

Pan-Cancer Study Detects Novel Genetic Risk Variants and Shared Genetic Basis in Two Large Cohorts

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Abstract

Deciphering the shared genetic basis of distinct cancers has the potential to elucidate carcinogenic mechanisms and inform broadly applicable risk assessment efforts. However, no studies have investigated pan-cancer pleiotropy within single, well-defined populations. We undertook novel genome-wide association studies (GWAS) and comprehensive evaluations of heritability and pleiotropy across 18 cancer types in two large, population-based cohorts: the UK Biobank (413,870 European ancestry individuals; 48,961 cancer cases) and the Kaiser Permanente Genetic Epidemiology Research on Adult Health and Aging (66,526 European ancestry individuals; 16,001 cancer cases). The GWAS detected 29 novel genome-wide significant risk variants. In addition, numerous cancer sites exhibited clear heritability. Investigations of pleiotropy identified 12 cancer pairs exhibiting genetic correlations and 43 pleiotropic loci. We also detected 158 variants associated with multiple cancers, including 21 with positive associations for some cancers and negative associations for others. Our findings demonstrate widespread pleiotropy and offer insight into the complex genetic architecture of cross-cancer susceptibility.

Introduction

The global burden of cancer is substantial, with an estimated 18.1 million individuals diagnosed each year and approximately 9.6 million deaths attributed to the disease.¹ Efforts toward cancer prevention, screening, and treatment are thus imperative, but they require a more comprehensive understanding of the underpinnings of carcinogenesis than we currently possess. While studies of twins,² families,³ and unrelated populations^{4–6} have demonstrated substantial heritability and familial clustering for many cancers, the extent to which genetic variation is unique versus shared across different types of cancer remains unclear.

Genome-wide association studies (GWAS) of individual cancers have identified loci associated with multiple cancer types, such as 1q32 (*MDM4*)^{7,8}; 2q33 (*CASP8-ALS2CR12*)^{9,10}; 3q28 (*TP63*)^{11,12}; 4q24 (*TET2*)^{13,14}; 5p15 (*TERT-CLPTMIL*)^{9,12}; 6p21 (HLA complex)^{15,16}; 7p15¹⁷; 8q24^{12,18}; 11q13^{18,19}; 17q12 (*HNF1B*)^{18,20}; and 19q13 (*MERIT40*)²¹. In addition, recent studies have tested single nucleotide polymorphisms (SNPs) previously associated with one cancer to discover pleiotropic associations with other cancer types.^{22–25} Consortia, such as the Genetic Associations and Mechanisms in Oncology, have looked for variants and pathways shared by breast, colorectal, lung, ovarian, and prostate cancers.^{26–30} Comparable studies for other cancers—including those that are less common—have yet to be conducted.

In addition to individual variants, recent studies have evaluated genome-wide genetic correlations between pairs of cancer types.^{4–6} One evaluated 13 cancer types and found shared heritability between kidney and testicular cancers, diffuse large B-cell lymphoma (DLBCL) and osteosarcoma, DLBCL and chronic lymphocytic leukemia (CLL), and bladder and lung cancers.⁴

Another study of six cancer types found correlations between colorectal cancer and both lung and pancreatic cancers.⁵ In an updated analysis with increased sample size, the same group identified correlations of breast cancer with colorectal, lung, and ovarian cancers and of lung cancer with colorectal and head/neck cancers.⁶ While these studies provide compelling evidence for shared heritability across cancers, they lack data on several cancer types (e.g., cervix, melanoma, and thyroid). Moreover, previous analyses aggregated numerous studies with case-control sampling and heterogeneous populations recruited during different time periods.

Here, we present analyses of genome-wide SNP data with respect to 18 cancer types, based on 475,312 individuals of European ancestry from two large, independent, and contemporary cohorts unselected for phenotype – the UK Biobank (UKB) and the Kaiser Permanente Genetic Epidemiology Research on Adult Health and Aging (GERA) cohorts. We sought to detect novel risk SNPs and pleiotropic loci and variants and to estimate the heritability of and genetic correlations between cancer types. Leveraging the wealth of individual-level genetic and clinical data from both cohorts allowed us to comprehensively interrogate the shared genetic basis of susceptibility to different cancer types, with the ultimate goal of better understanding common genetic mechanisms of carcinogenesis and improving risk assessment.

Results

Genome-wide Association Analyses of Individual Cancers

We found 29 novel, independent variants that attained genome-wide significance at $P < 5 \times 10^{-8}$ upon meta-analysis (**Table 1**) and an additional 12 genome-wide significant variants that were only genotyped or imputed in one cohort (**Supplementary Table 1**). There were a total of 202 novel, independent variants with a suggestive $P < 1 \times 10^{-6}$ across the 18 cancer types in the set of SNPs either meta-analyzed or analyzed in just one cohort (**Supplementary Table 1**). In addition, we detected and confirmed 280 independent signals with $P < 1 \times 10^{-6}$ identified by previous GWAS with $P < 5 \times 10^{-8}$ (**Supplementary Table 2**).

Of the 29 novel genome-wide significant associations from the meta-analyses, 13 were in known susceptibility regions for the cancer of interest but independent of previously reported variants. The remaining 16 were in regions not previously associated with the cancer of interest in individuals of European ancestry. Twenty-one of the 29 novel SNPs were in regions previously associated with at least one of the other cancer types evaluated in this study.

Genome-Wide Heritability and Genetic Correlation

Array-based heritability estimates across cancers ranged from $h^2 = 0.04$ (95% CI: 0.00-0.13) for oral cavity/pharyngeal cancer to $h^2 = 0.26$ (95% CI: 0.15-0.38) for testicular cancer (**Table 2**). For some of the cancers, our array-based heritability estimates were comparable to twin- or family-based heritability estimates but were more precise. Several were also similar to array-based heritability estimates from consortia comprised of multiple studies. For example, our estimate for testicular cancer closely matches the previous family-based estimate of heritability ($h^2 = 0.25$; 95%

CI: 0.15-0.37),³ as well as a previous estimate of array-based heritability ($h^2=0.30$; 95% CI: 0.08-0.51).⁴ For lung cancer, our estimated heritability of $h^2=0.15$ (95% CI: 0.10-0.20) approaches the twin-based estimate of $h^2=0.18$ (95% CI: 0.00-0.42),² exceeds the array-based estimate from a study using the same methodology ($h^2=0.08$; 95% CI: 0.05-0.10),⁶ and is comparable to an earlier array-based estimate using individual-level data ($h^2=0.21$; 95% CI: 0.14-0.27).⁴ For rectal ($h^2=0.11$; 95% CI: 0.07-0.16) and bladder ($h^2=0.08$; 95% CI: 0.04-0.12) cancers, our heritability estimates are close to those from twin/family studies – $h^2=0.14$ (95%CI: 0.00-0.50)² and $h^2=0.07$ (95% CI: 0.02-0.11),³ respectively. One of our highest heritability estimates was observed for thyroid cancer ($h^2=0.21$; 95% CI: 0.09-0.33), a cancer that has not been evaluated in other array-based studies.

Among pairs of cancers, only colon and rectal cancers ($r_g=0.85$, $P=5.33 \times 10^{-7}$) were genetically correlated at a Bonferroni corrected significance threshold of $P=0.05/153=3.27 \times 10^{-4}$ (**Figure 1a-b; Supplementary Table 3**). However, at a nominal threshold of $P=0.05$, we observed suggestive relationships between 11 other pairs. Seven pairs showed positive correlations: esophageal/stomach cancer was correlated with Non-Hodgkin's lymphoma (NHL; $r_g=0.40$, $P=0.0089$), breast ($r_g=0.26$, $P=0.0069$), lung ($r_g=0.44$, $P=0.0035$), and rectal ($r_g=0.32$, $P=0.024$) cancers; bladder and breast cancers ($r_g=0.22$, $P=0.017$); melanoma and testicular cancer ($r_g=0.23$, $P=0.028$); and prostate and thyroid cancers ($r_g=0.23$, $P=0.013$). The remaining four pairs showed negative correlations: endometrial and testicular cancers ($r_g=-0.41$, $P=0.0064$); esophageal/stomach cancer and melanoma ($r_g=-0.27$, $P=0.038$); lung cancer and melanoma ($r_g=-0.28$, $P=0.0048$); and NHL and prostate cancer ($r_g=-0.21$, $P=0.012$).

