Appraising causal relationships of dietary, nutritional and physical-activity exposures with overall and aggressive prostate cancer: two-sample Mendelian randomization study based on 79,148 prostate cancer cases and 61,106 controls

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Summary

Background: Prostate cancer is the second most common male cancer worldwide, but there is substantial geographical variation suggesting a potential role for modifiable risk factors in prostate carcinogenesis.

Methods: We identified previously reported prostate cancer risk factors from the World Cancer Research Fund's (WCRF) systematic appraisal of the global evidence (2018). We assessed whether each identified risk factor was causally associated with risk of overall (79,148 cases and 61,106 controls) or aggressive (15,167 cases and 58,308 controls) prostate cancer using Mendelian randomization (MR) based on genome wide association study (GWAS) summary statistics from the PRACTICAL and GAME-ON/ELLIPSE consortia. We assessed evidence for replication in UK Biobank (7,844 prostate cancer cases and 204,001 controls).

Findings: WCRF identified 57 potential risk factors, of which 22 could be instrumented for MR analyses using single nucleotide polymorphisms (SNPs). In MR analyses for overall prostate cancer, we identified evidence compatible with causality for the following risk factors (odds ratio [OR] per standard deviation increase; 95% confidence interval): accelerometer-measured physical-activity, OR=0.49 (0.33-0.72; p=0.0003); serum iron, OR=0.92 (0.86-0.98; p=0.007); body mass index (BMI), OR=0.90 (0.84-0.97; p=0.003); and mono-unsaturated fat, OR=1.11 (1.02-1.20; p=0.02). Findings in our replication analyses in UK Biobank were compatible with our main analyses (albeit with wide confidence intervals). In MR analysis, height was positively associated with aggressive prostate cancer risk: OR=1.07 (1.01-1.15; p=0.03).

Interpretation: The results for physical-activity, serum iron, BMI, mono-unsaturated fat and height are compatible with causality for prostate cancer but more research is needed to rule out violations of MR assumptions for some risk factors. The results suggest that interventions aimed at increasing physical activity may reduce prostate cancer risk, but the direction of effects of BMI, and iron are at odds with their effects on other diseases, so the overall public health impact of intervening on these need to be considered.

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Research in context

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Evidence before this study

- 3 The World Cancer Research Fund (WCRF) Continuous Update Project (CUP) has reported associations
- 4 between diet, nutrition and physical activity, and prostate cancer using observational studies. Establishing
- 5 causality is an important step in the development of prevention strategies but is challenging because inference
- 6 from observational studies is limited by often intractable biases, including measurement error, reverse causation
- 7 and residual, or unmeasured confounding. 8

Added value of this study

- 9 Mendelian randomization (MR) is a methodological approach to addressing measurement error, reverse
- 10 causation and confounding within observational studies, based on 'instrumental variable' (IV) analysis. We
- 11 systematically applied two-sample MR to appraise the evidence for a causal link between previously reported
- 12 lifestyle and anthropometric risk factors (from WCRF report) with overall prostate cancer risk and aggressive 13 prostate cancer risk.

Implications of all the available evidence

Our MR analyses showed that physical activity, BMI, and serum iron levels were inversely associated with overall prostate cancer risk and mono-unsaturated fat levels were positively associated and that these effects were likely to be causal. For these risk factors, the direction of association was consistent for aggressive prostate cancer risk. In addition, our MR analyses showed that height was positively associated with aggressive prostate cancer risk.

Introduction

In 2012, 1.1 million men were diagnosed with prostate cancer, making it the second most common male cancer worldwide^{1,2}. There is wide global variation in prostate cancer incidence, with almost 70% of cases occurring in more developed regions of the world². This variation is thought in part to be related to the intensity of prostatespecific antigen (PSA) based screening practices^{3,4}, although migration studies suggest an influence of environmental and lifestyle factors^{5,6}. Established risk factors include advanced age, ethnicity and family history of prostate cancer^{7,8}. In addition, several lifestyle and anthropometric factors have been hypothesised to play an aetiological role, and measures of adiposity a prognostic role⁸. However, the epidemiological evidence to support a causal role for these potentially modifiable factors is weak. This is because inference from observational studies is limited by residual or unmeasured confounding, and other biases such as reverse causation and detection bias^{9,10}.

Mendelian randomization (MR) is a methodological approach to addressing reverse causation and confounding within observational studies, based on long-established 'instrumental variable' (IV) principles¹¹. MR exploits the random assortment of alleles through meiotic cell division at conception, and can be thought of as a natural experiment that generates conditions equivalent to randomised controlled trials, where randomised treatment arms are analogous to randomly assigned genetic subgroups¹²⁻¹⁵. At the population level, individuals defined by specific genotypes should on average only differ with respect to that genotype and its phenotypic consequences if certain IV assumptions are met. These assumptions are that the instrument is: i) robustly associated with the exposure it is acting as a proxy for; ii) independent of confounders; and iii) independent of the outcome conditional on the exposure (i.e. 'no pleiotropy' where a single locus influences the outcome through biological pathways that are independent of the exposure of interest)¹⁶. If these assumptions can be shown to have been met, then the genetic polymorphism can be used in an IV framework (i.e. MR) to provide an unconfounded and unbiased estimate of the causal association between the potentially modifiable risk factor and outcome of interest¹⁷. An extension of this methodology - two-sample MR - derives estimates for the required genotypeexposure and genotype-outcome associations from separate and non-overlapping samples of the same representative population^{13,18}. Two-sample approaches exploit the rapidly growing availability of summary data from large consortia of genome-wide association studies (GWASs) and thus allow for greater sample sizes to improve statistical power¹⁹.

49 The aim of this study was to systematically apply two-sample MR analyses to appraise the evidence of a causal 50 link between previously reported lifestyle and anthropometric risk factors with overall and aggressive prostate 51 cancer.

