preliminary
420 hypothesis, we therefore propose that, for most polyploidies in animals and plants, the majority
421 of the purely neutral homoeolog losses occur prior to extensive species divergence in the
422 polyploid clade. A natural extension to this proposal would be that the post-polyploidy lag
423 represents this earlier period of neutral homoeolog loss, though the question of why speciation
424 events might be rare during such a period is still to be answered. A further implication would be
425 that later losses (including RGLs) would have occurred in homoeologous pairs that were initially
426 preserved to maintain dosage balance. They are then only lost when later mutations, such as
427 expression changes, release this dosage constraint and allow the loss of one of the copies
428 (Birchler, et al. 2005; Conant, et al. 2014). The higher selective constraint of genes with
429 surviving homoeologs is arguably also consistent with this hypothesis.

430 While the best-studied ancient polyploidy is in bakers’ yeast, it is hence atypical in a
431 number of respects. Biased fractionation is much less evident here (Emery, et al. 2018), losses
432 are not heavily biased toward the earliest phases of the polyploidy (Figure 3) and RGL is much

433 more prevalent. As mentioned above, one major source of these differences is likely the relative
434 timing of the post-polyploidy speciations: the yeasts had almost no lag between their polyploidy
435 and the first speciation (Supplemental Figure 4; Schranz, et al. 2012).

436 Other questions remain unanswered. The relative formation rates of allo- and
437 autopolyploids are uncertain. While recent polyploids appear to be approximately equally
438 divided between the two (Barker, et al. 2016), the potential selective advantages of being an
439 allopolyploid, and hence a hybrid (Alix, et al. 2017; Blanc-Mathieu, et al. 2017), could result in a
440 strong skew towards allopolyploids among the rare polyploidies that survive to become the
441 ancient events of the kind studied here (Barker, et al. 2016). The results here are consistent with
442 this hypothesis, but our sample of events is potentially biased by the available genome
443 sequences. Across all of the events, we find that the ubiquity of homoeolog fixation and (except
444 in paramecia) convergent homoeolog losses both speak to a common selective environment
445 acting to maintain certain homoeologs after all of these events. The most obvious candidate for
446 such a selective force is again the dosage balance hypothesis: it argues that highly interacting
447 genes tend to remain in multiple copies post-polyploidy to preserve the stoichiometry of those
448 interactions (Birchler, et al. 2005; Birchler and Veitia 2012; Tasdighian, et al. 2017). Whatever
449 the role of RGL in speciation, it is clear that all of these polyploid organisms possess a degree of
450 isolation due to it. The role of RGL in recurrent polyploidy is hence an important topic for future
451 research. Biology has a history of viewing “rules” as being more honored in the breach, but the
452 commonalities in post-polyploidy genome evolution across wide taxonomic distances are both
453 interesting in their own right and for the insight they give on other aspects of biology (Pires and
454 Conant 2016).

455

456 **Methods**

457 *Synteny block inference.*

458 Our three-step pipeline for inferring blocks of n -fold conserved synteny (NCS) produced
459 by polyploidy (Conant 2020) first uses GenomeHistory (Conant and Wagner 2002) to find all
460 pairs of homologous genes between each polyploid genome and a nonpolyploid outgroup (see
461 Supplemental Table 5 for genome details and Supplemental Table 6 for parameters). The second
462 step seeks to place these homologous genes into $N:1$ relationships between the polyploid genome
463 and the outgroup ($N=2$ for a WGD, $N=3$ for a hexaploidy and $N=4$ for an octoploidy). Using

