

1 **Using genetic variation to disentangle the complex relationship between food intake and**
2 **health outcomes.**

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44 **Abstract:**

45 **Despite food choices being one of the most important factors influencing health, efforts**
46 **to identify individual food groups and dietary patterns that cause disease have been**
47 **challenging, with traditional nutritional epidemiological approaches plagued by biases**
48 **and confounding. After identifying 302 individual genetic determinants of dietary**
49 **intake in 445,779 individuals in the UK Biobank study, we develop a statistical genetics**
50 **framework that enables us, to directly assess the impact of food choices on health**
51 **outcomes. We show that the biases which affect observational studies extend also to**
52 **GWAS, genetic correlations and causal inference through genetics, which can be**
53 **corrected by applying our methods. Finally, by applying Mendelian Randomization**
54 **approaches to the corrected results we identify some of the first robust causal**
55 **associations between eating patterns and cancer, heart disease, obesity, and several**
56 **other health related risk factors, distinguishing between the effects of specific foods or**
57 **dietary patterns.**

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67 **Introduction**

68 Given their profound impact on human well-being, diet is one of the most studied human
69 behaviours. Quality, quantity, and patterns of consumed foods are associated with a wide
70 range of medical conditions such as metabolic, inflammatory, or mental health diseases¹.
71 However, despite the growing number of studies reporting associations between diet and
72 health outcomes, it has been challenging to establish causal relationships due methodological
73 limitations such as measurement error, confounding, and reverse causation. To date, several
74 methods have been devised to try to account for intrinsic limitations in nutritional studies
75 such as calibration of food records² or the implementation of domiciled feeding studies (ie.
76 the PREDICT study³) in which participants are instructed to eat only the food provided by the
77 study. Although these methods have helped in addressing some the limitations related to food
78 consumption measurement, problems still remain especially when it comes to measure the
79 effects of food on health over a long period of time.

80 In this context genetics may represent an alternative approach through the use of Mendelian
81 Randomization. Mendelian Randomization (MR) is a methodological approach in which
82 genetic variants associated with a phenotype of interest are used as instrumental variables to
83 measure the “life-long effect of an exposure” to an outcome.⁴ To date, several MR studies
84 have been designed to investigate the associations between the consumption of single food
85 groups, such as alcoholic beverages⁵, coffee⁶, milk⁷⁻⁹ and specific health outcomes, but a
86 systematic study investigating the overall role of diet is missing. In addition, previous MR
87 studies have not accounted for the fact that genetic variants associated with reported dietary
88 intake may be primarily associated with other risk factors or social determinants of health
89 which may confound the causal estimates if used. In addition, previous studies on single food
90 groups have not accounted for inter-relationships between different foods thus limiting the
91 interpretability of the findings.

92 Given the complex number of factors that are driving the association between diet and health
93 outcomes, the present study was designed to initially identify the genetic variants associated
94 with reported food consumption, and then to leverage a causal inference statistical framework
95 to systematically investigate the causal effects of dietary factors on health outcomes, while
96 accounting for the effects that health determinants have on habitual dietary intake reporting.

97 **Methods**

98 **Study population and genome-wide association for dietary intake**

99 The UK Biobank¹⁰ is a large population-based cohort including 500 000 adults aged between
100 40 and 69 years at baseline across 22 assessments centers in the United Kingdom. Data were
101 collected based on clinical examinations, assays of biological samples, detailed information
102 on self-reported health characteristics, and genome-wide genotyping. Dietary intake in UK
103 Biobank was assessed using a food frequency questionnaire which included questions about
104 the frequency of consumption specific foods and beverages over the past year. The number of
105 samples used for each trait can be found in table S1 while a detailed description of the
106 phenotypes, can be found in the in the supplementary methods 1.2 and table S2.

107 We used the BOLT-LMM software¹¹ to assess the association between the genetic variants
108 across the human genome and 29 food phenotypes. Analyses were conducted on genetic data
109 release version 3 imputed to the HRC panel¹², as provided by the UK Biobank

110 ([http://www.ukbiobank.ac.uk/wp-](http://www.ukbiobank.ac.uk/wp-content/uploads/2014/04/UKBiobank_genotyping_QC_documentation-web.pdf)
111 [content/uploads/2014/04/UKBiobank_genotyping_QC_documentation-web.pdf](http://www.ukbiobank.ac.uk/wp-content/uploads/2014/04/UKBiobank_genotyping_QC_documentation-web.pdf)). Population
112 stratification was assessed using LD-score regression as implemented in LD Hub^{13,14} using
113 the LD scores provided with the software. Table S15 reports for each food trait the LD
114 regression intercept and heritability estimation using ldsc. Cluster analysis conducted on the
115 foods identified 5 main groups of traits (see additional online methods paragraph 1.8 and 2.2

116 for details of group definition) and we thus set the genome-wide significance threshold at
117 1×10^{-8} . Work within was conducted under UKB application 19655. Participants enrolled in
118 UK Biobank have signed consent forms. Replication analyses for identified signals
119 associated with food phenotypes were conducted independently by using genetic and dietary
120 data from the EPIC-Norfolk Study¹⁵ and the Fenland Study¹⁶. Details additional online
121 methods 1.4.