Methods

Selection of risk factors

- 55 The World Cancer Research Fund (WCRF) Continuous Update Project (CUP) is a rigorous and systematic
- 56 synthesis of the global scientific literature on diet, weight and physical-activity in relation to prostate cancer
- 57 risk, based largely on observational epidemiological studies. We selected all potential risk factors for prostate
- 58 cancer identified by the WCRF CUP reported in 20188.

Defining genetic instruments

- 60 We searched for each of the risk factors included in the WCRF report, using exact wording and synonyms, in
- 61 both the GWAS catalog (www.ebi.ac.uk/gwas) and the MR-Base repository of full GWAS association statistics
- 62 (www.mrbase.org)²⁰. This was done to identify any studies that reported associations between single nucleotide
- polymorphisms (SNPs) and the specific risk factor of interest²⁰. Further details of how we defined the genetic 63
- 64 instruments and their properties are provided in the supplemental material, including; selection criteria for the
- 65 SNP(s) to proxy each risk factor, how we standardised the beta coefficient and standard error (SE) for each
- 66 SNP-exposure association, and how we calculated the proportion of variance (R^2) in the risk factor explained by
- 67 the SNP(s), the strength of the instrument represented by the F-statistic, and the power to detect an odds ratio
- 68 (OR) of 1.2 (or conversely a protective OR of 0.80).

Outcome trait

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- 70 GWAS results for prostate cancer were obtained from fixed-effects meta-analyses based on individuals of
- European ancestry in the PRACTICAL and GAME-ON/ELLIPSE consortia (PRACTICAL: Prostate Cancer 71
- 72 Association Group to Investigate Cancer-Associated Alterations in the Genome; GAME-ON: Genetic
- 73 Associations and Mechanisms in Oncology; ELLIPSE: Elucidating Loci Involved in Prostate Cancer
- 74 Susceptibility). The summary data were derived from a GWAS of overall prostate cancer on 79,148 cases and
- 75 61,106 controls²¹, and a GWAS of aggressive prostate cancer involving 15,167 cases and 58,308 controls²¹.
- 76 Aggressive prostate cancer was defined as Gleason score ≥8, PSA >100 ng/mL, metastatic disease (M1), or
- 77 death from prostate cancer.

Two-sample MR analysis

- 79 We used the inverse-variance weighted (IVW) method as our primary MR analytical approach. The IVW 80 method estimates the effect of the exposure on the outcome from the slope of the relationship between β_{XG} 81 (SNP-exposure association) and β_{YG} (SNP-outcome association). This approach performs an inverse variance weighted meta-analysis of each Wald ratio²², effectively treating each SNP instrumenting a specific risk-factor 82 83 as a valid natural experiment. We used a random effects IVW model by default, unless there was
- 84 underdispersion in the causal estimates between SNPs, in which case a fixed effects model was used. The 85 estimates from the random and fixed effects IVW models are the same but the variance for the random effects
- 86 model is inflated to take into account heterogeneity between SNPs. For risk-factors that only had one SNP available as the instrument, we used the Wald ratio, which is equivalent to β_{YG} , β_{XG} (where Y= outcome, G= 87
- 88 gene and X=exposure).
- In sensitivity analyses, we applied weighted median²³, weighted mode²⁴ and MR-Egger regression²⁵ methods. 89
- 90 The weighted median has the advantage that only half the SNPs need to be valid instruments (i.e. exhibiting no
- 91 horizontal pleiotropy, no association with confounders, and a robust association with the exposure) for the 92 causal effect estimate to be unbiased. The mode-based estimator clusters the SNPs into groups based on
- 93 similarity of causal effects, and returns the causal effect estimate based on the cluster that has the largest number
- 94 of SNPs. The weighted mode introduces an extra element similar to IVW and the weighted median, weighting
- 95 each SNP's contribution to the clustering by the inverse variance of its outcome effect.
- 96 The MR-Egger method is similar to the IVW approach but relaxes the 'no horizontal pleiotropy' assumption.
- 97 MR-Egger regression allows a non-zero intercept in the relationship between multiple SNP-outcome and SNP-
- 98 exposure associations, where the intercept provides a formal statistical test for the presence of directional (bias
- 99 inducing) pleiotropy. The slope of the MR-Egger regression between multiple SNP-outcome and SNP-exposure
- 100 associations can be considered as an unbiased causal effect between the risk factors and prostate cancer,
- 101 assuming any horizontal pleiotropic effects are not correlated with the SNP-exposure effects (strength of the
- 102 instrument). The MR pleiotropy residual sum and outlier test (MR-PRESSO) was also implemented to identify
- outlying genetic variants and analyses were re-run after excluding these variants²⁶. Violations of the MR 'no horizontal pleiotropy' assumption were also assessed by visual inspection of funnel²⁷, forest, scatter and leave-103
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- one-out plots, and tests of heterogeneity²⁸ between SNPs making up a multi-allelic instrument²⁰ 105
- 106 All analyses were conducted using the TwoSampleMR and MRInstruments R packages, curated by MR-Base,
- 107 www.mrbase.org.

108 Replication

- 109 The risk factors that showed suggestive evidence of association (p<0.05) with overall prostate cancer were
- 110 assessed for replication among men in the UK Biobank prospective cohort of 7,844 prostate cancer cases and
- 111 204.001 controls, using two sample MR. The information on prostate cancer diagnosis was obtained from
- 112 National Cancer Registries, UK (http://biobank.ndph.ox.ac.uk/showcase/field.cgi?id=40006), based on ICD10
- 113 code for prostate cancer C61. The GWAS results of this study are available on www.mrbase.org.

