

Negative selection in humans and fruit flies involves synergistic epistasis

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Abstract

Negative selection against deleterious alleles produced by mutation is the most common form of natural selection, which strongly influences within-population variation and interspecific divergence. However, some fundamental properties of negative selection remain obscure. In particular, it is still not known whether deleterious alleles affect fitness independently, so that cumulative fitness loss depends exponentially on the number of deleterious alleles, or synergistically, so that each additional deleterious allele results in a larger decrease in relative fitness. Negative selection with synergistic epistasis must produce negative linkage disequilibrium between deleterious alleles, and therefore, underdispersed distribution of the number of deleterious alleles in the genome. Indeed, we detected underdispersion of the number of rare loss-of-function (LoF) alleles in eight independent datasets from modern human and *Drosophila melanogaster* populations. Thus, ongoing selection against deleterious alleles is characterized by synergistic epistasis, which can explain how human and fly populations persist despite very high genomic deleterious mutation rates.

Keywords: Epistasis, negative selection, evolution, sex, genetic recombination, mutation load, linkage disequilibrium

Negative selection plays a key role in evolution, preventing unlimited accumulation of deleterious mutations and establishing the mutation-selection equilibrium (1). The properties of negative selection are determined by the corresponding fitness landscape, the function which relates fitness to the "mutation burden" of a genotype. In the simplest case of equally deleterious mutations, mutational burden is the total number of mutant alleles in a genome. According to the null hypothesis of no epistasis, selection acts on different mutations independently, so that each additional mutation causes the same decline in relative fitness and fitness depends exponentially on their number. By contrast, if synergistic, or narrowing (2), epistasis between deleterious alleles is present, each additional mutation causes a larger decrement of relative fitness. Synergistic epistasis can reduce the mutation load under a given genomic rate of deleterious mutations (1, 3-4) and can produce the evolutionary advantage of sex and recombination (5). However, because neither the mutational burden nor fitness can be easily measured, data on fitness landscapes of negative selection remain inconclusive (6). Recent genome-wide investigations have found pervasive epistasis, but no consistent directionality of effect (6-8). Synergistic epistasis between deleterious mutations is more prevalent in organisms with complex genomes (7). Moreover, theoretical work suggests that narrowing epistasis may emerge as a result of pervasive pleiotropy, and modular organization of biological networks (9). This would lead to antagonism between beneficial mutations and synergism between deleterious mutations.

In this paper, we study the distribution of the mutation burden in human and *Drosophila melanogaster* populations. In the absence of epistasis, deleterious alleles independently contribute to the mutation burden (3). Thus, if mutant alleles

are rare, the mutation burden has Poisson distribution, so that its variance (σ^2) is equal to its mean (μ) (Fig. S1). More generally, the variance of the mutation burden is equal to the sum of variances at all deleterious loci, or the additive variance (V_A) (10), computed as $\sum_i 2p_i(1-p_i)$ for all deleterious loci i with mutant allele frequency p_i in the genome (Fig. 1). This is mathematically equivalent to the genome-wide nucleotide diversity of deleterious alleles (11). In contrast, epistatic selection creates dependencies between individual alleles, so total variance of the mutation burden is no longer equal to the additive variance. Selection with synergistic epistasis creates negative linkage disequilibria (LDs) between deleterious alleles. Due to the dependencies between individual alleles, variance of the mutation burden is reduced by a factor of ρ (<1), which is determined by the strength of selection and the extent of epistasis (13, 14, Fig. S2). Antagonistic epistasis, instead, creates positive linkage disequilibria (LDs) between deleterious alleles and increases variance of the mutation burden. Truncation selection, which represents the extreme mode of synergistic epistasis (4) leads to the smallest ρ . In the extreme example, if 50% of individuals with above average numbers of mutations would not contribute to the next generation, $\rho = 0.36$ if the average genomic number of mutations is high. Because free recombination halves LDs in the course of one generation, at the mutation-selection equilibrium $\sigma^2 = V_A/2 - \rho$, where V_A is the variance of the mutation burden under linkage equilibrium. More subtly, the difference between σ^2 and V_A is also a genome-wide estimate of the net linkage disequilibrium in fitness. For all pairs of loci i and j in the genome, $I = \sigma^2 - V_A = 4 \sum_{i,j} D_{i,j}$, where $D_{i,j}$ is the pair-wise linkage disequilibrium. Using data on multiple genotypes from a population, we utilized this statistical framework to create a