122 **Investigating the effect of health outcomes on reported food intake using MR.**

123 Univariable MR analyses were initially conducted to measure the causal effect of health
124 outcomes on food consumption using the TwoSampleMR¹⁷ R package. Exposures of interest
125 were selected amongst those for which nutritional advice is given and included body mass
126 index (BMI), low density lipoprotein cholesterol (LDLc), high density lipoprotein cholesterol
127 (HDLc), Total cholesterol, Triglycerides, Diastolic and Systolic blood pressure, Type 2
128 diabetes, and coronary artery disease. In addition, we included educational attainment as a
129 proxy of socio-economic status which is likely to affect food consumption. The full list of
130 studies from which the summary statistics were derived is detailed in Table S6. For each
131 exposure we selected all SNPs with $p < 5 \times 10^{-8}$ and $r^2 < 0.001$ to be used as instruments in the
132 MR analysis. After performing stepwise heterogeneity pruning we performed MR analysis
133 using the inverse variance method¹⁸. We then tested if the intercept from the MR-Egger¹⁹
134 regression was different from zero ($p < 0.05$). If this was the case, MR-Egger was used for the
135 analysis instead.

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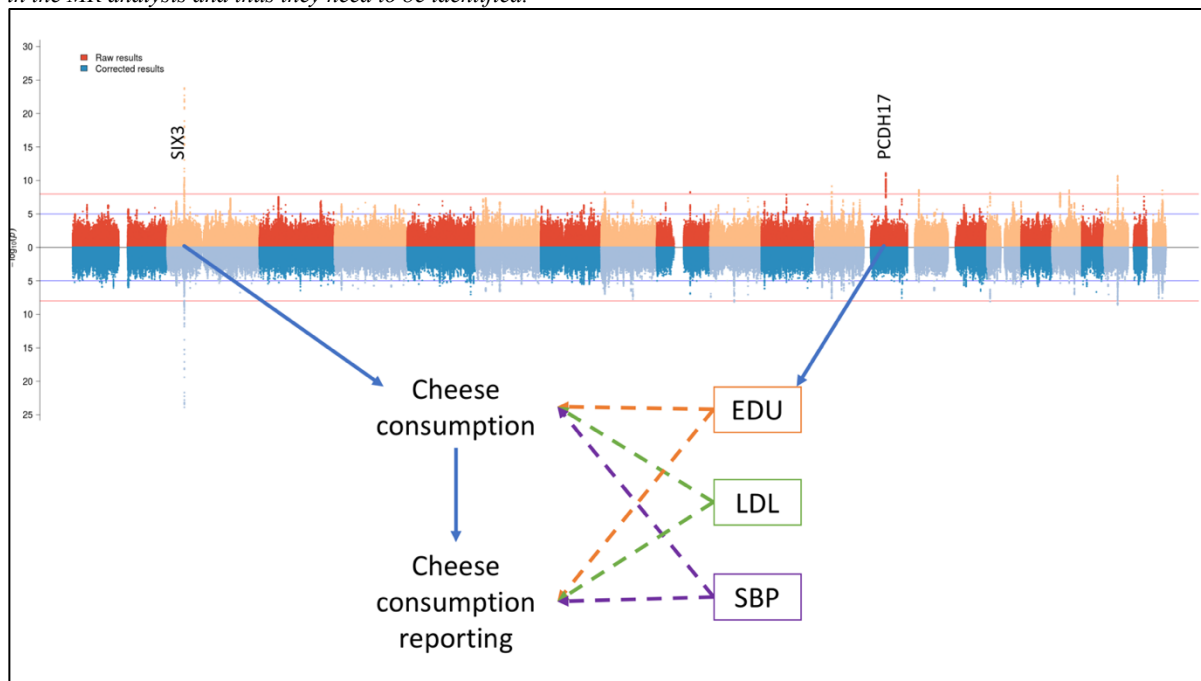
137 **Measuring the direct effects of food types on health outcomes and identifying genetic** 138 **variants with predominantly direct-effects**

139 One of the most important assumptions in MR is that the effect of the instrument on the
140 outcome must be mediated only through the exposure of interest (sometimes referred as
141 exclusion restriction criteria)²⁰. In this light the instruments whose effect on food is mediated
142 through the health outcomes or through educational attainment may violate this assumption
143 acting as confounders in the relationship between the exposure and the outcome. Moreover if
144 the mediating trait is acting on the reporting of food consumption and not food consumption
145 itself it would mean that the genetic variant is not truly associated to food consumption and it
146 would thus not be a valid instrument. It is thus important to estimate the direct effect (i.e., the
147 effect that acts directly on food intake rather than is mediated through other factors see Figure
148 1) the SNPs are exerting on actual food consumption in order to properly select the genetic
149 variants to be used as instrumental variables.

150 To this end we use a modified version of the method implemented in bGWAS²¹. This method
151 consists of a first step where the phenotype of interest (i.e., food consumption) is used as
152 outcome in multivariable MR. Next, exposures of interest are selected using a forward step
153 wise regression selection algorithm where each exposure is added until their p-value is less
154 than 0.05. The method provides a corrected estimate for each genetic variant of its effect on
155 the outcome trait once all mediated effects are removed. Further details can be found in
156 supplementary methods 1.6. In order to identify genetic variants with only a direct effect on
157 the phenotype of interest we defined the corrected to uncorrected ratio (CUR) as the ratio
158 between the corrected and the uncorrected effects (see additional methods 1.7 for a detailed
159 explanation).

