

1 Neutral competition within a long- 2 lived population of symmetrically 3 dividing cells shapes the clonal 4 composition of cerebral organoids

5 Florian G. Pflug^{1*}, Simon Haendeler¹, Christopher Esk², Dominik Lindenhofer², Jürgen A.
6 Knoblich², Arndt von Haeseler^{1,3}

7 ¹ Center for Integrative Bioinformatics Vienna (CIBIV), Max Perutz Laboratories, University of
8 Vienna and Medical University of Vienna, Vienna Bio Center (VBC), Vienna, Austria

9 ² Institute of Molecular Biotechnology of the Austrian Academy of Science (IMBA), Vienna Bio
10 Center (VBC), Vienna, Austria

11 ³ Bioinformatics and Computational Biology, Faculty of Computer Science, University of
12 Vienna, Vienna, Austria

13 *Correspondence: florian.pflug@univie.ac.at

14 Summary

15 Cerebral organoids model the development of the human brain and have become an
16 indispensable tool for studying neural development and neuro-developmental
17 diseases. Comprehensive whole-organoid lineage tracing has revealed the number
18 of progeny arising from each initial stem cell to be highly diverse, with lineage sizes
19 ranging from one to more than 20,000 cells. This variability exceeds what can be
20 explained by existing stochastic models of corticogenesis, which indicates that an
21 additional source of stochasticity must exist. We propose the quantitative SAN model
22 in which this additional source of stochasticity is neutral competition within a long-
23 lived population of symmetrically dividing cells. In this model, the eventual size of a
24 lineage is determined by its survival time within this population of symmetrically
25 dividing cells, which due to neutral competition varies widely between individual
26 lineages. We demonstrate the SAN model to explain the experimentally observed
27 variability of lineage sizes and use it to derive a formula that captures the
28 quantitative relationship between survival time and lineage size. Finally, we show
29 that our model implies the existence of a mechanism which keeps the size of the
30 population of symmetrically dividing cells approximately constant, and that it enables
31 this mechanism to be probed experimentally.

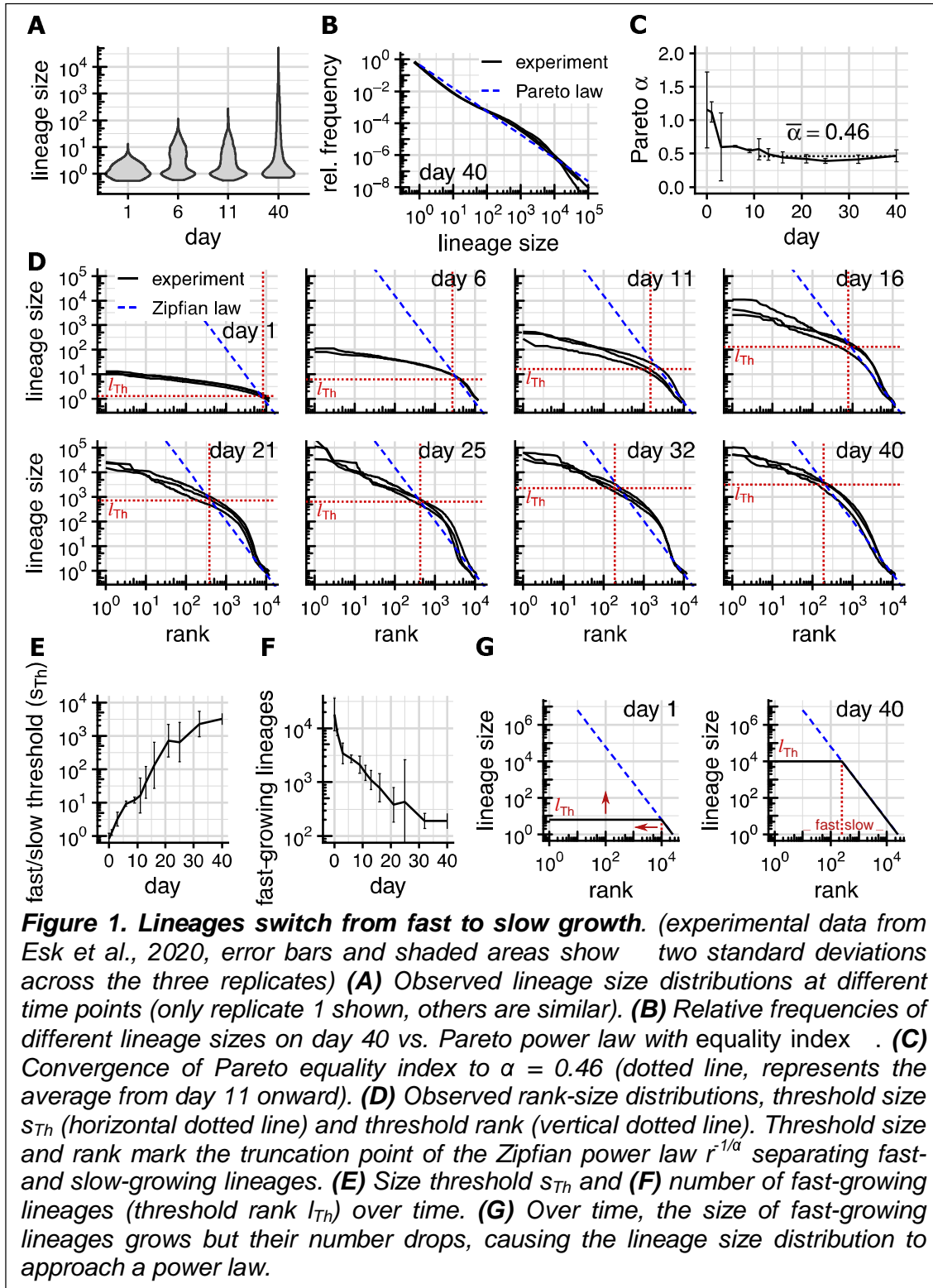
32 Introduction

33 The development and maintenance of the tissues and organs comprising complex
34 organisms rely on sophisticated genetic programs to coordinate the differentiation of
35 cells in both space and time. In many cases, this “program” does not consist of fully
36 deterministic decision chains but instead contains stochastic components; examples
37 include the development of the cortex (Klingler and Jabaudon, 2020; Llorca et al.,
38 2019) and stem cell homeostasis in intestinal crypts (Snippert et al., 2010).

39 During cortical development, neurons are produced (directly or indirectly) by
40 progenitor cells in the ventricular zone called radial glial cells (RGCs). In mice, the
41 neuronal output of individual RGCs was observed to vary by about one to two orders
42 of magnitude between seemingly identical progenitors, which suggests a stochastic
43 model of cortical neurogenesis (Llorca et al., 2019). Cerebral organoids grown from
44 human stem cells (Lancaster et al., 2017) show even stronger variability of offspring
45 numbers; comprehensive whole-organoid lineage tracing data shows the sizes of
46 individual lineages arising from each ancestral stem cells to vary over up to four to
47 five orders of magnitude (Esk et al., 2020). For RGCs, an alternative explanation of
48 the apparent stochasticity of their number of offspring are hidden variables (Zechner
49 et al., 2020) like transcriptional state within the seemingly homogenous population of
50 progenitors. But in organoids, we expect the pool of ancestral stem cells to be
51 homogenous, and thus conclude that lineage sizes vary predominantly due to
52 stochastic effects. While RGCs output may vary more widely in humans than in mice,
53 varying RGC output alone still cannot account for lineage sizes varying over 4 to 5
54 orders of magnitude in human cerebral organoids. There is thus likely an additional
55 source of stochasticity in organoid development beyond the stochastic model of
56 neurogenesis proposed by Llorca et al (2019).