Locus-Specific Pleiotropy

We identified 43 pleiotropic regions by our linkage disequilibrium (LD) clumping procedure (**Figure 1c; Supplementary Table 4**). Most were at known cancer pleiotropic loci: HLA (24 regions), 8q24 (10 regions), *TERT-CLPTMIL* (5 regions), and *TP53* (1 region). Twenty-two of the HLA regions were associated with both cervical cancer and NHL, and the remaining two were associated with cervical and prostate cancers and, separately, NHL and prostate cancer. Six regions in 8q24 were associated with prostate and colon cancers, and four were associated with prostate and breast cancers. Of the regions in *TERT-CLPTMIL*, two were associated with prostate cancer and, separately, breast and testicular cancers; two were associated with melanoma and, separately, breast and bladder cancers; and the last was associated with melanoma and cervical, lung, and pancreatic cancers. The one region in *TP53*, indexed by rs78378222, was associated with melanoma and lymphocytic leukemia. The remaining three pleiotropic regions were in loci previously associated with at least one cancer and were indexed by rs772695095 (*DIRC3* at 2q35; breast and thyroid cancers), rs11813268 (intergenic at 10q24.33; melanoma and prostate cancer), and rs6507874 (*SMAD7* at 18q21.1; colon and rectal cancers).

Variant-Specific Pleiotropy

We found 137 independent variants with at least two associated cancers, the same direction of effect for all associated cancers, and an overall pleiotropic $P < 1 \times 10^{-6}$ (**Supplementary Table 5**), among which 85 attained genome-wide significance ($P < 5 \times 10^{-8}$). Of the 137 one-directional pleiotropic SNPs, there were 45 for which the overall pleiotropic P was smaller than the P for each of the associated cancers, 17 of which attained genome-wide significance (**Figure 2**). While 134

of the total 137 were in regions that have previously been associated with cancer, 113 were associated with at least one new cancer.

We also considered omni-directional pleiotropic associations, wherein the same SNP allele was associated with increased risk for some cancers but decreased risk for others. We found 21 such variants with $P < 1 \times 10^{-6}$, all of which were independent from one another and from the one-directional effects (**Figure 3; Supplementary Table 6**). Fifteen attained genome-wide significance. There were 13 variants where the overall pleiotropic P was smaller than the P for the associated cancers, eight of which attained genome-wide significance. While 20 of the total 21 were in regions that have previously been associated with cancer, 10 were independent of known risk variants, and all 20 were associated with at least one new cancer. The SNP in a novel region (*SSPN* at 12q12.1) was rs10842692, which was associated with five cancers in one direction and nine cancers in the other.

The number of one- and omni-directional SNPs shared by cancer pairs ranged from four (bladder and esophagus/stomach; colon and NHL) to 32 (pancreas and prostate; leukemia and prostate) (**Figure 4; Supplementary Table 7**). For 19 cancer pairs, the shared associations had exclusively the same direction of effect (tabulating across both the one- and omni-directional analyses), including breast and pancreatic cancers (23 variants), and breast and lung cancers (14 variants). Some pairs, such as melanoma and lung cancer had predominantly one-directional associations (14 variants), with some some omni-directional effects (two variants, associated in opposite directions). For other cancer pairs, most of the shared variants were associated in opposite

directions, such as NHL and prostate cancer, with three one-directional SNPs and six SNPs in opposite directions.

We summarized the pathways represented by genes harbouring one- and omni-directional pleiotropic variants using the Reactome database.³¹ For the 45 one-directional variants with a smaller overall pleiotropic P than the P for each contributing cancer, we observed statistically significant (FDR $q < 0.05$) enrichment in 28 pathways that can be categorized into over 10 independent mechanisms (**Supplementary Table 8**). Most of the enriched pathways represent molecular functions that are known to be dysregulated in cancer, such as oncogene-induced senescence ($q = 2.24 \times 10^{-3}$), p53-dependent DNA damage response ($q = 3.40 \times 10^{-3}$), transforming growth factor (TGF)- β complex signaling ($q = 3.40 \times 10^{-3}$), and telomere stress-induced senescence ($q = 0.031$).

Discussion

In this first ever study to comprehensively evaluate cancer pleiotropy in cohorts unselected for phenotype, we offer multiple lines of evidence for a shared genetic basis of several cancer types. By characterizing pleiotropy at the genome-wide, locus-specific, and variant-specific levels for more cancer sites than have ever previously been examined, we generate considerable novel insights into cancer susceptibility.

Individual GWAS of 18 cancer types yielded 29 genome-wide significant novel risk variants, 16 of which are in regions that have never previously been associated with the cancer of interest and 13 of which are new independent variants in previously associated regions. One notable finding

from our cervical cancer GWAS was rs10175462 in *PAX8* on 2q13, which, to our knowledge, is the first genome-wide significant cervical cancer risk SNP identified outside of the HLA region in a European ancestry population.¹⁵ Two variants in LD with rs10175462 (rs1110839, $r^2=0.33$; rs4848320, $r^2=0.34$) were suggestively associated with cervical cancer in the same direction in a candidate SNP study of *PAX8* expression quantitative trait loci (eQTLs) in a Han Chinese population.³² Several GWAS findings also provided evidence of pleiotropy, in that novel risk variants for one cancer had been previously associated with one or more other cancers. For instance, rs9818780 was associated with melanoma and has been implicated in sunburn risk.³³ This intergenic variant is an eQTL for *LINC00886* and *METTL15P1* in skin tissue. The former gene has previously been linked to breast cancer,³⁴ and both genes have been implicated in ovarian cancer.³⁵ Beyond novel associations, our GWAS confirmed 280 independent associations from previous GWAS. We could thus feel confident using the summary statistics for subsequent analyses of cancer pleiotropy.

In evaluating pairwise genetic correlations between the 18 cancer types, we observed the strongest signal for colon and rectal cancers – an expected relationship consistent with findings from a twin study.³⁶ We also identified several novel cancer pairs for which the genetic correlations were nominally significant. One pair supported by previous evidence is melanoma and testicular cancer; some studies have found that individuals with a family history of the former are at an increased risk for the latter.^{37,38} Esophageal/stomach cancer was a component of five correlated pairs – with melanoma, NHL, and breast, lung, and rectal cancers. Despite some similarities between esophageal and stomach cancers, testing them as a combined phenotype may have inflated the number of correlated cancers. Our genetic correlation results contrast with previous consortia-

based findings;^{4–6} we did not find several correlations that they did and found others that they did not. The differences may be partly due to a smaller number of cases in our cohorts for some sites. However, we achieved comparable or higher cancer-specific heritability estimates for breast, colon, and lung cancers, which suggests that differences in study design may also play a role.

The assessment of pleiotropy at the locus level confirmed previously reported associations at 5p15.33, HLA, and 8q24.^{9,12,15,16,18} Out of the 43 pleiotropic loci that we identified, over half, all in the HLA locus, were associated with cervical cancer and NHL. The two cancers were weakly negatively correlated in the two cohorts combined and nominally significantly negatively correlated in the UKB alone (**Supplementary Table 9**). The difference may reflect better coverage and imputation of the HLA region in the UKB than in GERA. Other findings support a pleiotropic role of several loci previously associated with specific cancers in separate studies. For example, we validated previous results showing that *DIRC3* is associated with breast and thyroid cancers.^{14,39} Additionally, the intergenic region surrounding rs11813268 on 10q24.33 has not been previously associated with melanoma or prostate cancer (as it was in our study), although associations with other cancers have been reported, including kidney, lung, and thyroid cancers.^{35,40–43} *SMAD7* has been previously linked to colorectal cancer,⁴⁴ and we confirmed its association with colon and rectal cancers separately.

Variant-specific analyses provided further evidence of locus-specific cancer pleiotropy, including validation of previously reported signals at 1q32^{7,8} and 2q33^{9,10} (*ALS2CR12*). Interestingly, our lead 1q32 variant (rs1398148) maps to *PIK3C2B* and is in LD ($r^2 > 0.60$) with known *MDM4* cancer risk variants,^{7,8} suggesting that the 1q32 locus may be involved in modulating both p53- and PI3K-

mediated oncogenic pathways. The 158 independent SNPs identified (with an overall pleiotropic $P < 1 \times 10^{-6}$) mapped to a total of 78 genomic locations, which included all of the regions identified from the locus-specific analyses. Although 154 of the 158 SNPs showing one- or omni-directional pleiotropic associations are in regions previously associated with cancer, 133 of the 154 were associated with at least one new cancer. One SNP (rs10842692) associated with 14 cancers in the omni-directional analysis mapped to a novel region at 12p12.1 and is an eQTL for the *SSPN* gene in multiple tissue types, including adipose. *SSPN* has been linked to waist circumference,⁴⁵ suggesting that increased adiposity may be one plausible mechanism underlying the pleiotropic associations observed for this locus.