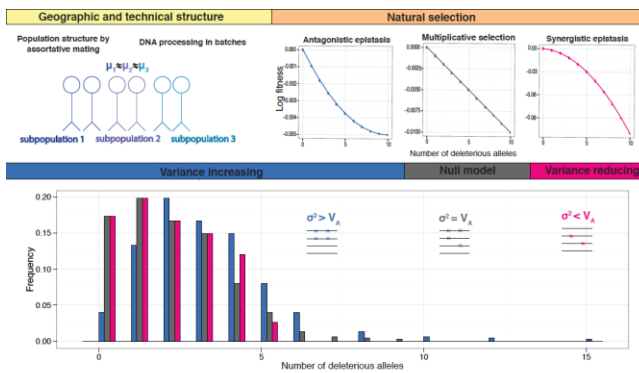


Fig. 1. Rare mutation burden under natural selection (orange, right) and population structure (yellow, left). In the absence of epistasis (grey), variance (σ^2) of the mutation burden (bottom panel) is equal to its additive variance (V_A). Overdispersion (blue) occurs due to natural selection with antagonistic epistasis, and due to population structure and technical heterogeneity during sequencing. Underdispersion (pink) occurs due to natural selection with synergistic epistasis.

test for synergistic epistasis without the need to measure fitness.

The ideal population for our test would be single ancestry, outbreeding, non-admixed and randomly mating. We analyzed three suitable European datasets – the Genome of the Netherlands (GoNL) (14), Alzheimer’s Disease Neuroimaging Initiative (ADNI), and Dutch controls from Project MinE, an amyotrophic lateral sclerosis (ALS) study. For each of these, we obtained whole-genome sequences of unrelated individuals. We obtained the same data for Zambian flies from Phase 3 of the *Drosophila* Population Genomics Project (DPGP3) (15). For each population, after applying stringent quality control filters (Tables S8-S12), we computed the mutation burden distribution for coding synonymous, missense, and loss-of-function or LoF, defined as splice site disrupting and nonsense variants. For all of these datasets, distribution of LoF singletons was underdispersed (nonsense variants in the MinE dataset, if considered separately, were the exception, although underdispersion was also observed for stop gain variants in this dataset at a slightly higher allele frequency threshold (Table S2)). On average, rare LoF variants displayed variance (σ^2) reduced by a factor of ~0.9, or underdispersion, compared to additive variance (V_A) (Table 1, Fig. 2). Thus, $\sigma^2 = 0.9V_A$ is consistent with $\rho = 0.89$ which appears, for example, after truncation of less than 2% of the population. In contrast, rare coding synonymous variants

	Net LD per pair	
	Mean	
<i>Genome of the Netherlands GoNL (495)</i>		
Coding		
synonymous	30.26	0.022
Missense	60.88	0.018
Nonsense	1.67	-0.039
Splice	0.90	-0.049
LoF	2.58	-0.029
<i>European ancestry ADNI (714)</i>		
Coding		
synonymous	38.99	0.028
Missense	77.98	0.013
Nonsense	2.10	-0.032
Splice	1.16	-0.104
LoF	3.26	-0.022
<i>Dutch ALS (601)</i>		
Coding		
synonymous	42.93	0.017
Missense	79.34	0.012
Nonsense	1.89	0.028
Splice	0.95	-0.033
LoF	2.83	-0.001
<i>Zambian DPGP3 (191)</i>		
Coding		
synonymous	3577.06	0.016
Missense	2051.52	0.008
Nonsense	10.21	-0.007
Splice	2.60	-0.020
LoF	12.81	-0.005

Table 1. Negative linkage disequilibrium (LD) between rare LoF variants in human and *D. melanogaster* genomes. For humans, only singletons are included (see Table S2 for other frequency cut-offs). For flies, alleles up to a minor allele count of 5 are included (see Table S3 for other cut-offs). The number of samples is given in parentheses for each dataset. LoF variants include splice site disrupting and nonsense variants. Net LD per pair of alleles (\hat{I}) is computed as the difference between the variance σ^2 and additive variance V_A of the rare mutation burden, normalized by the square of the mean mutation burden μ ($\hat{I} = I/\mu^2$). A p-value was obtained for each human dataset by permutation (Table S1), and a joint p-value for all 3 human datasets shown (GoNL, ADNI, ALS) was computed by meta-analysis using Stouffer’s method (coding synonymous $p=0.999$, missense $p < 1 \times 10^{-3}$, LoF $p = 0.002$).

showed σ^2 greater than V_A , or overdispersion. We replicated the same signal in three non-European populations from the

1000 genomes Phase I Project (16) (Table S1, Table S2) and an American *D. melanogaster* population from the *D. melanogaster* Genetic Reference Panel (DGRP, 17, Table S3). We proceeded to ask two questions: why were the synonymous variants overdispersed compared to their expectation under independence? Was the underdispersion in LoF variants significant?