57 In this study, we propose the source of this additional stochasticity to be neutral
58 competition within a long-lived population of roughly 10,000 symmetrically dividing
59 stem cells (S-cells). Neutral competition between stem cells has previously been
60 shown to shape the clonal composition of tissues in homeostasis (Snippert et al.,
61 2010; Corominas-Murtra et al., 2020), and to accurately predict the time until
62 monoclonality (the time until all but a single lineage has died out). We show that in
63 growing tissue like cerebral organoids, neutral competition does not lead to eventual
64 monoclonality. Instead, the tissue records the changing clonal composition of its

65 stem cell population, which causes the sizes of individual lineages to grow
 66 increasingly diverse over time. To quantify this effect and its dependence on the size
 67 of the S-cell population, we introduce the stochastic SAN model and show that



68 neutral competition within its S-cell population suffices to explain the observed
69 variation of lineage sizes over four to five orders of magnitude.

70 Results

71 Empirical lineage size distribution

72 In the experiment conducted by Esk *et al.* (Esk et al., 2020), cerebral organoids were
73 grown from roughly 24,000 stem cells, genetically identical except for a distinct
74 genetic barcode in each cell serving as a *lineage identifier* (LID). To determine the
75 contribution of each initial stem cell to organoids of different ages, organoids were
76 subjected to amplicon high-throughput sequencing. The sequencing reads (after
77 filtering and error-correction) corresponding to each LID were counted, and the per-
78 LID read counts normalized to an approximate number of cells comprising each
79 lineage (see *experimental procedures* for details).

80 The resulting *lineage size distribution* (figure 1A) shows, as expected, small
81 and equally sized lineages for organoids harvested at day 1 (lineages sizes around 1
82 cell). The distribution grows more uneven until day 11 (up to 30 cells/lineage) and
83 extends over 4 to 5 orders of magnitude (up to 100,000 cells/lineage) after 40 days.

84 A common mathematical model for distributions extending over multiple orders
85 of magnitude are so-called (Pareto) power laws where the frequency of objects of
86 size l or larger is proportional to $l^{-\alpha}$. Parameter α is called the (Pareto) equality
87 index because it determines how even (large α) or uneven (small α) object sizes are
88 distributed. In double-logarithmic frequency vs. size plots, power laws appear as
89 straight lines with slope α , which we find matches the lineage size distribution on day
90 40 well for $\alpha \approx 0.46$ (figure 1B). We remark that $\alpha \approx 0.46$ represents a small equality
91 index (i.e. diverse lineage sizes); in applications of Pareto distributions values of α
92 often lie between 1 and 2.

93 While the unevenness of the lineage size distribution grows considerably
94 between days 11 and 40 (figure 1A), the equality index stays close to $\alpha \approx 0.46$ from
95 day 11 onwards (figure 1C). The equality index thus fails to capture the large
96 increase in non-uniformity of lineage sizes between days 11 and 40.

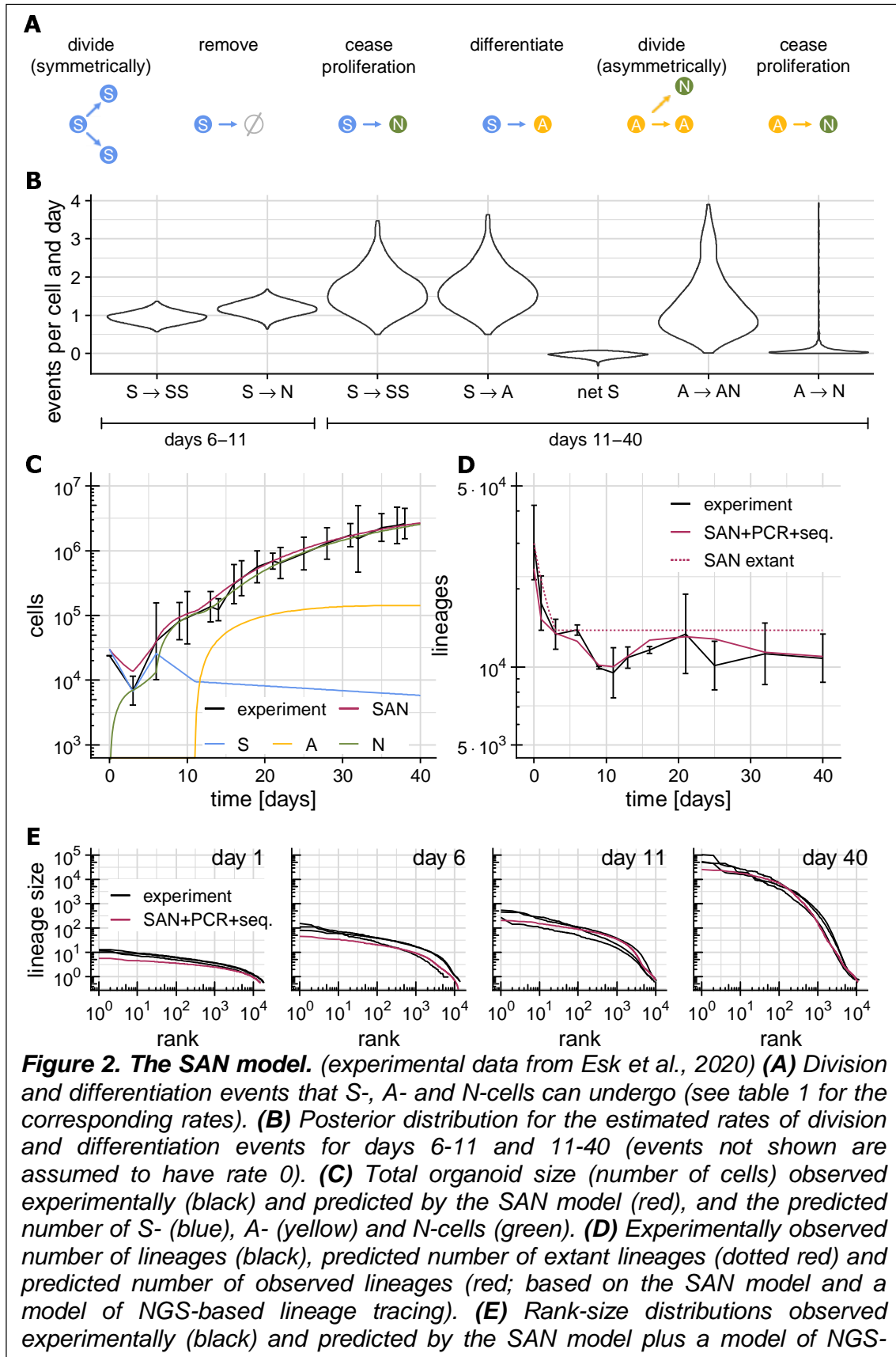
97 **Truncated Zipfian rank-size distribution**

98 To describe the evolution of lineage sizes over time, we thus instead rank lineages
99 by size (largest lineage first) and plot the resulting *rank-size distributions* (figure 1D).
100 For lineage sizes governed by a Pareto law with index α , the rank-size distribution
101 would be expected to be governed by a Zipfian power law, meaning lineage sizes
102 should decrease proportional to $r^{-1/\alpha}$ with increasing rank r (Adamic and
103 Huberman, 2002).