Out of 158 total SNPs identified from the variant-specific pleiotropy analyses, 20 were in 8q24 and 19 were in the HLA region. Different distributions of one- and omni-directional results highlight patterns of directional pleiotropy: of the 19 HLA SNPs, eight were omni-directional, while only three of the 20 variants in 8q24 were omni-directional. The HLA region is critical for innate and adaptive immune response and has a complex relationship with cancer risk. Heterogeneous associations with HLA haplotypes have been reported for different subtypes of NHL⁴⁶ and lung cancer,⁴⁷ suggesting that relevant risk variants are likely to differ within, as well as between, cancers. Studies have further demonstrated that somatic mutation profiles are associated with HLA class I⁴⁸ and class II alleles.⁴⁹ Specifically, mutations that create neoantigens more likely to be recognized by specific HLA alleles are less likely to be present in tumors from patients carrying such alleles. It is thus possible that some of the positive and negative pleiotropy we identified is related to mutation type. Overall, our results reinforce the importance of the immune system playing a role in cancer susceptibility.

In contrast to the HLA region, the majority of the 8q24 pleiotropic variants had the same direction of effect for all associated cancers, implying the existence of shared genetic mechanisms driving tumorigenesis across sites. The proximity of the well-characterized *MYC* oncogene makes it a compelling candidate for such a consistent, one-directional effect. It could work via regulatory elements, such as acetylated and methylated histone marks.⁵⁰ Consistent with this hypothesis, we observed heritability enrichment⁵¹ for variants with the H3K27ac annotation for breast ($P = 3.09 \times 10^{-4}$), colon ($P = 4.44 \times 10^{-4}$), prostate ($P = 2.74 \times 10^{-5}$), and rectal ($P = 0.036$) cancers – all of which share susceptibility variants in 8q24, according to our analyses and previous studies.⁵⁰ Considering one-directional pleiotropic effects more broadly, pathway analysis indicated that such variants are more likely to be part of canonical signaling pathways that control cell-cycle progression, apoptosis, and DNA replication stress. Dysregulation of such pathways is a hallmark of cancer.

It is important to acknowledge limitations of our analyses. First, counts for some of the cancer types were limited. Therefore, we grouped esophageal and stomach cancers and, separately, oral cavity and pharyngeal cancers. While it might have been preferable to analyze them independently, we felt that the sites had enough in common to be combined and still facilitate new discoveries. Second, due to the complexity of the LD structure in the HLA region, it is possible that we overestimated the number of distinct, independent signals. Slight overestimation, however, does not affect our overall conclusions regarding the pleiotropic nature of this region. Finally, we focused solely on individuals of European ancestry. Further analyses are needed to accurately assess patterns of pleiotropy in non-Europeans.

The characterization of pleiotropy is fundamental to understanding the genetic architecture of cross-cancer susceptibility and its biological underpinnings. The availability of two large, independent cohorts provided an unprecedented opportunity to efficiently evaluate the shared genetic basis of many cancers, including some not previously studied. The result was a multifaceted assessment of common genetic factors implicated in carcinogenesis, and our findings illustrate the importance of investigating different aspects of cancer pleiotropy. Broad analyses of genetic susceptibility and targeted analyses of specific loci and variants may both contribute insights into different dimensions of cancer pleiotropy. Future studies should consider the contribution of rare variants to cancer pleiotropy and aim to elucidate the functional pathways mediating associations observed at pleiotropic regions. Such research, combined with our findings, has the potential to inform drug development, risk assessment, and clinical practice toward reducing the burden of cancer.

Methods

Study Populations and Phenotyping

The UKB is a population-based prospective cohort of 502,611 individuals in the United Kingdom. Study participants were aged 40 to 69 at recruitment between 2006 and 2010, at which time all participants provided detailed information about lifestyle and health-related factors and provided biological samples.⁵² GERA participants were drawn from adult Kaiser Permanente Northern California (KPNC) health plan members who provided a saliva sample for the Research Program on Genes, Environment and Health (RPGEH) between 2008 and 2011. Individuals included in this study were selected from the 102,979 RPGEH participants who were successfully genotyped as part of GERA and answered a survey concerning lifestyle and medical history.^{53,54}

Cancer cases in the UKB were identified via linkage to various national cancer registries established in the early 1970s.⁵² Data in the cancer registries are compiled from hospitals, nursing homes, general practices, and death certificates, among other sources. The latest cancer diagnosis in our data from the UKB occurred in August 2015. GERA cancer cases were identified using the KPNC Cancer Registry, and we included all diagnoses captured through June 2016. Following Surveillance, Epidemiology, and End Results Program (SEER) standards, the KPNC Cancer Registry contains data on all primary cancers (i.e., cancer diagnoses that are not secondary metastases of other cancer sites; excluding non-melanoma skin cancer) diagnosed or treated at any KPNC facility since 1988.

In both cohorts, individuals with at least one recorded diagnosis of a borderline, in situ, or malignant primary cancer were defined as cases for our analyses. For the UKB, all diagnoses

described by International Classification of Diseases (ICD)-9 or ICD-10 codes were converted into ICD-O-3 codes; the KPNC Cancer Registry already included ICD-O-3 codes. We then classified cancers according to organ site using the SEER site recode paradigm.⁵⁵ We grouped all esophageal and stomach cancers and, separately, all oral cavity and pharyngeal cancers to ensure sufficient statistical power. The 18 most common cancer types (except non-melanoma skin cancer) were examined. Testicular cancer data were obtained from the UKB only due to the small number of cases in GERA.

Controls were restricted to individuals who had no record of any cancer in the relevant registries, who did not self-report a prior history of cancer (other than non-melanoma skin cancer), and, if deceased, who did not have cancer listed as a cause of death. For analyses of sex-specific cancer sites (breast, cervix, endometrium, ovary, prostate, and testis), controls were restricted to individuals of the appropriate sex.

Quality Control

For the UKB population, genotyping was conducted using either the UKB Axiom array (436,839 total; 408,841 self-reported European) or the UK BiLEVE array (49,747 total; 49,746 self-reported European).⁵² The former is an updated version of the latter, such that the two arrays share over 95% of their marker content. UKB investigators undertook a rigorous quality control (QC) protocol.⁵² Genotype imputation was performed using the Haplotype Reference Consortium as the main reference panel and the merged UK10K and 1000 Genomes phase 3 reference panels for supplementary data.⁵² Ancestry principal components (PCs) were computed using *fastPCA*⁵⁶ based on a set of 407,219 unrelated samples and 147,604 genetic markers.⁵²

For GERA participants, genotyping was performed using an Affymetrix Axiom array (Affymetrix, Santa Clara, CA, USA) optimized for individuals of European race/ethnicity. Details about the array design, estimated genome-wide coverage, and QC procedures have been published previously.^{54,57} The genotyping produced high quality data with average call rates of 99.7% and average SNP reproducibility of 99.9%. Variants that were not directly genotyped (or that were excluded by QC procedures) were imputed to generate genotypic probability estimates. After pre-phasing genotypes with SHAPE-IT v2.5,⁵⁸ IMPUTE2 v2.3.1 was used to impute SNPs relative to the cosmopolitan reference panel from 1000 Genomes.^{59–61} Ancestry PCs were computed using Eigenstrat v4.2, as previously described.⁵³

For both cohorts, analyses were limited to self-reported European ancestry individuals for whom self-reported and genetic sex matched. To further minimize potential population stratification, we excluded individuals for whom either of the first two ancestry PCs fell outside five standard deviations of the mean of the population. Based on a subset of genotyped autosomal variants with minor allele frequency (MAF) ≥ 0.01 and genotype call rate $\geq 97\%$, we excluded samples with call rates $< 97\%$ and/or heterozygosity more than five standard deviations from the mean of the population. With the same subset of SNPs, we used KING⁶² to estimate relatedness among the samples. We excluded one individual from each pair of first-degree relatives, first prioritizing on maximizing the number of the cancer cases relevant to these analyses and then maximizing the total number of individuals in the analyses. Our study population ultimately included 413,870 UKB participants and 66,526 GERA participants. We excluded SNPs with imputation quality score < 0.3 , call rate $< 95\%$ (alternate allele dosage required to be within 0.1 of the nearest hard call

to be non-missing; UKB only), Hardy-Weinberg equilibrium P among controls $<1 \times 10^{-5}$, and/or MAF <0.01 , leaving 8,876,519 variants for analysis for the UKB and 8,973,631 for GERA.

Genome-Wide Association Analyses

We used PLINK^{63,64} to implement within-cohort logistic regression models of additively modeled SNPs genome-wide, comparing cases of each cancer type to cancer-free controls. All models were adjusted for age at specimen collection, sex (non-sex-specific cancers only), first ten ancestry PCs, genotyping array (UKB only), and reagent kit used for genotyping (Axiom v1 or v2; GERA only). Case counts ranged from 471 (pancreatic cancer) to 13,903 (breast cancer) in the UKB and from 162 (esophageal/stomach cancer) to 3,978 (breast cancer) in GERA (**Supplementary Table 10**). Control counts were 359,825 (189,855 female) and 50,525 (29,801 female) in the UKB and GERA, respectively. After separate GWAS were conducted in each cohort, association results for the 7,846,216 SNPs in both cohorts were combined via meta-analysis. For variants that were only examined in one cohort (22% of the total 10,003,934 SNPs analyzed), original summary statistics were merged with the meta-analyzed SNPs to create a union set of SNPs for use in downstream analyses.

