Interacting Phenotypes and the Coevolutionary Process

Running title: Trait interactions and coevolution

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3 Abstract Coevolution occurs when species interact to influence one another's fitness. 4 resulting in reciprocal evolutionary change. In many coevolving lineages, trait expression in 5 one species is modified by the genotypes and phenotypes of the other, forming feedback 6 loops reminiscent of models of intraspecific social evolution. Here, we adapt the theory of 7 within-species social evolution, characterized by indirect genetic effects and social selection 8 imposed by interacting individuals, to the case of interspecific interactions. In a trait-based 9 model, we derive general expressions for multivariate evolutionary change in two species and 10 the expected between-species covariance in evolutionary change across a selection mosaic. 11 We show that reciprocal interspecific indirect genetic effects can dominate the 12 coevolutionary process and drive patterns of correlated evolution beyond what is expected 13 from direct selection alone. In extreme cases, interspecific indirect genetic effects can lead to 14 coevolution when selection does not covary between species or even when one species lacks 15 genetic variance. Moreover, our model indicates that interspecific indirect genetic effects 16 may interact in complex ways with cross-species selection to determine the course of coevolution. Importantly, our model makes empirically testable predictions for how different 17 18 forms of reciprocal interactions contribute to the coevolutionary process and influence the 19 geographic mosaic of coevolution.

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21 Key words: Coevolution, cross-species selection, interspecific indirect genetic effects,

22 geographic mosaic, quantitative genetics, species interactions

23 Introduction

24 Coevolution occurs when interacting lineages evolve reciprocally in response to one another 25 (Janzen 1980, Thompson 1982). Although the concept of coevolution may be applied to 26 lineages that share genes, such as males and females of the same species (Arnovist and Rowe 27 2002), it was originally invoked to explain patterns of correlated evolution between 28 interacting species (Ehrlich and Raven 1964). In some cases, coevolution can result in tightly 29 integrated mutualisms or spectacular arms races that drive the evolution of exceptional 30 phenotypes (Brodie et al. 2002, Pellmyr 2003, Johnson and Anderson 2010). Yet even 31 beyond these striking cases, coevolution is likely important for a wide range of interacting 32 lineages, including consumers and their resources, hosts and their pathogens, competitors, 33 and mutualists (Thompson 1982, 1994). Although coevolution has clearly played a major role in the origins of diversity, much is still unknown about when and how species 34 interactions generate reciprocal evolutionary change. 35

36 Theoretical models of coevolution typically focus on the fitness effects of trait interactions between coevolving species and how selection imposed by one species manifests 37 evolutionary change in an interacting species (Nuismer 2017). The interaction between 38 39 species in coevolving lineages is often intimate, with one species spending a greater part of 40 its life cycle in close contact with the other. Thus, coevolution bears a striking resemblance to intraspecific social evolution (Stearns 2012). Like social evolution, coevolution often 41 42 includes interacting or extended phenotypes, which arise when traits can only be understood within the context of interactions with others (Dawkins 1982, Moore et al. 1997). 43

In quantitative genetic models of intraspecific social evolution, social interactants
influence one another via two pathways, each of which has a counterpart in coevolutionary
theory. First, the phenotype of one individual may cause fitness effects in a social partner,
leading to a form of selection known as social selection (West-Eberhard 1979, West-

Eberhard 1983, West-Eberhard 1984, Wolf et al. 1999). At the heart of all coevolutionary 48 49 models is a form of reciprocal fitness interaction that resembles social selection, where the 50 fitness of individuals in one species is influenced by traits in an interacting species (Brodie and Ridenhour 2003, Ridenhour 2005, Nuismer 2017). Second, models of social evolution 51 52 may also include indirect genetic effects, which occur when the phenotype of one individual 53 depends on the genotype of an interacting partner (Moore et al. 1997, Wolf et al. 1998). A 54 classic example of indirect genetic effects is maternal effects, in which offspring phenotype is 55 a function of both their own genes (a direct genetic effect) and maternal phenotypes such as 56 litter size and provisioning (an indirect genetic effect) (Kirkpatrick and Lande 1989, 57 Mousseau and Fox 1998, McAdam et al. 2002). When indirect genetic effects are reciprocal, feedback effects may inflate the genetic variance available for response to selection, 58 drastically accelerating the rate of evolution (Moore et al. 1997). Although most models of 59 indirect genetic effects do not extend beyond species boundaries, indirect genetic effects may 60 61 also be common in species interactions between mutualists, competitors, parasites and hosts, and predators and prev. Phenotypic plasticity in response to an interacting species is 62 63 common (Agrawal 2001), and when these influences on trait expression have a genetic basis 64 they may represent interspecific indirect genetic effects (IIGEs; Shuster et al. 2006). Although IIGEs have received some attention in the context of community genetics (Shuster 65 66 et al. 2006, Witham et al. 2020), their potential role in driving trait coevolution has been mostly unexplored. 67

To date, most explorations of IIGEs have been studies providing empirical support for their likely existence and their contribution to trait variation. Examples of interspecific phenotypic manipulation are common in nature (Table 1), and many of these cases can be argued to be putative cases of IIGEs. One possible example occurs in arbuscular mycorrhizae, where the genotypes of fungal mutualists can alter root traits in the plants they

Interaction	IIGE synonym	Diffuse/individual interaction	Example References
host- parasite	host manipulation	both	(Thomas et al. 2012)
predator- prey	trait-mediated indirect interaction	both	(Peacor and Werner 2001)
plant- animal	inducible direct defense host-plant manipulations	both	(Weis and Abrahamson 1986, Chen 2008)
plant- microbe	joint trait microbially mediated trait	diffuse	(Friesen et al. 2011, O'Brien et al. 2021)
host- microbiome	host control host-microbiome interaction	diffuse	(Stappenbeck and Virgin 2016, Foster et al. 2017)
consumer- resource	toxin sequestration	both	(Züst et al. 2018)

Table I. Coevolutionary inte	eractions where	e IIGEs may	v be prevalent.
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73 inhabit (Gianinazzi-Pearson et al. 2007). In host-parasite systems, parasite manipulation of 74 host traits (such as behavior) and reciprocal host manipulation of parasite traits (such as growth rate, via immune response) are key features of species interactions (Thomas et al. 75 76 2012). Importantly, in many host-parasite systems, both the host and the parasite experience sustained interactions with a small number of individuals of the other species, often over key 77 periods of the life history. For example, helminth parasites excrete a variety of 78 79 immunomodulatory products that suppress or misdirect the immune system of their 80 individual host (Damian 1997, Schmid-Hempel 2008, Oladiran and Belosevic 2012). Thus, host immune response to infection is controlled by the genotype of both the host and parasite 81 82 (e.g., Barribeau 2014). As a specific example, some populations of threespine stickleback (Gasterosteus aculeatus) initiate a strong immune response to infection by the cestode 83 84 Schistocephalus solidus (Fig. 1A) involving granulocyte proliferation and fibrosis, which

effectively suppress cestode growth and viability (Weber et al. 2017). Other stickleback
populations do not exhibit this response and allow rapid cestode growth, perhaps representing
a tolerance strategy. Cestode growth is thus an indirect genetic effect of its host's genotype.
Conversely, the cestode has been shown to secrete compounds that suppress this host
response (Scharsack et al. 2004, Scharsack et al. 2007, Scharsack et al. 2013) and downregulate sticklebacks' pro-fibrotic gene expression (Fuess et al. 2020), suggesting reciprocal
indirect genetic effects.

92 In goldenrod (Solidago; Fig. 1B), size of galls produced by the gall fly Eurosta are 93 determined by genotypes of both fly and plant, and evolution of gall size is influenced in part 94 by cross-species selection imposed on *Eurosta* larvae by species at other trophic levels (Weis and Abrahamson 1986, Weis et al. 1992, Abrahamson and Weis 1997). This type of 95 interaction is common across gall-forming insects and their plant hosts; in *Hormaphis* aphids, 96 97 variation the *bicycle* gene has been linked to variation in gall size (Korgaonkar et al. 2021). 98 In many herbivore-plant interactions, physical damage to leaves induces upregulation of defensive compounds to deter further herbivory, which can be countered by matching 99 100 physiological changes in the herbivore (Ohgushi 2005). For example, Littorina snail 101 herbivory changes foliar chemistry of the brown seaweed Ascophyllum nodosa (increased phlorotannin concentrations), which in turn reduces snail movement and consumption rates 102 103 (Borell et al. 2004).