104 Instead, we observe *truncated* Zipfian laws in which lineages sizes obey a
105 Zipfian law ($\alpha \approx 0.46$) only up to a certain threshold size l_{Th} above which lineages
106 are multiple orders of magnitudes smaller and more uniform than the Zipfian law
107 would predict (figure 1D). The threshold size l_{Th} grows more than 1,000-fold (figure
108 1E) over 40 days, while the ratio between threshold size and largest lineage size
109 grows only by a factor of 2 (from 8.5 to 21; figure 1D); lineages above the threshold
110 therefore grow roughly uniformly. Lineages below the threshold, in contrast, show no
111 overall shift towards larger lineages sizes over time, indicating that growth has
112 mostly ceased for these lineages.

113 **Lineages switch from fast to slow growth**

114 The size threshold l_{Th} thus partitions lineages according to their growth regime into
115 *fast-growing* and *slow/non-growing*. Of the (on average) 10,851 lineages that
116 contribute to the final organoid 8,389 lineages fall into the fast-growing category on
117 day 1; but on day 11 their number has dropped to 1,496, and on day 40 only 191
118 (about 2%) fast-growing lineages remain (figure 1F). Lineages thus start out fast-
119 growing, and one by one switch to a regime of slow/no growth as time progresses.
120 The later that switch occurs for a particular lineage, the bigger it has become before
121 its growth ceases, leading to larger and larger lineages in the slow-growing regime
122 and consequently to l_{Th} increasing as time progresses. In this coarse approximation,
123 lineages are assumed to have the same size as long as they are fast-growing (figure
124 1G); experimentally we observe a spread of 1.5 orders of magnitude within the sizes
125 of fast-growing lineages versus a spread of 3.5 orders of magnitude within the slow-
126 growing regime (figure 1D).



128 This proposed lineage-specific switching from fast to slow growth can also
129 quantitatively reproduce the observed truncated Zipfian with $\alpha = 0.46$, with one
130 mathematically simple example being fast-growing lineages growing exponentially
131 with rate γ and the number of fast-growing lineages declining exponentially with rate
132 $\sigma = \alpha\gamma$. But while this simple example assumes an unspecified biological mechanism
133 behind the lineage-specific growth regime switches, we show in the following that no
134 such mechanism is in fact necessary. Instead, we show that such growth regime
135 switches emerge naturally from a cellular model of organoid growth.

136 **SAN model**

137 In the SAN model of organoid growth we distinguish between three types of cells
138 based on the proliferation behavior they exhibit (figure 2A). Cells are either
139 *symmetrically* dividing (S-cells), *asymmetrically* dividing (A-cells) or *non-dividing* (N-
140 cells). In this model, S-cells have the ability to self-renew indefinitely through
141 symmetric division and can thus be considered stem cells. They form the initial cell
142 population of an organoid, and apart from dividing symmetrically they differentiate
143 into either A- or N-cells or are removed permanently. A-cells are cells that have
144 committed to a differentiation trajectory and produce N-cells through asymmetric
145 division, while N-cells do not further divide. We emphasize that S, A and N refer
146 solely to a cell's proliferation behavior, not its functional cell type.

147 All these division and differentiation events occur randomly and independently
148 for each cell with specific time-dependent rates (table 1); from a single-lineage
149 perspective, the SAN model is thus stochastic in nature. Any difference between the
150 trajectories of the lineages arising from different ancestral cells is thus assumed to
151 be purely the result of random chance, not of cell fate decisions or spatial
152 configuration. From a whole-organoid perspective, on the other hand, the SAN
153 model is deterministic, because random effects average out over the roughly 10,000
154 lineages comprising an organoid.

155 **Division and differentiation rates**

156 To find the rates of cell division and differentiation, we split the organoid
157 development into four time intervals (days 0-3, 3-6, 6-11 and 11-40) according to the
158 main phases of the protocol of Lancaster *et al.* (Lancaster *et al.*, 2017). Until day 6,
159 formation of embryoid bodies (EBs) is still ongoing, and organoid development thus

160 does not reflect development *in vivo*. For these time intervals we manually chose
161 rates of S-cell division ($S \rightarrow S S$), removal ($S \rightarrow \emptyset$) and death ($S \rightarrow N$; dead cells
162 present in the EB are still counted by NGS-based lineage tracing) for which predicted
163 and observed numbers of cells, lineages, and lineage sizes match (table 1). After
164 day 6, EB formation is complete, and no further cells are removed from the organoid.
165 Until day 11 S-cells are then assumed to either divide symmetrically ($S \rightarrow S S$) or
166 cease proliferation ($S \rightarrow N$), but to not produce A-cells yet. After embedding the
167 organoids into Matrigel droplets on day 11, organoid growth enters the asymmetric
168 division phase where S-cells are assumed to multiply ($S \rightarrow S S$) and to differentiate
169 into A-cells ($S \rightarrow A$), which then produce N-cells through asymmetric division ($A \rightarrow A$
170 N) before they eventually cease to proliferate ($A \rightarrow N$).

171 From day 6 onwards organoid development reflects development *in vivo* and
172 we hence desired to identify the range of likely rates for each event in addition to a
173 single most-likely value. We thus adopted a Bayesian model comprising log-normally
174 distributed measurement inaccuracies on top of the SAN model, and used Markov
175 chain Monte Carlo (MCMC) sampling to find 1,000 likely rate combinations and their
176 (posterior) probabilities (figure 2B). To arrive at a single set of most-likely values for
177 the rates to be estimated, we then computed MAP (maximum a-posteriori) estimates
178 (table 1) from this posterior distribution.

179 Both the posterior distribution (figure 2B) and the MAP estimates (table 1) show
180 the *net* rate of S-cell proliferation (the rate with which the S-cell population grows or
181 shrinks, i.e. the difference between the rates of $S \rightarrow S S$ and $S \rightarrow A$) to lie close to
182 zero. From this, we conclude that the size of the S-cell population changes only
183 slowly from day 11 onwards. The posterior distributions of the individual rates are, on
184 the other hand, much broader. While the MAP estimates are thus arguably the single
185 most likely set of rates, other combinations of rates are possible as well.

186 **Model validation**

187 For the MAP rate estimates (table 1), the predicted organoid sizes between day 0
 188 and 40 agree well with the experimentally determined number of cells (figure 2C).
 189 Similarly, the lineage size distribution predicted by the SAN model matches the
 190 observed lineage size distribution both in the original data of Esk et al. (figure 2E) as
 191 well as in independent replicate experiments (figure S1). In particular, the predictions
 192 show the same truncated Zipfian distributions as the experimental data recapitulate
 193 the spread over 4 – 5 orders of magnitude. The SAN model predicts the number of
 194 *extant* lineages (lineages containing at least one S-, A or N-cell) to drop to about
 195 $\approx 13,700$ on day 3 where it then remains. This drop in the number of extant lineages
 196 is caused by lineages that do not make it into the organoid during EB formation.
 197 While the predicted number of remaining lineages slightly exceeds the experimental
 198 observation ($\approx 10,900$ on day 40 on average), the numbers match closely once we
 199 account for non-observed lineages due to the stochastic nature of sequencing (figure
 200 2D).