464 simulated annealing (Kirkpatrick, et al. 1983), this step proposes sets of ordered pillars, each of
465 which contains a single gene from the nonpolyploid outgroup (G) and no more than N of the
466 homologs of that gene from the polyploid genome. The annealing algorithm then seeks a
467 combination of these assignments and a relative ordering of the m outgroup genes $G_1..G_m$ that
468 maximizes the number of synteny relations. We define two genes to be in synteny if they are
469 neighbors in the genome, ignoring any genes without homologs to the compared genome. In the
470 third step, these NCS blocks for each polyploid genome are merged across all of the polyploid
471 genomes. In this merging, only pillars where we have at least one homologous and syntenic gene
472 from each polyploid genome are included. With the set of merged pillars, a further simulated
473 annealing search is undertaken to infer a global pillar order that minimizes the number of
474 synteny breaks. While not strictly an ancestral genome inference (Sankoff and Blanchette 1998),
475 it is helpful to think of this optimal ordering as approximating the order of the genes just prior to
476 the polyploidy. Our previous work has shown that this inference approach is highly specific, with
477 no apparent cases of paralogous genes not created by the polyploidies in question being included
478 in the pillars (Emery, et al. 2018; Conant 2020).

479

480 *Modeling polyploidies with POInT*

481 At each pillar, POInT calculates the probability of the observed gene presence-absence
482 data conditional upon all possible orthology relationships and a tree. It carries this uncertainty in
483 orthology through its likelihood computations using a hidden Markov model that resembles the
484 Lander-Green approach for constructing linkage maps on a pedigree (Lander and Green 1987).
485 The parameter θ_i corresponds to the probability that the inferred orthology relationships change
486 between syntenic neighbors at pillars $i-1$ and i . When a pair of pillars are separated by a synteny
487 break, their orthology relationships are independent (i.e., $\theta_i=1/2$). Otherwise, $\theta_i= \theta$, a global
488 parameter estimated from the data by maximum likelihood.

489 This modeling framework allows for testing hypotheses about post-polyploidy gene
490 losses. For tetraploidies, we analyzed three phenomena: duplicate fixation, biased fractionation
491 and overly frequent parallel losses of the same homoeolog on independent branches of the
492 phylogeny (Supplemental Table 1). For the triplication events, we focused on differences in
493 homoeolog loss rates between the three subgenomes (Supplemental Table 2). We further allowed
494 the root branch to have separate values of the model parameters to account for the two-step

495 nature of hexaploidy formation (Figure 2; Tang, et al. 2012). For the Paramecium and vertebrate
496 2R octoploidies, we used a null model (WGQ_n; Figure 2) where losses occur equally from all
497 four subgenomes, but where the loss rate from triplicated and duplicated loci can differ from that
498 seen in quadruplicated loci. We also added an octoploid formation step to this model, with all
499 pillars starting in state D_{1,3} and then either experiencing a loss followed by the second tetraploidy
500 (transitioning to D_{1,2} or D_{3,4}) or becoming quadruplicated (Q).

501

502 *Analyzing nested genome duplications with POInT.*

503 The vertebrates and ciliates experienced two sequential genome duplications relative to
504 the outgroup genome to which they were compared. They hence present a challenge because the
505 POInT computation for such an octoploidy with n genomes scales as $O(24^{2n})$. As a result, it is
506 only computationally feasible to analyze two such octoploid genomes. However, if the
507 consecutive whole-genome doublings were sufficiently separated in time, POInT can separate
508 them using the two-step model just described. This model assumes each locus starts as a
509 *duplicated* one and then may either remain duplicated until the second polyploidy (and hence
510 become quadruplicated) or experience a gene loss prior to that event, meaning that the second
511 event only produces a *duplicate* gene pair (Figure 2). We thus sought to phase regions from both
512 octoploidies into pairs of regions created by the most recent genome doubling. For the ciliate
513 genomes, we were able to phase the quadruplicated loci into 11,683 pairs of duplicated loci with
514 at least one gene from each genome and where our orthology assignment confidence for
515 assigning extant genes to one of the two subgenomes from the *first* polyploidy event was $\geq 99\%$.
516 For the vertebrate 2R events, a model that attempts to phase the 2R duplicates fit the data no
517 better than did the null model ($P=0.1$, likelihood ratio test with 1 *d.f.*) and so no further phasing
518 was attempted.