Host-parasite, host-parasitoid, and some plant-herbivore interactions can entail
intimate long-term associations between individuals. In contrast, predator-prey interactions
tend to be more diffuse (Brodie and Brodie 1999). Prey sense predation risk through
chemical, auditory, or visual cues and change their morphology, physiology, or behavior in



Figure 1. Examples of traits potentially mediated by IIGEs. IIGEs are a common feature of species interactions, particularly in host-parasite systems where prolonged contact between single individuals of both species determines fitness and trait expression for both. Here we highlight four examples where such IIGEs are likely contributing to trait expression and possibly patterns of among-population (co)variation. Panel A shows threespine stickleback fish (Gasterosteus aculeatus) and their cestode parasite Schistocephalus. Freshwater stickleback are the intermediate host for Schistocephalus, and each cestode acquires most of its lifetime resource pool while living inside a single host fish. Schistocephalus impose substantial reproductive and survival costs on hosts, and hosts have evolved an inducible (by the cestode) defense that suppresses cestode growth. Substantial among population-variation in this trait and in infection rates suggest these IIGEs may be mediating coevolution. Panel B shows goldenrod (Solidago) and a gall induced by the larvae (visible inside the gall) of the specialist gall-forming fly Eurosta. Gall size is induced by the genotype of the Eurosta larvae, and past work has shown that gall expression is a complex interaction between plant and fly genotype. Panel C shows the fungus *Ophiocordyceps unilateralis*, which manipulates behavior of its ant host prior to emergence of its fruiting body (Anderson et al. 2009). Panel D shows the water flea Daphnia lumholtzi, which induces growth of protective spines in response to chemical cues released by predatory fish (Agrawal 2001). Panel A and B main photos: S. De Lisle. Inset (B) photo: SriMesh / CC BY-SA (https://creativecommons.org/licenses/by-sa/3.0); C photo: "File:Ophiocordyceps unilateralis.png" by David P. Hughes, Maj-Britt Pontoppidan / CC BY 2.5; Panel D: "Water

flea (Daphnia lumholtzi)" by Frupus / CC BY-NC 2.0.

ways that mitigate their risk of predator encounter (Werner and Peacor 2003, Preisser et al. 108 109 2005). Perhaps the most prominent example is the tendency of some Daphnia genotypes to 110 grow spines (Fig. 1D), or to migrate to other depths, when they detect scent cues 111 (kairomones) from predatory fish (Weber and Declerck 1997, Boersma et al. 1998). These 112 antipredator responses lead to systemic changes in gene expression and morphology (Tams et 113 al. 2019), which are controlled in part by fish traits (e.g., production of a scent cue). These 114 scents can themselves be variable and genetic, as illustrated by differences in how Daphnia 115 respond to cues from landlocked versus anadromous alewife (Walsh and Post 2011). The 116 Daphnia traits are thus IIGEs controlled by both Daphnia and fish genotypes. Unlike the intimate host-parasite interactions, however, prey may be responding to diffuse cues from a 117 predator population as a whole. 118

Although IIGEs would appear to play a central role in trait interactions between many 119 coevolving species (broad categories summarized in Table 1), little is known about how these 120 121 effects influence the dynamics of coevolution (Scheiner et al. 2015). Past approaches, which have included variance-partitioning models of community assembly (Shuster et al. 2006, 122 Whitham et al. 2020) and models of "joint traits" expressed together by interacting species 123 124 (Queller 2014, O'Brien et al. 2021), are suggestive of an important role for IIGEs in species 125 interactions. However, we lack a general understanding of how reciprocal IIGEs may affect the coevolution of interacting phenotypes. The fact that indirect genetic effects within a 126 species can create feedback loops and other complex evolutionary dynamics, including in the 127 context of within-species coevolution (Drown and Wade 2014), suggests that IIGEs could 128 129 play a major role in mediating trait coevolution between species. Coopting concepts from intraspecific social evolution theory, where trait-based IGEs are well developed, thus 130 131 provides a natural way to understand the trait interactions that drive coevolution. 132 Importantly, this trait-based approach allows the contribution of indirect genetic effects to the

coevolutionary process to be fully explored. Our goal in this paper is to develop such models
to provide a comprehensive theoretical assessment of how IIGEs contribute to coevolution
between species.

Here we adapt the trait-based theory of intraspecific social evolution (Moore et al. 136 1997, Wolf et al. 1999, McGlothlin et al. 2010) to the case of two interacting species. Our 137 138 model applies to both pairwise interactions between individuals (e.g., stickleback and 139 cestodes) as well as diffuse interactions between species mean values (e.g., predatory fish and 140 Daphnia). Importantly, our model accommodates both IIGEs, where genes in individuals of 141 one species influence trait expression in individuals of another species, and cross-species 142 selection, where phenotypes of individuals in one species influence fitness of individuals of another species (Figs. 1, 2; analogous to social selection in intraspecific social evolution 143 models). In addition to describing the contribution of these interspecific interactions to 144 evolutionary change, we develop expressions for the among-population evolutionary 145 146 covariance between traits of two interacting species; that is, the expected covariance in trait means across populations between two coevolving species. In the process, we formalize an 147 148 interspecific analog of Zeng's (1988) quantitative genetic model of among-population trait 149 covariation, which we expand to incorporate interacting phenotypes. Our analysis shows that 150 IIGEs may have a central role in driving and mediating coevolution.

151

152 Reciprocal evolutionary change in interacting species

To model coevolution in two interacting species, we first decompose trait expression into three components: direct genetic effects, environmental effects, and indirect effects mediated by the phenotype of an interacting species. The phenotypic interface of coevolution involves traits with interacting effects across individuals of two species *x* and *y* and can be written as

$$\mathbf{z}_x = \mathbf{a}_x + \mathbf{e}_x + \mathbf{\Phi}_{xy} \mathbf{z}'_y \tag{1a}$$

$$\mathbf{z}_{y} = \mathbf{a}_{y} + \mathbf{e}_{y} + \mathbf{\Phi}_{yx}\mathbf{z}_{x,}^{\prime}$$
(1b)

158

where \mathbf{z}_i is a column vector of *m* traits expressed in an individual of species *i* (*m* is not 159 160 necessarily equal in each species), \mathbf{a}_i is the corresponding column vector of direct genetic 161 effects, and \mathbf{e}_i is an uncorrelated vector of residual environmental effects. The matrix $\mathbf{\Phi}_{xy}$ quantifies the effect of traits in an interacting individual of species $y(\mathbf{z}'_{v})$ on the expression of 162 traits in a focal individual of species x (Fig. 2, Table 2), while the matrix Φ_{vx} quantitifies 163 164 such effects in the opposite direction. Thus, Φ_{ii} is an interspecific analog of the matrix of conspecific indirect genetic effects, Ψ (Moore et al. 1997). The individual elements of Φ_{ij} , 165 which we write as ϕ_{ij}^{kl} , represent partial regression coefficients of trait l in species j on trait k 166 in species *i*. When traits are standardized to the same scale, these coefficients will typically 167 be limited to a range of -1 to 1. Although we focus on interactions between pairs of 168 169 individuals for simplicity of notation, this model could be easily extended to describe interactions with multiple individuals (cf. McGlothlin et al. 2010). For more diffuse 170 interactions (e.g., allwife and *Daphnia*), the elements of Φ_{ij} could represent a weighted 171 average of effects from integrating across the phenotype distribution of the interacting 172 173 population.