201 **A-cell output**

202 According to the SAN model (table 1) a single A-cell has two options, either to divide
 203 asymmetrically (probability $r_{A \rightarrow AN} / (r_{A \rightarrow AN} + r_{A \rightarrow N}) = 91\%$) or to cease proliferation
 204 (probability $r_{A \rightarrow N} / (r_{A \rightarrow AN} + r_{A \rightarrow N}) = 9\%$). The likely range of additional N-cells
 205 produced over the lifetime of an A-cell is thus 0 to 30 (95% quantile), with an
 206 average of $r_{A \rightarrow AN} / r_{A \rightarrow N} = 10$. This slightly exceeds (about 3x) the observed
 207 stochasticity of the neuronal output of RGCs in mice (Llorca et al., 2019) and is
 208 consistent with a previously observed increase of intermediate progenitor divisions in
 209 humans.

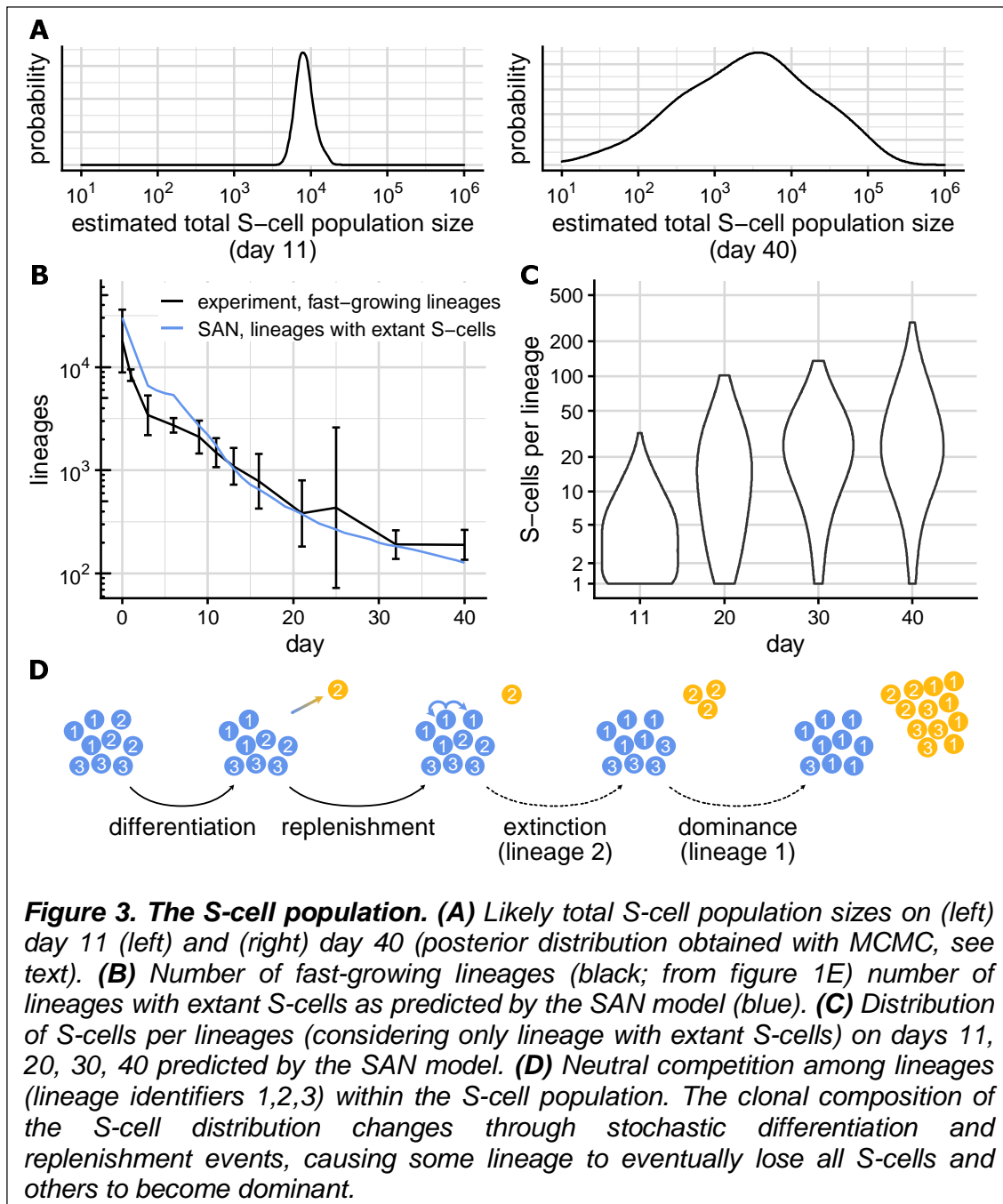
210 **Predicted S-cell population size**

211 The MAP rate estimates (table 1) predict that organoids contain $\approx 9,500$ S-cells on
 212 day 11 and still $\approx 5,800$ S-cells on day 40. To take the inherent ambiguity of the MAP
 213 estimate due to the broadness of the posterior distribution into account, we

day	S \rightarrow S S	S \rightarrow \emptyset	S \rightarrow N	S \rightarrow A	A \rightarrow A N	A \rightarrow N	phase
0-3	-	0.35	0.15	-	-	-	EB formation
3-6	0.6	-	0.15	-	-	-	EB formation
6-11	0.94	-	1.14	-	-	-	neural induction
11-40	1.68	-	-	1.69	0.71	0.07	asymmetric division

Table 1. Division and conversion rates in the SAN model. Rates specify the number of expected events per cell and day

214 computed the posterior distribution of these population sizes (figure 3A). We find that
 215 the total S-cell population size on day 11 is well-defined up to a factor of at most 2
 216 around 10,000 cells. On day 40, the estimates are more dispersed, owing to the
 217 large cumulative effect that rates have over 30 days; yet while the exact population
 218 size is difficult to estimate, the finding that organoids contain a significant number of
 219 S-cells on day 40 is robust.



220 **Fast-growing lineages contain S-cells**

221 While the total size of the S-cell population changes only slowly, its clonal
 222 composition changes rapidly. From the $\approx 13,700$ lineages comprising the organoid
 223 from day 3 forward, $\approx 1,700$ lineages still contain S-cells on day 11, and until day 40
 224 that number has dropped to ≈ 100 (figure 3B). This drop in the number of lineages
 225 with extant S-cells is offset by an increase in the number of S-cells each of these
 226 lineages contains (figure 3C; average grows from ≈ 5 cells/lineage on day 11 to ≈ 36
 227 cells/lineage on day 40).

