

1 **Life-history predicts global population responses to the** 2 **weather in the terrestrial mammals**

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14

15 **Abstract**

16 With the looming threat of abrupt ecological disruption due to a changing climate, predicting which
17 species are most vulnerable to environmental change is critical. The life-history of a species is a
18 promising candidate for explaining differences in climate-change responses, but we now need data
19 linking population change, weather and life-history to explore these predictions. Here, we use long-
20 term abundance records from 157 species of terrestrial mammals to investigate the link between weather
21 and annual population growth rates. Overall, we found no consistent effect of temperature or
22 precipitation anomalies on annual population growth rates, but there was variability in weather
23 responses for populations within a species. Crucially, however, long-lived mammals with smaller litter
24 sizes had responses with a reduced absolute magnitude compared to their shorter-living counterparts
25 with larger litters. These results highlight the role of species-level life-history in driving responses to
26 the environment.

27

28 **Introduction**

29 Climate change is one of the greatest challenges we face in the twenty first century¹. Although habitat
30 loss and direct exploitation are currently the greatest drivers of extinction in the natural world^{1,2}, changes
31 to the climate are predicted to cause widespread declines to global biodiversity in the coming decades³⁻
32 ⁵. For mammals and birds, temperature increases are already associated with declining population
33 trends⁶ and many endangered species have already been negatively impacted by climate change in at
34 least part of their range⁷. Perhaps more worryingly, abrupt ecological disruption due to runaway climate
35 change has been predicted to have large negative impacts on biodiversity, with tropical ecosystems
36 being affected as early as 2030⁸. Furthermore, these future impacts will likely be exacerbated by
37 synergism between the climate and other drivers of extinction such as habitat loss^{9,10}. Research
38 highlighting the species and ecosystems that are most vulnerable to climate change impacts will
39 therefore provide crucial knowledge to prevent future losses to global biodiversity.

40 Not all species are equally vulnerable to climate change. Species vary in their climatic niches
41 and in their behavioural, physiological, and demographic responses to environmental change and we
42 therefore expect there to be both climate ‘winners’ and ‘losers’^{11–13}. At a macro-scale, species
43 occupancy data highlight that geographic range shifts are the key response associated with climate
44 change across taxa, resulting in changes to community composition, but not necessarily population
45 decline^{14,15}. Recent broad assessments of biodiversity change metrics from species assemblage and
46 abundance records mirror this paradigm with both occupancy and abundance trend patterns suggesting
47 a balanced frequency of winners and losers^{15,16}. In the Marine realm, species richness increases were
48 associated with warming temperatures, consistent with the expectation that warming marine ecosystems
49 will receive an influx of species tracking suitable temperatures¹². However, there was no consistent
50 temperature-related biodiversity change effect on land¹². Changes in species richness and other
51 biodiversity metrics however do not necessarily equate to population declines. For birds and mammals
52 average abundance trends were negatively associated with rates of climate warming⁶. Studies unpicking
53 how changes in weather patterns cause population change are therefore vital¹⁷, and a growing body of
54 literature is exploring the relationship between the climate and the demographic processes driving
55 population decline^{18–22}. Applying these concepts at a comparative scale and assessing finer-scale
56 population changes with respect to changes in the weather, and their relationship to species traits, will
57 aid in illuminating consistent or disparate climate change responses across the tree of life^{21,23}.

58 Variation in the demographic responses of organisms climate change suggests that life-history is
59 a promising target for explaining responses to environmental change. The timing of demographic events
60 relating to the key demographic rates of survival and recruitment are evolved responses to the
61 environment, and characteristics relating to both ‘slow’ and ‘fast’ life-histories are therefore adaptive
62 in different environmental contexts²⁴. Indeed, Pacifici et al.⁷ concluded that intrinsic traits, including
63 habitat specialisation and aspects of life-history, were associated with negative climate-mediated
64 population effects reported in mammals and birds. Life-history differences between three amphibian
65 species in Western Europe drove differences in survival and reproduction in response to the North
66 Atlantic Oscillation²⁵. Generally, we expect that organisms with slower life-histories are better-adapted
67 to cope with fluctuations in the environment. Longer-lived organisms have a reduced relative effect of
68 variability in vital rates, variability which is expected during environmental change, on population
69 growth rates²⁶ and long-lived plants have weaker absolute demographic responses to weather²³.
70 Furthermore, the recently developed concept of demographic resilience uses demographic rates
71 characterising the life cycle of an organism to quantify their resilience to perturbations²⁷. Comparative
72 approaches linking life-history traits to climate change responses may therefore provide a crucial
73 predictive link to improve our understanding of climate vulnerability.

74 In this study, we investigated annual population responses to temperature and precipitation in
75 populations of terrestrial mammals across the world’s ecosystems. Importantly, we tested whether life-
76 history predicts population responses to the weather, and therefore its utility in assessing vulnerability

77 to climate change. We addressed these questions using 486 long-term (≥ 10 years) abundance records
78 from 157 species of terrestrial mammal obtained from the Living Planet Database⁴, by implementing a
79 two-step meta-regression framework. First, for each abundance record, we assessed how observed
80 annual population growth rates were influenced by weather anomalies (annual deviation from long-
81 term average weather patterns) using autoregressive additive models that accounted for temporal
82 autocorrelation in abundance records and overall abundance trends. Then, we used a phylogenetically
83 controlled Bayesian meta-regression with weather effect coefficients as the response variable to address
84 three key questions: 1) Are there consistent temperature and precipitation effects on abundance change
85 across the terrestrial mammals? 2) How are these patterns influenced by covariance both within and
86 between species, and are there vulnerable biomes? 3) Can species-level life-history traits predict the
87 magnitude of population responses to the weather? The terrestrial mammals are an ideal study system
88 to explore the predictors of population responses to climate change because they are a well-studied
89 group with a combination of intensive abundance monitoring across the globe⁴, detailed life-history
90 information for hundreds of species^{28,29} and a highly-resolved phylogeny to facilitate phylogenetic
91 comparative analyses³⁰. Furthermore, there is growing evidence from the mammals of the mechanistic
92 links between the climate, demography, and population dynamics^{17,20-22}.