To determine independent signals in our union set of SNPs, we implemented the LD clumping procedure in PLINK^{63,64} based on genotype hard calls from a reference panel comprised of a downsampled subset of 10,000 random UKB participants. For each cancer separately, LD clumps were formed around index SNPs with the smallest P not already assigned to another clump. To identify all potential signals, in each clump, index SNPs had a suggestive association based on $P < 1 \times 10^{-6}$, and SNPs were added if they were marginally significant with $P < 0.05$, were within

500kb of the index SNP, and had $r^2 > 0.1$ with the index SNP. Index variants were determined to be novel if they were not in LD with previously reported risk variants in European ancestry populations (as described below).

To identify SNPs previously associated with each cancer type, we abstracted all genome-wide significant SNPs from relevant GWAS published through June 2018. We determined that a SNP was potentially novel if it had LD $r^2 < 0.1$ with all previously reported SNPs for the relevant cancer based on both the UKB reference panel and the 1000 Genomes EUR superpopulation via LDlink.⁶⁵ To confirm novelty and consider pleiotropy, an additional literature review was conducted to investigate whether these SNPs had previously been reported for the same or other cancers, including those not attaining genome-wide significance and those in non-GWAS analyses. For this additional review, we used the PhenoScanner database⁶⁶ to search for SNPs of interest and variants in LD in order to comprehensively scan previously reported associations. We then supplemented with more in-depth PubMed searches to determine if the genes in which novel SNPs were located had previously been reported for the same or other cancers.

Genome-Wide Heritability and Genetic Correlation

We used LD score regression (LDSC) of summary statistics from the union set of SNPs across both cohorts to calculate the genome-wide liability-scale heritability of each cancer type and the genetic correlation between each pair of cancer types.^{67,68} Internal LD scores were calculated using the aforementioned downsampled subset of UKB participants. To convert to liability-scale heritability, we adjusted for lifetime risks of each cancer based on SEER 2012-2014 estimates

(Supplementary Table 11).⁶⁹ LDSC was unable to estimate genetic correlations for testicular cancer with both oral cavity/pharyngeal and pancreatic cancers.

Locus-Specific Pleiotropy

Pleiotropic regions comprised of variants in LD with one another and associated with multiple cancers were constructed by LD clumping our union set of SNP-based summary statistics across all cancers. Regions were formed around index SNPs associated with any cancer with $P < 1 \times 10^{-6}$ to assess whether suggestive regions for one cancer might also be informative for another. SNPs were added to a region if they were associated with any cancer with $P < 1 \times 10^{-6}$, were within 500kb of the index SNP, and had LD $r^2 > 0.5$ with the index SNP. We used a larger threshold for assessing pleiotropic regions ($r^2 > 0.5$) than for identifying truly independent signals ($r^2 > 0.1$; above) to ensure that all SNPs within a region were in LD. If all SNPs in a region were associated with the same cancer, the region was not considered pleiotropic.

Variant-Specific Pleiotropy

We quantified one-directional and, separately, omni-directional variant-specific pleiotropy via the R package ASSET (association analysis based on subsets), which explores all possible subsets of traits for the presence of association signals in either the same or opposite directions.⁷⁰ Independent variants were determined via LD clumping, where index SNPs had a suggestive $P < 1 \times 10^{-6}$, and SNPs were clumped if they had $P < 0.05$, were within 500kb of the index SNP, and had $r^2 > 0.1$ with the index SNP. We used a suggestive significance threshold to comprehensively assess all potentially pleiotropic variants. A SNP was determined to have a one-directional pleiotropic association if the overall P was $< 1 \times 10^{-6}$, included at least two cancers, and was not in LD with a

SNP with an omni-directional pleiotropic association with a smaller overall P . A SNP was determined to have an omni-directional pleiotropic association if the overall P was $<1 \times 10^{-6}$, the P for each direction was <0.05 , and the SNP was not in LD with a SNP with a one-directional pleiotropic association with a smaller overall P . We deconstructed omni-directional associations into cancers with risk-increasing effects and cancers with risk-decreasing effects.

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References

1. Bray, F. *et al.* Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. *CA. Cancer J. Clin.* **68**, 394–424 (2018).
2. Mucci, L. A. *et al.* Familial Risk and Heritability of Cancer Among Twins in Nordic Countries. *JAMA* **315**, 68–76 (2016).
3. Czene, K., Lichtenstein, P. & Hemminki, K. Environmental and heritable causes of cancer among 9.6 million individuals in the Swedish Family-Cancer Database. *Int. J. Cancer* **99**, 260–266 (2002).
4. Sampson, J. N. *et al.* Analysis of Heritability and Shared Heritability Based on Genome-Wide Association Studies for 13 Cancer Types. *J Natl Cancer Inst* **107**, (2015).
5. Lindström, S. *et al.* Quantifying the Genetic Correlation between Multiple Cancer Types. *Cancer Epidemiol Biomark. Prev* **26**, 1427–1435 (2017).
6. Jiang, X. *et al.* Shared heritability and functional enrichment across six solid cancers. *Nat. Commun.* **10**, 431 (2019).
7. Couch, F. J. *et al.* Genome-wide association study in BRCA1 mutation carriers identifies novel loci associated with breast and ovarian cancer risk. *PLoS Genet.* **9**, e1003212 (2013).
8. Eeles, R. A. *et al.* Identification of 23 new prostate cancer susceptibility loci using the iCOGS custom genotyping array. *Nat. Genet.* **45**, 385–391, 391e1-2 (2013).
9. Barrett, J. H. *et al.* Genome-wide association study identifies three new melanoma susceptibility loci. *Nat. Genet.* **43**, 1108–1113 (2011).

10. Broeks, A. *et al.* Low penetrance breast cancer susceptibility loci are associated with specific breast tumor subtypes: findings from the Breast Cancer Association Consortium. *Hum. Mol. Genet.* **20**, 3289–3303 (2011).
11. Ellinghaus, E. *et al.* Identification of germline susceptibility loci in ETV6-RUNX1-rearranged childhood acute lymphoblastic leukemia. *Leukemia* **26**, 902–909 (2012).
12. Rothman, N. *et al.* A multi-stage genome-wide association study of bladder cancer identifies multiple susceptibility loci. *Nat. Genet.* **42**, 978–984 (2010).
13. Eeles, R. A. *et al.* Identification of seven new prostate cancer susceptibility loci through a genome-wide association study. *Nat. Genet.* **41**, 1116–1121 (2009).
14. Michailidou, K. *et al.* Large-scale genotyping identifies 41 new loci associated with breast cancer risk. *Nat. Genet.* **45**, 353–361 (2013).
15. Bahrami, A. *et al.* Genetic susceptibility in cervical cancer: From bench to bedside. *J. Cell. Physiol.* **233**, 1929–1939 (2017).
16. Smedby, K. E. *et al.* GWAS of follicular lymphoma reveals allelic heterogeneity at 6p21.32 and suggests shared genetic susceptibility with diffuse large B-cell lymphoma. *PLoS Genet.* **7**, e1001378 (2011).
17. Jin, G. *et al.* Genetic variants at 6p21.1 and 7p15.3 are associated with risk of multiple cancers in Han Chinese. *Am. J. Hum. Genet.* **91**, 928–934 (2012).
18. Eeles, R. A. *et al.* Multiple newly identified loci associated with prostate cancer susceptibility. *Nat. Genet.* **40**, 316–321 (2008).
19. Purdue, M. P. *et al.* Genome-wide association study of renal cell carcinoma identifies two susceptibility loci on 2p21 and 11q13.3. *Nat. Genet.* **43**, 60–65 (2011).

20. Spurdle, A. B. *et al.* Genome-wide association study identifies a common variant associated with risk of endometrial cancer. *Nat. Genet.* **43**, 451–454 (2011).
21. Couch, F. J. *et al.* Common variants at the 19p13.1 and ZNF365 loci are associated with ER subtypes of breast cancer and ovarian cancer risk in BRCA1 and BRCA2 mutation carriers. *Cancer Epidemiol Biomark. Prev* **21**, 645–657 (2012).
22. Setiawan, V. W. *et al.* Cross-cancer pleiotropic analysis of endometrial cancer: PAGE and E2C2 consortia. *Carcinogenesis* **35**, 2068–2073 (2014).
23. Rafnar, T. *et al.* Sequence variants at the TERT- CLPTM1L locus associate with many cancer types. *Nat. Genet.* **41**, 221–227 (2009).
24. Cheng, I. *et al.* Pleiotropic effects of genetic risk variants for other cancers on colorectal cancer risk: PAGE, GECCO and CCFR consortia. *Gut* **63**, 800–807 (2014).
25. Jones, C. C. *et al.* Cross-Cancer Pleiotropic Associations with Lung Cancer Risk in African Americans. *Cancer Epidemiol Biomark. Prev* **28**, 715–723 (2019).
26. Hung, R. J. *et al.* Cross Cancer Genomic Investigation of Inflammation Pathway for Five Common Cancers: Lung, Ovary, Prostate, Breast, and Colorectal Cancer. *J. Natl. Cancer Inst.* **107**, (2015).
27. Qian, D. C. *et al.* Identification of shared and unique susceptibility pathways among cancers of the lung, breast, and prostate from genome-wide association studies and tissue-specific protein interactions. *Hum. Mol. Genet.* **24**, 7406–7420 (2015).
28. Fehringer, G. *et al.* Cross-Cancer Genome-Wide Analysis of Lung, Ovary, Breast, Prostate, and Colorectal Cancer Reveals Novel Pleiotropic Associations. *Cancer Res.* **76**, 5103–5114 (2016).