228 The number of lineages with extant S-cells matches the number of lineages
 229 classified as fast-growing by our fast-slow model well (figure 3B). This highlights S-

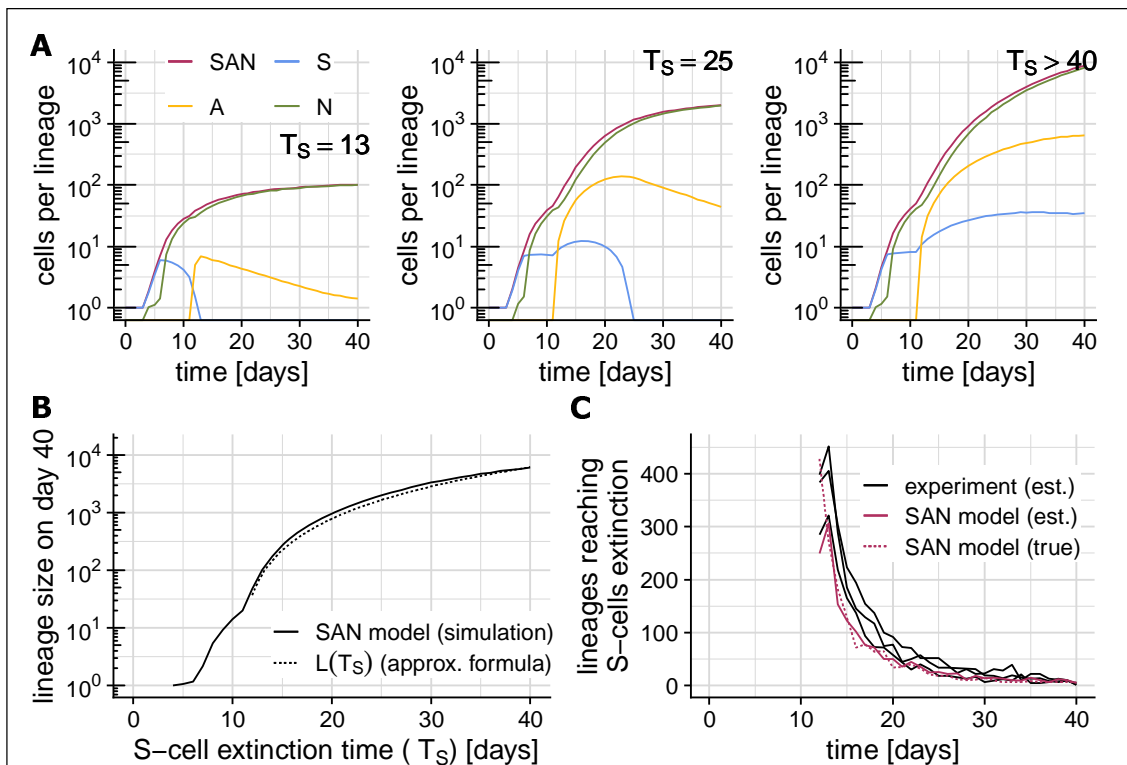


Figure 4. Lineage-specific S-cell extinction time determines lineage size. (Error bars show two standard deviations across the three replicates of Esk et al, shaded areas show the range between the 2.5% and 97.5% quantile across 5 billion simulations). (A) Lineage-specific growth trajectories under the SAN model stratified by the lineage's S-cell extinction time T_s . Plot show the most likely total lineage size (red), and number of S- (blue), A- (yellow) and N- (green) cells comprising the lineage. (B) S-cell extinction time T_s vs. final lineage size on day 40. Plot shows simulation results (solid) and the analytical approximation $L(T_s)$ (dotted). (C) Recovering S-cell extinction times from lineage sizes on day 40. Plot shows the number of lineages reaching S-cell extinction on each day estimated using $L(T_s)$ from experimental data (black) and simulated data (red), and the true number of such lineages according to the SAN model.

230 cells as being the main driver of lineage growth; as stated above a single
231 differentiating S-cells eventually on average produces 10 additional N-cells, and
232 once a lineage contains no more S-cells its growth will thus slow down and
233 eventually cease.

234 **Neutral competition shapes S-cell clonal composition**

235 Over time, not only does the average number of S-cells found within lineages with
236 extant S-cells grow, but so does the spread between the lineages' S-cells counts
237 (from about 1 - 30 S-cells per lineage on day 11 to about 1 – 300 S-cells per lineage
238 on day 40). The clonal composition of an organoid's S-cell population thus grows
239 more and more non-uniform over time. Under the SAN model, this change is the
240 result of *neutral competition* (figure 3D) amongst S-cells, a term introduced to
241 describe the population dynamics of stem cells within intestinal crypts (Snippert et
242 al., 2010).

243 Qualitatively, the dynamics of an organoid's S-cell population under neutral
244 competition mimic the population-genetic Moran model (Moran, 1958) in which
245 individuals (cells in our case) carrying different neutral alleles (lineage identifiers in
246 our case) are randomly removed (differentiate) and are replaced (through symmetric
247 division) by offspring of another randomly selected individual. Once the last S-cell of
248 a particular lineage has differentiated, the lineage cannot reappear within the
249 organoid's S-cell population. The observed disappearance (figure 3B) of lineages
250 from the organoids S-cell population is thus a result of more S-cells differentiating
251 than dividing due to random chance. Similarly, the observed growth of the remaining
252 lineages (figure 3C) results from more symmetric divisions than differentiations,
253 again due to random chance.

254 Using the SAN model, we now study the effects of neutral competition between
255 S-cells on the clonal composition quantitatively.

256 **Lineage-specific S-cell extinction times determine final lineage sizes**

257 Under the population-genetic Moran model, alleles eventually either disappear from
258 a population or become fixed. Tissue homeostasis driven by a stem cell population
259 under neutral competition likewise leads to eventual monoclonality, i.e. to all extant
260 cells being eventually derived from a single ancestral stem cell. In growing neural
261 tissue like cerebral organoids however, the lack of constant cell turn-over restricts

262 eventual monoclonality to S-cells. The clonal composition of the N-cell population
263 instead records the evolution of the S-cell's clonal composition over time; lineages
264 whose last S-cell was lost later and/or which contained more S-cells will contribute
265 more N-cells than lineages which die out quickly from the S-cell population.

266 To study the effects of S-cell extinction on lineage sizes quantitatively, we
267 stratified simulated lineage growth trajectories according to their *S-cell extinction*
268 *time* (T_S ; the time at which a particular lineage loses the last S-cell). Lineages whose
269 S-cell population goes extinct at day $T_S = 13$ (figure 4A left) respectively day $T_S = 25$
270 (figure 4A middle) show diminished growth and a declining number of A-cells after
271 losing their S-cells at time T_S . In contrast, lineages whose S-cell population survives
272 past day 40 (figure 4A right) grow considerably faster and reach a considerably
273 larger size. Comparing the variations in lineage sizes on day 40 between S-cell
274 extinction time strata shows the variation due to T_S to dominate the variations within
275 each stratum (figure 4B). Thus, while other random factors have some influence,
276 their influence on a lineage's sizes on day 40 is negligible compared to the time the
277 lineage loses its last S-cell.

278 Mathematical analysis of the SAN model yields the approximate expression

(Eq. 1)
$$L(\Delta T_S) = \left(\frac{1}{3} s_0 \Delta T_S + \frac{r_{S \rightarrow SS}}{2\sqrt{6}} \Delta T_S^2 \right) r_{S \rightarrow A} \left(1 + \frac{r_{A \rightarrow AN}}{r_{A \rightarrow N}} \right)$$

279 for the final lineage size of a lineage comprising s_0 S-cells on day 11 ($s_0 \approx 5$ for
280 rates in table 1; and we assume $r_{S \rightarrow SS} \approx r_{S \rightarrow A}$) and whose S-cell population goes
281 extinct ΔT_S days later. We note that *final lineage size* here does not refer to the size
282 on day 40 (or any other particular point in time), but rather to the eventual size a
283 lineage will have reached when its growth ceases. While this does not exactly match
284 our simulation setup (we only simulate up to day 40) the approximate final lineage
285 sizes $L(\Delta T_S)$ still matches the simulation results well (figure 4B).