Results

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There were 57 potential risk factors for prostate cancer considered by the WCRF 2018 report8(Supplementary 116 117

File 1; Table S1 and Table S2). The WCRF reported strong evidence that being obese (BMI, waist

118 circumference and WHR) increases the risk of aggressive prostate cancer (Supplementary File 1; Table S2) and 119

height increases the risk of prostate cancer (Supplementary File 1; Table S1 and S2). There was limited

120 evidence that consumption of dairy products, diets high in calcium, and low plasma alpha-tocopherol and low 121

plasm selenium concentrations increased prostate cancer risk. The evidence was too weak to draw any

122 conclusions for the remaining risk factors. Of these 57 exposures, 22 could be analysed using MR because they 123

had at least one SNP that was strongly associated with them (Table 1).

Mendelian randomization results

Of the 22 potential risk factors examined in our study, only four showed evidence of an association with overall prostate cancer risk (Figure 1, Supplementary File 1; Table S3). Physical-activity, assessed as 'average acceleration' (OR per SD change: 0.49; 95% CI: 0.33, 0.72; P=0.0003), serum iron levels (OR: 0.92; 95% CI: 0.86, 0.98; P=0.007), and BMI (OR: 0.90; 95% CI: 0.84, 0.97; P=0.003) were inversely associated with overall prostate cancer risk. Circulating mono-unsaturated fat (OR: 1.11; 95% CI: 1.02, 1.20; P=0.02) was positively associated with overall prostate cancer risk. Compared to results from observational studies⁸ for overall prostate cancer risk (highest versus lowest total physical-activity; risk ratio (RR): 0.97; 95% CI: 0.9, 1.04), the estimate obtained by MR for the physical-activity measure (average acceleration) was much more strongly supportive of a protective effective (Figure 1). The WCRF reported strong evidence of association for increased body fatness (marked by BMI (RR: 1.08; 95% CI: 1.04, 1.12), waist circumference (RR: 1.12; 95% CI: 1.04, 1.21) and WHR (RR: 1.15; 95% CI: 1.03, 1.28)) with aggressive prostate cancer risk but these results were inconsistent with our MR results, which found evidence of protective association of BMI with overall prostate cancer risk. In the WCRF report (2018), the observational analysis reported no association between intake of mono-unsaturated fat and overall prostate cancer risk (RR: 1.00; 95% CI: 0.99, 1.01) but our MR-analysis found a positive association between circulating mono-unsaturated fat and overall prostate cancer risk.

139 140 None of the risk factors we examined showed strong evidence of association with aggressive prostate cancer 141 although height showed weak evidence of increasing risk; OR: 1.07; 95% CI: 1.01, 1.15; P=0.03 (Figure 2, 142 Supplementary File 1; Table S4). The observational studies also reported positive association of height with 143 overall prostate cancer (OR: 1.04.; 95% CI: 1.03, 1.05; P=1.3x10⁻¹⁵) and aggressive prostate cancer⁸ (OR: 1.04; 144 95% CI: 1.02, 1.06; P=6.4×10⁻⁵). However, our MR analysis did not find evidence of association between height 145 and overall prostate cancer risk. There was weak evidence for iron, and average acceleration, with effect-146 estimates being similar to those observed for overall prostate cancers. In the WCRF report, observational studies 147 have reported positive effects of dairy products, calcium and low selenium concentration on overall prostate 148 cancer, but we did not find strong evidence of associations with these in our MR analyses (Figure 1). However, 149 the confidence intervals (CIs) for these risk factors were overlapping between observational and MR analyses. 150 The power to detect the observationally reported effect size for these risk factors was >74%). Low alpha-151 tocopheral concentration was positively associated with prostate cancer risk in WCRF report but due to lack of 152 an instrument it was not possible to conduct MR analyses.

There were only two SNPs available for the MR analysis of physical-activity (average acceleration) so we could not perform extensive sensitivity analyses. The direction of association for both SNPs was consistent (Supplementary File 1; Table S5) and the p-value for heterogeneity test was 0.99. These SNPs were on different chromosomes so represent independent associations. After exploring MRBASE-PheWAS (http://phewas.mrbase.org/), we found that these two SNPs were associated with anthropometric traits other than physical activity (Supplementary File 1; Table S6). The results for the effect of serum iron, and increasing BMI on overall prostate cancer were consistent across the various sensitivity analyses (Supplementary File 1; Figure S1-S3). The test for directional horizontal pleiotropy by MR-Egger (serum iron: intercept: 0.0005, P=0.97 and BMI: intercept: -0.0002, P=0.89) didn't find evidence of pleiotropy. Mono-unsaturated fat showed results in the opposite direction using MR-Egger regression (OR: 0.85; 95% CI: 0.55, 1.31; P=0.51) compared to other MRmethods (Supplementary File 1; Figure S1 and S4). However, there was no strong evidence for directional pleiotropy for mono-unsaturated fat using MR-Egger (test for directional horizontal pleiotropy by MR-Egger: intercept: 0.04, P=0.31). The MR-Egger tests for both iron and mono-unsaturated fat had low power due to the small number of SNPs used, however all 5 SNPs for iron and 4/5 SNPs for mono-unsaturated fats showed associations with prostate cancer in the same direction (Supplementary File 1; Table S7-S8). At MRBASE-PheWAS, five SNPs of iron were associated with haemoglobin concentration and blood cells count (Supplementary File 1; Table S9) and SNPs of mono-unsaturated fat were associated with lipids (Supplementary File 1; Table S10). The MR results for single SNP analyses of BMI (overall prostate cancer) and height (aggressive prostate cancer) are provided in Supplementary File 1; Table S11-S12.

Replication

The MR analyses were repeated using prostate cancer summary data generated from UK Biobank for physical-activity, iron, BMI, and mono-unsaturated fat (Figure 3). The point-estimates showed consistent directions of association for physical-activity (OR: 0.37; 95% CI: 0.13, 1.06; P=0.07), and BMI (OR: 0.83, 95% CI: 0.74, 0.94; P=0.002). The point-estimates for iron (OR: 1.07; 95% CI: 0.96, 1.20; P=0.24) and mono-unsaturated fat (OR: 0.89; 95% CI: 0.73, 1.07; P=0.20) were in the opposite direction, but the power to detect an effect with these risk factors in UK Biobank was low and confidence intervals for the replication analysis overlapped with our main analysis for all risk factors.