29. Toth, R. *et al.* Genetic Variants in Epigenetic Pathways and Risks of Multiple Cancers in the GAME-ON Consortium. *Cancer Epidemiol. Biomarkers Prev.* **26**, 816–825 (2017).
30. Karami, S. *et al.* Telomere structure and maintenance gene variants and risk of five cancer types. *Int. J. Cancer* **139**, 2655–2670 (2016).
31. Yu, G. & He, Q.-Y. ReactomePA: an R/Bioconductor package for reactome pathway analysis and visualization. *Mol. Biosyst.* **12**, 477–479 (2016).
32. Han, J. *et al.* Expression quantitative trait loci in long non-coding RNA PAX8-AS1 are associated with decreased risk of cervical cancer. *Mol. Genet. Genomics* **291**, 1743–1748 (2016).
33. Kichaev, G. *et al.* Leveraging Polygenic Functional Enrichment to Improve GWAS Power. *Am. J. Hum. Genet.* **104**, 65–75 (2019).
34. Michailidou, K. *et al.* Association analysis identifies 65 new breast cancer risk loci. *Nature* **551**, 92–94 (2017).
35. Phelan, C. M. *et al.* Identification of 12 new susceptibility loci for different histotypes of epithelial ovarian cancer. *Nat. Genet.* **49**, 680–691 (2017).
36. Graff, R. E. *et al.* Familial Risk and Heritability of Colorectal Cancer in the Nordic Twin Study of Cancer. *Clin Gastroenterol Hepatol* **15**, 1256–1264 (2017).
37. Hemminki, K. & Chen, B. Familial risks in testicular cancer as aetiological clues. *Int. J. Androl.* **29**, 205–210 (2006).
38. Zhang, L. *et al.* Familial Associations in Testicular Cancer with Other Cancers. *Sci. Rep.* **8**, 10880 (2018).
39. Gudmundsson, J. *et al.* Discovery of common variants associated with low TSH levels and thyroid cancer risk. *Nat. Genet.* **44**, 319–322 (2012).

40. Scelo, G. *et al.* Genome-wide association study identifies multiple risk loci for renal cell carcinoma. *Nat. Commun.* **8**, (2017).
41. Chahal, H. S. *et al.* Genome-wide association study identifies 14 novel risk alleles associated with basal cell carcinoma. *Nat. Commun.* **7**, (2016).
42. McKay, J. D. *et al.* Large-scale association analysis identifies new lung cancer susceptibility loci and heterogeneity in genetic susceptibility across histological subtypes. *Nat. Genet.* **49**, 1126–1132 (2017).
43. Gudmundsson, J. *et al.* A genome-wide association study yields five novel thyroid cancer risk loci. *Nat. Commun.* **8**, 14517 (2017).
44. Broderick, P. *et al.* A genome-wide association study shows that common alleles of SMAD7 influence colorectal cancer risk. *Nat. Genet.* **39**, 1315–1317 (2007).
45. Graff, M. *et al.* Genome-wide physical activity interactions in adiposity - A meta-analysis of 200,452 adults. *PLoS Genet.* **13**, e1006528 (2017).
46. Wang, S. S. *et al.* HLA Class I and II Diversity Contributes to the Etiologic Heterogeneity of Non-Hodgkin Lymphoma Subtypes. *Cancer Res.* **78**, 4086–4096 (2018).
47. Ferreira-Iglesias, A. *et al.* Fine mapping of MHC region in lung cancer highlights independent susceptibility loci by ethnicity. *Nat. Commun.* **9**, 3927 (2018).
48. Marty, R. *et al.* MHC-I Genotype Restricts the Oncogenic Mutational Landscape. *Cell* **171**, 1272-1283.e15 (2017).
49. Marty Pyke, R. *et al.* Evolutionary Pressure against MHC Class II Binding Cancer Mutations. *Cell* **175**, 416-428.e13 (2018).
50. Grisanzio, C. & Freedman, M. L. Chromosome 8q24-Associated Cancers and MYC. *Genes Cancer* **1**, 555–559 (2010).

51. Finucane, H. K. *et al.* Partitioning heritability by functional annotation using genome-wide association summary statistics. *Nat. Genet.* **47**, 1228–1235 (2015).
52. Bycroft, C. *et al.* The UK Biobank resource with deep phenotyping and genomic data. *Nature* **562**, 203–209 (2018).
53. Banda, Y. *et al.* Characterizing Race/Ethnicity and Genetic Ancestry for 100,000 Subjects in the Genetic Epidemiology Research on Adult Health and Aging (GERA) Cohort. *Genetics* **200**, 1285–1295 (2015).
54. Kvale, M. N. *et al.* Genotyping Informatics and Quality Control for 100,000 Subjects in the Genetic Epidemiology Research on Adult Health and Aging (GERA) Cohort. *Genetics* **200**, 1051–1060 (2015).
55. Site Recode ICD-O-3/WHO 2008 Definition. Available from URL:
https://seer.cancer.gov/siterecode/icdo3_dwhohome/index.html [accessed June 30, 2017].
56. Galinsky, K. J. *et al.* Fast Principal-Component Analysis Reveals Convergent Evolution of ADH1B in Europe and East Asia. *Am. J. Hum. Genet.* **98**, 456–472 (2016).
57. Hoffmann, T. J. *et al.* Next generation genome-wide association tool: Design and coverage of a high-throughput European-optimized SNP array. *Genomics* **98**, 79–89 (2011).
58. Delaneau, O., Marchini, J. & Zagury, J.-F. A linear complexity phasing method for thousands of genomes. *Nat. Methods* **9**, 179–181 (2012).
59. Howie, B. N., Donnelly, P. & Marchini, J. A Flexible and Accurate Genotype Imputation Method for the Next Generation of Genome-Wide Association Studies. *PLOS Genet.* **5**, e1000529 (2009).
60. Howie, B., Marchini, J. & Stephens, M. Genotype Imputation with Thousands of Genomes. *G3 Genes Genomes Genet.* **1**, 457–470 (2011).

61. Howie, B., Fuchsberger, C., Stephens, M., Marchini, J. & Abecasis, G. R. Fast and accurate genotype imputation in genome-wide association studies through pre-phasing. *Nat. Genet.* **44**, 955–959 (2012).
62. Manichaikul, A. *et al.* Robust relationship inference in genome-wide association studies. *Bioinformatics* **26**, 2867–2873 (2010).
63. Purcell, S. *et al.* PLINK: A Tool Set for Whole-Genome Association and Population-Based Linkage Analyses. *Am. J. Hum. Genet.* **81**, 559–575 (2007).
64. Chang, C. C. *et al.* Second-generation PLINK: rising to the challenge of larger and richer datasets. *GigaScience* **4**, (2015).
65. Machiela, M. J. & Chanock, S. J. LDlink: a web-based application for exploring population-specific haplotype structure and linking correlated alleles of possible functional variants. *Bioinformatics* **31**, 3555–3557 (2015).
66. Staley, J. R. *et al.* PhenoScanner: a database of human genotype-phenotype associations. *Bioinformatics* **32**, 3207–3209 (2016).
67. Bulik-Sullivan, B. K. *et al.* LD Score regression distinguishes confounding from polygenicity in genome-wide association studies. *Nat. Genet.* **47**, 291–295 (2015).
68. Bulik-Sullivan, B. *et al.* An atlas of genetic correlations across human diseases and traits. *Nat. Genet.* **47**, 1236–1241 (2015).
69. Howlader, N. *et al.* SEER Cancer Statistics Review, 1975-2014, National Cancer Institute. Bethesda, MD, https://seer.cancer.gov/csr/1975_2014/, based on November 2016 SEER data submission, posted to the SEER web site, April 2017.

70. Bhattacharjee, S. *et al.* A Subset-Based Approach Improves Power and Interpretation for the Combined Analysis of Genetic Association Studies of Heterogeneous Traits. *Am. J. Hum. Genet.* **90**, 821–835 (2012).