160 **Fig. 1 Direct and indirect SNP effects.** The plot shows the causal path of exemplar genes identified for cheese consumption.
161 In the multivariable MR model cheese consumption is causally influenced by educational attainment (EDU), low density
162 lipoprotein cholesterol levels (LDL) and systolic blood pressure (SBP). The effect of PDCH17 is mediated through
163 educational attainment, while SLX3 has a direct effect on cheese consumption. The mediated effects cannot be used reliably
164 as MR instruments as they could be affecting either consumption or its reporting. Moreover, they could act as confounders

165 *in the MR analysis and thus they need to be identified.*



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167 The threshold to define genetic variants with non-mediated effects ($CUR=1\pm 0.05$) is based on
168 simulations provided in the supplementary note 2.1 and on the genetic variants with known
169 biological function (ie. bitter receptors). We defined as “non-mediated” those SNPs whose
170 CUR fell within the defined ranges while “uncertain” the others. We applied bGWAS to all
171 29 food phenotypes. As potential mediators, we used the same cardiometabolic phenotypes as
172 before except total cholesterol to avoid collinearity issues with LDL and HDL cholesterol,
173 and we added summary statistics from Crohn’s disease and ulcerative colitis as they are
174 likely to affect dietary patterns. A Detailed discussion of this approach can be found in
175 supplementary methods 1.6.

176 **Genome-wide genetic correlations between corrected dietary intake and health**
177 **outcomes.**

178 We used LD-score regression implemented in LD Hub^{13,14} to estimate genome-wide genetic
179 correlations between dietary intake phenotypes and 844 health outcomes and intermediary
180 phenotypes. Genetic correlations were estimated both with the corrected and uncorrected
181 GWAS summary statistics using the bivariate LD-score regression model. Stratified LD-

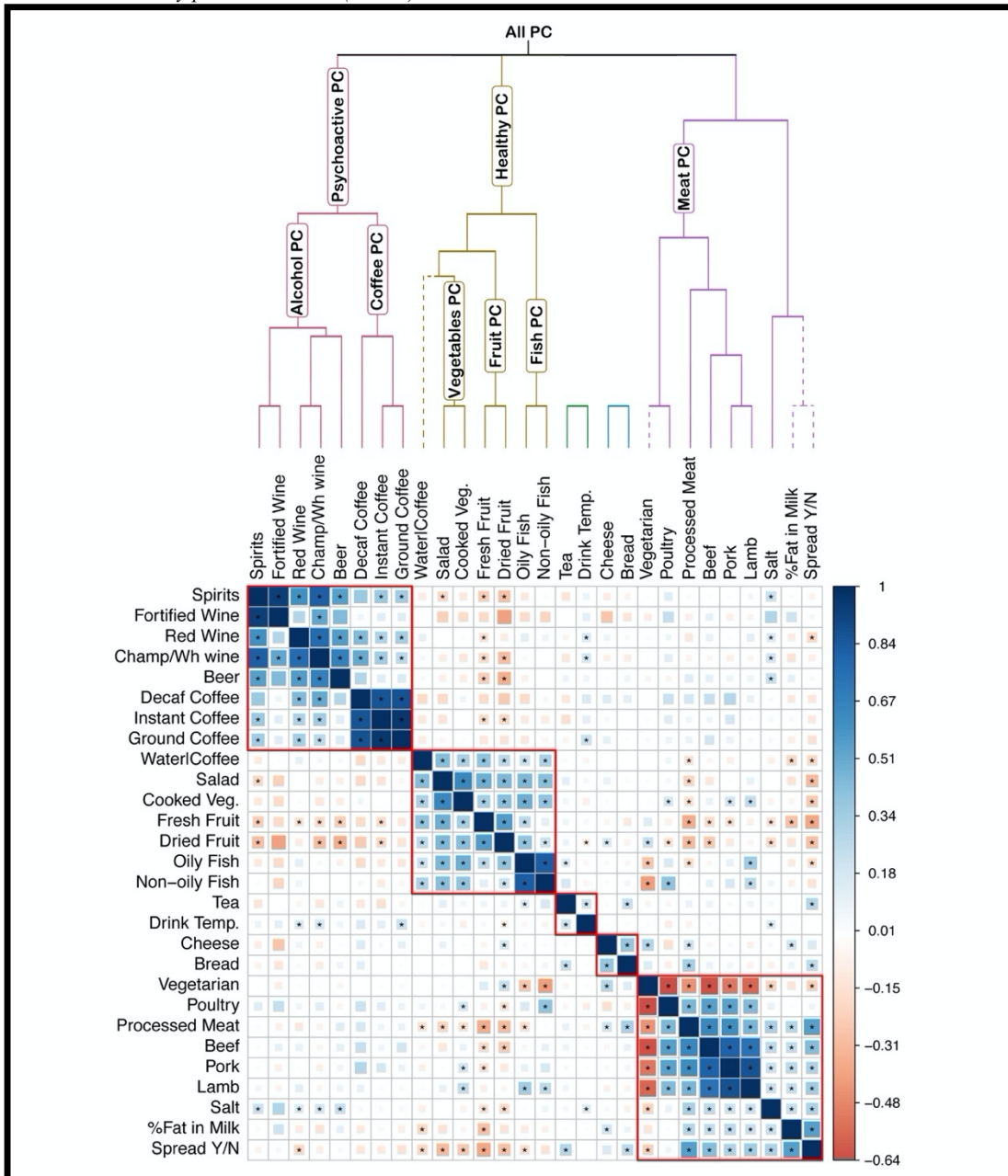
182 score regression²² analyses were implemented using ldsc and the annotation files available on
183 the ldsc website.

184 **Definition of food group variables**

185 In order to define measures of dietary patterns we first performed cluster analysis of the 29
186 food items applying iCLUST²³ to the corrected genetic correlation matrix between the
187 different foods. iCLUST clusters items in different groups based on a hierarchical structure
188 (Details additional methods 1.8). Figure 2 shows the resulting dendrogram and its
189 comparison with the genetic correlation matrix.