Discussion

We found consistent evidence that physical-activity (assessed as 'average accelerations', but not other measures of physical activity), and BMI have an inverse effect on overall prostate cancer risk. There was also evidence of an inverse effect of iron and a positive effect of circulating mono-unsaturated fat in our initial analyses, the effect sizes were in opposite direction in UK Biobank study however power for the replication analysis was low and their CIs overlapped with the PRACTICAL study. There was weak evidence for physical-activity (average accelerations), and iron had a similar effect on aggressive prostate cancer to that seen for overall prostate cancer. We found little evidence that any of the other risk factors studied have a causal role in overall or aggressive prostate cancer, but height showed a positive association with aggressive prostate cancer. The CIs of MR results overlapped with those seen in the observational analyses for all risk factors except for average acceleration and mono-unsaturated fat (overall prostate cancer). In fact, confidence intervals for our MR analysis of aggressive prostate cancer were wide and the power for these analyses was low for many risk factors.

The WCRF report meta-analysed self-reported physical-activity which was assessed in different studies by various methods (i.e. occupational, recreational and total physical-activity) as highest versus lowest level of total physical-activity, a relatively crude dichotomy that may have masked associations. Our MR analysis which proxied fraction accelerations was the most similar to the observational analyses and we did not find evidence of an association of this measure with prostate cancer. However, we did find an association with average acceleration, which is a different measure and could be high if someone is consistently engaging in light-intensity activity across most of the waking day (vs lots of sitting and a 30 minutes bout of MVPA). Indeed, there is little genetic or phenotypic correlation between the two measures in the UKBiobank population²⁹. The mechanism for our association between average accelerations and prostate cancer is unclear, although this could be through improved insulin sensitivity or reduced insulin-like growth factor-1 (IGF-1)³⁰, reduced levels of testosterone and dihydrotestosterone³¹, alterations in the antioxidant defence system³¹, or improvements in the immune system through enhanced natural killer cell activity³¹.

Whilst not consistent with observational analyses, our results for BMI are concordant with other lines of evidence. We have previously shown weak evidence that higher BMI is associated with a reduced prostate cancer risk, in a smaller sample from the PRACTICAL consortium³². A study examined childhood and adult body size in relation to total incident prostate cancer in a prospective cohort of 47,491 US men³³. High BMI at age 21 was inversely associated with total prostate cancer, with fatal and advanced disease. The association for late adult BMI was more complex and differed by age this could represent confounding by other factors for example physical activity, and diet etc.

Despite showing a protective effect of BMI on prostate cancer risk, we did not find any strong evidence of waist circumference or WHR with risk. However, the results for all these risk factors were in the same direction. The point estimates for BMI, waist circumference and WHR were in the opposite direction to findings from observational results for aggressive prostate cancer. For aggressive prostate cancer, power calculations suggested that we would have good power to detect an effect of BMI (96%) but very low power (33% and 23%) to detect an OR of 1.20 (or, conversely a protective OR of at least 0.80) for waist circumference and WHR respectively. An increased estrogen production has been observed in obese men³⁴. The sensitivity of prostate cancer to sex hormones has been exploited for therapeutic purposes for many years. Androgen-deprivation therapy is a common treatment in prostate cancer³⁵, as is the therapeutic use of estrogen for patients with metastatic prostate cancer in both the US³⁶ and Europe³⁷, though not popular due to cardiovascular and other side effects. Hence the clinical prediction would be that obesity would be associated with a lower risk of prostate cancer. However, we cannot rule out detection bias³⁸ arising from delayed diagnosis and therefore a more advanced stage at diagnosis in obese men. This could arise due to lower accuracy of digital rectal examination in obese men or lower PSA values caused by obesity-related traits.

The analysis of iron as a risk factor for prostate cancer found that the evidence was too limited to draw conclusions in the WCRF Second Expert Report and this was not updated in the Third Expert Report due to a lack of new evidence^{8,39}. Population studies that have examined the associations between serum iron and cancer outcomes are limited and have reported discordant findings. A prospective cohort study with 15-16 years follow up time reported higher serum iron concentrations increased non-skin cancer risk overall but conversely, in men,

higher serum iron concentrations decreased the risk of non-skin cancer⁴⁰. A Swedish cohort reported increased serum iron concentrations were not associated with overall cancer risk except for a slightly higher risk of postmenopausal breast cancer⁴¹. Previous findings have suggested that, although higher circulating iron concentrations may potentially increase the risk of cancer in women it may be protective against cancer in men⁴⁰ which is in accordance with our MR findings. Although again we were unable to replicate these findings in UK Biobank due to low power.

Our MR-analysis investigated the association of circulating levels of mono-unsaturated fat on prostate cancer, circulating levels of this nutrient have been shown to be poorly correlated with mono-unsaturated fat intake measured by questionnaire⁴². The inclusion of an objective measure in our MR-analysis versus questionnaire data in observational studies could be the reason for the discordant results. Although discordance could also be due to negative confounding in the former studies resulting in the effect estimate being closer to the null (e.g., through other dietary, lifestyle, or molecular factors. Instruments for mono-unsaturated fat are correlated with instruments for saturated fat and other lipids. It could be possible that saturated fat or other lipids cause prostate cancer and this result might reflect that. More research is needed into the independence of mono-unsaturated fat of other lipid factors.

Our MR results showed positive association between height and aggressive prostate cancer and the results was consistent with observational studies. Adult height is associated with the rate of growth during fetal life and childhood childhood childhood may affect the age of sexual maturity. These processes are mediated by changes in the hormonal microenvironment that may have both short and long term effects on circulating levels of growth factors, insulin and other endocrine or tissue specific mediators that may impact cancer risk 45.