93 In light of previous findings from studies exploring biodiversity change metrics and life-history
94 theory in the context of environmental change, we make the following predictions related to our three
95 key questions. First, reflecting the complexity of overall biodiversity change and the paradigm of
96 climate winners and losers^{12,15,16}, we predict that across the mammals there will be no consistent
97 responses to temperature and precipitation. Second, recent evidence has highlighted phylogenetic
98 covariance between demographic traits³¹, which may also scale to population responses. However,
99 spatial variation between study sites, heterogeneity across a species range and the complex interaction
100 of demographic traits to drive population dynamics may reduce phylogenetic covariance. Finally, we
101 predict that organisms with slow life-histories will be more resilient to weather effects, reflecting an
102 evolutionary adaptation to buffer against fluctuating environments²⁶. We therefore expect responses to
103 weather to have a higher magnitude in short-living organisms with larger reproductive output compared
104 to their longer-lived counterparts.

105

106 **Results**

107 We assessed population responses to weather in 486 long-term abundance time-series records of 157
108 species of terrestrial mammals from across the world's ecosystems (Fig. S1). The time-series records
109 ranged in duration from 10 years to 35 years, with mean and median record lengths across records of
110 15.7 and 14 years, respectively (Fig. S1). The records were distributed across 13 terrestrial biomes,
111 including both tropical and temperate regions, but were generally biased towards north western Europe
112 and North America. We had records from 12 of 27 mammalian orders recognised by the IUCN Red
113 List for threatened species³², but most densely in the Artiodactyla ($n = 172$), Carnivora ($n = 127$) and

114 Rodentia (n = 82) (Fig. S1). The number of records for each species ranged from 1-17, with a mean of
115 3.1 and median of 2 records per species.

116

117 *No consistent population response to weather*

118 Overall, we found no consistent effect of either temperature or precipitation anomalies on annual
119 population growth rates in the terrestrial mammals (Fig.1). The raw weather effects on population
120 growth rates, ω , varied across species and records but were centred close to 0, with a mean temperature
121 coefficient of $-0.08 (\pm 0.73 \text{ SD})$; Fig. 1a) and mean precipitation coefficient of $0.02 (\pm 0.36 \text{ SD})$; Fig. 1b).
122 Furthermore, 95% of records had temperature and precipitation coefficients between -1.87 - 1.38 and $-$
123 0.59 - 0.83 , respectively. Nevertheless, approximately 10% (n = 51) of temperature effects and 3% of
124 precipitation were greater than 1 or less than -1, indicating that small clusters of populations experienced
125 more extreme annual responses to the weather (Fig.1). Our Bayesian meta-regression, controlling for
126 both within species variance, phylogenetic covariance and differences in sample size (number of years)
127 between records, mirrored the lack of consistent weather effects on population growth. The posterior
128 mean global intercept, $\bar{\alpha}$, for temperature effects was $0.07 [-0.17-0.33]$ (95% credible intervals) and for
129 precipitation effects was $-0.06 [-0.28-0.14]$ (Fig.1a and Fig.1b). There was, however, a small positive
130 effect of the number of years of population data for a record and the response to temperature, with a
131 linear slope, β_N of $0.09 [0.0-0.18]$. Together with the results of the global intercept $\bar{\alpha}$, this suggests that
132 shorter records were associated with more negative temperature effects. Overall, these results highlight
133 the paradigm of the existence of both winners and losers in weather responses, but no clear effect across
134 the Mammalia.

135

136 *Spatial effects and variation between species*

137 We tested whether there were differences in weather responses among ecological biomes because biome
138 effects may be indicative of more extreme responses to weather in some habitats. Using leave-one-out
139 cross-validation, we compared the predictive performance of the model including the effect of biome
140 relative to the base model, and we found no evidence for an influence of biome on either temperature
141 ($\Delta \text{elpd} = -0.67$ relative to base model) or precipitation ($\Delta \text{elpd} = -0.73$) effects (see Fig. S15-16 for more
142 information). Furthermore, we explored the role of spatial autocorrelation at driving differences in
143 weather coefficients across records using Morans I tests and spatially explicit meta-regressions but did
144 not find evidence for spatial autocorrelation in weather effects (Figs. S18-S20). We also incorporated
145 both phylogenetic covariance (σ_{PHYLO}^2) and species-level variance (σ_{SPECIES}) to capture both among-
146 and within-species variation. Interestingly, we found far greater levels of within-species variation in
147 temperature responses compared to among-species variance (Fig. 1c). The posterior mean for species-
148 level variance in temperature effects was $0.2 [0.01-0.4]$ which was 20 times greater than the posterior
149 estimate of $0.01 [0.0-0.03]$ for phylogenetic covariance (Fig. 1c). Similarly, for precipitation the

150 posterior mean for species-level variance was five times greater than for phylogenetic covariance, with
151 a value 0.05 [0.0-0.15] compared to 0.01 [0.0-0.02] (Fig. 1c). These patterns are reflected in the
152 temperature and precipitation coefficients, for which large variation can be seen among records of the
153 same species. For example, *Myodes glareolus* (bank vole) in the Rodentia had 9 population records,
154 and a range of temperature/precipitation effects of -1.64-3.59 and -1.27-1.44 respectively, compared to
155 -3.32-5.67 and -1.27-1.44 across Rodentia as a whole (Fig.1). This result highlights the potential
156 importance of within-species variability in population responses to environmental change.