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233 **Fig2 Clustering of the food traits and definition of measures of dietary patterns.** The plot reports the genetic correlation
 234 plot amongst the food traits after applying the correction. The stars report the Bonferroni-corrected significant correlations.
 235 The dendrogram and the boxes represent the clustering according to the ICLUS algorithm. The labels on the dendrogram
 236 branches show the traits used to define each measure of dietary pattern. The dashed line represents the traits excluded from
 237 the estimation of the dietary patterns traits. The “Vegetarian” trait was excluded from the “Meat PC” trait but was included
 238 in the overall dietary pattern measure (All PC).



239 We then defined based on the resulting structure several measures of dietary pattern at
 240 different levels of the dendrogram as shown in Figure 2. For each measure we performed
 241 principal component analysis of the items which participated to each group. The rotation
 242 matrix was derived from the eigen decomposition of the correlation matrix of the foods in the
 243 PC trait of interest. For example for the Coffee PC measure we performed principal
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245 component analysis of “Ground Coffee”, “Instant Coffee” and “Decaf Coffee”. Once the
246 rotation matrix was estimated for each SNP its effect on the new measure was estimated as
247 the linear combination of the effect on each food trait using as weights the loadings on each
248 PC. A correlation plot of the loadings of each item onto the PC traits can be found in figure
249 S3.

250 **MR analyses to assess causal relationships between food intake and health outcomes**

251 MR analyses were conducted to estimate the effects of the food phenotypes on 79 health
252 related phenotypes (see table S17 for details) available in MR-base.¹⁷ Genetic instruments for
253 each exposure of interest included independent genetic variants ($p < 5 \times 10^{-8}$ and pruning for
254 LD ($r^2 < 0.001$)). For dietary patterns exposures SNPs were selected as outlined in additional
255 methods 1.12. For the main analysis we restricted the genetic instruments to those that only
256 had evidence of a direct effect (i.e., not affecting the main exposure through a different
257 pathway; CUR 1 ± 0.05). Discussion of the relationship with other methods can be found in
258 supplementary note 2.7. Weights for the genetic instruments were based on the uncorrected
259 effects. To verify the effects of using only direct effect only SNPs on MR, all the analyses
260 were also conducted without applying the CUR filtering.

261 After selecting the genetic instruments, exposure and outcome data were harmonised. The
262 MR estimates were tested for heterogeneity and outliers were removed using the MR-Radial
263 method.²⁴ MR analyses were based on the inverse variance weighted method, which
264 estimates the causal effect of an exposure on an outcome by combining ratio estimates using
265 each variant. A random effect model was used if significant heterogeneity between the
266 different estimates was detected. We then tested for the presence of directional pleiotropy
267 using the intercept from the MR-Egger regression. MR median and MR-Raps were used as
268 sensitivity analyses. All results have been made available through an online app (
269 https://npirastu.shinyapps.io/Food_MR/) and can be found in additional table S18.

270 Patient and public involvement

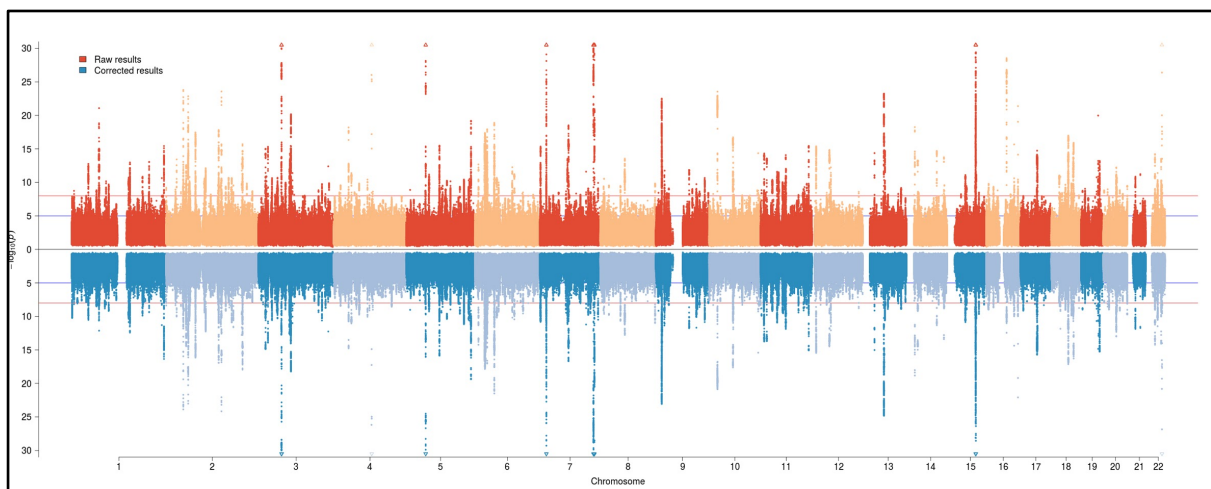
271 This research did not involve patients or the public as it uses data from the UK Biobank study
272 that were previously obtained from a cohort of people who had already been recruited. As
273 such, no patients or member of the public were involved in the design or implementation of
274 this study or the research questions addressed.

275 Results

276 Genetic variants associated with food intake

277 In a GWAS of 29 food phenotypes we identified 414 genetic associations in 260 independent
278 loci (Fig 3 and additional table S4) at Bonferroni corrected level of significance ($P < 1 \times 10^{-8}$).

279 **Fig. 3 302 independent genomic loci associate with food choices.** Results for both univariate (260 loci) and multivariate
280 (additional 42 loci see paragraph S2.3) analyses are included. For each SNP the lowest p-value for all traits was plotted.
281 The upper panel represents the unadjusted GWAS associations while the lower panel represents the association with food
282 choices, after adjustment for mediating traits, such as health status.