Figure 2. Path model of selection when traits and fitness are determined by interactions with heterospecifics. Individual trait values of two interacting species, x (blue) and y (red), are represented by z. Traits directly influence individual relative fitness (w) of the species that express them via natural selection (β_{xx} and β_{yy} , black arrows). Traits can also reciprocally influence expression of traits of heterospecific social partners, via interspecific indirect genetic effects Φ . Traits can also influence fitness of heterospecific social partners, via cross-species selection (β_{xy} and β_{yx} , grey arrows). Social effects are illustrated in grey arrows, direct effects in black.

Table 2. Definition of key parameters and expressions.

Expression / parameter	Biological definition
$\mathbf{Z}_{\chi}, \mathbf{Z}_{\mathcal{Y}}$	Vectors of individual trait values for species <i>x</i> and <i>y</i> , with a prime when present denoting partner traits.
$\mathbf{a}_x, \mathbf{a}_y$	Vectors of individual breeding values for species <i>x</i> and <i>y</i> .
$\mathbf{\Phi}_{xy},\mathbf{\Phi}_{yx}$	Interspecific indirect effects (in matrix form). The effect of traits in individuals of species <i>y</i> on trait expression in species <i>x</i> , and vice-versa.
$\boldsymbol{\beta}_{xx}, \boldsymbol{\beta}_{yy}$	Vectors of within-species natural selection gradients for each species.
$\boldsymbol{\beta}_{xy}, \boldsymbol{\beta}_{yx}$	Vectors of cross-species selection gradients for each species. The effects of traits in species y on fitness of individual species x, and vice- versa.
$\mathbf{G}_{xx},\mathbf{G}_{yy}$	Genetic covariance matrices of traits for species <i>x</i> and <i>y</i> .
\mathbf{G}_{xy} , \mathbf{G}_{yx}	Covariance matrices for breeding values of interacting individuals of species <i>x</i> and <i>y</i> within a population.
$\operatorname{Cov}(\Delta \bar{z}_x, \Delta \bar{z}_y)$	Coevolutionary covariance, defined as the covariance in evolutionary change between interacting species, across populations. When standardized by the variances in evolutionary response for each species, becomes a scale-free correlation.
$\operatorname{Cov}(\beta_{xx},\beta_{yy}),$	Between-species, among-population covariance in selection gradients (shown here for within- species selection)

Assuming that the mean residual environmental effect is zero, the population meanphenotype vector for each species is

176

$$\bar{\mathbf{z}}_x = \bar{\mathbf{a}}_x + \Phi_{xy} \bar{\mathbf{z}}_y \tag{2a}$$

$$\bar{\mathbf{z}}_y = \bar{\mathbf{a}}_y + \mathbf{\Phi}_{yx} \bar{\mathbf{z}}_x. \tag{2b}$$

177

178 Noting the change in the mean additive genetic value in each species is $\Delta \bar{\mathbf{a}} = \text{Cov}(\mathbf{a}, w)$

179 (Robertson 1966, Price 1970, 1972) where w is relative fitness, we can now define

180 evolutionary change in the multivariate mean phenotype as

181

$$\Delta \bar{\mathbf{z}}_{x} = \operatorname{Cov}(\mathbf{a}_{x}, w_{x}) + \mathbf{\Phi}_{xy} \Delta \bar{\mathbf{z}}_{y}$$
(3a)

$$\Delta \bar{\mathbf{z}}_y = \operatorname{Cov}(\mathbf{a}_y, w_y) + \mathbf{\Phi}_{yx} \Delta \bar{\mathbf{z}}_x.$$
(3b)

182

In equation (3), the first term on each right-hand side describes direct change due to natural selection on the focal species, and the second term describes indirect change due to the product of IIGEs and phenotypic change in the interacting species. It can be seen from equation (3) that species will coevolve whenever there are IIGEs in both species, because the change in phenotypic mean in species x depends upon the change of mean in species y and vice versa whenever both $\Phi_{ij} \neq 0$. Expanding equation (3) by substitution results in explicit equations for evolutionary change in each species,

190

$$\Delta \bar{\mathbf{z}}_{x} = \left(\mathbf{I}_{x} - \mathbf{\Phi}_{xy}\mathbf{\Phi}_{yx}\right)^{-1} [\operatorname{Cov}(\mathbf{a}_{x}, w_{x}) + \mathbf{\Phi}_{xy}\operatorname{Cov}(\mathbf{a}_{y}, w_{y})]$$
(4a)

$$\Delta \bar{\mathbf{z}}_{y} = (\mathbf{I}_{y} - \mathbf{\Phi}_{yx} \mathbf{\Phi}_{xy})^{-1} [\operatorname{Cov}(\mathbf{a}_{y}, w_{y}) + \mathbf{\Phi}_{yx} \operatorname{Cov}(\mathbf{a}_{x}, w_{x})].$$
(4b)

where I_x and I_y are identity matrices with dimensionality equal to the number of traits in the 192 two species and the multiplier $(\mathbf{I} - \mathbf{\Phi}_{ij}\mathbf{\Phi}_{ji})^{-1}$ quantifies the feedback effect of reciprocal 193 IIGEs. Whenever IIGEs occur in both species such that $\Phi_{ij} \Phi_{ji} \neq \mathbf{0}$, this multiplier alters the 194 total amount of evolutionary change in both species. In order for such an effect to arise, there 195 196 must be a feedback loop in phenotypic expression. The simplest of these arises when there 197 are two traits with reciprocal IIGEs, such that trait k in species i affects trait l in species j and trait *l* in turn influences the expression of trait *k*. In general, when ϕ_{ij}^{kl} and ϕ_{ji}^{lk} are of the 198 same sign, the magnitude of evolutionary change will be enhanced, and when they are of 199 opposite signs, evolutionary change will be diminished (Fig. 3). 200

201 To expand equation (4), we define the fitness of each interacting species using the202 linear equations

203

$$w_x = \alpha_x + \mathbf{z}_x^{\mathrm{T}} \boldsymbol{\beta}_{xx} + \mathbf{z}'_y^{\mathrm{T}} \boldsymbol{\beta}_{xy} + \varepsilon_x$$
(5a)

$$w_{y} = \alpha_{y} + \mathbf{z}_{y}^{\mathrm{T}} \boldsymbol{\beta}_{yy} + \mathbf{z}_{x}^{\prime \mathrm{T}} \boldsymbol{\beta}_{yx} + \varepsilon_{y}$$
(5b)

204

where w is individual relative fitness, α is an intercept, ε is an error term, and the superscript 205 T denotes transposition. Directional selection in each species is partitioned into two selection 206 gradients. First, the within-species selection gradients (β_{xx} and β_{yy}) describe the direct 207 effects of an individual's traits on its own fitness (Fig. 2; Lande and Arnold 1983). The 208 cross-species selection gradients (β_{xy} and β_{yx}) relate the fitness of a focal individual of one 209 species (x or y) to the traits of individuals of the other coevolving species (y or x) (Fig. 2). 210 211 Both of these are linear components of selection; we consider nonlinear terms (interactions for fitness between focal and partner traits) below (see Incorporating Specific Fitness 212 213 Models). The cross-species selection gradient is analogous to the directional social selection

gradient in within-species models (Wolf et al. 1999) and if pairs or groups of interacting

individuals can be identified in a natural population, can be estimated in a similar way (cf.