157

158 *Life-history predicts population responses to weather*

159 Across the terrestrial mammals, we found that both maximum longevity and mean litter size at the
160 species-level predicted the magnitude of population responses to weather. We tested a set of Gamma
161 models incorporating univariate, multivariate and 2-way interactions between maximum longevity,
162 litter size, and adult body mass and their influence on the absolute magnitude of
163 temperature/precipitation effects using model selection and leave-one-out cross-validation (Table S1 &
164 S2). As with our Gaussian models of overall weather effects, we found that the sample size of a record
165 had strong negative impact on the absolute temperature and precipitation responses, with posterior
166 estimates on the linear predictor scale of $\beta_N = -0.26$ [-0.35- -0.16] and $\beta_N = -0.35$ [-0.46- -0.24],
167 respectively (Fig. S17). Namely, shorter records were associated with temperature and precipitation
168 responses of a larger magnitude. We found no association between adult body mass and either
169 temperature ($\beta_{BODYMASS} = -0.01$ [-0.19-0.16]) or precipitation responses ($\beta_{BODYMASS} = 0.03$ [-0.20-
170 0.25]). Furthermore, we found no strong evidence for any two-way interactions between life-history
171 variables (Table S1 & S2). For both temperature and precipitation effects, the most competitive model
172 was the univariate model including mean litter size ($\Delta\text{elpd} = 3.03$ and $\Delta\text{elpd} = 3.00$, compared to the
173 base model for temperature and precipitation, respectively; Table S1 & S2). However, univariate
174 models including maximum longevity alone also had a higher predictive performance than the base
175 model ($\Delta\text{elpd} = 0.94$ and $\Delta\text{elpd} = 1.67$ for temperature and precipitation, respectively). In both cases
176 the second-best predictive model was the one that included univariate effects for longevity, bodymass
177 and litter size ($\Delta\text{elpd} = 2.88$ and $\Delta\text{elpd} = 2.94$; Table S1 & S2). Therefore, we selected the models
178 including all univariate life-history effects.

179 For both temperature and precipitation, our results highlight that shorter-living mammals with
180 greater litter sizes experienced weather effects of a significantly greater magnitude than longer-living,
181 slowly reproducing mammals (Fig. 2). The magnitude of weather responses was negatively associated
182 with longevity, with posterior means on the linear predictor scale of $\beta_{LONGEVITY} = -0.19$ [-0.44-0.07]
183 and $\beta_{LONGEVITY} = -0.24$ [-0.57-0.07] for temperature and precipitation, respectively (Fig. 2a & 2c).
184 Thus, a maximum longevity change from 10 months (*Akodon azarae*) to 80 years (*Loxodonta africana*)
185 was associated with a 2.3-fold and 2.9-fold decrease in the predicted magnitude of responses to

186 temperature and precipitation. So, for every additional 5 years of life, there was a 14% decrease in the
187 magnitude of responses to temperature and 18% decrease in the magnitude of responses to precipitation.
188 An organism's longevity is strongly correlated to their body mass, but the effect of longevity held
189 irrespective of whether adult body mass was also included in the model. In contrast, but also following
190 key predictions from life-history theory, the magnitude of weather responses had a strong positive
191 association with litter size, with posterior means of $\beta_{LITTER} = 0.33$ [0.13-0.52] and $\beta_{LITTER} = 0.30$ [0.06-
192 0.54] for temperature and precipitation, respectively (Fig. 2b & 2d). In other words, mammals bearing
193 more offspring in a single litter had greater responses to temperature and precipitation. A change in
194 litter size from 1 (monotocous species, various) to 17 (*Thylamys elegans*) was associated with a 4.17-
195 fold and 3.71-fold increase in the predicted magnitude of temperature and precipitation responses. For
196 every additional offspring invested into at the litter stage, there is a 24% increase in the magnitude of
197 temperature responses and 22% increase in the magnitude of precipitation responses.

198

199 **Discussion**

200 Our results provide an important empirical link between a species' life-history and its population
201 responses to environmental change. While we found no consistent patterns of responses to temperature
202 and precipitation anomalies across the mammals, life-history traits relating to the pace of life were
203 associated responses to weather. Namely, shorter-living species with increased litters sizes, or species
204 characterised with 'fast' life-history traits, responded with a greater magnitude compared to those with
205 'slow' life-history traits. While it has long been theorised that an organism's life-history traits evolve
206 in response to, and as an adaptation to, environmental conditions²⁴, rarely has this theory been tested at
207 a global scale. We find strong support for the hypothesis that longevity, and 'slow' life-history
208 characteristics more generally, buffer organisms against variability in the environment²⁶ and add to a
209 small number of studies linking population demography and the climate^{21,23}. Critically, our results
210 highlight the potential utility of life-history traits for predicting species vulnerability to climate change.
211 Based on our findings, abrupt ecological disruption from climate change will disproportionately impact
212 shorter-lived species with higher reproductive output.