Tables and Figures

Table 1. Novel genome-wide significant loci from meta-analysis of UK Biobank (UKB) and Genetic Epidemiology Research on Adult Health and Aging (GERA) SNPs for each cancer site.

| Cancer Site | SNP | Chromosome | Position | Gene | REF/ALT* | MAF** | | Odds Ratio (OR) | | | Meta P |
|----------------------|----------------------------|------------|-----------|---------------------------|----------|-------|-------|-----------------|------|------|------------------------|
| | | | | | | UKB | GERA | UKB | GERA | Meta | |
| Bladder | rs76088467 ^{†,+} | 6 | 21795787 | <i>CASC15</i> | G/A | 0.025 | 0.030 | 0.67 | 0.59 | 0.64 | 2.34x10 ⁻⁸ |
| Breast | rs6752414 ^{†,+} | 2 | 121425339 | Intergenic | T/C | 0.077 | 0.083 | 0.89 | 0.87 | 0.88 | 1.81x10 ⁻⁹ |
| Breast | rs8027730 ^{§,+} | 15 | 49872585 | <i>FAM227B</i> | A/C | 0.48 | 0.48 | 1.06 | 1.08 | 1.06 | 2.68x10 ⁻⁸ |
| Cervix | rs10175462 ^{§,+} | 2 | 113988492 | <i>PAX8</i> | A/G | 0.36 | 0.37 | 1.16 | 1.08 | 1.15 | 7.71x10 ⁻¹⁴ |
| Cervix | rs2856437 ^{†,+} | 6 | 32157364 | <i>PBX2</i> | A/G | 0.063 | 0.047 | 0.76 | 0.88 | 0.77 | 1.24x10 ⁻¹⁵ |
| Colon | rs71518872 ^{§,+} | 8 | 103561978 | Upstream of <i>ODF1</i> | G/C | 0.015 | 0.017 | 0.65 | 0.61 | 0.64 | 1.27x10 ⁻⁸ |
| Colon | rs8114643 [§] | 20 | 7833046 | Intergenic | G/A | 0.14 | 0.15 | 0.83 | 0.84 | 0.83 | 2.10x10 ⁻⁹ |
| Esophagus/Stomach | rs75460256 ^{†,+} | 2 | 106687838 | <i>C2orf40</i> | G/A | 0.024 | 0.022 | 0.52 | 0.67 | 0.53 | 1.04x10 ⁻⁸ |
| Kidney | rs112248293 ^{§,+} | 15 | 61500352 | <i>RORA</i> | A/G | 0.024 | 0.025 | 0.53 | 0.62 | 0.55 | 3.36x10 ⁻⁹ |
| Lung | rs10863899 [§] | 1 | 211666218 | 5'UTR of <i>RD3</i> | G/A | 0.42 | 0.42 | 1.23 | 1.09 | 1.18 | 1.91x10 ⁻⁸ |
| Lung | rs146099759 [§] | 5 | 12883592 | Intergenic | A/G | 0.024 | 0.028 | 0.69 | 0.57 | 0.64 | 3.50x10 ⁻⁸ |
| Lung | rs12543486 ^{†,+} | 8 | 13012376 | <i>DLC1</i> | C/T | 0.17 | 0.16 | 1.30 | 1.18 | 1.26 | 3.51x10 ⁻⁸ |
| Lymphocytic Leukemia | rs114490818 [†] | 3 | 126099101 | Intergenic | A/G | 0.022 | 0.011 | 0.48 | 0.53 | 0.48 | 2.86x10 ⁻⁸ |
| Lymphocytic Leukemia | rs61965473 ^{§,+} | 13 | 95571786 | Intergenic | T/C | 0.023 | 0.023 | 0.52 | 0.44 | 0.49 | 3.95x10 ⁻⁸ |
| Lymphocytic Leukemia | rs78378222 ^{†,+} | 17 | 7571752 | 3'UTR of <i>TP53</i> | G/T | 0.012 | 0.014 | 0.44 | 0.34 | 0.40 | 1.89x10 ⁻⁹ |
| Melanoma | rs9818780 ^{§,+} | 3 | 156492758 | Intergenic | C/T | 0.49 | 0.48 | 0.92 | 0.89 | 0.91 | 3.16x10 ⁻⁸ |
| Melanoma | rs12186662 [§] | 5 | 90356197 | <i>ADGRV1</i> | G/A | 0.32 | 0.36 | 0.90 | 0.89 | 0.90 | 1.09x10 ⁻⁸ |
| Melanoma | rs55797833 ^{†,+} | 9 | 21995044 | 5'UTR OF <i>CDKN2A</i> | G/T | 0.023 | 0.021 | 1.71 | 1.72 | 1.71 | 6.71x10 ⁻¹² |
| Melanoma | rs145268882 [†] | 16 | 90009717 | Upstream of <i>DEF8</i> | A/G | 0.062 | 0.050 | 1.29 | 1.24 | 1.28 | 2.96x10 ⁻⁸ |
| Melanoma | rs78378222 ^{†,+} | 17 | 7571752 | 3'UTR of <i>TP53</i> | G/T | 0.012 | 0.014 | 0.70 | 0.63 | 0.67 | 1.18x10 ⁻⁸ |
| Prostate | rs13114291 ^{†,+} | 4 | 74486834 | Upstream of <i>RASSF6</i> | G/A | 0.058 | 0.059 | 1.19 | 1.23 | 1.20 | 2.88x10 ⁻⁸ |
| Prostate | rs11781162 ^{†,+} | 8 | 128110896 | Intergenic | G/A | 0.015 | 0.019 | 0.72 | 0.76 | 0.73 | 8.24x10 ⁻⁹ |
| Prostate | rs12678349 ^{†,+} | 8 | 128198564 | <i>CASC19</i> | T/C | 0.092 | 0.096 | 1.14 | 1.20 | 1.16 | 4.61x10 ⁻⁸ |
| Prostate | rs78449170 ^{†,+} | 8 | 128445011 | <i>CASC8</i> | C/T | 0.054 | 0.061 | 1.25 | 1.32 | 1.27 | 3.58x10 ⁻¹¹ |
| Prostate | rs12549761 ^{†,+} | 8 | 128540776 | <i>CASC8</i> | G/C | 0.12 | 0.13 | 1.23 | 1.26 | 1.23 | 3.39x10 ⁻¹⁸ |
| Prostate | rs3862792 ^{§,+} | 11 | 69462856 | <i>CCND1</i> (synonymous) | T/C | 0.026 | 0.026 | 0.77 | 0.87 | 0.79 | 2.86x10 ⁻⁸ |
| Prostate | rs7407025 ^{†,+} | 17 | 36080810 | <i>HNF1B</i> | A/G | 0.24 | 0.25 | 1.13 | 1.10 | 1.12 | 1.06x10 ⁻¹¹ |

| | | | | | | | | | | | |
|---------|-------------------------------|---|-----------|----------------|-------------|-------|-------|------|------|------|-----------------------|
| Rectum | rs145503185 [§] | 9 | 23455764 | Intergenic | C/T | 0.013 | 0.018 | 0.57 | 0.50 | 0.55 | 4.36x10 ⁻⁸ |
| Thyroid | 2:173859846_TA_T [§] | 2 | 173859846 | <i>RAPGEF4</i> | T/TA | 0.25 | 0.26 | 1.45 | 1.15 | 1.36 | 3.49x10 ⁻⁸ |

*REF is reference allele and ALT allele is effect allele; bold allele is minor allele

**MAF = minor allele frequency calculated in all controls

§ Indicates SNPs in loci not previously associated with the cancer of interest in European ancestry

† Indicates SNPs in known susceptibility loci for cancer of interest in European ancestry but independent of previously reported variants (LD $r^2 < 0.1$ in Europeans)

⁺ Indicates SNPs in loci previously associated with at least one of the other cancers evaluated in this study in European ancestry

Table 2. Heritability estimates (h^2) and 95% confidence intervals (CIs) for each cancer based on the union set of UK Biobank (UKB) and Genetic Epidemiology Research on Adult Health and Aging (GERA) SNPs, compared with previous estimates from other array-based and twin/family-based studies.