216 Ridenhour 2005).

We now substitute our definitions of fitness into equation (4) and expand, yieldingpredictive equations for evolutionary change in each species:

219

$$\Delta \bar{\mathbf{z}}_{x} = \left(\mathbf{I}_{x} - \boldsymbol{\Phi}_{xy} \boldsymbol{\Phi}_{yx}\right)^{-1} [\operatorname{Cov}(\mathbf{a}_{x}, \mathbf{z}_{x}^{\mathrm{T}}) \boldsymbol{\beta}_{xx} + \operatorname{Cov}(\mathbf{a}_{x}, \mathbf{z}_{y}^{'\mathrm{T}}) \boldsymbol{\beta}_{xy} + \boldsymbol{\Phi}_{xy} \operatorname{Cov}(\mathbf{a}_{y}, \mathbf{z}_{y}^{\mathrm{T}}) \boldsymbol{\beta}_{yy} + \boldsymbol{\Phi}_{xy} \operatorname{Cov}(\mathbf{a}_{y}, \mathbf{z}_{x}^{'\mathrm{T}}) \boldsymbol{\beta}_{yx}]$$
(6a)

220

$$\Delta \bar{\mathbf{z}}_{y} = \left(\mathbf{I}_{y} - \boldsymbol{\Phi}_{yx} \boldsymbol{\Phi}_{xy}\right)^{-1} [\operatorname{Cov}(\mathbf{a}_{y}, \mathbf{z}_{y}^{\mathrm{T}}) \boldsymbol{\beta}_{yy} + \operatorname{Cov}(\mathbf{a}_{x}, \mathbf{z}_{x}^{\mathrm{T}}) \boldsymbol{\beta}_{yx} + \mathbf{\Phi}_{yx} \operatorname{Cov}(\mathbf{a}_{x}, \mathbf{z}_{x}'^{\mathrm{T}}) \boldsymbol{\beta}_{xy}].$$

$$(6b)$$

221

Each of these equations consists of four terms representing four components of the total response to selection. The first term $(Cov(\mathbf{a}_x, \mathbf{z}_x^T)\boldsymbol{\beta}_{xx})$ represents response to within-species selection $(\boldsymbol{\beta}_{xx} \text{ or } \boldsymbol{\beta}_{yy})$, the second $(Cov(\mathbf{a}_x, \mathbf{z}_x'^T)\boldsymbol{\beta}_{xy})$ represents response to cross-species linear selection $(\boldsymbol{\beta}_{xy} \text{ or } \boldsymbol{\beta}_{yx})$, and the last two terms $(\boldsymbol{\Phi}_{xy}Cov(\mathbf{a}_y, \mathbf{z}_y^T)\boldsymbol{\beta}_{yy} +$

226 Φ_{xy} Cov $(\mathbf{a}_y, \mathbf{z}'_x^{T}) \boldsymbol{\beta}_{yx}$) represent the component of change caused by the change in mean of 227 the interacting species. Note that the third and fourth terms of the first part of equation (6a) 228 equal the first two terms of (6b) multiplied by the IIGE coefficient Φ_{xy} . Equation (6) also 229 shows that the change in response to within-species selection depends on the covariance of 230 additive genetic value with the phenotype of the same species, while change in response to 231 cross-species selection depends on the covariance of additive genetic value with the 232 phenotype of the opposite species.

233 To determine the components that give rise to these covariances, we expand our

234 definition of the phenotypes in equation (2) and rearrange, yielding

235

236
$$\Delta \overline{\mathbf{z}}_{x} = \left(\mathbf{I}_{x} - \mathbf{\Phi}_{xy}\mathbf{\Phi}_{yx}\right)^{-1} \left[\left(\mathbf{G}_{xx} + \mathbf{G}_{xy}\mathbf{\Phi}_{xy}^{\mathrm{T}}\right)\left(\mathbf{I}_{x} - \mathbf{\Phi}_{yx}^{\mathrm{T}}\mathbf{\Phi}_{xy}^{\mathrm{T}}\right)^{-1}\mathbf{\beta}_{xx} + (7a)$$

237
$$(\mathbf{G}_{xy} + \mathbf{G}_{xx} \mathbf{\Phi}_{yx}^{\mathrm{T}}) (\mathbf{I}_{y} - \mathbf{\Phi}_{xy}^{\mathrm{T}} \mathbf{\Phi}_{yx}^{\mathrm{T}})^{-1} \mathbf{\beta}_{xy} +$$

238
$$\Phi_{xy} (\mathbf{G}_{yy} + \mathbf{G}_{yx} \Phi_{yx}^{\mathrm{T}}) (\mathbf{I}_{y} - \Phi_{xy}^{\mathrm{T}} \Phi_{yx}^{\mathrm{T}})^{-1} \boldsymbol{\beta}_{yy} +$$

239
$$\Phi_{xy} (\mathbf{G}_{yx} + \mathbf{G}_{yy} \Phi_{xy}^{\mathrm{T}}) (\mathbf{I}_{x} - \Phi_{yx}^{\mathrm{T}} \Phi_{xy}^{\mathrm{T}})^{-1} \boldsymbol{\beta}_{yx}]$$

240

241
$$\Delta \bar{\mathbf{z}}_{y} = \left(\mathbf{I}_{y} - \boldsymbol{\Phi}_{yx} \boldsymbol{\Phi}_{xy}\right)^{-1} \left[\left(\mathbf{G}_{yy} + \mathbf{G}_{yx} \boldsymbol{\Phi}_{yx}^{\mathrm{T}}\right)\left(\mathbf{I}_{y} - \boldsymbol{\Phi}_{xy}^{\mathrm{T}} \boldsymbol{\Phi}_{yx}^{\mathrm{T}}\right)^{-1} \boldsymbol{\beta}_{yy} + (7b)$$

242
$$(\mathbf{G}_{yx} + \mathbf{G}_{yy} \mathbf{\Phi}_{xy}^{\mathrm{T}}) (\mathbf{I}_{x} - \mathbf{\Phi}_{yx}^{\mathrm{T}} \mathbf{\Phi}_{xy}^{\mathrm{T}})^{-1} \mathbf{\beta}_{yx} +$$

243
$$\Phi_{yx} (\mathbf{G}_{xx} + \mathbf{G}_{xy} \Phi_{xy}^{\mathrm{T}}) (\mathbf{I}_{x} - \Phi_{yx}^{\mathrm{T}} \Phi_{xy}^{\mathrm{T}})^{-1} \boldsymbol{\beta}_{xx} +$$

244
$$\Phi_{yx}(\mathbf{G}_{xy} + \mathbf{G}_{xx}\Phi_{yx}^{\mathrm{T}})(\mathbf{I}_{y} - \Phi_{xy}^{\mathrm{T}}\Phi_{yx}^{\mathrm{T}})^{-1}\boldsymbol{\beta}_{xy}],$$

245

where \mathbf{G}_{xx} and \mathbf{G}_{yy} represent within-species genetic (co)variance matrices and \mathbf{G}_{xy} and \mathbf{G}_{yx} 246 represent cross-species genetic covariance. Note that in a model of pairwise interactions, 247 $\mathbf{G}_{yx} = \mathbf{G}_{xy}^{\mathrm{T}}$. The cross-species genetic covariance represents a covariance of breeding values 248 249 between the species within the population of interest. This covariance may arise via any mechanism that leads to nonrandom genetic assortment between species x and y, including 250 251 behavioral preference for certain traits in heterospecific partners, fine-scale population 252 structure, and habitat preference. For example, a nonzero G_{xy} could arise if a phytophagous 253 insect showed genetic variance in host-plant choice based on genotypes of the plant. We elaborate on the contribution of G_{xy} and Φ to phenotypic assortment between individuals of 254 interacting species (C_{xy}) in equation (A1). 255

Response to within-species selection depends on the sum of within-species genetic 256 variance, \mathbf{G}_{yy} or \mathbf{G}_{xx} , and $\mathbf{G}_{xy} \mathbf{\Phi}_{xy}^{\mathrm{T}}$ or $\mathbf{G}_{yx} \mathbf{\Phi}_{yx}^{\mathrm{T}}$, which represent an interaction between IIGEs 257 and cross-species genetic covariance. Response to cross-species selection depends on the 258 sum of cross-species genetic variance and $\mathbf{G}_{xx} \mathbf{\Phi}_{yx}^{\mathrm{T}}$ or $\mathbf{G}_{yy} \mathbf{\Phi}_{xy}^{\mathrm{T}}$, which describes the genetic 259 variance created by IIGEs. Each term in equation (7) also contains an additional feedback 260 multiplier, $(\mathbf{I}_{v} - \mathbf{\Phi}_{xv}^{T} \mathbf{\Phi}_{vx}^{T})^{-1}$, which further enhances the response to selection when there 261 are reciprocal IIGEs. Because of this additional multiplier, when IIGEs are reciprocal and of 262 the same sign, their effects on can be massive, mirroring the effects observed for within-263 species IGEs (Moore et al. 1997, McGlothlin et al. 2010). As in previous equations, the last 264 two terms in equation (7) represent a sort of evolutionary feedback that occurs across 265 generations and is only present when there are IIGEs. These effects of Φ and cross-species 266 267 selection on evolutionary response are illustrated in Fig. 3.