213 Demography has a vital role to play in predicting population declines in the Anthropocene and
214 in highlighting targets for conservation management^{28,33}. Our study emphasises this role, demonstrating
215 the predictive power of demographic life-history traits when investigating responses to environmental
216 change. However, there are limitations and barriers to the utility of demography in conservation. Only
217 1.3% of tetrapods globally have sufficient demographic information with which to estimate population
218 dynamics²⁸. Here, we used summary traits that are available for many species (maximum longevity and
219 mean litter size), but in particular maximum recorded longevity, while sufficient as a broad indicator,
220 is strongly influenced by sampling variance and a flawed measure of longevity differences between
221 taxa³⁴. Ideally, lifetables with mortality and reproduction trajectories across the lifecycle can be
222 combined data on external drivers to investigate detailed patterns in population dynamics^{35,36}. The

223 recent development of the demographic resilience framework, which uses demographic data across the
224 lifecycle to simulate how a population may respond to perturbations²⁷, has excellent potential in
225 extending these findings to explore demographic relationships with climate responses in detail.
226 Unfortunately, however, detailed (st)age-specific demographic information is not currently available
227 for a majority of species, but growing in availability rapidly³⁷. Therefore, there is a need to continue to
228 increase the collection of demographic data (and other traits) for many more species than are currently
229 available²⁸, so that we may predict population changes with respect to environmental change. Achieving
230 this target may revolutionise the way we quantify species vulnerability to climate change^{12,15,16,21},
231 helping to prevent extinctions before they occur.

232 In line with recent global assessments of biodiversity in the face of climatic change²¹, we did
233 not find an overall consistent effect of weather anomalies on population growth rates. This may in part
234 reflect the fact that abundance changes are a higher-order process determined by complex interactions
235 between demographic processes that counteract each other^{16,21}. However, our results contrast with
236 findings of linear associations between mammal abundance and temperature change⁶. These differences
237 may reflect our approach to investigate annual changes, rather than long-term trends. Significant
238 population trends from long time-series are detectable from smaller component time-series even when
239 sampling is incomplete³⁸, and thus responses detected in trends may reflect broader changes in response
240 to the climate that are not detected in models of annual change. Furthermore, we estimated linear, annual
241 effects of weather on population growth rates, where population responses may actually be more
242 complex non-linear patterns or lagged effects. However, the detection of climate effects on average
243 trends may also be confounded by effects of other (sometimes more dominant) drivers (e.g. habitat
244 loss)². Nevertheless, our findings can be explained in light of recent studies from the Living Planet
245 Database that have found that the large majority of records do not exhibit population declines¹⁶.

246 Interestingly, we did not find evidence for phylogenetic covariance in weather responses
247 between species. Recent evidence from birds indicated strong phylogenetic covariance in vital rates,
248 particularly in adult survival, and the incorporation of phylogenetic information greatly improved
249 predictive performance when imputing vital rates³¹. Therefore, as with overall patterns, our findings
250 may reflect the trade-offs between vital rates, which cancel one another out when scaling up to
251 population-level processes such as population growth rates in response to the weather²¹. Furthermore,
252 for long-term time-series, there may also be temporal trade-offs in vital rates, where for example
253 investing heavily into survival in one year (in response to climate) may impact subsequent reproduction
254 for several years, decreasing the magnitude of population growth rates. The extent of phylogenetic
255 covariance in vital rate responses and trade-offs remains unknown, understanding how the climate
256 impacts demographic rates across species may provide a useful tool for imputing population responses
257 to the climate across the tree of life³¹.

258 We highlight the importance of variation in population responses to climate within a species
259 range. Sampling heterogeneity has recently been shown to have broad implications for metrics of

260 population dynamics, where demographic rates are poorly correlated among sampling sites for the same
261 species^{39,40}. Therefore, inferences obtained from monitoring single populations or studies may not
262 accurately portray species-level variability. This has broad implications for macroecology, particularly
263 for population viability assessments (PVA) and species-distribution modelling. First, as well as
264 suffering from data quality issues in their parameterisation⁴¹, our findings suggest that PVAs based on
265 data from a single population may not accurately reflect population viability across a species'
266 geographic range. Therefore, incorporating detailed demographic data, and investigating differences in
267 population responses across a range, could greatly improve our perspective on population viability³⁶.
268 Second, presence-only models of species distributions that do not account for the fact that responses to
269 the environment within a species range do not accurately represent species distributions⁴². Moving
270 towards trait-based monitoring and explicitly including demographic processes with mechanistic links
271 to appropriate drivers into species distribution models could greatly improve predictions of climate
272 change impacts on the biosphere⁸.

273 Ultimately, improving our predictions of how humans are influencing the natural world is
274 paramount to prevent rapid declines to global biodiversity⁴³. This however requires a large shift towards
275 both broad and detailed monitoring of global biodiversity. We show that linking species traits such as
276 life-history to changes in the environment may equip us with tools to predict and prevent future losses.

277

278 **Methods**

279 To assess the effects of weather on population growth rates we collated information on global weather
280 and the abundance, life-history and phylogeny of the terrestrial mammals. All analyses were carried out
281 using R version 4.0.4⁴⁴. For all data on the terrestrial mammals, taxonomies were resolved using the
282 *taxize* package version 0.9.98⁴⁵ and matched using the Global Biodiversity Information Facility
283 database (<https://www.gbif.org/>). All code used in the current study and full descriptions of the analyses
284 are archived in the Zenodo repository (doi:10.5281/zenodo.4707232), which was created from the
285 following GitHub repository https://github.com/jjackson-eco/mammal_weather_lifecycles.