The results from replication analyses were compatible with the main findings for BMI, and physical activity (albeit with wide confidence intervals for physical activity), which increases the likelihood that these findings are real. The UK Biobank sample size was however, smaller, and the effects were estimated with less precision and whilst the results for iron and mono-unsaturated fats did not appear to replicate, due to low power in the replication study we cannot rule out causal effects of these nutrients.

This study's major strength is the use of MR, which is less susceptible to problems of measurement error, confounding and reverse causation in comparison to conventional observational studies. The use of two-sample MR enabled the use of the largest GWAS of prostate cancer²¹ to date. We were also able to make use of the largest GWASs on the risk factors of interest, to increase the precision of the SNP-exposure estimates, which should reduce impact of weak instruments bias, which in turn increase statistical power assuming the SNP-exposure estimates are unbiased and risk factor/outcome samples come from the same population.

The study also has some limitations. We had only two SNPs for physical-activity assessed as average acceleration. If there were many independent SNPs available the causal inference could have been strengthened because a) each variant represents an independent natural experiment, and a more precise overall causal estimate (i.e. tighter CIs) can be obtained by meta-analysing the single estimates from each instrument; and b)potential bias arising from the violation of the assumptions can be detected or corrected by evaluating the consistency of effects across instruments ^{16,24,28,46,47}. For many of the risk factors reported in the WCRF 2018 report, for example alpha-tocopheral, vitamin A, vitamin C etc, we did not find genetic instruments to conduct MR analyses. For the majority of the risk factors in overall prostate cancer, MR analyses were sufficiently powered to detect effect sizes of a modest magnitude (OR of 1.20 or 0.80) except for physical-activity traits (overall acceleration average, fraction of accelerations >425 milli-gravities, and sedentary behaviour), thus failure to detect strong evidence of effects for these risk factors could be due to low power to detect smaller effect sizes. Further identification of independent genetic variants that influence these risk factors will help to improve statistical power for future analyses.

In conclusion, we found evidence that physical-activity, serum iron, and BMI may be causally and inversely related to and circulating mono-unsaturated fat and height may be causally and positively related to, prostate cancer risk. Further studies should investigate the mechanisms by which these factors may lead to prostate cancer and investigate the potential to intervene to reduce risk.

Contributors

SL, RMM and NK conceived and designed the study. NK conducted the analyses. NK wrote the manuscript with input from all authors. Correspondence and material requests should be addressed to SL (s.j.lewis@bristol.ac.uk).

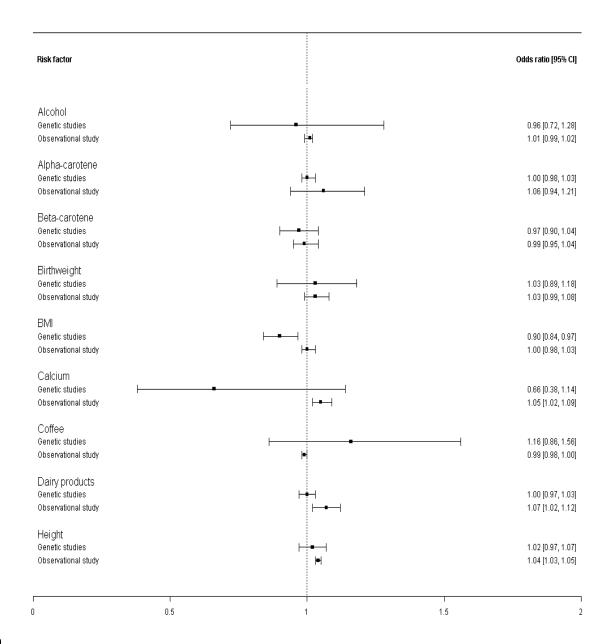
Declaration of interests

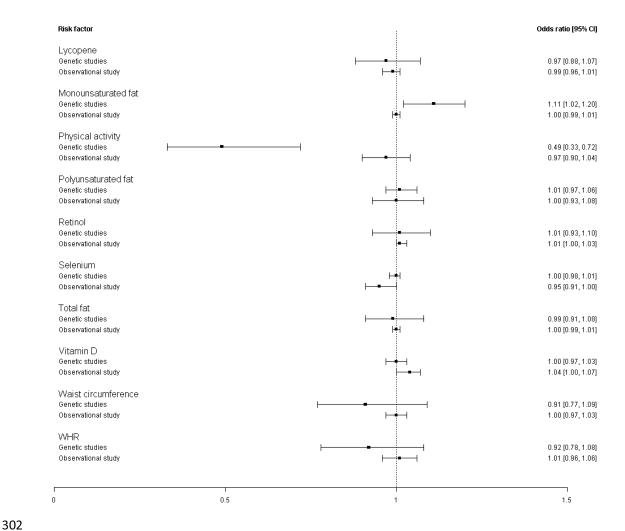
None

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Figure 1. Comparison between observational and MR estimates for the risk factors and overall prostate cancer risk





Only those risk factors are plotted whose observational estimates were reprted in WCRF Second or Third Exper Report^{8,39}.

The effect estimate for physical activity (Genetic studies) represent average acceleration.