286 **Recovering S-cell extinction times from final lineage sizes**

287 By solving the equation $L(\Delta T_S) = L_i$, the time at which a lineage lost its last S-cell
288 can be estimated from the final size (L_i) of that lineage. To gauge the reliability of
289 this approach, we applied it to a simulated lineage size distribution for day 40, and
290 found that it recovers the number of lineages that reached S-cell extinction on a
291 particular day well (figure 4C). When applied to the experimentally observed lineage

292 sizes on day 40, the estimated number of lineages reaching S-cell extinction lies
293 close to the SAN model prediction, but slightly exceeds it up to about day 30.

294 **Emergence of a Zipfian law**

295 If A- and N-cells are disregarded, the SAN model is equivalent to the well-studied
296 birth-death process, and in particular, the distribution of the S-cell extinction time ΔT_S
297 is known (Feller, 1939). By translating this distribution via $L(\Delta T_S)$ into the
298 corresponding distribution of lineage sizes, the (approximate) distribution of final
299 lineage sizes (i.e. of lineages which have ceased growth) can be found. If we
300 consider only sufficiently large S-cell extinction times ΔT_S , the probability of ΔT_S is
301 (approximately) proportional to $1/\Delta T_S^2$ and translation into lineage sizes via $L(\Delta T_S)$
302 yields a Zipfian law with $\alpha = 0.5$. This theoretical prediction matches the empirical
303 observation that lineage sizes approach a Zipfian law with $\alpha \approx 0.46$.

304 **Discussion**

305 We have empirically observed lineages in cerebral organoids to initially grow fast
306 and roughly uniformly until some lineage-specific stopping time at which growth
307 slows down significantly or ceases altogether; and have found the size of slow or
308 non-growing lineages to follow a Zipfian power law with exponent $-1/\alpha$, $\alpha = 0.46$.
309 While the destructive nature of NGS-based lineage tracing prevents us from directly
310 observing lineages as they switch their growth regime, alternative hypotheses would
311 necessarily involve either very early fate decisions, or lineage-specific proliferation
312 rates to explain the large diversity of observed lineage sizes. Both alternative models
313 seem unlikely given that organoids are grown from a homogenous population of
314 stem cells.

315 To study the cause of the apparently random and lineage-wide switch of growth
316 regime we introduced the cellular SAN model. This model accurately recapitulates all
317 experimental data and shows that observed lineage growth dynamics to emerge
318 from neutral competition within a proposal long-lived population of roughly 10,000
319 symmetrically dividing stem cells (S-cells). Under the SAN model the apparently
320 lineage-wide switch of growth regime occurs despite the lack of either direct or
321 indirect (e.g., through spatial colocation) lineage-wide events. Instead, growth of a
322 lineage slows down and eventually ceases as the result of the lineage vanishing

323 from the S-cell population through neutral competition; lineage survival time within
324 the organoid's S-cell population is thus the major determinant of lineage size.

325 The relationship between a lineage's survival times within the organoids S-cell
326 population and the size it eventually attains can be expressed by a formula. Inverting
327 this formula allows the history of the organoids S-cell population that was recorded
328 within its clonal composition to be read; doing so we found for days 11-30 a slight
329 excess of lineages reaching S-cell extinction in the experimental data compared to
330 the SAN model. We hypothesize that this might point to gradual reduction of division
331 and differentiation rates in organoids; Since the SAN model assumes constant rates
332 between days 11 and 40, a gradual reduction of rates would cause the model to
333 appear to fall behind at first, and then to catch up once the true rates have fallen
334 below the model's rates.

335 While we found that we cannot estimate the rates of most division and
336 differentiation events in the SAN model precisely, we could robustly determine the
337 rates of S-cell division and differentiation to be almost identical. This implies that the
338 population of symmetrically dividing cells in cerebral organoids is long-lived, and in
339 particular that organoids still contain a population of symmetrically dividing cells after
340 40 days.

341 Furthermore, the similarity of the S-cell division and differentiation rates implies
342 the existence of a mechanism that controls the S-cell population size by linking S-cell
343 differentiation and subsequent replenishment through symmetric divisions. Yet that
344 link must be stochastic in nature; if S-cells simply divided asymmetrically to produce
345 A-cells, or if after symmetric division exactly one offspring always differentiated, no
346 neutral competition between S-cells would occur and the observed large variability of
347 lineage size would remain unexplained. In the terminology of Simons and Clevers
348 (2011), the mechanism must thus be of the *population asymmetric* type.

349 Given the similarity between the population-level link of S-cell division and
350 differentiation and the dynamics within stem cell niches in intestinal crypts (Snippert
351 et al., 2010), we conjecture that similar structures located within proliferation centers
352 called *neural rosettes* (Esk et al., 2020). might be responsible for balancing division
353 and differentiation of S-cells.

354 To study the mechanism controlling S-cell population size in more detail, it
355 needs to be probed experimentally by perturbing organoids at specific points in time
356 and observing their response. If a fraction of cells is killed, different mechanism

357 would respond differently: A mechanism that relies on spatial constraints (i.e. stem
358 cells being pushed out of a niche) would be expected to show a reduced rate of
359 differentiations until the population has recovered. A regulatory mechanism which
360 more directly links S-cell division to differentiation would respond differently; there we
361 might expect the S-cell population to never reach its original size, but to instead
362 increase its overall cell turn-over to make up for lost S-cells.

363 Since the SAN model accurately predicts the lineage sizes observed for
364 cerebral organoids grown from wildtype cells, it is also useful both when planning
365 organoid-based perturbation screens, and when analyzing the resulting data. During
366 the planning phase the model makes it possible to judge the effect of proliferation
367 phenotypes on final lineage size, and thus to estimate the statistical power of
368 different screen designs. During statistical analysis of screening data, the model
369 provides a baseline (null model), against which the sizes of (genetically) perturbed
370 lineages can be compared.

371 To facilitate the adoption of the SAN model, we offer an implementation of the
372 model both as an interactive online service (URL to be determined) as well as a R
373 package (<http://github.com/Cibiv/SANjar>).

374 **Funding**

375 This work was supported by the Austrian Science Fund (FWF) project number F78.

376 Methods

377 Total organoid sizes

378 For days 0 through 21, organoid sizes were measured using fluorescence-activated
379 cell sorting (FACS). For days 11 through 40, organoid volumes were estimated from
380 microscopy images, and translated into cell counts using the average number of
381 cells per volume for days 11 through 21 where both FACS and volume
382 measurements were available.

383 NGS data processing

384 The lineage tracing data of Esk et al. (2020) was obtained from GEO (accession
385 GSE151383, supplementary file GSE151383_LT47.tsv.gz), and organoids “H9-
386 day06-03” and “H9-day09-01” removed as outliers. Based on the assumption that in
387 all samples the most common lineage size is 1 cell, we located the mode of the log-
388 transformed read count distribution for every sample and used it to normalize relative
389 lineage sizes (reads) to absolute cell counts. The validity of the underlying
390 assumption is confirmed by the good agreement the sum of absolute lineage sizes
391 and the FACS and area-derived estimates of total organoid size.