286

287 *Time-series abundance data*

288 Annual time-series abundance data from across the terrestrial mammals were obtained from the Living
289 Planet Database found at https://livingplanetindex.org/data_portal. This database was developed by the
290 World Wildlife Fund and the Zoological Society of London as a tool to monitor global biodiversity,
291 and contains over 20,000 population records for over 4000 species of vertebrate⁴. The records measure
292 annual abundance in a variety of ways (e.g. full population counts, density, indices), and contain
293 information on the location, realm, biome and taxonomy of the species in the record. First, we included
294 only data for the terrestrial mammals that had species-level life-history information and coordinate
295 locations, which referred to either specific or more general locations for each population (accounted for
296 using weather data from a buffered radius around each location). We natural-log-transformed the raw

297 abundance data to ensure that weather coefficients were comparable across records and abundance
298 measure types. Because our analyses were focussed on estimating weather effects on annual population
299 growth rates using regression models with several variables, we included only long-term records with
300 10 or more consecutive years of abundance data, and only for years in which there was also weather
301 data (1979-2013). In one record (for *Bettongia penicillata*), there were two blocks with ≥ 10 years of
302 data, which were analysed separately. We also removed records with a high proportion ($>32\%$) and
303 consecutive occurrences of 0 in the raw abundance time-series. Our final dataset contained 486 records
304 from 157 terrestrial mammal species, which was used in all subsequent analyses (Fig. S1).

305

306 *Global weather data*

307 We used version 1.2.1 of the CHELSA monthly gridded temperature and precipitation timeseries
308 dataset, which is at a spatial resolution of 30 arc seconds ($\sim 1\text{km}^2$) for all months between 1979-2013
309 across the globe's land surface. The CHELSA datasets can be obtained from Karger et al.⁴⁶. Raster files
310 of the raw monthly mean temperature and total precipitation data were processed using the *raster*, *rgeos*,
311 and *sf* packages⁴⁷⁻⁴⁹. Using the Living Planet Database record coordinate locations as a centroid, we
312 averaged the monthly weather data for a buffered radius of 5km around each record location to account
313 for the lack of specificity in record locations using the *exactextractr* package⁵⁰. Averaged weather
314 variables and weather effects for alternate buffer radii (50m and 50km) were highly correlated (Fig. S2;
315 Fig. S7). Generally, we expect that organisms will respond to deviations in the weather compared to the
316 average values, as opposed to raw weather variables. Furthermore, across the globe's surface the
317 variance in weather variables changes substantially, which may influence population responses. Thus,
318 we explored population responses for the key weather variable of standardised annual anomalies, and
319 then validated our approach using annual weather variance. For the anomalies, we decomposed z-
320 transformed averaged monthly weather data for each location for the full timeseries (1979-2013) using
321 a Seasonal-Trend Decomposition by Loess (STL). We used a seasonal window of 7 and trend window
322 of 1000 for the decomposition and extracted the anomaly component, which describes the remainder
323 when accounting for the trend and seasonal components of the timeseries. We then used annual mean
324 temperature and precipitation anomalies as the key weather variables in subsequent analyses. These
325 weather anomalies are the average deviation of the temperature and precipitation from expected values
326 in a given year. We also investigated population responses to weather using annual weather variance,
327 which was calculated as the annual variance of monthly mean temperature and total precipitation values.

328

329 *Species-level life-history and phylogeny*

330 We used three key traits that broadly characterise species-level life-history that are available for a large
331 number of species: maximum longevity, litter size and adult body mass. We collected these traits from
332 the compendium developed by Conde et al.²⁸, combining information from three primary database
333 sources: The Amniote Life-History Database²⁹, PanTHERIA⁵¹ and AnAge⁵² databases. Adult body

334 mass data was obtained exclusively from the Amniote Life-History Database²⁹. Where multiple records
335 were available for a single species, we took the largest maximum longevity value and the mean litter
336 size/adult body mass. We removed erroneous raw litter size data for *Hydrochoerus hydrochaeris* and
337 *Marmota broweri*. For analysis, we z-transformed the natural-logarithm of raw life-history trait data,
338 and verified that the life-history variables were represented across the range of weather anomaly
339 variables in the raw data (Fig. S3). The mammal phylogeny was obtained from Upham et al.³⁰, which
340 uses a ‘backbone-and-patch’ Bayesian approach for a newly assembled 31-gene supermatrix and is part
341 of the Vertlife project (<https://vertlife.org/>). We used the maximum clade credibility tree in analysis,
342 which was processed using the *ape* package⁵³. *Loxodonta cyclotis* (African forest elephant) was
343 considered as *Loxodonta Africana* (African elephant) for analysis so that the abundance record and
344 phylogenetic data matched.

345

346 *Weather effects on annual population growth rates*

347 To assess comparative population responses to weather in the terrestrial mammals we used a two-step
348 meta-regression approach. First, for each record we estimated the effect of annual weather anomalies
349 (and weather variance) on population growth rates. We calculated the realised, proportional population
350 growth rate in year t as

$$351 \quad \lambda_t = \frac{X_{t+1}}{X_t}, \quad (1)$$

352 where X is the natural logarithm of abundance in year t . Then, with λ_t as the response variable, we
353 estimated the effect of temperature and precipitation anomalies on population growth using generalised
354 additive mixed models (GAMMs) fit using the *gamm* function of the *mgcv* package⁵⁴. Changes in
355 abundance are influenced by several drivers of population dynamics including habitat loss² and
356 population processes such as density dependence⁵⁵, which may confound any influence of the weather
357 on abundance. Therefore, because we aimed to assess the isolated impact of weather anomalies,
358 accounting for these trends in abundance and temporal autocorrelation was crucial. We initially
359 explored the extent of autocorrelation in abundance patterns using timeseries analysis and found
360 evidence for lag 1 autocorrelation in abundance, but not for greater lags (Fig. S4-5). Furthermore, we
361 tested the potential impact of density dependence on estimating environmental effects using an
362 autoregressive timeseries simulation and found that environmental effects were robust to density
363 dependence even for short timeseries (Fig. S6). Thus, for each record, we model population growth rate
364 in each year as

$$365 \quad \lambda_t = \beta^0 + \omega W_t + f(y_t), \quad (2)$$

366 where β^0 is the intercept and ωW_t is a linear parametric term with coefficient ω for the weather W
367 (temperature or precipitation anomaly) in year t . Here, positive coefficients indicate that positive
368 weather anomalies i.e. hotter/wetter years, were associated with population increases, and *vice versa*.
369 Identical additive regression models were run using weather variances as the weather variable W . The