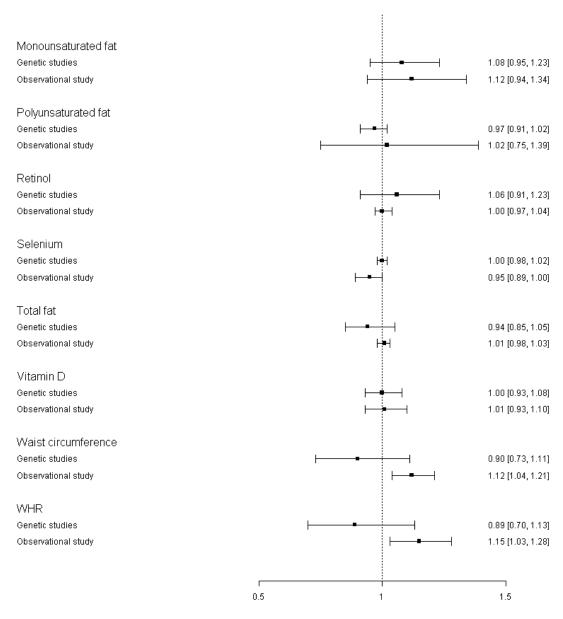
Figure 2. Comparison between observational and MR estimates for the risk factors and aggressive prostate cancer risk

Risk factor		Odds ratio [95% CI]
Alcohol		
Genetic studies	-	0.94 [0.66, 1.35]
Observational study	 • 	1.00 [0.96, 1.03]
Beta-carotene		
Genetic studies	<u> </u>	0.94 [0.83, 1.06]
Observational study	` 	0.97 [0.85, 1.12]
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Birthweight		
Genetic studies	├	1.12 [0.92, 1.35]
Observational study	 ■ 	1.09 [0.97, 1.22]
BMI		
Genetic studies	 ■ 	0.94 [0.85, 1.03]
Observational study	 - 	1.08 [1.04, 1.12]
Calcium		
Genetic studies	-	0.94 [0.36, 2.42]
Observational study	├	1.02 [0.93, 1.12]
Dairy products		
Genetic studies	H ≢ -l	0.99 [0.94, 1.05]
Observational study	├ ■	0.97 [0.91, 1.05]
11.5.14		
Height		4 07 14 04 4 4 5
Genetic studies		1.07 [1.01, 1.15]
Observational study	= 1	1.04 [1.02, 1.06]
Lycopene		
Genetic studies	 ■	0.96 [0.81, 1.14]
Observational study	 	0.98 [0.93, 1.03]
0	1 2	3

Risk factor

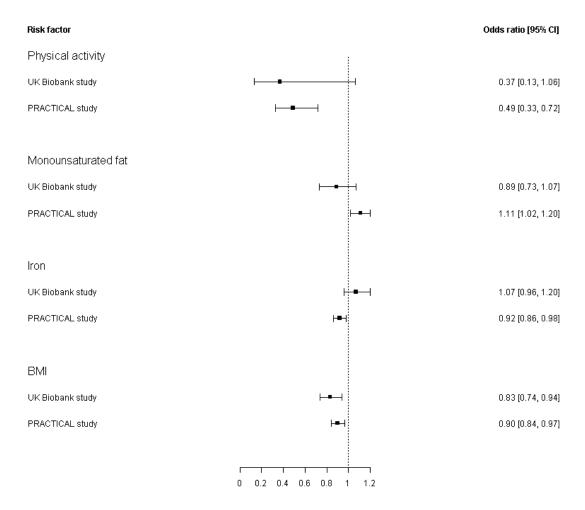
311 312

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Only those risk factors are plotted whose observational estimates were reported in WCRF Second or Third Exper Report^{8,39}.

Figure 3. Comparison of MR estimates (OR) from the main and replication analyses for the risk factors that showed evidence of association (p<0.05) with overall prostate cancer in PRACTICAL



Risk factor	PubMed ID	N	# SNPs1	# SNPs	² Units	SD	\mathbb{R}^2	F	Power ¹	Power ²
Anthropometrics and other measures			-11		"				l	
Birth weight	27680694	143677	46	45	g	1	1.69	53.54	>99	53
BMI	30124842	681275	535	525	kg/m ²	1	5.66	76.39	>99	96
Height	25282103	253288	433	426	m	1	12.01	79.71	>99	>99
Waist circumference	25673412	232101	45	44	cm	1	0.95	49.60	94	33
WHR	25673412	212244	31	28	ratio	1	0.65	44.95	83	23
Circulating macro- and micro-nutrients									1	
Sugar/Sucrose (glucose)	22885924	133010	15	15	mmol/l	0.73	0.52	46.65	74	21
Monounsaturated fat	27005778	13535	5	5	mmol/l	1	2.01	55.62	>99	61
Polyunsaturated fat	27005778	13549	19	19	mmol/l	1	13.79	113.93	>99	>99
Total fat	27005778	13505	12	12	mmol/l	1	4.62	54.51	>99	92
Alpha-carotene	28002826	433	3	3	μmol/l	0.23	22.08	40.53	>99	>99
Beta-carotene	19185284	3932	1	1	μmol/l	0.67	2.63	106.36	>99	73
Calcium	24068962	61079	5	5	mg/dl	1	0.65	79.85	83	25
Iron	25352340	72958	5	5	μmol/l	1	2.20	328.85	>99	66
Lycopene	26861389	441	1	1	N/A	N/A	8.34	39.93	>99	>99
Phosphorous	20558539	16264	4	4	mg/dl	0.49	1.52	62.82	99	50
Retinol	21878437	5006	2	2	μg/1	0.22	1.37	34.79	99	46
Selenium	25343990	9639	2	2	μg/L	0.18	1.44	70.33	>99	48
Vitamin D	23393431	42024	4	4	ng/ml	10	2.35	253.15	>99	68
Consumption of foods and drinks										
Alcohol	30643251	941280	77	76	drinks per week	1	0.56	68.35	77	22
Coffee	25288136	91462	4	4	cups/day	1.96	0.45	103.79	68	19
Dairy products (milk intake)	29071499	74241	1	1	glasses/week	1.7	0.52	388.87	74	21
Physical-activity	1		-1	1				I		
Average acceleration	29899525	91084	2	2	milli gravities (mg)	8.14	0.10	44.68	20	7
Sedentary behaviour	30531941	91105	2	2	proportion units	1	0.08	34.91	17	6
Fraction accelerations	29899525	90667	1	1	inverse- normalized fraction of accelerations	1	0.04	39.06	11	5
Sleep duration	30531941	91105	7	7	proportion units	1	0.38	49.29	60	16

Table 1. Details of the instruments used to proxy risk factors for prostate cancer risk