519

520 *POInT and topological inference.*

521 For the legume WGD, the grass ρ event, the Paramecium tetraploidy, the nematode
522 triploidy and the salmonid WGD, we used POInT to infer the maximum likelihood phylogeny
523 under the WGD_{bfc-nb} or WGT_{G3} models and an exhaustive tree search (Supplemental Figure 4).
524 For the Brassica WGT, we assumed that *B. rapa* and *B. oleracea* were sister taxa and tested all

525 three rooted topologies consistent with this constraint. The topology for the yeast WGD was
526 taken from Kurtzman and Robnett (2003), for the TGD from Near et al., (2012) and for At- α
527 from Huang et al., (2016). The vertebrate 2R topology is trivial.

528 For the salmonid WGD, the inferred topology differs significantly from others that have
529 been published. We therefore fit the full POInT model under the topology published by Crespi
530 and Fulton (2004). The orthology estimates and model parameters are largely unaffected by this
531 topology change: the orthology relationships of only 106 (0.7%) pillars with posterior probability
532 >80% differ when the topology is changed, and 91 of these changes simply swap the identities of
533 the more and less fractionated genomes. The corresponding figures for 95% confidence are 9 and
534 7 pillars.

535

536 *Orthology inferences and inference of synonymous distances.*

537 Using high confidence orthologs estimated with POInT, we computed the mean
538 synonymous divergence for every branch for each polyploidy. The nematode triploidy and
539 vertebrate 2R events were omitted from this analysis due to their fragmented synteny blocks. For
540 the tetraploidies, we considered “nearly fully duplicated” pillars: i.e., pillars with at most one
541 missing gene copy from each of the two gene trees produced by the genome duplication (two
542 total losses) for all events except the TGD and yeast WGDs, where we allowed two losses from
543 each subtree (four total losses). For the Brassica hexaploidy, we analyzed only fully triplicated
544 pillars. At each such pillar, we aligned amino acid sequences for the genes in question with T-
545 coffee (Notredame, et al. 2000). We fit the Goldman and Yang codon model of evolution
546 (Goldman and Yang 1994) to the corresponding codon-preserving alignments and mirrored gene
547 trees and extracted the estimated synonymous divergence (K_s) for each branch from this codon
548 model as described by these authors.

549 With the possible exception of the salmonids (Allendorf and Thorgaard 1984; Braasch
550 and Postlethwait 2012), all of the events studied are believed to be allopolyploids. For a given
551 pillar in set of allopolyploid taxa, the mean synonymous divergence observed along this root
552 branch ($\overline{K_s^R}$; Figure 2D) should represent the sum of the pre-polyploidy divergence of the diploid
553 progenitors as well as the divergence that occurred after the polyploidy but before the first
554 speciation event among the polyploid taxa. However, recombination events could, through
555 genetic drift, result in the replacement of alleles from one of the progenitors with those from the

556 other (Wolfe 2001). These recombinations, or homoeologous exchanges (HE; Gaeta and Chris
557 Pires 2010) are reasonably common in neopolyploid plants (Doyle, et al. 2008; Chalhoub, et al.
558 2014; Zhang, et al. 2020), but it is not clear whether they are frequent enough to effect the
559 divergence seen along these root branches. We extracted the coding sequences for each pillar
560 that had every homoeologous gene preserved. Post-polyploidy homoeolog displacement (Gaut
561 and Doebley 1997; Wolfe 2001) will erase the divergence between the progenitor genomes,
562 leaving only the post-displacement divergence to be observed. In such a case, we might expect to
563 observe two modes in synonymous divergence, a larger value for homoeologs that did not
564 experience displacement and a smaller one (lacking the progenitor divergence) for homoeologs
565 that did. To test this hypothesis, we fit the set of estimated synonymous divergences (K_s) along
566 the root branches to either one or two log-normal distributions using the R package *mclust*
567 (Scrucca, et al. 2016) with the best-fit model (i.e., one or two distributions) chosen with the
568 Bayesian information criterion (BIC; Schwarz 1978). Values of K_s less than 5×10^{-3} or greater
569 than 2.0 were omitted from these analyses as representing either no synonymous divergence or
570 saturated synonymous divergence, respectively. When two distributions were fit, a “weighting” p
571 reflecting the mixing proportion of each component was also estimated. For a few root branches,
572 a bimodal distribution is preferred. However, in most cases this bimodality is not consistent
573 across different collections of pillars and, even when it is, the proportion of pillars belonging to
574 one of the “modes” is generally very small (Supplemental Table 7). We hence see little
575 suggestion of HE in these data.