370 term $f(y_t)$ captures the effect of year y_t as a non-linear trend, where the function f is a thin plate
 371 regression spline with a basis dimension of five⁵⁶. The function f was also fitted with an order 1
 372 autoregressive (AR(1)) correlation structure, as specified in the *nlme* package⁵⁷. Thus, the term $f(y_t)$
 373 incorporates both the non-linear trend in abundance and temporal autocorrelation. Finally, we validated
 374 our additive model approach by testing other models to calculate weather effects, including linear
 375 regressions both including and excluding temporal trends or density dependence, and a temporally
 376 autocorrelated model fit using the *glmmTMB* package⁵⁸ (Fig. S8-S9). Weather coefficients ω generated
 377 using linear year effects were positively correlated to those from additive models (Fig. S10).

378

379 *Bayesian meta-regression*

380 Second, with the weather effects ω from each record as the response variable, we explored comparative
 381 patterns in population responses to weather using a Bayesian meta-regression framework implemented
 382 in the *brms* package⁵⁹. Separate models were fit for temperature and precipitation. Bayesian meta-
 383 regression was used to address three key questions: 1) Were there consistent population responses to
 384 weather across the terrestrial mammals? 2) How did population responses vary within and between
 385 species and were there spatial patterns across biomes? 3) Does life-history predict the magnitude of
 386 population responses? To address questions 1 and 2, we used Gaussian models controlling for both
 387 phylogenetic and species-level covariance. The full model for record i and species j is given by equation

388 3

389

$$\begin{array}{l}
 \text{Linear model} \\
 \omega \sim \text{MVNormal}(\mu, \mathbf{S}) \\
 \mu_i = \alpha_{\text{SPECIES}[j]} + \beta_{\text{BIOME}[i]} + \beta_N N_i \\
 \text{Varying effects} \\
 \mathbf{S} = \sigma_{\text{PHYLO}}^2 \mathbf{V} \\
 \alpha_j \sim \text{Normal}(\bar{\alpha}, \sigma_{\text{SPECIES}}) \\
 \text{Priors} \\
 \bar{\alpha} \sim \text{Normal}(0, 0.3) \\
 \beta_{\text{BIOME}} \sim \text{Normal}(0, 0.15) \\
 \beta_N \sim \text{Normal}(0, 0.5) \\
 \sigma_{\text{PHYLO}}^2 \sim \text{Exponential}(8) \\
 \sigma_{\text{SPECIES}} \sim \text{Exponential}(8)
 \end{array}$$

391

392

393 where the weather effect ω (z-transformed for analyses), is given by a multivariate normal distribution
 394 with mean μ and phylogenetic covariance matrix \mathbf{S} . The global intercept is given by $\bar{\alpha}$, which estimates
 395 overall patterns in weather effects across records, addressing question 1. We incorporated phylogenetic
 396 covariance using a Brownian motion model, with the correlation matrix given by \mathbf{V} (calculated from
 397 the maximum clade-credibility tree) and variance factor σ_{PHYLO}^2 , from which between-species variance
 398 was estimated. We incorporated an intercept-only varying effect for species with the term $\alpha_{\text{SPECIES}[j]}$,
 399 from which within-species variance was estimated with σ_{SPECIES} . The term β_{BIOME} gives the spatial

406 effect of biome on weather responses. Thus, estimating within ($\sigma_{SPECIES}$) species variance, between
 407 (σ_{PHYLO}^2) species variance and the spatial effect of biome (β_{BIOME}), we explored question 2. All meta-
 408 regression models also included the linear effect of record length N (scaled number of years in the
 409 record) on weather effects, which was estimated using β_N . For all meta-regression models, we used
 410 regularising priors obtained from prior predictive simulations of the slope, intercept and exponential
 411 variance terms^{60,61}, to reflect the constraints in the raw data across species (see S1 and Fig. S11-14 for
 412 details). Gaussian meta-regression models were also fit for weather effects calculated using the annual
 413 weather variance, and the results obtained were largely identical to those obtained for weather anomalies
 414 (Fig. S21).

415
 416 For question 3, although on average we expect that species life-history influences the magnitude of
 417 responses to the environment, we have little evidence to suggest that life-history *per se* influences the
 418 directionality of responses²⁶. Thus, to address this question we explored how maximum longevity, litter
 419 size and adult body mass influenced the absolute magnitude of weather responses, $|\omega|$, using Gamma
 420 regression models with a log link. The full model for record i and species j is given by equation 4
 421

$$\begin{aligned}
 &\underline{\text{Linear model}} \\
 &|\omega| \sim \text{Gamma}(\eta, \mu) \\
 &\log \mu_i = \alpha_{SPECIES[j]} + \gamma_{PHYLO[j]} + LH + \beta_N N_i \\
 &\underline{\text{Varying effects}} \\
 &\alpha_j \sim \text{Normal}(\bar{\alpha}, \sigma_{SPECIES}) \\
 &\gamma_j \sim \text{MVNormal}(0, \mathbf{S}) \\
 &\mathbf{S} = \sigma_{PHYLO}^2 \mathbf{V} \\
 &\underline{\text{Priors}} \\
 &\bar{\alpha} \sim \text{Normal}(0, 0.3) \\
 &\beta_{LH} \sim \text{Normal}(0, 0.2) \\
 &\beta_N \sim \text{Normal}(0, 0.3) \\
 &\sigma_{PHYLO}^2 \sim \text{Exponential}(11) \\
 &\sigma_{SPECIES} \sim \text{Exponential}(2) \\
 &\eta \sim \text{Gamma}(2, 0.6)
 \end{aligned}$$