Even for a set of independent alleles, overdispersion in the mutation burden is observed if genome-wide positive LD is present due to population structure, which can also be seen as deviations from Hardy-Weinberg equilibrium for the entire genome (Fig. 1, 18, 19). If the population has a cline in average mutation burden (μ) (20) due to, for example, a south-to-north expansion (14)(21) followed by assortative mating, this translates into an excess of σ^2 over V_A (Fig. S3, Fig. S4, Table S4). Overdispersion can also be caused by technical reasons (Fig. 1). When DNA samples are sequenced or processed in different batches, the heterogeneity introduced can result in a clustering effect similar to that of geographic structure. Using GoNL samples, for which we had detailed geographic and technical information, we showed that a large proportion of the overdispersion in rare mutation burden could be attributed to geographic origin and sequencing batch (Fig. S5, Table S4). We also showed that by calculating mutation burden across allele frequencies, where no strong differences in μ between human populations have been detected (20), we effectively observed an independent distribution of mutation burden for all variants (Table S2).

We next proceeded to investigate the contrast between coding synonymous and LoF variants. Having uncovered the primary sources for overdispersion of rare mutation burden, we realized that overdispersion scaled with the mean (μ) of the mutation burden distribution (Table S5). We, therefore, generated an empirical null distribution for each dataset by resampling coding synonymous variants at the same mean (μ) and allele frequency as our test set of LoF variants (1000 resamples for each dataset, Fig. 2). Meta-analyzing across all datasets using Stouffer's method (Tables S1-S3), we showed that the deleterious mutation burden for LoF variants was significantly underdispersed in humans ($p = 0.0003$) and flies ($p = 9.43 \times 10^{-6}$). We also tested significance in humans using an alternative approach (Table S1). Permuting functional consequences across variants, we confirmed the significance of our underdispersion signal in deleterious mutation burden (missense $p < 1 \times 10^{-3}$, LoF $p = 0.002$). Furthermore, we showed that the underdispersion signal persists after correcting

raw metrics for population structure and other confounding factors (Table S5)

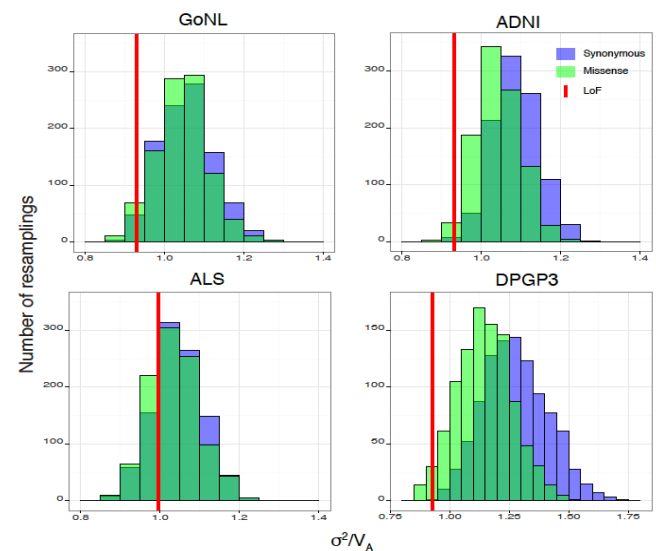


Fig. 2. Resampling distributions of σ^2 to V_A ratio (n) for LoF rare mutation burden in humans and *D. melanogaster*. Synonymous (purple) and missense (green) variants were resampled at the same mean and allele frequency as LoF variants to obtain empirical null distributions for the variance σ^2 to additive variance V_A ratio for each dataset. For humans, only singletons are included (see Table S2 for other frequency cut-offs). For flies, alleles up to a minor allele count of 5 are included (see Table S3 for other cut-offs). A synonymous p-value for the σ^2 to V_A ratio of the rare LoF mutation burden (red) was obtained for each dataset (Table S1), and a joint p-value for all 3 human datasets shown (GoNL, ADNI, ALS) was computed by meta-analysis using Stouffer's method ($p = 0.0003$).