| Cancer Site | Array-Based Heritability | | | Twin/Family-Based Heritability |
|------------------------|--------------------------|----------------------------------|------------------------------------|-----------------------------------|
| | Current Study | Jiang <i>et al.</i> ^a | Sampson <i>et al.</i> ^c | Mucci <i>et al.</i> ^d |
| Bladder | 0.08 (0.04-0.12) | | 0.12 (0.09-0.16) | 0.07 (0.02-0.11) ^e |
| Breast | 0.10 (0.08-0.13) | 0.14 (0.12-0.16) | 0.10 (0.00-0.20)** | 0.31 (0.11-0.51) |
| Cervix | 0.07 (0.02-0.12) | | | 0.13 (0.06-0.15) ^{e,+} |
| Colon | 0.07 (0.04-0.10) | 0.09 (0.07-0.11)* | | 0.15 (0.00-0.45) |
| Endometrium | 0.13 (0.07-0.18) | | 0.18 (0.09-0.27) | 0.27 (0.11-0.43) |
| Esophagus/Stomach | 0.14 (0.07-0.21) | | 0.38 (0.17-0.59)*** | 0.22 (0.00-0.55) ⁺⁺ |
| Kidney | 0.09 (0.04-0.15) | | 0.15 (0.02-0.27) | 0.38 (0.21-0.55) |
| Lung | 0.15 (0.10-0.20) | 0.08 (0.05-0.10) | 0.21 (0.14-0.27) | 0.18 (0.00-0.42) |
| Lymphocytic Leukemia | 0.14 (0.05-0.23) | | 0.22 (0.16-0.28) [†] | 0.09 (0.09-0.16) ^{e,+++} |
| Melanoma | 0.08 (0.04-0.11) | | | 0.58 (0.43-0.73) |
| Non-Hodgkin's Lymphoma | 0.13 (0.03-0.23) | | 0.09 (0.04-0.15) ^{††} | 0.10 (0.08-0.10) ^e |
| Oral Cavity/Pharynx | 0.04 (0.00-0.13) | 0.10 (0.05-0.14) | | 0.09 (0.00-0.60) |
| Ovary | 0.07 (0.01-0.13) | 0.03 (0.02-0.05) | | 0.39 (0.23-0.55) |
| Pancreas | 0.06 (0.00-0.18) | 0.05 (0.00-0.10) ^b | 0.10 (0.04-0.16) | |
| Prostate | 0.16 (0.13-0.20) | 0.18 (0.14-0.22) | 0.38 (0.24-0.51) | 0.57 (0.51-0.63) |
| Rectum | 0.11 (0.07-0.16) | | | 0.14 (0.00-0.50) |
| Testis | 0.26 (0.15-0.38) | | 0.30 (0.08-0.51) | 0.25 (0.15-0.37) ^e |
| Thyroid | 0.21 (0.09-0.33) | | | 0.53 (0.52-0.53) ^e |

* Colorectal

** Estrogen receptor negative (ER-)

*** For esophageal in Asian population (stomach in Asian population: $h^2 = 0.25$ [0.00-0.52])

† For chronic lymphocytic leukemia

†† For diffuse large B-cell lymphoma

+ For in situ (invasive: $h^2 = 0.22$ [0.14-0.27])

++ Stomach

+++ Age >15 years

^a Taken from Jiang, et al. (2019),⁶ 95% CI calculated from provided standard error

^b Taken from Lindström, et al. (2017)⁵

^c Taken from Sampson, et al. (2015)⁴

^d Taken from Mucci, et al. (2016),² except where not included in analysis or 95% CI range was > 0.60; remaining taken from Czene, et al. (2002),³ as marked

^e Taken from Czene, et al. (2002),³ family-based not twin

Figure 1. Cross-cancer genetic correlations (r_g) calculated via LD-score regression (LDSC) and associated cancers from the locus-specific pleiotropy analysis. (a) Cancer pairs are connected if the genetic correlation had $P < 0.05$, width of the line is proportional to magnitude of the point estimate, and shading is proportional to strength of association according to P , where the Bonferroni-corrected threshold is $0.05/153 = 3.27 \times 10^{-4}$; (b) genetic correlation, 95% confidence interval (CI), and P for all cancer pairs with $P < 0.05$; (c) cancer pairs are connected by a line (each line represents one region) if a region contains any SNPs associated with either cancer, where regions are formed around index SNPs with $P < 1 \times 10^{-6}$ for any cancer and SNPs are added if they have $P < 1 \times 10^{-6}$ for any cancer, are within 500kb of the index SNP, and have LD $r^2 > 0.5$ with the index SNP.