Figure 3. Reciprocal IIGEs and cross-species selection change evolutionary response in a single species. Panels show the separate effects on evolutionary response in species x of indirect genetic effects (Panel A) and cross-species selection with genetic assortment (Panel B). Panel A shows the effects of reciprocal IIGEs holding all other evolutionary parameters constant, and assuming no cross-species selection. Panel B shows the effects of cross-species selection imposed by species y on species x, in combination with the genetic assortment between interactants of the different species, and assuming IIGEs are absent. In both panels, $G_x = G_y = 1$, $\beta_{yy} = \beta_{xx} = 1$.

268 Coevolution across a selection mosaic: the coevolutionary covariance

269 When response to selection in one species depends on trait evolution in another species, the 270 species can coevolve. When selection varies through space or time due to varying ecological 271 conditions, so too does the position of each population of each species in trait space. For 272 coevolving lineages imposing reciprocal selection on each other, we expect correlated 273 evolution of interacting traits both through time and across replicate populations of both 274 species in space. In this case, local ecological drivers of natural selection may generate a 275 mosaic of varying selection across populations, mediating where each population falls on an 276 axis of correlated evolution and generating a "geographic mosaic" of coevolution (Thompson 1994, 2005). In a coevolutionary arms race, for example, this is expected to yield 277 278 populations where each interacting lineage has elevated trait values ("hot spots"), as well 279 other populations where the race appears deescalated ("cold spots") (e.g., elevated toxin levels and toxin resistance across prey and their predators; Brodie et al. 2002, Hanifin et al. 280 281 2008), and still other populations where one lineage appears to have the upper hand (Arnqvist and Rowe 2002, Zangerl and Barenbaum 2003, Toju and Sota 2005, Perry and Rowe 2011, 282 283 Hague et al. 2020). Thus, a key feature of coevolving lineages is correlated evolution across 284 populations subject to varying ecological conditions, creating a geographic mosaic of 285 (co)varying mean phenotypes. Here, we seek to understand how the equations of selection response can be used to understand how selection and IIGEs contribute to this shared among-286 population divergence. 287

We can explore the contribution of interspecific social effects to correlated evolution between interacting species by solving for the covariance in evolutionary change between interacting species, $Cov(\Delta \bar{z}_x, \Delta \bar{z}_y)$, which we call the coevolutionary covariance. Such a covariance represents the expected pattern of coevolution through time in a single pair of populations of two species experiencing fluctuating selection, or perhaps more importantly,

across a set of populations in space under varying selection pressures. Although
mathematically equivalent, we focus on the latter scenario for its relevance to understanding
geographic variation among populations.

The coevolutionary covariance reflects the degree to which divergence among 296 population means in two species has occurred jointly. High absolute values of the 297 298 coevolutionary covariance indicate tightly-coupled coevolutionary change between the two species, whereas values around zero indicate that evolutionary change occurs independently. 299 When scaled by the total amount of population divergence (variances in $\Delta \bar{z}_x$ and $\Delta \bar{z}_y$), this 300 becomes a scale-free correlation describing the proportion of total divergence shared between 301 302 the two species. Importantly, we expect the coevolutionary covariance to be related to the covariance in population means, $Cov(\bar{z}_x, \bar{z}_y) \propto Cov(\Delta \bar{z}_x, \Delta \bar{z}_y)$. Thus, expanding 303 $Cov(\Delta \bar{z}_x, \Delta \bar{z}_y)$ allows the possibility to assess how selection and IIGEs contribute to patterns 304 of geographic variation in species mean phenotypes. 305

306 To simplify our analysis, we use models of a single trait in each species. Developing a full equation including all sources of covariance quickly becomes cumbersome, so we focus 307 on three instructive special cases that illustrate the explicit impacts of considering cross-308 species selection and interspecific interacting phenotypes. The simplest case occurs when 309 there are no IIGEs ($\phi_{xy} = \phi_{yx} = 0$) and heterospecific interactions occur at random within 310 populations ($G_{xy} = G_{yx} = 0$). If we make the simplifying assumption that genetic variance 311 does not differ among populations, the only source of covariance in selection response is 312 covariance in direct within-species selection, or 313

314

$$\operatorname{Cov}(\Delta \bar{z}_{x}, \Delta \bar{z}_{y}) = G_{xx}G_{yy}\operatorname{Cov}(\beta_{xx}, \beta_{yy}).$$
(8)

Equation (8) represents what is usually thought of as the source of a coevolutionary mosaic: 316 317 predictable variation in the type of selection occurring across space. Empirical data suggest 318 that selection varies substantially in magnitude across space (Siepielski et al. 2013) and that spatially autocorrelated biotic selection plays a substantial role in driving divergence in trait 319 320 means (Urban et al. 2011). Importantly, not just any variance in selection will do. In order to 321 create covariance in evolutionary change, selection must covary between the interacting 322 species. Cross-species selection, however, does not play a role in equation (8) because it 323 does not contribute to an evolutionary response to selection in the absence of IIGEs and 324 cross-species genetic covariance. Equation (8) is analogous to the results of Zeng's (1988) 325 model of correlated trait evolution under directional selection, although for the case of traits expressed in different species. 326

327 To see an effect of cross-species selection, we first add cross-species genetic 328 covariance but no IIGEs. For a one-trait (per species) model, $G_{xy} = G_{yx}$. Adding this effect 329 leads to three new sources of covariance among populations:

330

$$Cov(\Delta \bar{z}_{x}, \Delta \bar{z}_{y}) = G_{xx}G_{yy}Cov(\beta_{xx}, \beta_{yy}) + G_{xy}^{2}Cov(\beta_{yx}, \beta_{xy}) + G_{xx}G_{xy}Cov(\beta_{xx}, \beta_{yx}) + G_{yy}G_{xy}Cov(\beta_{yy}, \beta_{xy}).$$
(9)

331

The second term in equation (9) $(G_{xy}^2 \text{Cov}(\beta_{yx}, \beta_{xy}))$ shows that cross-species selection will contribute to the coevolutionary covariance whenever there is cross-species genetic covariance. Such covariances occur if, for example, parasite and host genotypes associate nonrandomly and have traits that influence their partner's fitness. The third and fourth terms $(G_{xx}G_{xy}\text{Cov}(\beta_{xx},\beta_{yx}) + (G_{yy}G_{xy}\text{Cov}(\beta_{yy},\beta_{xy}))$ represent a relationship between the effect of species on itself and on its heterospecific partner. These will be nonzero if populations with strong within-species selection also exhibit strong cross-species selection.

The most interesting effects on the coevolutionary covariance occur when we add IIGEs (ϕ_{xy} and ϕ_{yx}). With $G_{xy} = 0$, the covariance among populations in evolutionary response to selection is

342

$$Cov(\Delta \bar{z}_{x}, \Delta \bar{z}_{y}) = \mathcal{U}[G_{xx}G_{yy}Cov(\beta_{xx}, \beta_{yy}) + G_{xx}^{2}\phi_{yx}Var(\beta_{xx}) + G_{yy}^{2}\phi_{xy}Var(\beta_{yy}) + G_{xx}G_{yy}\phi_{xy}\phi_{yx}Cov(\beta_{xx}, \beta_{yy}) + G_{xx}G_{yy}\phi_{xy}\phi_{yx}Cov(\beta_{xy}, \beta_{yx}) + G_{xx}G_{yy}\phi_{xy}\phi_{yx}Cov(\beta_{xy}, \beta_{yx}) + G_{xx}G_{yy}\phi_{yx}Cov(\beta_{yy}, \beta_{xy}) + 2G_{xx}^{2}\phi_{yx}^{2}Cov(\beta_{yy}, \beta_{yx}) + \mathcal{V}],$$
(10)