Notably, the detected signal varies between human and fly populations. The underdispersion signal in deleterious mutation burden is stronger in flies compared to humans (Fig. 2). First, this is because recombination, which opposes the reduction in genetic variance caused by negative LDs, is weaker in flies compared to humans. The observed effect decreases with the harmonic mean c_H of the recombination frequencies among the sites involved (22). Flies, with only four pairs of chromosomes and no crossing over in males, has an estimated c_H of 0.1, significantly lower than a c_H close to 0.4 for humans (23). Second, in industrialized human populations, after the second demographic transition, selection

due to pre-reproductive mortality is deeply relaxed (24, 25). Thus, recombination would rapidly destroy linkage disequilibria between deleterious alleles.

While several factors – geographic structure, technical issues, antagonistic epistasis – can lead to an overdispersed mutation burden distribution, only synergistic epistasis can lead to an underdispersed distribution for unlinked loci. We further partitioned the underdispersion signal into within- and between-chromosome components, and demonstrated that the deleterious mutation burden was underdispersed due to multi-locus associations both within and between different chromosomes (Table S6, Table S7). While the negative linkage disequilibria from linked regions may be attributed to Hill-Robertson interference (26, 27) the majority of our underdispersion signal comes from unlinked pairs of loci, even for pairs of loci within the same chromosome (Fig. S6, Fig. S7). Having identified and controlled for other sources of LD, we thus invoke synergistic epistasis as a significant contributor to the underdispersion signal in deleterious mutation burden that we observe in humans and flies.

Thirty years ago, Neel posed the question: "The amount of silent DNA is steadily shrinking. The question of how our species accommodates such [high deleterious] mutation rates is central to evolutionary thought (28)." Indeed, a newborn receives ~70 *de novo* mutations (29). Although estimates for the target size for deleterious alleles (fraction of the genome that is "functional") vary, an overwhelming majority suggest that about 10% of the human genome sequence is functionally significant and selectively constrained (30, 31). Thus, the average human individual is expected to carry at least seven *de novo* deleterious mutations, which is incompatible with the long-term population survival if selection is non-epistatic. Moreover, regardless of epistasis, at the mutation-selection equilibrium, the sum of coefficients of selection against mutant alleles present in an average genotype must equal the genome-wide deleterious mutation rate (32). Recently Henn et al (20) independently estimated this sum in humans to be 15. Without epistatic selection, this suggests a mutation load that is inconsistent with the existence of the population ($1 - e^{-15} > 0.999$). Thus, synergistic epistasis is the only way for humans to survive, and in a sense, our findings are not unexpected. In industrialized human populations, while selection due to pre-reproductive mortality is deeply relaxed, there is still a substantial opportunity for selection due to differential fertility (33). Also, only ~30% of human conceptions result in live births (34), indicating a substantial

opportunity for prenatal selection. Thus, our results suggest that epistatic negative selection in humans is ongoing.

Conflict of Interest:

The authors declare no conflicts of interest.