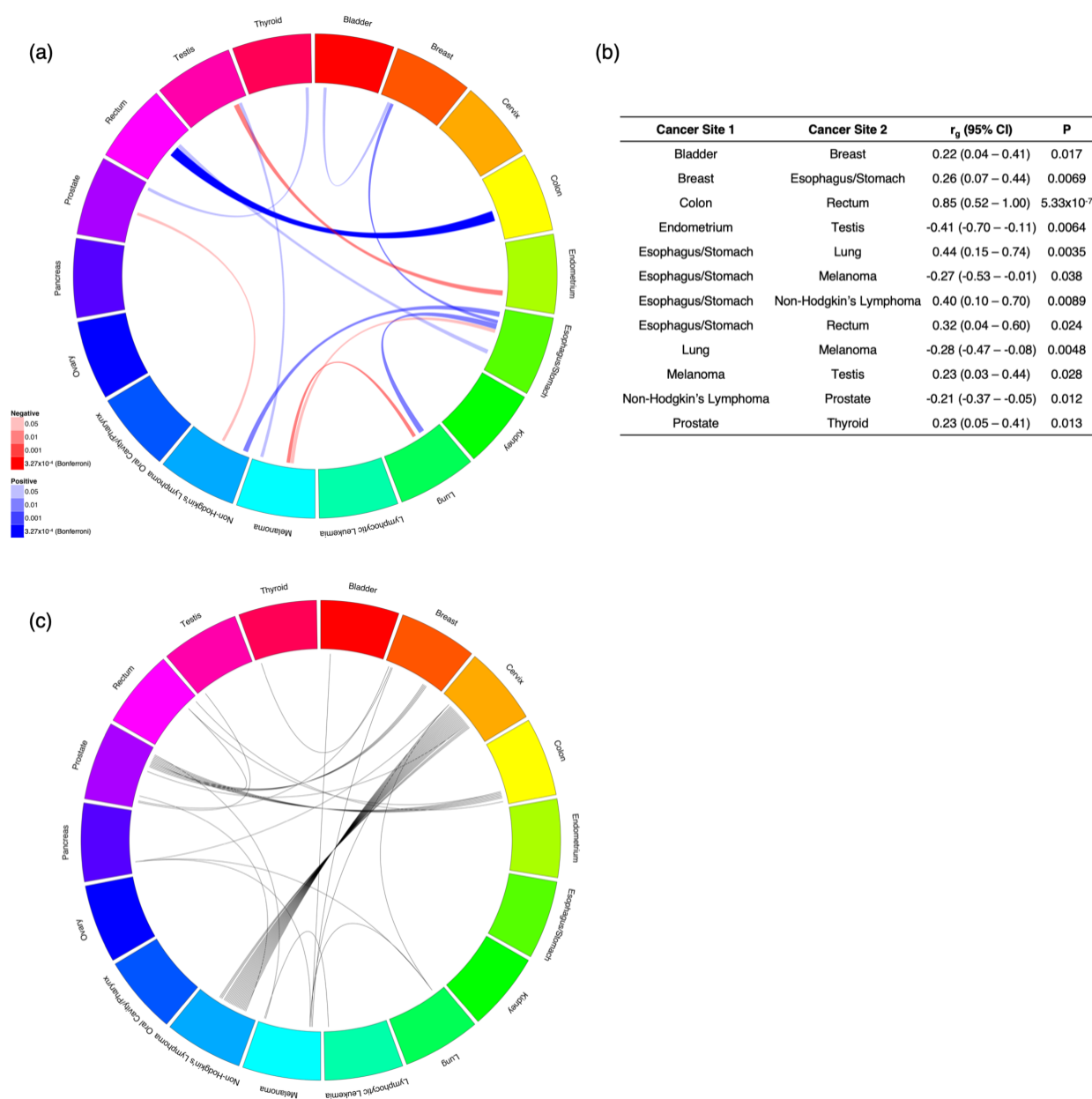
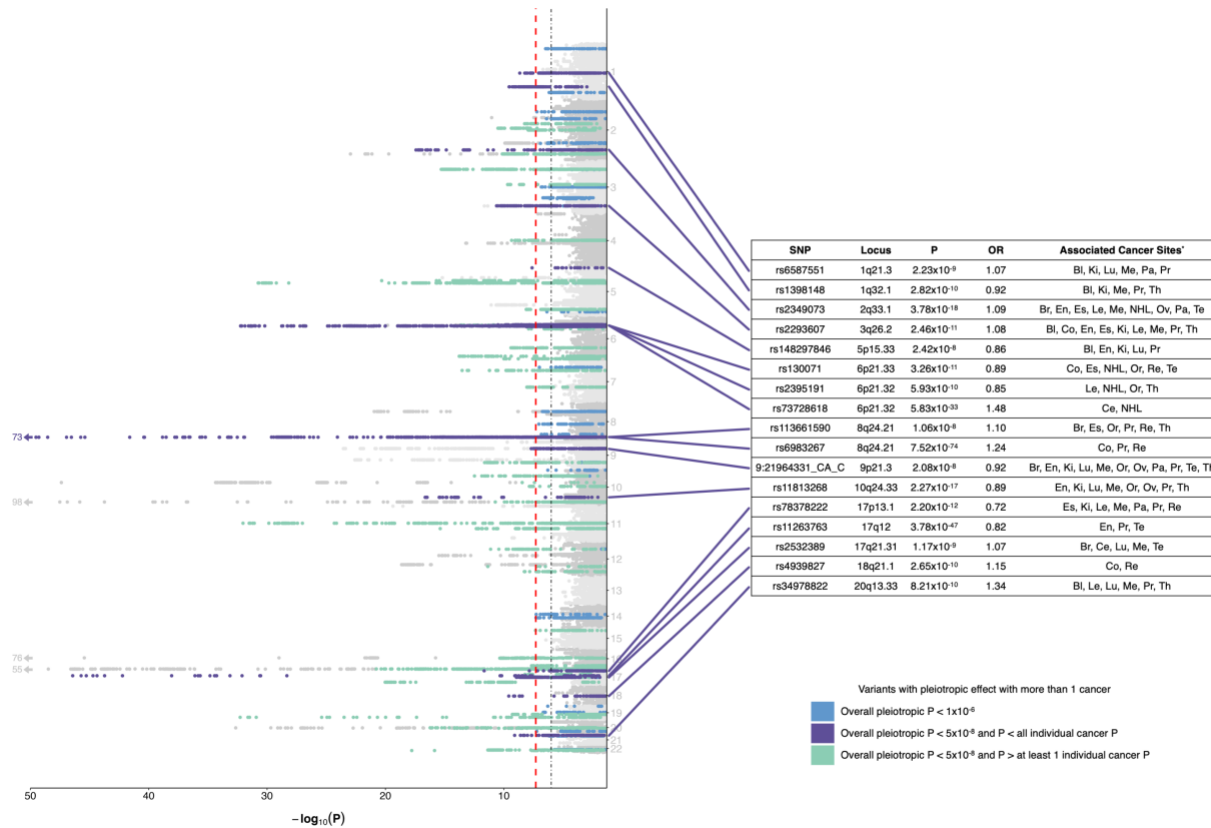
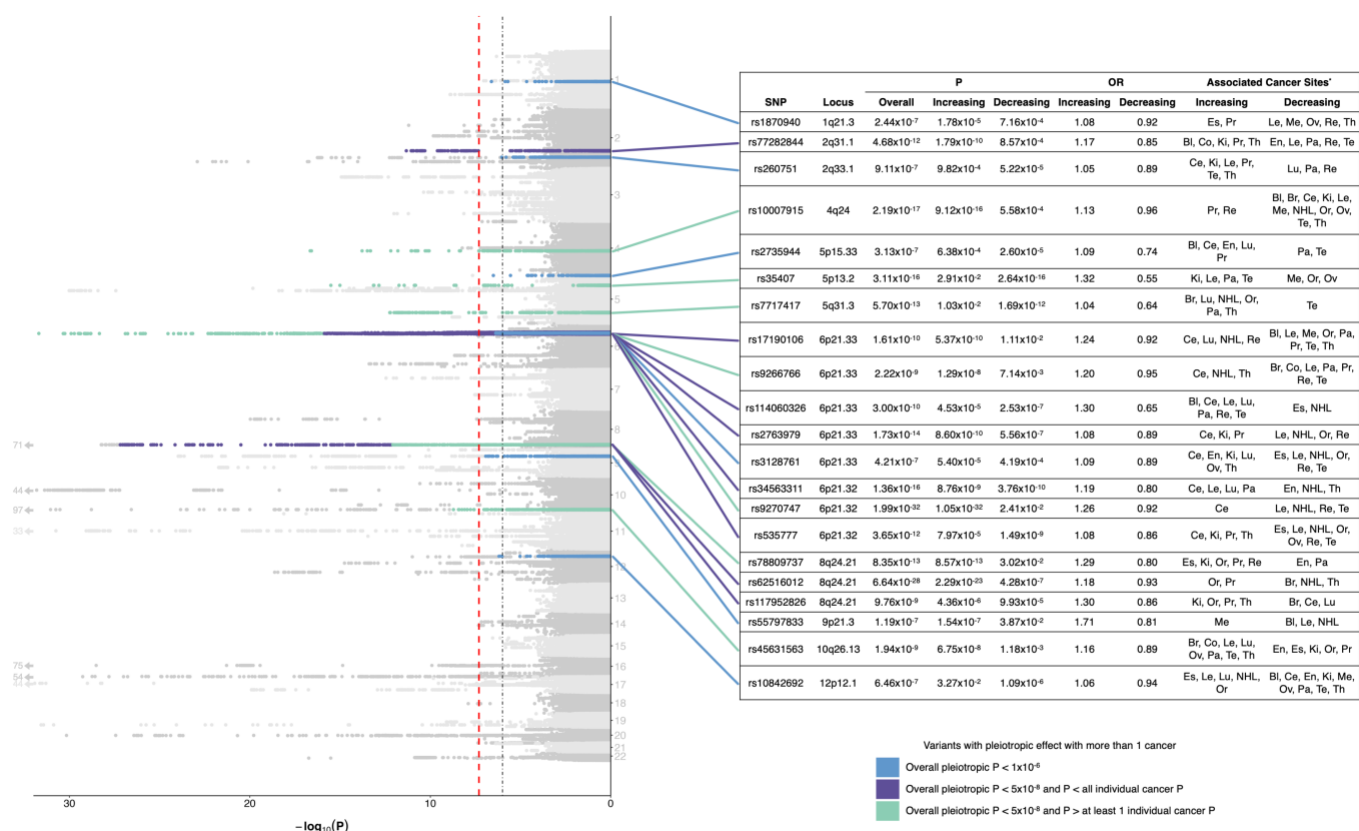


Figure 2. Manhattan plot displaying one-directional variant-specific pleiotropy, where the effect is maximized across all possible subsets of 18 cancers and assumes the same direction of effect across all selected cancers. The red dashed line represents the genome-wide significance threshold ($P < 5 \times 10^{-8}$), and the black dotted line represents a suggestive threshold ($P < 1 \times 10^{-6}$). Highlighted in purple are genome-wide significant loci where the overall pleiotropic P is less than all individual P for the selected cancers; details for the strongest signal at each locus are provided in the table, including the overall P and odds ratio (OR). Highlighted in green are the genome-wide significant loci where the overall pleiotropic P is greater than at least one of the individual P for the selected cancers, and highlighted in blue are loci with overall pleiotropic $P < 1 \times 10^{-6}$.



* Bl=Bladder, Br=Breast, Ce=Cervix, Co=Colon, En=Endometrium, Es=Esophagus/Stomach, Ki=Kidney, Le=Lymphocytic Leukemia, Lu=Lung, Me=Melanoma, NHL=Non-Hodgkin's Lymphoma, Or=Oral Cavity/Pharynx, Ov=Ovary, Pa=Pancreas, Pr=Prostate, Re=Rectum, Te=Testis, Th=Thyroid

Figure 3. Manhattan plot displaying omni-directional variant-specific pleiotropy, where the effect is maximized across all possible subsets of 18 cancers and allows for different directions of effect for selected cancers. The red dashed line represents the genome-wide significance threshold ($P < 5 \times 10^{-8}$), and the black dotted line represents a suggestive threshold ($P < 1 \times 10^{-6}$). Highlighted are loci with overall pleiotropic $P < 1 \times 10^{-6}$, the two directional $P < 0.05$, and not in LD with a one-directional SNP with smaller P ; loci in purple are genome-wide significant loci where the overall pleiotropic P is less than all individual P for the selected cancers, loci in green are genome-wide significant loci where the overall pleiotropic P is greater than at least one of the individual P for the selected cancers, and loci in blue have $P < 1 \times 10^{-6}$. Details for the strongest signal at each highlighted locus are provided in the table, including overall P and odds ratio (OR).



* Bl=Bladder, Br=Breast, Ce=Cervix, Co=Colon, En=Endometrium, Es=Esophagus/Stomach, Ki=Kidney, Le=Lymphocytic Leukemia, Lu=Lung, Me=Melanoma, NHL=Non-Hodgkin's Lymphoma, Or=Oral Cavity/Pharynx, Ov=Ovary, Pa=Pancreas, Pr=Prostate, Re=Rectum, Te=Testis, Th=Thyroid

Figure 4. The number of pleiotropic variants (of the 158 one- and omni-directional variants with overall pleiotropic $P<1\times10^{-6}$) associated with each pair of cancers by type of pleiotropic effect for select cancer pairs. Variants were counted for a cancer pair if they were associated with both cancers using ASSET. For omni-directional variants, effects in the same direction and in opposite directions were tabulated separately. Included are all cancer pairs with any of the following cancer sites: breast, cervix, lung, melanoma, Non-Hodgkin's Lymphoma (NHL), pancreas, and prostate.