422
 423 where η is a shape parameter that was fit with a Gamma prior, and LH refers to a set of linear life-
 424 history terms ($\beta_1 x_1 + \dots \beta_k x_k$) that were explored using model selection. Specifically, for the three life-
 425 history traits, we explored a set of models incorporating univariate, multivariate and 2-way interaction
 426 terms, as well as a base model excluding all life-history effects. For the full set of ten candidate models
 427 please refer to the supplementary information (table S1 & S2). All life-history effects were fit with the
 428 same Normal prior, with mean 0 and standard deviation 0.3 (S1; Fig. S13). We assessed the predictive
 429 performance of candidate models using leave-one-out cross-validation implemented in the *loo*
 430 package⁶². Models were compared using the Bayesian LOO estimate of out-of-sample predictive
 431 performance, or the expected log pointwise predictive density (elpd)⁶². All final meta-regression models

432 were run over 3 Markov chains, with 4000 total iterations and 2000 warmup iterations per chain. Model
433 convergence was assessed by inspecting Markov chains, and the degree of mixing between chains using
434 \hat{R} .

435

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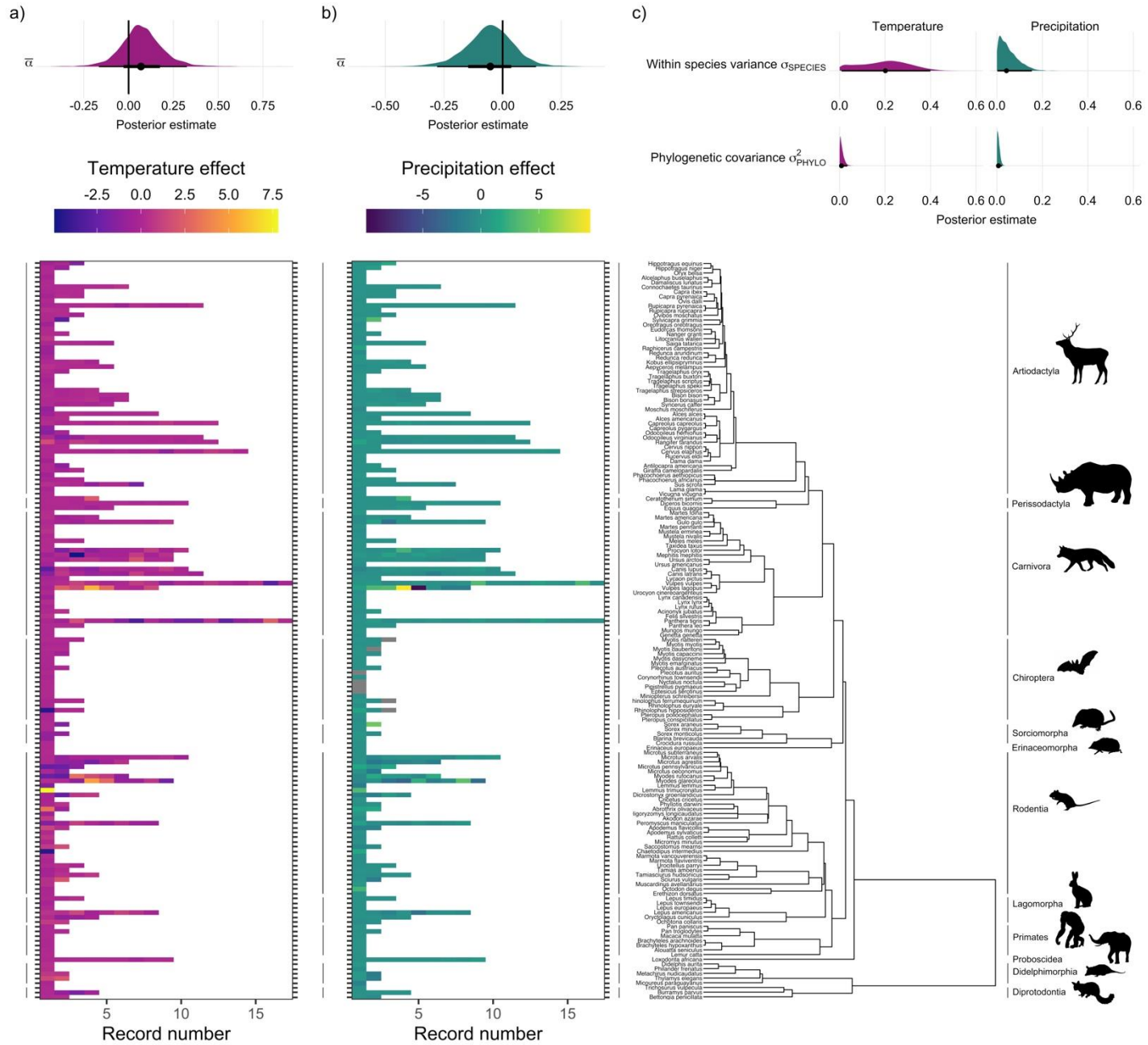
447