343

where feedback effects of IIGEs are described by $\mathcal{U} = (1 - \phi_{xy}\phi_{yx})^{-4}$ and \mathcal{V} collects negligible third- and fourth-order ϕ terms:

346

$$\mathcal{V} = G_{xx}^2 \phi_{yx}^3 \operatorname{Var}(\beta_{xy}) + G_{yy}^2 \phi_{xy}^3 \operatorname{Var}(\beta_{yx}) + G_{xx}G_{yy}\phi_{xy}\phi_{yx}^2 \operatorname{Cov}(\beta_{yy},\beta_{xy}) + G_{xx}G_{yy}\phi_{xy}^2 \phi_{yx} \operatorname{Cov}(\beta_{yx},\beta_{xx}) + (11)$$
$$G_{xx}G_{yy}\phi_{xy}^2 \phi_{yx}^2 \operatorname{Cov}(\beta_{yx},\beta_{xy}).$$

347

The first term in equation (10) is identical to equation (8) and represents covariance in direct within-species selection. The second $(G_{xx}^2 \phi_{yx} \operatorname{Var}(\beta_{xx}))$ and third $(G_{yy}^2 \phi_{xy} \operatorname{Var}(\beta_{yy}))$ terms show that in the presence of IIGEs, simple variance in within-species selection across populations can generate a coevolutionary covariance (or correlation, Fig. 4A; corresponding



Figure 4. IIGEs drive and mediate coevolution between interacting species. Panels show the coevolutionary correlation between two interacting species as a function of the indirect genetic effect parameters ϕ , from equation (7) standardized by the evolutionary rates. In the absence of social interaction effects, correlated evolution (coevolution) between species is driven entirely by covariance in natural selection between the species (see text). Reciprocal IIGEs can generate coevolution even when there is not covariance in natural selection (Panel A), and can even drive coevolution even when one species lacks genetic variance (Panel B). IIGEs modify observed coevolutionary patterns when natural selection does covary (Cov(β_{xx},β_{yy}) = 0.5; Panel C). When cross-species and IIGEs act together, coevolutionary patterns are a complex third order polynomial (Panel D; Cov(β_{xy},β_{yx}) = 0.5, Var(β_{xy}) = Var(β_{yx}) = 1). When natural and cross-species selection both act (Var(β_{xy}) = Var(β_{yx}) = Var(β_{yx}) = 1) and covary positively (0.5), (Panel E), effects of IIGEs become stronger in comparison to the case of (co)variance in natural selection alone. When covariance between natural and cross-species selection is negative (-0.8) a ridge is observed (Panel F).

evolutionary rates and covariances are plotted in Figs. S1 and S2, respectively). This
covariance in evolutionary response occurs as a necessary consequence of the dependence of
the trait mean of one species on that of the other. Thus, in the presence of IIGEs, a
coevolutionary covariance may occur even when selection is uncorrelated between the
species, and in extreme cases, when only one species varies in selection or when only species
has genetic variation (Fig. 4B).

359 The fourth and fifth terms $(G_{xx}G_{yy}\phi_{xy}\phi_{yx}Cov(\beta_{xx},\beta_{yy}) +$

 $(G_{xx}G_{yy}\phi_{xy}\phi_{yx}\text{Cov}(\beta_{xy},\beta_{yx}))$ show the effects of reciprocal IIGEs. When traits in the two 360 species exist in a feedback loop, the effect of covariance in within-species selection is 361 amplified if ϕ_{xy} and ϕ_{yx} are of the same sign and diminished if they are of opposite signs. 362 363 The combined effects of the second, third, and fourth terms lead to a complex relationship 364 between IIGEs and the coevolutionary covariance when within-species selection varies (Fig. 4C). In some cases, IIGEs can even reverse the sign of the correlation that would be 365 366 expected in their absence (Fig. 4C). In addition, reciprocal IIGEs may cause covariance in cross-species selection to contribute to the coevolutionary covariance, causing an even more 367 complex relationship (Fig. 4D). 368

The last four terms in equation (10) show that in the presence of IIGEs, covariance 369 370 between within-species and cross-species selection may contribute to the coevolutionary 371 covariance (Fig. 4E-F). Because IIGEs inherently tie together the evolutionary responses of 372 the two species through their effects on the coevolutionary covariance, this can occur both when there is covariance in gradients across species (terms 6 and 7) and when there is 373 374 covariance in gradients in the same species (terms 8 and 9). The total effect of IIGEs may be quite complex when cross-species selection varies, with subtle changes in ϕ resulting in 375 dramatic changes in the expected among population correlation (Fig. 4D, E-F). 376

377 An equation incorporating both IIGEs and G_{xy} quickly becomes unwieldy, but we 378 present a compact form containing 10 (co)variance terms as equation (A2), which illustrates 379 that when present, non-random assortment and IIGEs together interact in complex ways to 380 influence coevolution.

381

382 Incorporating specific fitness models

Coevolutionary models often posit complex relationships between interacting phenotypes and 383 384 fitness (Nuismer 2017). Although the selection model we present here is linear, more complex relationships can be incorporated by translating specific fitness functions to 385 386 selection gradients. The adaptive landscape represents the theoretical relationship between a population's mean fitness and its phenotypic mean (Arnold et al. 2001). Selection gradients 387 represent the partial slope of the adaptive landscape with respect to a given phenotype, and 388 389 thus the multivariate selection gradient may be calculated using a vector of partial derivatives 390 if the adaptive landscape can be written as a differentiable function (Lande 1979, Lande and Arnold 1983). In many cases, the multivariate selection gradient may be calculated using 391 partial derivatives of the individual fitness function as well (Lande and Arnold 1983, Abrams 392 et al. 1993, McGlothlin et al. 2021). Once selection gradients have been calculated for a 393 394 given model, they may be substituted into equations (6-7) to explore the effects of a given 395 fitness function on selection response and the coevolutionary covariance (see also Brodie and 396 Ridenhour 2003). This exercise also allows us to explore the effects of adding IIGEs to an 397 existing coevolutionary model.

First, we consider a "phenotypic difference," or trait-matching, model of ecological trait interaction (Nuismer et al. 2007, Nuismer 2017), where absolute fitness is a function of the difference between interacting traits, $W_x \propto \exp(z_x - z_y)$ and $W_y \propto \exp(z_y - z_x)$.

401 Using the logarithm of the fitness function to calculate relative fitness (Lande and Arnold

402 1983),

403

$$w_x = \alpha_x + b_x (z_x - z_y) + \varepsilon_x$$

$$w_y = \alpha_y + b_y (z_y - z_x) + \varepsilon_y.$$
(12)

404

where b_x and b_y are constants that may vary across populations. Differentiating, the selection gradients are then $\beta_{xx} = -\beta_{xy} = b_x$ and $\beta_{yy} = -\beta_{yx} = b_y$. Substituting these selection gradients into equation (7) is trivial. However, is worth noting that under this fitness model, the coevolutionary covariance simplifies to a function of just three (co)variance components,

$$Cov(\Delta \bar{z}_{x}, \Delta \bar{z}_{y}) = \mathcal{U}[G_{xx}G_{yy}(1 - \phi_{xy} - \phi_{yx} + 2\phi_{xy}\phi_{yx})Cov(b_{x}, b_{y}) + G_{xx}^{2}\phi_{yx}(1 - 2\phi_{yx})Var(b_{x}) + G_{yy}^{2}\phi_{xy}(1 - 2\phi_{xy})Var(b_{y}) + \mathcal{V}],$$
(13)

410

411 which corresponds to a relationship with ϕ as in Figs. 4F, S1F, and S2F. This analysis 412 illustrates that reciprocal IIGEs have greatest impact on mediating coevolutionary outcomes 413 in a trait-matching models when IIGEs are similar in both sign and magnitude. Biologically, 414 such a situation corresponds to a scenario where, for example, trait expression is reciprocally 415 escalated in response to heterospecific partners.