BMI = body mass; WHR = waist-hip ratio; N is the sample size for the GWAS used to define the instruments; # SNPs represents the number of SNPs used within the instrument for each risk factor after clumping, harmonization and extraction of data from a GWAS of prostate cancer (# SNPs¹ for overall prostate cancer risk and # SNPs² for aggressive prostate cancer risk); Units and SD represent the analysis scale and standard deviation scale for betas and SE of SNPs for each risk factor(1 = the GWAS results were already on SD scale otherwise the SD of population mean), respectively; R² represents the variance explained in the risk factor by the instrument; F indicates strength of the instrument used for each risk factor (a strong instrument is sometimes defined as an F-statistic >10); Power¹ represents the power to detect an odds ratio of 1.2 for an association of the risk factor with overall prostate cancer at a significance level (P) of 0.05; Power² represents the power to detect an odds ratio of 1.2 for an association of the risk factor with aggressive prostate cancer at a significance level (P) of 0.05.

References

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- 333 1. Stewart B, Wild CP. International Agency for Research on Cancer, World Cancer Report 2014,
- ed. International Agency for Research on Cancer 2014.
- 335 2. Ferlay J, Soerjomataram I, Dikshit R, et al. Cancer incidence and mortality worldwide:
- sources, methods and major patterns in GLOBOCAN 2012. *International journal of cancer* 2015;
- 337 **136**(5): E359-86.
- 338 3. Center MM, Jemal A, Lortet-Tieulent J, et al. International variation in prostate cancer
- incidence and mortality rates. European urology 2012; **61**(6): 1079-92.
- 4. Wong MC, Goggins WB, Wang HH, et al. Global Incidence and Mortality for Prostate Cancer:
- Analysis of Temporal Patterns and Trends in 36 Countries. European urology 2016; 70(5): 862-74.
- 342 5. Haenszel W, Kurihara M. Studies of Japanese migrants. I. Mortality from cancer and other
- diseases among Japanese in the United States. *Journal of the National Cancer Institute* 1968; **40**(1):
- 344 43-68.
- 345 6. McCredie M, Williams S, Coates M. Cancer mortality in East and Southeast Asian migrants to
- 346 New South Wales, Australia, 1975-1995. British journal of cancer 1999; **79**(7-8): 1277-82.
- 347 7. Gann PH. Risk factors for prostate cancer. Reviews in urology 2002; 4 Suppl 5: S3-s10.
- 348 8. **Diet, nutrition, physical activity and prostate cancer.** World Cancer Research fund/American
- 349 Institute For Cancer Research Continuous Update Project Expert Report 2018.
- 350 9. Lawlor DA, Davey Smith G, Kundu D, Bruckdorfer KR, Ebrahim S. Those confounded vitamins:
- 351 what can we learn from the differences between observational versus randomised trial evidence?
- 352 *Lancet (London, England)* 2004; **363**(9422): 1724-7.
- 353 10. Davey Smith G, Ebrahim S. Epidemiology--is it time to call it a day? *International journal of*
- 354 *epidemiology* 2001; **30**(1): 1-11.
- 355 11. Angrist JD, Imbens GW, Rubin DB. Identification of Causal Effects Using Instrumental
- 356 Variables. Journal of the American Statistical Association 1996; 91(434): 444-55.
- 357 12. Davey Smith G ES. "Mendelian randomisation": can genetic epidemiology contribute to
- understanding environmental determinants of disease? . Int J Epidemiology 2003; 32: 1-22.
- 359 13. Davey Smith G HG. Mendelian randomization: genetic anchors for causal inference in
- 360 epidemiology studies. *Hum Mol Genet* 2014.
- 361 14. Lawlor DA HR, Sterne JAC, Timpson NJ, Davey Smith G. Mendelian randomization: using
- 362 genes as instruments for making causal inferences in epidemiology. Stat Med 2008; 27: 1133-63.
- 363 15. Haycock PC, Burgess S, Wade KH, Bowden J, Relton C, Davey Smith G. Best (but oft-
- forgotten) practices: the design, analysis, and interpretation of Mendelian randomization studies.
- 365 The American Journal of Clinical Nutrition 2016.
- 366 16. Bowden J, Davey Smith G, Haycock PC, Burgess S. Consistent Estimation in Mendelian
- 367 Randomization with Some Invalid Instruments Using a Weighted Median Estimator. Genetic
- 368 *epidemiology* 2016; **40**(4): 304-14.
- 369 17. Timpson NJ, Wade KH, Smith GD. Mendelian randomization: application to cardiovascular
- disease. Current hypertension reports 2012; **14**(1): 29-37.
- 371 18. Pierce BL, Burgess S. Efficient Design for Mendelian Randomization Studies: Subsample and
- 372 2-Sample Instrumental Variable Estimators. American Journal of Epidemiology 2013; 178(7): 1177-
- 373 84.