576
577 *Filtering for extreme instances of gene conversion.*

578 Because gene conversion among homoeologs (as seen in yeasts; Evangelisti and Conant
579 2010; Scienski, et al. 2015) could confound our K_s estimates, we sought to filter out pillars that
580 showed strong evidence of having experienced it. We created “gene conversion gene trees” for
581 each pillar where each homoeologous gene was forced to be sister to its paralog(s). Any pillars
582 where the likelihood of the sequence alignment under these gene conversion trees was higher
583 than that seen in the mirrored species trees was omitted from our estimates of synonymous
584 divergence (Supplemental Figure 5).

585

586 *Comparing duplicate loss rates to estimated synonymous divergence.*

587 Using the K_s inferences made above for each branch, we compared POInT's maximum
588 likelihood estimate (MLE) of the rate of homoeolog loss (e.g., its estimated branch length, αt in
589 Supplemental Figure 4) to each branch's mean synonymous divergence, $\overline{K_s}$, to see if the number
590 of losses on any particular branch was unusually large or small. Estimating confidence intervals
591 for these ratios of $\alpha \cdot t / \overline{K_s}$ is challenging. We treated the numerators and denominators of these
592 ratios as being normally distributed and independent random variables. The maximum likelihood
593 estimates (MLEs) of αt in the numerators should have asymptotically normal distributions with
594 means that are equal to the true parameter values. The variances of these normal distributions
595 were approximated by evaluating the inverse of the observed Fisher information (i.e., the
596 Hessian of the negative log-likelihood; see Kendall and Stuart 1973). We estimated the observed
597 Fisher information values via a single-dimension finite difference approximation that ignored
598 covariances between the αt parameter and other parameters.

599 For each branch of the phylogeny, the K_s estimates that are in the denominator of the
600 ratio $\alpha \cdot t / \overline{K_s}$ are obtained via a sample mean of the K_s estimates from the sequences of
601 individual pillars (i.e., $\overline{K_s}$). Due to the Central Limit Theorem, this sample mean should be
602 approximately normally distributed with mean equal to the true parameter value and with
603 variance being approximately the sample variance among individual K_s estimates divided by the
604 number of individual K_s estimates.

605 To infer confidence intervals for the ratio of $\alpha \cdot t / \overline{K_s}$ on each branch, we independently
606 sampled from the aforementioned normal distributions that are used to approximate the
607 uncertainty of αt and $\overline{K_s}$ estimates in the ratio. For each branch, we calculated the ratio of these
608 sampled values for 1000 pairs of randomly sampled values. We then sorted the resulting ratios
609 and set 95% confidence intervals by finding the ratio value that defined the lower and upper
610 2.5% of the sorted values.

611 Because the inclusion of fixation in our loss models can give rise to long tip branches
612 (effectively the model suggests that all surviving duplicates in some genomes are now fixed), we
613 present data using a model with convergent losses and biased fractionation but no fixation
614 (WGD_{bc-nb}). However, our results are very similar with using the full WGD_{bfc-nb} model
615 (Supplemental Figure 6).