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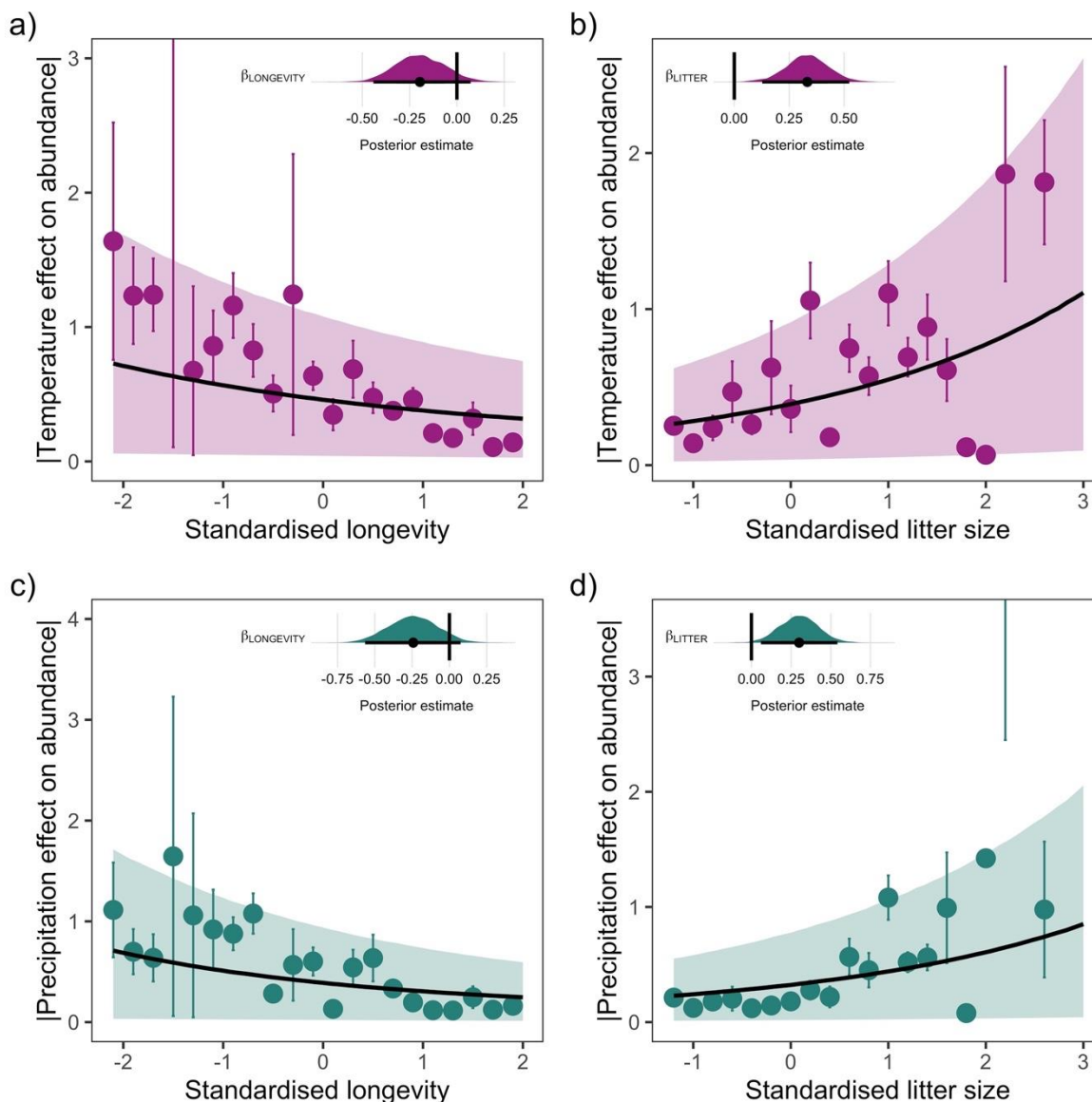
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563 **Figure 1. Global population responses to weather in the terrestrial mammals.** Heatmaps for
564 population responses to temperature (a) and precipitation (b) for 157 species of terrestrial mammal.
565 Each row of the heatmap is a species, and coloured rectangles are the population records. The colour
566 denotes the coefficient of temperature/precipitation effects derived from autoregressive additive
567 models. Here, positive numbers indicate that positive temperature/precipitation anomalies
568 (hotter/wetter than average in a given year) were associated with increases in population size, and *vice*
569 *versa*. The distribution half-eye plots in (a) and (b) (top) are summaries of the posterior distribution for
570 the global intercept ($\bar{\alpha}$) of temperature and precipitation responses across records, fit with a Gaussian
571 Bayesian meta-regression. The points give the approximate posterior mean and the error bar is
572 calculated using a cumulative distribution function. Bayesian models were fit incorporating
573 phylogenetic covariance using the maximum clade credibility tree from Upham et al.⁵⁹, which is plotted
574 on the right with annotations indicating the mammal order⁶⁰. The distribution half-eye plots in (c) are
575 the posterior distribution summaries for phylogenetic covariance and within-species variance from the
576 Gaussian Bayesian meta-regression.



577

578 **Figure 2. Life-history predicts population responses to weather in the terrestrial mammals.** Each

579 panel presents the mean absolute effect of temperature (a & b) and precipitation on population growth

580 rates, $|\omega|$, for standardised maximum longevity (a & c) and standardised mean litter size (b & d).

581 Standardisation was performed using a z-transformation of natural-log of the raw life-history trait. The

582 values on each x-axis are split into equal bins of 0.2 from the minimum to the maximum life-history

583 value. Points are coefficient means, with standard error bars. The black lines are the mean posterior

584 predictions from the best predictive model, where predictions were calculated averaging over all other

585 covariates and varying effects in the model. The shaded intervals are the 80% prediction intervals

586 calculated in the *rethinking* package⁶⁰. Panel insets give posterior distribution summaries for the slope

587 terms presented in each panel. Two points are omitted from the plotting panel due to large mean

588 coefficient values and high standard errors, which are visible on the plot.

589