- 374 19. Yarmolinsky J, Wade KH, Richmond RC, et al. Causal Inference in Cancer Epidemiology: What
- 375 Is the Role of Mendelian Randomization? Cancer epidemiology, biomarkers & prevention: a
- 376 publication of the American Association for Cancer Research, cosponsored by the American Society of
- 377 Preventive Oncology 2018; **27**(9): 995-1010.
- 378 20. Hemani G, Zheng J, Elsworth B, et al. The MR-Base platform supports systematic causal
- inference across the human phenome. *eLife* 2018; **7**.
- 380 21. Schumacher FR, Al Olama AA, Berndt SI, et al. Association analyses of more than 140,000
- men identify 63 new prostate cancer susceptibility loci. *Nature genetics* 2018; **50**(7): 928-36.
- 382 22. Johnson T. Efficient calculation for Multi-SNP genetic risk scores. American Society of Human
- 383 Genetics Annual Meeting 2012.
- 384 23. Bowden J, Davey Smith G, Haycock PC, Burgess S. Consistent Estimation in Mendelian
- 385 Randomization with Some Invalid Instruments Using a Weighted Median Estimator. Genetic
- 386 *Epidemiology* 2016; **40**(4): 304-14.
- 387 24. Hartwig FP, Davey Smith G, Bowden J. Robust inference in summary data Mendelian
- randomization via the zero modal pleiotropy assumption. International journal of epidemiology
- 389 2017; **46**(6): 1985-98.
- 390 25. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments:
- 391 effect estimation and bias detection through Egger regression. International Journal of Epidemiology
- 392 2015.
- 393 26. Verbanck M, Chen CY, Neale B, Do R. Detection of widespread horizontal pleiotropy in causal
- 394 relationships inferred from Mendelian randomization between complex traits and diseases. Nature
- 395 *genetics* 2018; **50**(5): 693-8.
- 396 27. Sterne JA, Sutton AJ, Ioannidis JP, et al. Recommendations for examining and interpreting
- funnel plot asymmetry in meta-analyses of randomised controlled trials. BMJ (Clinical research ed)
- 398 2011: **343**: d4002
- 399 28. Bowden J, Del Greco MF, Minelli C, Davey Smith G, Sheehan N, Thompson J. A framework for
- 400 the investigation of pleiotropy in two-sample summary data Mendelian randomization. Statistics in
- 401 medicine 2017; **36**(11): 1783-802.
- 402 29. Klimentidis YC, Raichlen DA, Bea J, et al. Genome-wide association study of habitual physical
- 403 activity in over 377,000 UK Biobank participants identifies multiple variants including CADM2 and
- 404 APOE. International journal of obesity (2005) 2018; **42**(6): 1161-76.
- 405 30. Ngo TH, Barnard RJ, Tymchuk CN, Cohen P, Aronson WJ. Effect of diet and exercise on serum
- 406 insulin, IGF-I, and IGFBP-1 levels and growth of LNCaP cells in vitro (United States). Cancer causes &
- 407 control: CCC 2002; 13(10): 929-35.
- 408 31. Friedenreich CM, Thune I. A review of physical activity and prostate cancer risk. Cancer
- 409 causes & control: CCC 2001; **12**(5): 461-75.
- 410 32. Davies NM, Gaunt TR, Lewis SJ, et al. The effects of height and BMI on prostate cancer
- 411 incidence and mortality: a Mendelian randomization study in 20,848 cases and 20,214 controls from
- the PRACTICAL consortium. Cancer causes & control: CCC 2015; 26(11): 1603-16.
- 413 33. Moller E, Wilson KM, Batista JL, Mucci LA, Balter K, Giovannucci E. Body size across the life
- 414 course and prostate cancer in the Health Professionals Follow-up Study. International journal of
- 415 cancer 2016; **138**(4): 853-65.
- 416 34. Schneider G, Kirschner MA, Berkowitz R, Ertel NH. Increased estrogen production in obese
- 417 men. The Journal of clinical endocrinology and metabolism 1979; **48**(4): 633-8.

- 418 35. Wong YN, Ferraldeschi R, Attard G, de Bono J. Evolution of androgen receptor targeted
- therapy for advanced prostate cancer. *Nature reviews Clinical oncology* 2014; **11**(6): 365-76.
- 420 36. Mohler JL, Kantoff PW, Armstrong AJ, et al. Prostate cancer, version 2.2014. Journal of the
- 421 National Comprehensive Cancer Network: JNCCN 2014; 12(5): 686-718.
- 422 37. Heidenreich A, Bastian PJ, Bellmunt J, et al. EAU guidelines on prostate cancer. Part II:
- 423 Treatment of advanced, relapsing, and castration-resistant prostate cancer. European urology 2014;
- 424 **65**(2): 467-79.
- 425 38. Fowke JH, Motley S, Dai Q, Concepcion R, Barocas DA. Association between biomarkers of
- 426 obesity and risk of high-grade prostatic intraepithelial neoplasia and prostate cancer--evidence of
- 427 effect modification by prostate size. *Cancer letters* 2013; **328**(2): 345-52.
- 428 39. World Cancer Research Fund/ American Institute of Cancer Research. Food, Nutrition,
- 429 Physical Activity, and the Prevention of Cancer: a Global Perspective. Washington DC: AICR, 2007
- 430 (Second Expert Report). 2007.
- 431 40. Chua AC, Knuiman MW, Trinder D, Divitini ML, Olynyk JK. Higher concentrations of serum
- iron and transferrin saturation but not serum ferritin are associated with cancer outcomes. The
- 433 American journal of clinical nutrition 2016; 104(3): 736-42.
- 434 41. Gaur A, Collins H, Wulaningsih W, et al. Iron metabolism and risk of cancer in the Swedish
- 435 AMORIS study. Cancer causes & control: CCC 2013; 24(7): 1393-402.
- 436 42. Marchioni DM, de Oliveira MF, Carioca AAF, et al. Plasma fatty acids: Biomarkers of dietary
- 437 intake? Nutrition (Burbank, Los Angeles County, Calif) 2019; 59: 77-82.
- 438 43. Barker DJ, Thornburg KL. Placental programming of chronic diseases, cancer and lifespan: a
- 439 review. *Placenta* 2013; **34**(10): 841-5.
- 44. Rolland-Cachera MF. Rate of growth in early life: a predictor of later health? Advances in
- experimental medicine and biology 2005; **569**: 35-9.
- 442 45. Le Roith D, Bondy C, Yakar S, Liu JL, Butler A. The somatomedin hypothesis: 2001. Endocrine
- 443 reviews 2001; **22**(1): 53-74.
- 444 46. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments:
- effect estimation and bias detection through Egger regression. *International journal of epidemiology*
- 446 2015; **44**(2): 512-25.

- 447 47. Hemani G, Tilling K, Davey Smith G. Orienting the causal relationship between imprecisely
- measured traits using GWAS summary data. PLoS genetics 2017; 13(11): e1007081.