392 Pareto index and fast-slow threshold estimation

393 For each organoid, we used the observed lineage sizes l_1, \dots, l_n to estimate the
394 Pareto equality index α and minimal lineage size m with the maximum-likelihood
395 estimator

$$\hat{m} = \min_i l_i, \quad \hat{\alpha} = n \left(\sum_i \log \frac{l_i}{\hat{m}} \right)^{-1},$$

396 and computed the steady-state average $\bar{\alpha}$ from the alpha estimates of all organoids
397 sequenced on day 11 or later. To find the fast-slow threshold l_{Th} for a particular
398 organoid, we first found intersect d^{Pareto} such that the Pareto-induced rank-size
399 powerlaw $\log_{10} L^{\text{Pareto}}(r) = -\bar{\alpha}^{-1} \log_{10} r + d^{\text{Pareto}}$ fits the size of the smallest observed
400 lineage, and determined the smallest rank R for which the actual lineage size $l_{(r)}$
401 matches or exceeds the power law $L^{\text{Pareto}}(r)$. We then fit a separate log-log-linear
402 model $\log_{10} L^{\text{Large}}(r) = k \log_{10} r + d^{\text{Large}}$ to lineages with ranks $1, \dots, \sqrt{R}$ (which we

403 assume are surely not governed by the Pareto law), and set l_{Th} to the size at which
 404 the two laws intersect (meaning $l_{Th} = L^{Pareto}(r) = L^{Large}(r)$).

405 SAN model simulation

406 The total number of S-, A- and N-cells that an organoid is predicted to comprise at
 407 time t is computed based on the deterministic SAN model (with rates $r_{S \rightarrow SS}$, $r_{S \rightarrow \emptyset}$,
 408 $r_{S \rightarrow A}$, $r_{S \rightarrow N}$, $r_{A \rightarrow AN}$, $r_{A \rightarrow N}$ of these events occurring per cell and per day). The
 409 deterministic SAN model is described by the ordinary differential equations (ODE),

$$\begin{aligned} s(0) &= 30,000, & a(0) &= 0, & n(0) &= 0, \\ \dot{s} &= (r_{S \rightarrow SS} - r_{S \rightarrow \emptyset} - r_{S \rightarrow A} - r_{S \rightarrow N})s, \\ \dot{a} &= r_{S \rightarrow A}s - r_{A \rightarrow N}a, \\ \dot{n} &= r_{S \rightarrow N}s + (r_{A \rightarrow AN} + r_{A \rightarrow N})a, \end{aligned}$$

410 which can be solved analytically (for the time-homogenous case) and is then
 411 evaluated separately for each time interval within which rates are constant (The
 412 initial number $s(0)$ of S-cells is set to 30,000 instead of 24,000 to account for a slight
 413 excess in the number of observed lineages on day 0, likely due to a combination of
 414 multiple labelling and sequencing artefacts).

415 To find the predicted lineage size distribution at time t , the stochastic SAN
 416 model is simulated independently for each of the 30,000 lineages in an organoid.
 417 The simulation proceeds in discrete time steps Δt , which are chosen small enough to
 418 make the probability of a single cell undergoing two events negligible ($< 10^{-3}$).
 419 Given the numbers $S_i(t)$, $A_i(t)$, $N_i(t)$ of S-, A-, N-cells comprising lineage i at time t ,
 420 the number of cells Δ_e undergoing event e is chosen from a Poisson distribution.
 421 Specifically,

$$\begin{aligned} \Delta_{S \rightarrow SS} &\sim \text{Poisson}(r_{S \rightarrow SS}S_i(t)\Delta t), & \Delta_{S \rightarrow \emptyset} &\sim \text{Poisson}(r_{S \rightarrow \emptyset}S_i(t)\Delta t), \\ \Delta_{S \rightarrow A} &\sim \text{Poisson}(r_{S \rightarrow A}S_i(t)\Delta t), & \Delta_{S \rightarrow N} &\sim \text{Poisson}(r_{S \rightarrow N}S_i(t)\Delta t), \\ \Delta_{A \rightarrow AN} &\sim \text{Poisson}(r_{A \rightarrow AN}A_i(t)\Delta t), & \Delta_{A \rightarrow N} &\sim \text{Poisson}(r_{A \rightarrow N}A_i(t)\Delta t), \end{aligned}$$

422 and the number of S-, A-, N-cells at time $t + \Delta t$ is then set to be

$$\begin{aligned} s_i(t + \Delta t) &= s_i(t) + \Delta_{S \rightarrow SS} - \Delta_{S \rightarrow \emptyset} - \Delta_{S \rightarrow A} - \Delta_{S \rightarrow N}, \\ a_i(t + \Delta t) &= a_i(t) + \Delta_{S \rightarrow A} - \Delta_{A \rightarrow N}, \\ n_i(t + \Delta t) &= n_i(t) + \Delta_{S \rightarrow N} + \Delta_{A \rightarrow AN} + \Delta_{A \rightarrow N}. \end{aligned}$$

423 Finally, the lineage size distribution $l_1, \dots, l_{30,000}$ at time t is found by summing up the
 424 number of S-, A- and N-cells, $l_i(t) = s_i(t) + a_i(t) + n_i(t)$.

425 Technical noise simulation

426 The effect of PCR amplification and sequencing on the observed lineage sizes was
 427 simulated using a stochastic model of PCR amplification and sequencing (Pflug and
 428 von Haeseler, 2018) with parameters *PCR efficiency* and *average reads per*
 429 *molecule* (in our case per *lineage*). For every sampling time t , we simulated one
 430 read count (normalized to one read per cell on average) per lineage; parameters
 431 were *PCR efficiency* 35% (estimated from the day 0 data) and *average reads per*
 432 *lineage* $W l_i / \sum_i l_i$ for a lineage comprising l_i cells (W is the median experimental
 433 library size for time t). The simulated read counts were then normalized to cells by
 434 division by the average number of reads per cell ($W / \sum_i l_i$).

435 SAN rate estimation

436 For days 0-3 and 3-6, rates which replicate the experimental data well were found by
 437 trial and error. For the remaining 6 biologically relevant rates (of $S \rightarrow S$ S and $S \rightarrow N$
 438 between 6 and 11, and $S \rightarrow S$ S , $S \rightarrow A$, $A \rightarrow A$ N and $A \rightarrow N$ between days 11 and
 439 40) we computed the posterior distribution given experimentally observed total
 440 organoid sizes $\hat{c}^{(day t)}$ (on days $t \in \mathcal{D} = \{0, 3, 6, 9, 10, 13, 14, 16, 17, 19, 21, 22, 25,$
 441 $28, 31, 32, 35, 37, 38\}$) and ranked lineage sizes $\hat{l}_{(r)}^{(day 11)}$, $\hat{l}_{(r)}^{(day 40)}$ (on days 11 and 40,
 442 for ranks $r \in \mathcal{R} = \{1, 2, 5, 10, 15, 25, 40, 60, 100, 150, 250, 400, 600, 1000, 1500,$
 443 $2500, 4000, 6000, 10000, 15000, 25000\}$). To account for biological differences
 444 between replicates we assumed that experimental observations are log-normally
 445 distributed around the SAN model predictions $c^{(day t)}$ and $l_{(r)}^{(day 11)}$, $l_{(r)}^{(day 40)}$; the
 446 likelihood of the rate vector θ (comprising the 6 rates mentioned above) given the
 447 experimental data is thus

$$l(\theta) = -\frac{1}{2} \sum_{t \in \mathcal{D}} \left(\frac{\mu[\hat{c}^{(day t)}] - c^{(day t)}}{\sigma[\hat{c}^{(day t)}]} \right)^2 - \frac{1}{2} \sum_{t \in \{11, 40\}} \sum_{r \in \mathcal{R}} \left(\frac{\mu[\hat{l}_{(r)}^{(day t)}] - l_{(r)}^{(day t)}}{\sigma[\hat{l}_{(r)}^{(day t)}]} \right)^2$$