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284 Replication was sought in two additional UK-based cohorts including up to 32,779
285 participants. Despite relatively limited power in replication cohorts, concordant direction of
286 effect was observed for 82% of the signals ($p=7.82 \times 10^{-35}$, Binomial test; Table S5), and
287 nominal significance was achieved by 32% of the signals ($p=9.47 \times 10^{-54}$). Gene prioritization
288 is described in supplementary methods 1.10 while biological annotation, network analysis
289 and tissue enrichment analysis are discussed in additional paragraphs 1.11, 2.4 and 2.5.

290 Several of the identified loci have been previously associated with BMI. However, contrary
291 to our expectations, the BMI-raising allele was consistently associated with lower reported
292 consumption of energy-dense foods such as meat or fat, and higher reported intake of low-
293 calorie foods.

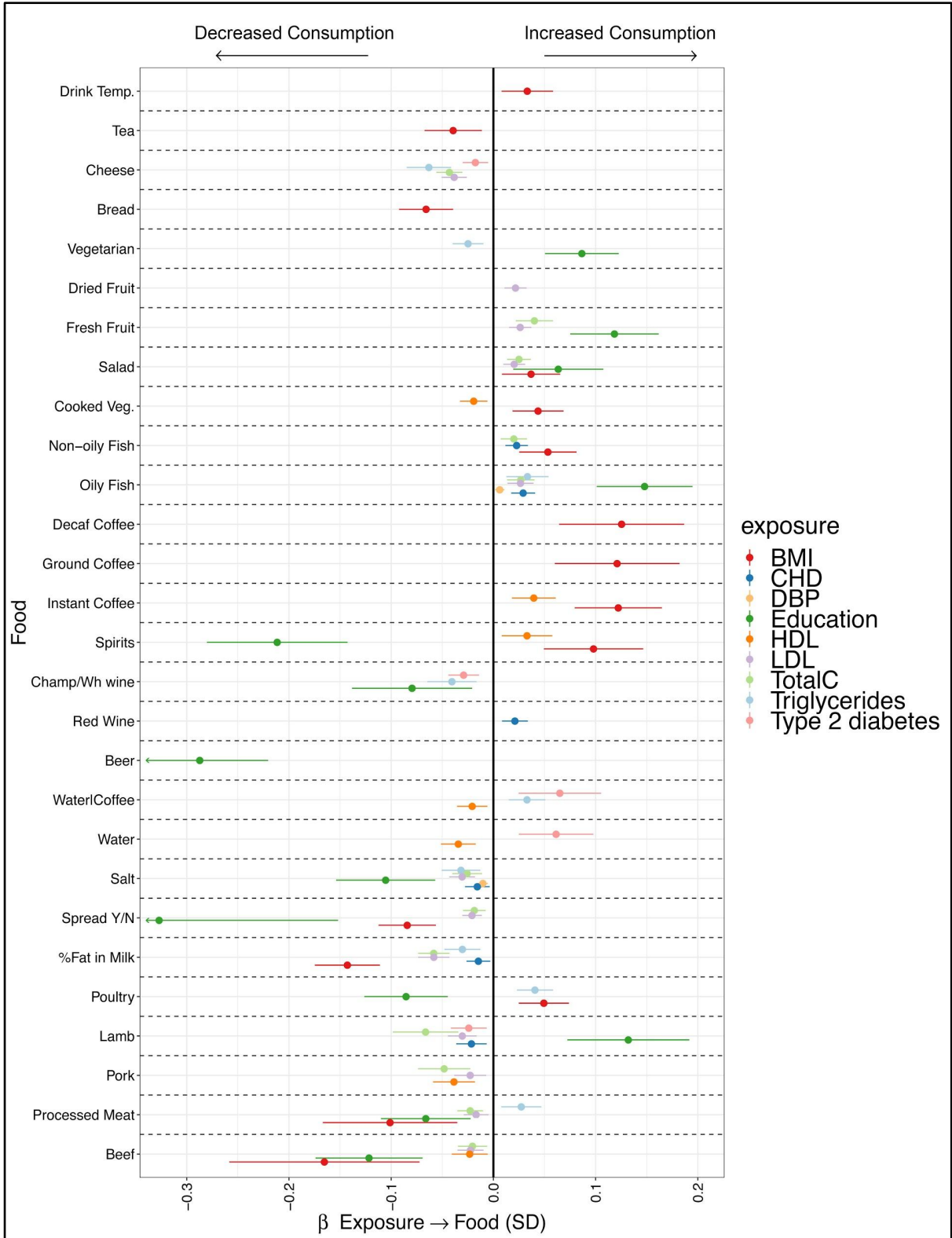
294 **Genetic variants associated with food intake are strongly influenced by other**
295 **phenotypes**

296 In univariable MR we identified 81 instances in which health-related traits significantly
297 influencing food intake (Fig. 4 additional table S7). In particular BMI and Educational
298 attainment influenced more than 50% of the food traits. Similar effects extend to a broad
299 range of traits, for example LDL and triglycerides influenced 15 and 18 traits respectively.
300 Higher genetically-determined CAD associates with higher consumption of fish and red
301 wine, and lower consumption of whole milk, salt and lamb. These findings suggest that some
302 of the signals identified in GWAS for reported food phenotypes are not directly associated
303 with food intake but are mediated through a wide range of potential confounders.

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323 **Fig 4. Health status influences reported food choices.** The plot reports only the univariable MR results which were
 324 significant at $FDR < 0.05$. For each food outcome the effect estimate (β) is reported in standard deviations of the exposure
 325 trait, together with 95% confidence intervals. Each colour represents a different exposure. BMI, body mass index; CHD,
 326 coronary heart disease; DBP, diastolic blood pressure; HDL, high density lipoprotein cholesterol; LDL, low density
 327 lipoprotein cholesterol; TotalC, total cholesterol. Champ/Wh wine, champagne, white wine. Temp, temperature.