Another important fitness model to consider is the case of nonlinear fitness
interactions. Nonlinearity is fundamental to many coevolutionary models that invoke
epistatic interactions across species' genomes, such as the matching-allele or gene-for-gene

419 models of host-parasite coevolution (Dybdahl et al. 2014). To simplify, consider a fitness

420 model where relative fitness depends solely on the product of the two phenotypes:

421

$$w_x = \alpha_x + b_{xy} z_x z_y + \varepsilon_x$$

$$w_y = \alpha_y + b_{yx} z_y z_x + \varepsilon_y.$$
(14)

422

423 Under this fitness model, all selection gradients become dependent on the population means424 of species *x* and *y*:

425

$$\beta_{xx} = b_{xy} \bar{z}_y$$

$$\beta_{xy} = b_{xy} \bar{z}_x$$

$$\beta_{yy} = b_{yx} \bar{z}_x$$

$$\beta_{yx} = b_{yx} \bar{z}_y$$
(15)

426

427 This relationship may cause selection to (co)vary across populations even when b_{xy} and b_{yx} 428 are homogeneous. In the absence of IIGEs, this simplest case would lead to a coevolutionary 429 covariance defined by

$$\operatorname{Cov}(\Delta \bar{z}_{x}, \Delta \bar{z}_{y}) = G_{xx}G_{yy}b_{xy}b_{yx}\operatorname{Cov}(\bar{z}_{x}, \bar{z}_{y}).$$
(16)

Thus, any initial covariance in the population means leads to a covariance in the response to
selection across species, lending a runaway aspect to the coevolutionary covariance. Adding
IIGEs, this becomes

434

$$Cov(\Delta \bar{z}_{x}, \Delta \bar{z}_{y}) = \mathcal{U}[G_{xx}G_{yy}b_{xy}b_{yx}Cov(\bar{z}_{x}, \bar{z}_{y}) + G_{xx}^{2}\phi_{yx}b_{xy}^{2}Var(\bar{z}_{y}) + G_{yy}^{2}\phi_{xy}b_{yx}^{2}Var(\bar{z}_{x}) + 2G_{xx}G_{yy}\phi_{xy}\phi_{yx}b_{xy}b_{yx}Cov(\bar{z}_{x}, \bar{z}_{y})$$

$$+G_{xx}G_{yy}\phi_{xy}b_{xy}b_{yx}Var(\bar{z}_{y}) + G_{xx}yG_{y}\phi_{yx}b_{xy}b_{yx}Var(\bar{z}_{x}) + 2G_{xx}^{2}\phi_{yx}^{2}b_{xy}^{2}Cov(\bar{z}_{x}, \bar{z}_{y}) + 2G_{yy}^{2}\phi_{xy}^{2}b_{yx}^{2}Cov(\bar{z}_{x}, \bar{z}_{y}) + \mathcal{V}].$$

$$(17)$$

435

Although this equation is complex, it shows that when IIGEs are present any variance across populations in the trait mean of either species leads to a cross-species covariance in the response to selection (terms 2, 3, 5, and 6). Cross species covariance in selection response is also mediated further by covariance in trait means in this model (terms 1, 4, 7, and 8) as well as higher order products of IIGEs captured in V (equation 11). These effects may lead trait means to become correlated across populations in future generations, further intensifying the geographic mosaic (Thompson 1994).

443

450

444 Discussion

Our model adapts the theory of trait-based intraspecific social evolution to the phenotypic
interface between two coevolving species. We show that two forms of interspecific
interaction, interspecific indirect genetic effects (IIGEs) and cross-species selection
(analogous to within-species social selection), both contribute to correlated evolution
between interacting species. Our analysis shows that reciprocal IIGEs modulate selection

response, suggesting that IIGEs may play a major role in generating and mediating patterns

451 of correlated evolution between species. Further, we show that constant (across space) IIGEs 452 can generate a coevolutionary covariance (and thus a geographic mosaic) even in the absence 453 of covariance in selection, or even the absence of genetic variance in one species. When 454 selection does covary between species across populations, reciprocal IIGEs will promote 455 changes in the magnitude of coevolution and even reversals in the expected among-456 population covariance. IIGEs also allow cross-species selection, which we model as the 457 effect of the traits of one species on the fitness of another, to influence evolutionary response. 458 Such a response may also be mediated by cross-species genetic assortment between 459 interacting individuals. When IIGEs and cross-species selection act together, effects on the 460 coevolutionary mosaic can be complex, with dramatic changes in the expected sign and magnitude of correlated evolution occurring with subtle changes in these parameters of 461 interspecific social interaction. Our results indicate that whenever coevolving species 462 socially interact to modify expression of one another's phenotypes, these interspecific social 463 464 interactions are key to understanding coevolution.

IIGEs represent a scenario where phenotypes in individuals of one species influence 465 466 trait expression of individuals of another species. Thus, IIGEs are a specific type of 467 environmental effect (Moore et al. 1997, Drown and Wade 2014) where the environment is the phenotypic value of the interspecific individual(s) with which an organism interacts. To 468 our knowledge, this type of environmental effect on between-species coevolution has been 469 470 considered in only two other theoretical studies (Scheiner et al. 2015, O'Brien et al. 2021; but see Shuster et al. 2006, Witham et al. 2020 for a variance-partitioning approach). Scheiner et 471 472 al. (2015) consider a special case of our model, where IIGEs are present in only one of the two interacting species. In this non-reciprocal model, they show a much more limited role 473 for IIGEs in coevolution. Our results are broadly consistent with this conclusion, in that 474 IIGEs in only a single species do not generate the reciprocal effects that lead to massive 475

inflation of evolutionary response. However, IIGEs in only a single species (e.g., $\Phi_{xy} = 0$, 476 $\Phi_{\nu x} \neq 0$) still play a role in mediating response to interspecific social selection whenever 477 interspecific social selection is a function of individual trait values (as opposed to the 478 479 population mean, as modeled by Scheiner et al. 2015). We also note that our fully multivariate model accommodates the possibility that reciprocal IIGEs act across different 480 481 types or numbers of traits in the two interacting species. More recently, O'Brien et al. (2021; see also Queller 2014) developed a model of coevolution between host plant and microbial 482 symbionts. Their parameterization differed from ours in that they consider evolution of a 483 484 single joint trait governed by genetic variation in host and symbiont, and so is most directly applicable to plant-microbe systems or other intimate interactions. Nonetheless, their model 485 shows an important role for reciprocal fitness feedbacks, consistent with the conclusions of 486 487 our trait-based model.

Interspecific indirect genetic effects, or at least the potential for a prevalence of such 488 489 effects, appear to be commonplace in many biological systems. In Table 1, we provide in a 490 breakdown of types of biological interaction in which there is a large literature suggesting 491 importance of IIGE-like phenomena. These types of effects on trait expression across 492 species, widely appreciated in their own specific contexts (Weis and Abrahamson 1986, 493 Peacor and Werner 2001, Werner and Peacor 2003, Chen 2008, Thomas et al. 2012, O'Brien 494 et al. 2021), have taken on a variety of different forms. We suggest that these disparate 495 biological phenomena may nonetheless share a commonality-reciprocal effects on trait 496 expression across interspecific partners—that we have shown can affect the coevolutionary 497 process in dramatic, and in some cases predictable, ways.

498 Cross-species selection features prominently in verbal descriptions of the
499 coevolutionary process (Thompson 1982), and we show that such selection is especially
500 important in the presence of interspecific indirect genetic effects. When individual trait

values of one species affect individual fitness of another, focal species, this cross-species 501 502 selection can manifest evolutionary change in the focal species when there is phenotypic 503 assortment between interspecific interactants. This assortment, analogous to that required for 504 evolutionary response to social selection within species (Wolf et al. 1999, McGlothlin et al. 505 2010, Brodie et al. submitted), can be generated directly by a non-random genetic assortment, 506 or via IIGEs. Examples of processes that could generate direct genetic assortment between 507 interacting individuals of two different species include shared genetic structure. Such shared 508 genetic structure could arise through shared features of the environment that limit gene flow 509 and panmictic mating in both species, or alternatively, through variation in habitat preference 510 across individuals of both species. Direct genetic assortment could also arise through behavioral preference for certain trait values in heterospecific partners. Such preferences 511 may be especially common in predator-prey interactions, where, for example, predator body 512 513 size may be expected to coevolve with behavioral preference for prey size (Troost et al. 514 2008).