616
617 *Comparisons of selective constraint for different classes of polyploid loci*

618 We examined the inferred average selective constraint (K_a/K_s , estimated as described
619 above) for five classes of polyploid loci (e.g., pillars) across the seven WGD events: 1) Pillars
620 that are single copy in all taxa and have a high probability of having returned to single-copy
621 along the root branch, 2) Pillars that are completely single copy but where the genes did not
622 return to single-copy on the root branch (e.g., where alternative copies of the duplicated genes
623 are preserved in different genomes), 3) pillars with duplicates surviving in only a single species,
624 4) pillars where all but one species maintains the duplication and 5) pillars where all species
625 maintain duplicate copies. Confidence intervals for these mean K_a/K_s estimates were estimated
626 with the approach of described above.

627

628 *Identifying reciprocal gene losses (RGLs) between polyploid taxa.*

629 For a pair of single-copy genes from distinct genomes, the probability that these genes
630 represent RGLs is simply the sum of the probabilities of the orthology relationships, estimated
631 with POInT, that place them as paralogs rather than orthologs. We computed, for each pair of
632 extant taxa in each polyploidy, the set of RGLs that we could identify with a confidence of $\geq 95\%$
633 (Figure 4A). To avoid spurious inferences, we restricted our identification of RGL pairs to
634 single-copy genes in each genome where either: a) both the gene and the “hole” corresponding to
635 its lost homoeolog were in synteny with genes on either side or b) the single-copy gene in
636 question was the only homolog of the outgroup gene used for the inference of the NCS blocks. In
637 the first case, this filter corresponds to a clear absence of a corresponding homoeolog in the
638 paralogous synteny block, in the second to the absence of a gene that could be the “missing”
639 homoeolog. We then used TBLASTX (Altschul, et al. 1997) to search the non-coding regions of
640 each genome for putative homoeologous copies of the inferred RGL gene that were missed in the
641 genome annotations (e.g., the inference of RGL was spurious due to an annotation artifact). In
642 Case “a” above, this search was restricted to the non-coding regions in the “hole” between the
643 neighboring syntenic genes; in Case “b,” we searched the entire genome for the potentially
644 unannotated homoeolog. Only RGL genes with no such matching noncoding regions at an E-
645 value cutoff of $\leq 10^{-10}$ were considered “true” RGLs. These secondary filters were not applied for
646 the yeast WGD because those data were taken from the manually curated Yeast Genome Order
647 Browser (YGOB, Byrne and Wolfe 2005).

648 Data on gene knockouts producing lethal phenotypes from zebrafish, *A. thaliana* and
649 bakers' yeast were taken from ZFIN (Howe, et al. 2013; Conant 2020); a set of 510 “embryo-
650 defective” genes identified by Meinke (2020); and Steinmetz et al., (2002), respectively. The
651 proportion of RGLs in these “essential gene” lists was compared to the proportion of all other
652 single-copy genes from the same organism in the list using Fisher's exact test (Sokal and Rohlf
653 1995). For these same three species, we used GeneOntology data (Gene Ontology Consortium
654 2015) and Panther Overrepresentation Tests (Release 20200728; Mi, et al. 2019) to ask if there
655 were terms from the GO-Slim Biological Process, Cellular Compartment or Molecular Function
656 ontologies that differed in their frequency between the RGL genes and other single-copy genes.
657 After FDR correction (Benjamini and Hochberg 1995), no such terms were found for any of the
658 three ontologies across any of the three genomes (FDR-corrected P -value > 0.05).

659
660
661
662

663 **Data availability:**

664 All underlying data are available from the POInT browser (wgd.statgen.ncsu.edu) and from
665 figshare (DOI: <https://doi.org/10.6084/m9.figshare.12750992.v4>); the POInT package is
666 available from GitHub (<https://github.com/gconant0/POInT>)

667

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674

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