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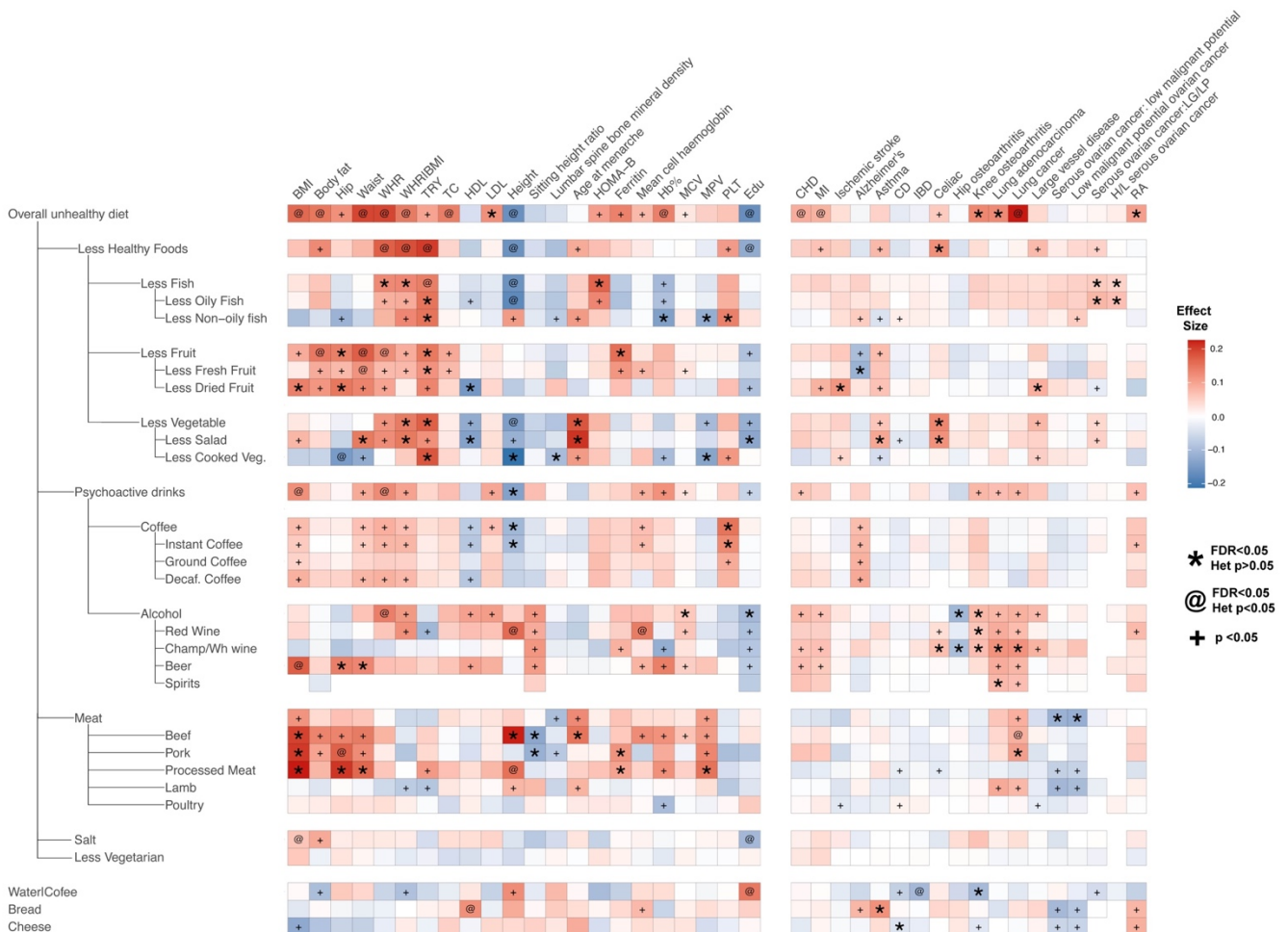
329 The Multivariable MR confirmed the univariable MR results (Supplementary Fig S4 panel A
330 and Supplementary Table S8). The percentage of genetic variance for the reported food
331 phenotypes explained by health determinants ranged from 42% for cheese to ~0% for
332 fortified wine and white wine/champagne (Supplementary Fig S3 panel B and Supplementary
333 Table S16). We systematically compared the estimated effect sizes of each genetic variants
334 influencing food consumption before and after correcting for the effect of health determinants
335 and showed that in many loci the variant initially identified for food phenotypes changed
336 dramatically after taking into account the effect of health factors (Fig. 3, see Supplementary
337 file 1 for trait-specific plots). For example, the effect size of the lead *FTO* variant
338 (rs55872725, $p=2 \times 10^{-29}$) on milk fat percentage chosen decreased three-fold after accounting
339 for the mediated effects. To further explore the magnitude of this indirect effect on food
340 intake phenotypes, we compared the correlation patterns between the 29 food phenotypes and
341 832 phenotypes present in the LD hub¹⁴ database identifying great differences. For example,
342 low fat milk intake was correlated with a beneficial effect on body fat percentage ($r_G = -0.43$)
343 but this association diminished to near zero ($r_G = -0.04$) after accounting for indirect effects
344 (Supplementary Data 2.2 and additional table S10). The effects of the correction procedure
345 on the genetic correlation amongst the traits and with the 844 health traits are discussed in
346 supplementary note 2.2 while full results can be found at in table S9 and browsed at
347 https://npirastu.shinyapps.io/rg_plotter_2/. These findings highlight the relevance of biases
348 and confounding in genetic correlation studies, and provide the framework to study complex
349 physiological relationships.