Our results also indicate that nonlinear effects on cross-species selection can 515 516 contribute to coevolution even in the absence of genetic assortment or IIGEs. This form of 517 selection corresponds to an interaction between focal and interspecific-partner trait values for 518 focal individual fitness. The effect of this form of cross-species selection on evolutionary response in a population depends on the mean genotype of the other species, and thus 519 520 represents a diffuse effect of population mean phenotype of the coevolving species. Such interspecific interactions are potentially less intimate, for example diffuse predator chemical 521 522 cues in aquatic environments, than the individual level interactions (e.g., of host and parasite) required to generate response from linear cross-species selection. Across populations, 523 nonlinear cross-species selection contributes to coevolution via covariance in mean genetic 524 525 values and/or linear selection between the species. This result is consistent with past models

of coevolution, verbal and mathematical, that indicate trait interactions for fitness are a key feature of coevolution (Thompson 1982, 1994, 2005, Nuismer 2017), and in our model, such interactions lead to a dependence between selection in one species and the mean trait value of another. By defining these interaction terms in the framework of social evolution, our model adds to past work by indicating that reciprocal IIGEs can substantially increase the degree to which trait interactions for fitness contribute to reciprocal evolutionary change.

A unique feature of our model, compared to other models of interspecific coevolution, 532 533 is our development of a formal expression for the expected covariance in evolutionary response, $Cov(\Delta \bar{z}_x, \Delta \bar{z}_y)$. This coevolutionary covariance is expected to be a key contributor 534 535 to generating among-population covariation in species mean trait values, a major focus in many studies of coevolution (Thompson 1994, 1995, Zangerl and Berenbaum 2003, Toju and 536 537 Sota 2005, Hanifin et al. 2008, Hague et al. 2020). Importantly, similar to existing withinspecies models of among-population quantitative genetic variation (Zeng 1988, Chenoweth et 538 al. 2010), defining this coevolutionary covariance illustrates how selection, IIGEs, and 539 540 genetics may contribute to patterns of trait variation across a geographic mosaic.

541 Our model subsumes mechanistic detail into broad statistical descriptions of species interactions and thus provides a general description of how IIGEs and cross-species selection, 542 543 when present, contribute to reciprocal evolutionary change and correlated evolution across 544 populations. In contrast to our approach, some models of coevolution have focused instead 545 on specific ecological mechanisms that generate trait-fitness relationships between interacting 546 species (reviewed in Nuismer 2017). By highlighting the key parameters that contribute to 547 coevolution—covariance in natural selection, covariance in cross-species selection, and 548 IIGEs—our model indicates various pathways through which specific ecological mechanisms may affect coevolution. Our framework can be tailored to specific scenarios by substituting 549 550 different fitness models into the general equations we present here.

Our model generates quantitative predictions for the shape of coevolution that are 551 552 directly testable with empirical data because it focuses on estimable statistical effects of 553 underlying ecological mechanisms rather than the mechanisms themselves, which are often unknown (Wade and Kalisz 1990). For example, using an empirical estimate of Φ (which 554 555 could be measured using methods analogous to those used to measure within-species indirect 556 genetic effects; Bleakley and Brodie 2009, McGlothlin and Brodie 2009), one could use 557 matrix comparison of covariances among population means and the covariance terms presented here to quantitatively test the contribution of IIGEs to among-population 558 559 covariance in selection response between two interacting species (e.g., see Chenoweth et al. 560 2010 for a within-species test of the predictions of Zeng's 1988 model). More generally, 561 Week and Nuismer (2019) have shown how datasets of among-population variation in trait 562 means can be used to test for conformation to expectations from coevolutionary models. Concomitantly, our models show how environmental effects can be partitioned into terms 563 564 describing genotypes of other species in the ecological community, which could be useful in 565 understanding when and why evolutionary response fails to conform to predictions arising from the standard breeder's equation. 566

Social interactions between individuals of the same species play a central role in the
evolutionary process. Within a single lineage, indirect genetic effects and social selection
fundamentally change selection response, the expression of genetic variance, and together
determine the course of social evolution (Moore et al. 1997, Wolf et al. 1998, Wolf et al.
1999, McGlothlin et al. 2010). We have shown that these effects of interactions among
individuals transcend species boundaries, and profoundly impact the dynamics of coevolution
between interacting lineages.

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Figure S1. IIGEs accelerate evolutionary rate in a single species. Panels show the evolutionary rate, $Var(\Delta \bar{z}_x)$, of species *x* as a function of the indirect genetic effect parameters ϕ , under the same parameter values as in Figure 3. Reciprocal IIGEs between interacting species generally accelerate evolutionary rate. Note that in the absence of any other effects, the evolutionary rate is equal to the variance in natural selection, which is unity in Panels A, C, E, and F. In panel D, evolutionary rate is driven entirely by cross-species selection and IIGEs. In panel B, where $G_{xx} = 0$, evolutionary rate in species *x* is driven entirely by reciprocal IIGEs and evolutionary change in species *y*.



Figure S2. Among population covariances. Panels show the coevolutionary covariance between two interacting species as a function of the indirect genetic effect parameters ϕ , from equation (7) unstandardized. Parameter values are as in Figs. 4 and S1. For all panels, genetic variances were set to unity.

781 Appendix

782 A1. Phenotypic covariance between interacting species

For a single trait, we can solve for the covariance between z_y and z_x to partition the

784 phenotypic covariance between individuals of two interacting species into terms describing

the contribution of IIGEs and terms describing non-random genetic assortment. Assuming

- 786 cross-species environmental covariance is zero,
- 787

$$C_{xy} = (1 - \phi_{yx}\phi_{xy})^{-2} [(1 + \phi_{yx}\phi_{xy})G_{xy} + \phi_{yx}(G_x + E_x) + \phi_{xy}(G_{yy} + E_x)]$$

788

where E_x and E_y represent within-species environmental variance and $G_{xy} = G_{yx}$. When IIGEs are absent, non-random genetic assortment G_{xy} is the sole contributor to the phenotypic association between individuals of coevolving species. When IIGEs are present, they can substantially change this phenotypic association.

793

794 A2. Covariance in selection response with nonzero IIGEs and genetic assortment

795 We expand the covariance in evolutionary response in two species when IIGEs are present

and constant and when genetic assortment $G_{xy} = G_{yx}$ is present and constant,

797

798
$$\operatorname{Cov}(\Delta \bar{z}_x, \Delta \bar{z}_y) = \mathcal{U}[\mathcal{A}^2 \phi_{yx} \operatorname{Var}(\beta_{xx}) + \mathcal{B}^2 \phi_{yx} \operatorname{Var}(\beta_{xy}) + \mathcal{C}^2 \phi_{xy} \operatorname{Var}(\beta_{yy})$$

799
$$+ \mathcal{D}^2 \phi_{xy} \operatorname{Var}(\beta_{yx}) + \mathcal{ABECov}(\beta_{xx}, \beta_{xy}) + \mathcal{ACECov}(\beta_{xx}, \beta_{yy})$$

800 +
$$\mathcal{ADECov}(\beta_{xx}, \beta_{yx}) + \mathcal{BCECov}(\beta_{xy}, \beta_{yy}) + \mathcal{BDECov}(\beta_{xy}, \beta_{yx})$$

801 + $\mathcal{CDECov}(\beta_{yy}, \beta_{yx})$]

803 where

- $\mathcal{U} = (1 \phi_{xy}\phi_{yx})^{-4}$
- $\mathcal{A} = G_{xx} + G_{xy}\phi_{xy}$
- $\mathcal{B} = G_{xy} + G_{xx}\phi_{yx}$
- 808 $\mathcal{C} = G_{yy} + G_{xy}\phi_{yx}$
- 809 $\mathcal{D} = G_{xy} + G_{yy}\phi_{xy}$

$$\mathcal{E} = 1 + \phi_{xy}\phi_{xy}$$