448 where $\mu[\dots]$ and $\sigma[\dots]$ denote the mean respectively standard deviation across
 449 biological replicates. Rates were restricted to lie between 0 and 4 and *a priori*
 450 assumed to be equally probable; the posterior probability of θ is thus proportional to
 451 $l(\theta)$. To find this posterior distribution, we sampled 1,000 random rate vectors
 452 according to their likelihoods by simulating 1,000 Markov chains using pseudo-
 453 marginal Metropolis-Hastings Markov chain Monte Carlo sampling (Beaumont 2003;

454 Andrieu & Roberts, 2009; Warne et al., 2020). We then computed the maximal mode
455 of the (joint) posterior distribution with the mean-shift algorithm to obtain the MAP
456 estimates (table 1).

457 **Mathematical Analysis**

458 If we consider only S-cells, the SAN model corresponds to the well-known birth-
459 death process (Feller, 1939). We consider the diffusion approximation of this process
460 and restrict our mathematical treatment to day 11 and later where the rates of
461 symmetric division ($r_{S \rightarrow SS}$; birth) and of differentiation ($r_{S \rightarrow A}$; amounts to death since
462 we consider only S-cells) are similar enough to be considered identical ($r_{S \rightarrow SS} =$
463 $r_{S \rightarrow A} = \lambda/2$). The number of S-cells within a lineage at time t (where $t = 0$ represents
464 day 11) is then governed by the stochastic differential equation (SDE)

$$ds(t) = \sqrt{\lambda s(t)} dW(t).$$

465 Using Onsager-Machlup theory (Onsager & Machlup, 1953; Dürr & Bach 1978) we
466 find the most probably trajectory of a lineage that contains s_0 cells at $t = 0$ and loses
467 its last S-cell ΔT_S days later,

$$s_{\text{ext}}(t | s_0, \Delta T_S) = s_0 \left(1 - \frac{t}{\Delta T_S}\right) \left(1 + \rho \frac{t}{\Delta T_S}\right) \quad \text{where} \quad \rho = \frac{\Delta T_S}{s_0} \lambda \sqrt{\frac{3}{8}} - 1.$$

468 On average, a lineage grows by $\lambda/2$ A-cells per S-cell and per day, and over its
469 lifetime every A-cell will eventually produce $r_{A \rightarrow AN}/r_{A \rightarrow N}$ additional N-cells through
470 asymmetric division. Eventually, a lineage that starts out with s_0 S-cells and loses its
471 last S-cell ΔT_S days later will thus approximately grow to size

$$L(\Delta T_S) = \frac{\lambda}{2} \left(1 + \frac{r_{A \rightarrow AN}}{r_{A \rightarrow N}}\right) \int_0^{\Delta T_S} s_{\text{ext}}(t | s_0, \Delta T_S) dt.$$

472 Integration of this expression yields Eq. (1).

473 **References**

- 474 Adamic, L. A., and Huberman, B. (2002). Zipf's law and the Internet. *Glottometrics* 3, 143–150.
- 475 Andrieu, C., and Roberts, G. (2009). The Pseudo-Marginal Approach for Efficient Monte Carlo Computations.
476 *The Annals of Statistics* 37 (2).
- 477 Corominas-Murtra, B., Scheele, C.L.G.J., Kishi, K., Ellenbroek, S.I.J., Simons, B.D., van Rheenen, J., and
478 Hannezo, E. (2020). Stem cell lineage survival as a noisy competition for niche access. *Proceedings of the*
479 *National Academy of Sciences* 117, 16969–16975.

- 480 Beaumont, M. (2003). Estimation of Population Growth or Decline in Genetically Monitored Populations. *Genetics*
481 164 (3), 1139-1160.
- 482 Dürr, D., and Bach, A. (1978). The Onsager-Machlup function as Lagrangian for the most probable path of a
483 diffusion process. *Communications in Mathematical Physics* 60, 153–170.
- 484 Esk, C., Lindenhofer, D., Haendeler, S., Wester, R.A., Pflug, F., Schroeder, B., Bagley, J.A., Elling, U., Zuber, J.,
485 von Haeseler, A., et al. (2020). A human tissue screen identifies a regulator of ER secretion as a brain-size
486 determinant. *Science* 370, 935–941.
- 487 Feller, W. (1939). Die Grundlagen der Volterraschen Theorie des Kampfes ums Dasein in
488 wahrscheinlichkeitstheoretischer Behandlung. *Acta Biotheoretica* 5, 11–40.
- 489 Klingler, E., and Jabaudon, D. (2020). Do progenitors play dice? *ELife* 9, e54042.
- 490 Lancaster, M.A., Corsini, N.S., Wolfinger, S., Gustafson, E.H., Phillips, A.W., Burkard, T.R., Otani, T., Livesey,
491 F.J., and Knoblich, J.A. (2017). Guided self-organization and cortical plate formation in human brain organoids.
492 *Nature Biotechnology* 35, 659–666.
- 493 Llorca, A., Ciceri, G., Beattie, R., Wong, F.K., Diana, G., Serafeimidou-Pouliou, E., Fernández-Otero, M.,
494 Streicher, C., Arnold, S.J., Meyer, M., et al. (2019). A stochastic framework of neurogenesis underlies the
495 assembly of neocortical cytoarchitecture. *ELife* 8, e51381.
- 496 Moran, P.A.P. (1958). Random processes in genetics. *Mathematical Proceedings of the Cambridge Philosophical*
497 *Society* 54, 60–71.
- 498 Onsager, L., and Machlup, S. (1953). Fluctuations and irreversible processes. *Physical Review* 91, 1505–1512.
- 499 Pflug, F.G. and von Haeseler, A. (2018). TRUmiCount: Correctly counting molecules using unique molecular
500 identifiers. *Bioinformatics* 34-18, 3137–3144.
- 501 Simons, B. D. and Clevers, H. (2011). Strategies for homeostatic stem cell self-renewal in adult tissues. *Cell* 145,
502 851–862.
- 503 Snippert, H.J., van der Flier, L.G., Sato, T., van Es, J.H., van den Born, M., Kroon-Veenboer, C., Barker, N.,
504 Klein, A.M., van Rheenen, J., Simons, B.D., et al. (2010). Intestinal Crypt Homeostasis Results from Neutral
505 Competition between Symmetrically Dividing Lgr5 Stem Cells. *Cell* 143, 134–144.
- 506 Warne, D. J., Baker, R. E., and Simpson M. J. (2020). A Practical Guide to Pseudo-Marginal Methods for
507 Computational Inference in Systems Biology. *Journal of Theoretical Biology* 496.
- 508 Zechner, C., Nerli, E., and Norden, C. (2020). Stochasticity and determinism in cell fate decisions. *Development*
509 147, dev181495.

511 Supplemental Information

512