350 **Causal inference analyses for diet phenotypes and health outcomes**

351 A total of 230 out of 414 genetic variants initially associated with food phenotypes
352 (corresponding to 169/260 loci) were categorized as “non-mediated” associations (Table S3).
353 The balance of uncertain to non-mediated genetic associations varied by food group, ranging

354 from none uncertain for tea, spirits and processed meat, to all uncertain for percentage fat in
 355 milk and adding spread to bread (Table S3).
 356 In two-sample MR analyses we found 141 significant associations between food phenotypes
 357 and health outcomes after multiple test correction (pFDR < 0.05, Table S18).
 358 Of these 89 showed no sign of heterogeneity amongst the estimates (heterogeneity test p
 359 >0.05). Figure 5 reports full results for all significant food exposure trait outcome pairs.

360 **Fig 5. Significant effects of food choice on disease related traits.** The heatmap reports the results for all significant food
 361 trait exposure trait outcome. Only dietary pattern exposures summarising the overall group consumption (PC1) have been
 362 reported. All exposures have been aligned to have a positive loading onto the “overall unhealthy diet” measure. Significant
 363 food/trait association are indicated with * if they show no sign of heterogeneity while @ if they show significant
 364 heterogeneity. To facilitate meaningful visualisation and maximise the appearance of signal rather than noise, we applied a
 365 shrinkage method - imposing a bayesian prior assumption on the distribution of beta (mean 0, SD 0.1), and conjugating that
 366 with the likelihood of our results and then taking mean beta from the resulting distribution, thus shrinking estimates with
 367 larger SEs more towards 0. Abbreviations: BMI Body Mass Index, WHR Waist to Hip Ratio, TRY tryglicerides, TC total
 368 cholesterol, HDL HDL cholesterol, LDL LDL cholesterol, Hb% Haemoglobin percentage, MCV Mean Corpuscular Volume,
 369 MPV Mean Platelet Volume, PLT Platelet count, Edu Educational attainment, CHD Coronary Heart Disease, MI
 370 Myocardial Infarction, CD Chron’s Disease, IBD Inflammatory Bowel Disease, Serous ovarian cancer:LG/LP low grade
 371 low potential. H/L serous ovarian cancer High and Low grade serous ovarian cancer, RA Rheumatoid Arthritis.



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373 Overall we found evidence supporting the beneficial effect of a healthy diet on health
374 outcomes. For example, for obesity/adiposity outcomes, genetically-determined unhealthy
375 diet leads to very similar effects across, increasing obesity measurements. For lipid-related
376 outcomes, the overall unhealthy diet is associated with higher levels of LDLc with no
377 significant heterogeneity, but no association with any of the other dietary traits. The overall
378 unhealthy diet was also strongly associated with Lung adenocarcinoma (OR 1.4xSD CI 1.2-
379 1.9) which seemed to be driven mostly by alcoholic beverages.

380 We identified 51 instances in which we would have not detected a significant result without
381 filtering out the non-direct effect instruments such as the effect of increased fruit
382 consumption on triglycerides levels (estimated uncorrected effect= -0.03 (SE=0.05) vs.
383 estimated corrected effect = -0.17 (SE=0.05)) or the effect of increased beef consumption on
384 height (uncorrected effect = -0.02 (-0.17, 0.13) vs corrected effect = -0.52 (0.29, 0.74). In
385 addition, we found 124 food/trait relationships which were not significant after applying
386 CUR filtering, showing that either confounding effects or reduced power explain the lack of
387 association (see additional note 2.6). For example, red wine consumption was initially
388 associated with increased BMI (uncorrected effect =0.22 (SE 0.05)) and waist circumference
389 (uncorrected beta= 0.26 (SE 0.07), but after correcting for CAD liability, both effects
390 disappeared (corrected effect for BMI 0.05 (SE 0.06), corrected effect for WC 0.005 (0.08)).

391 On the flip side, we showed that the effect of red wine on mean corpuscular volume remains
392 substantially unchanged when applying the filtering approach (beta 0.07 (SE 0.02)
393 uncorrected and 0.065 (SE 0.02) corrected), suggesting that our approach could precisely
394 identify relevant biological relationships.

395 A full description of our findings are found in table S18 and have been made available
396 through an online app (https://npirastu.shinyapps.io/Food_MR/).

397

398 **Discussion**

399 In this study we have provided quantitative data about the complex interplay between diet
400 and health outcomes showing that the causal path from food intake to adverse health
401 outcomes is not unidirectional and may be influenced by reverse causation and confounding
402 even when MR is used. We showed that genetic correlations and causal inference can be
403 improved by leveraging statistical approaches that take into account this mediated effects and
404 identify genetic variants that have a only non-mediated effects on the exposure of interest.
405 This information allowed us to perform causal inference analyses that helped identifying
406 more reliable potential causal effects of food on health outcomes.

407 **Results in context**

408 Previous MR studies have mainly focused on specific food groups such as coffee, alcohol and
409 milk consumption while none has comprehensively investigated the role of different food
410 groups on health outcomes. Our results support previous observations such as the effect of
411 alcohol consumption on coronary artery diseases reported in previous MR studies. In
412 addition, we were able to confirm similar previous results detecting no evidence of an effect
413 on IBD and CD²⁵, ovarian cancer²⁶ or rheumatoid arthritis²⁷.