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References

1. H. J. Muller, Our load of mutations. *J. Hum. Genet.* **2**, 111–176 (1950).
2. E. E. Shnol, A. S. Kondrashov, The effect of selection on the phenotypic variance. *Genetics*. **134**, 995–6 (1993).
3. T. Kimura, Motoo and Maruyama, The mutational load with epistatic gene interactions in fitness. *Genetics*, 1337–1351 (1966).
4. J. F. Crow, M. Kimura, Efficiency of truncation selection. *Proc. Natl. Acad. Sci. U. S. A.* **76**, 396–9 (1979).
5. A. S. Kondrashov, Deleterious mutations and the evolution of sexual reproduction. *Nature*. **336**, 435–440 (1988).
6. J. A. G. M. de Visser, S. F. Elena, The evolution of sex: empirical insights into the roles of epistasis and drift. *Nat. Rev. Genet.* **8**, 139–49 (2007).
7. R. Sanjuán, S. F. Elena, Epistasis correlates to genomic complexity. *Proc. Natl. Acad. Sci. U. S. A.* **103**, 14402–14405 (2006).
8. R. D. Kouyos, O. K. Silander, S. Bonhoeffer, Epistasis between deleterious mutations and the evolution of recombination. *Trends Ecol. Evol.* **22**, 308–15 (2007).
9. H.-C. Chiu, C. J. Marx, D. Segrè, Epistasis from functional dependence of fitness on underlying traits. *Proc. Biol. Sci.* **279**, 4156–64 (2012).
10. M. Bulmer, *The Mathematical Theory of Quantitative Genetics* (1980).
11. M. Nei, Analysis of gene diversity in subdivided populations. *Proc. Natl. Acad. Sci. U. S. A.* **70**, 3321–3323 (1973).
12. A. S. Kondrashov, Dynamics of unconditionally deleterious mutations: Gaussian approximation and soft selection. *Genet. Res.* **65**, 113–121 (1995).
13. B. Charlesworth, Mutation-selection balance and the evolutionary advantage of sex and recombination. *Genet. Res.* **55**, 199–221 (1990).
14. L. C. Francioli *et al.*, Whole-genome sequence variation, population structure and demographic history of the Dutch population. *Nat. Genet.* **46**, 818–25 (2014).
15. J. B. Lack *et al.*, The drosophila genome nexus: A population genomic resource of 623 *Drosophila melanogaster* genomes, including 197 from a single ancestral range population. *Genetics*. **199**, 1229–1241 (2015).
16. G. A. McVean *et al.*, An integrated map of genetic variation from 1,092 human genomes. *Nature*. **491**, 56–65 (2012).
17. T. F. C. Mackay *et al.*, The *Drosophila melanogaster* Genetic Reference Panel. *Nature*. **482**, 173–8 (2012).
18. R. C. Yang, Zygotic associations and multilocus statistics a nonequilibrium diploid population. *Genetics*. **155**, 1449–1458 (2000).
19. Rong-Cai Yang, Gametic and Zygotic Associations. *Genetics*. **452**, 451–452 (2003).
20. B. M. Henn, L. R. Botigué, C. D. Bustamante, A. G. Clark, S. Gravel, Estimating the mutation load in human genomes. *Nat. Publ. Gr.* **16**, 1–11 (2015).
21. O. Lao *et al.*, Correlation between genetic and geographic structure in Europe. *Curr. Biol.* **18**, 1241–8 (2008).
22. B. Charlesworth, D. Charlesworth, *Elements of evolutionary genetics* (2010).
23. M. G. Bulmer, Linkage disequilibrium and genetic variability. *Genet. Res.* **23**, 281–289 (1974).
24. H. M. E. Hed, Trends in Opportunity for Natural Selection in the Swedish Population During the Period 1650–1980. *Hum. Biol.* **59**:5, 785–797 (1987).
25. A. Courtiol *et al.*, The demographic transition influences variance in fitness and selection on height and BMI in rural Gambia. *Curr. Biol.* **23**, 884–889 (2013).
26. W. G. Hill, A. Robertson, Linkage disequilibrium in finite populations. *Theor. Appl. Genet.* **38**, 226–31 (1968).
27. N. H. Barton, S. P. Otto, Evolution of recombination due to random drift. *Genetics*. **169**, 2353–2370 (2005).
28. J. V. Neel *et al.*, The rate with which spontaneous mutation alters the electrophoretic mobility of polypeptides. *Proc. Natl. Acad. Sci. U. S. A.* **83**, 389–93 (1986).
29. S. Besenbacher *et al.*, Novel variation and de novo mutation rates in population-wide de novo assembled Danish trios. *Nat. Commun.* **6**, 5969 (2015).

30. C. M. Rands, S. Meader, C. P. Ponting, G. Lunter, 8.2% of the Human Genome Is Constrained: Variation in Rates of Turnover across Functional Element Classes in the Human Lineage. *PLoS Genet.* **10** (2014), doi:10.1371/journal.pgen.1004525.
31. L. D. Ward, M. Kellis, Evidence of abundant purifying selection in humans for recently acquired regulatory functions. *Science* (80-.). **337** (2012).
32. N. Morton, J. Crow, H. Muller, An estimate of the mutational damage in man from data on consanguineous marriages. *PNAS.* **42**, 855–863 (1956).
33. S. C. Stearns, S. G. Byars, D. R. Govindaraju, D. Ewbank, Measuring selection in contemporary human populations. *Nat. Rev. Genet.* **11**, 611–622 (2010).
34. E. C. Larsen, O. B. Christiansen, A. M. Kolte, N. Macklon, New insights into mechanisms behind miscarriage. *BMC Med.* **11**, 154 (2013).