414 Findings from this study also suggest that the same biases that affect measures of food
415 consumption such as reporting bias, confounding and reverse causation are reflected also in
416 studies focusing on genetic associations. We have shown that these issues extend beyond
417 obesity and socio-economic status including a broader range of intermediate factors. For
418 example blood LDL and triglycerides concentration influence a wide variety of food traits
419 thus being important factors to be considered as potential sources of bias, yet to our
420 knowledge this is the first time this has been reported. For our analyses we have used UK
421 biobank in which participants were aged between 40 and 60 at the time of the questionnaire,

422 it is likely that a younger cohort will suffer less from some of these (ie. LDL cholesterol or
423 blood pressure) as it is unlikely that they will display pathological level of these traits.
424 Our results are in contradiction to some previous studies in which no evidence of reverse
425 causation influencing genetic susceptibility for dietary patterns was reported.^{28,29} We believe
426 that this difference is due to our novel approach, which is not based on using the potential
427 mediators as covariates, but rather exploits MR, which should be able to distinguish the
428 forward and reverse effects when the causal relationship is bidirectional. We have thus shown
429 that it is possible, through the use of available data and methods, to disentangle these
430 different colliding effects and to select the instrumental variables which show a non-mediated
431 effect, thus enabling the use of MR for the assessment of causal relationships between food
432 and health.

433 Many studies have looked at the relationship between nutritional composition and health
434 outcomes. One of the most salient examples is the relationship between saturated fat intake
435 and cardiovascular disease and all-cause mortality, in which recent studies suggest that food
436 sources of saturated fatty acids are more important than saturated fat content per se^[Citation error].
437 Our study provide a new angle on the importance of food sources by providing evidence that
438 foods with similar nutrient profile, for example cheese and meat, which are both relatively
439 high in saturated fat and protein, have opposite effects on some metabolic risk factors such as
440 BMI (Figure S24 A) but there is no difference in other phenotypes such as blood lipids. A
441 similar conclusion can be drawn if we look at the foods which have the greatest effect on
442 triglycerides, fruit, vegetables and fish; all with very similar lowering effects (Figure S24 B),
443 which have relatively different macronutrient compositions. While the findings require
444 further investigations in mechanisms and related behaviours, our genetic evidence lends
445 support for the importance of studying foods in their complexity and not as a mere mixture of
446 nutrients. This approach, in fact, does not consider that the sources of the nutrients are not

447 equal due to the food matrix, the different preparations and that foods are seldom consumed
448 by themselves but in patterns which are likely to modify the effects on health.

449 Our findings illustrate that the effect of diet on health outcomes is complex, and components
450 of specific food groups have a differential association with health. In this case, although fish
451 and fruit and vegetables have a very different macronutrient composition it was impossible to
452 separate their effect on triglyceride concentrations. This suggests that at least in this case the
453 macronutrient composition is not as important as the an overall tendency to eat certain foods
454 and it highlights the importance of always including the assessment of dietary patterns before
455 claiming health effects of single foods or nutrients.

456 Some of the effects we have identified are more complex to explain and will need different
457 sources of evidence to be understood. For example we have found that the overall unhealthy
458 diet is associated to a higher risk of both lung adenocarcinoma and lung cancer. When
459 looking more closely to which of food explain this association the most we can see that
460 Alcohol seems to be driving the overall effect. One possibility is that this relationship is
461 confounded by smoking through a common tendency to addictive behaviours. However a
462 recent GWAS on cigarette smoking in Japan Biobank³⁰ reported a strong association between
463 the ALDH2 gene and number of cigarettes per day smoked which has also been associated to
464 differences in alcohol consumption³¹, suggesting a causal effect of increased alcohol
465 consumption on increased smoking thus predisposing to lung cancer. Regardless of the
466 interpretation this example shows how complex the interpretation of MR results are when
467 behavioural traits are involved as they influence each other constantly creating a complex net
468 of interrelationships. This also points to the need of extreme care when claiming beneficial
469 health effects of food and multiple sources of evidence and approaches should always be
470 used before translating these findings into public policies.

471 Our study has several potential limitations. First, the number of items available in the dietary
472 questionnaire in the UK BioBank is limited, and therefore it limited our ability to capture
473 overall diet or specific food groups not detailed. The inclusion of white and relatively healthy
474 and educated participants from UK Biobank may have limited the generalisability of our
475 findings. Estimated effect sizes could be inflated because of the underestimation of the SNP
476 effects on the actual food trait consumption, rather than its self-report, if so, this will have
477 inflated our estimates of the effects of food on health, due to the noise in the questionnaire
478 responses, and warrants further statistical investigations. Even so, our method should not
479 have falsely identified a causal effect or reversed its direction, but further studies are needed
480 to assess the precise effect sizes.

481 In conclusion, our findings show that overall what is generally considered a healthy diet leads
482 to many favourable health outcomes and to reducing a wide range of risk factors broadly
483 agreeing with current guidelines aimed at reducing meat and alcohol consumption while
484 increasing fruit vegetables and fish. We also show that some of these effects are mostly
485 reductable to specific food or group of foods which however are not characterized by
486 common nutrient composition thus adding granularity to our knowledge on the effect of diet
487 on health. This information can be useful to inform the design and implementation of future
488 studies to reduce the burden of diet-related diseases.

489 **Author Contributions**

490 NP,JFW,JRBP,ZK,EJG,FRD,KKO contributed to the study design.

491 JFW,TE,JRBP,AR,TG,FI,KKO,FRD contributed data.

492 NP,CMD,EJG,NM,FI,JZ,NT,KAK,MPC, performed the statistical analyses. NP, JFW,ZK,

493 JRBP, JM,TE, NT,KF,CMD,LR,EJG,FI,KKO,FRD contributed to the interpretation of the

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518

519 **Data Availability**

520 All GWAS results will be made available through GWAS catalog at the time of publication.

521 All results from the MR analyses have been shared in the additional tables.

522

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