

1 **Sex-dependent polygenic effects on the clinical progressions of Alzheimer's disease**

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30 **Abstract**

31 Sex differences in the manifestations of Alzheimer's disease (AD) are under intense
32 investigations^{1,2}. Despite the emerging importance of polygenic predictions for AD³⁻⁸, the sex-
33 dependent polygenic effects have not been demonstrated. Here, using a sex crossover analysis,
34 we show that sex-dependent autosomal genetic effects on AD can be revealed by characterizing
35 disease progress via the hazard function. We first performed sex-stratified genome-wide
36 associations, and then applied derived sex-dependent weights to two independent cohorts. Sex-
37 matched polygenic hazard scores (PHS) have significantly stronger associations with age-at-
38 disease-onset, clinical progressions, amyloid depositions, neurofibrillary tangles, and composite
39 neuropathological scores, than sex-mismatched PHS, independent of apolipoprotein E. Models
40 without using hazard weights, i.e. polygenic risk scores (PRS), have lower predictive power than
41 PHS and show no evidence for sex differences. Our results indicate revealing sex-dependent
42 genetic architecture requires the consideration of temporal processes of AD. This has strong
43 implications not only for the genetic underpinning of AD but also for how we estimate sex-
44 dependent polygenic effects for clinical use.

45

46 Sex, as both an endogenous and an exogenous factor modulating human biology, has a
47 ubiquitous impact on the pathogenesis of complex diseases⁹. Evidence on sex-dependent
48 clinicopathological progressions of Alzheimer's disease (AD) is just beginning to emerge².
49 Compared to men, women show later manifestation of verbal memory deficits¹⁰, faster decline
50 after disease onset¹¹, and some differences in neuropathological characteristics, such as tau
51 tangle density^{1,12}. Results from studies on incidence rate and prevalence are less consistent^{13,14},
52 yet women are often reported to have increased incidence of AD in older ages¹⁵ and higher
53 prevalence¹⁶. Given this unmet need for better understanding of sex differences in AD, we
54 wanted to investigate the potential for a sex-dependent genetic architecture of AD. Despite
55 evidence suggesting that AD is highly polygenic, with a heritability as high as 79 percent¹⁷, so
56 far only apolipoprotein E (*APOE* e4) has been found to have a differential impact on age-at-onset
57 between men and women^{18,19}. Sex-dependent differences in polygenic effects remain unresolved.

58 This sex-agnostic *status quo* is particularly problematic for disease prediction based on
59 polygenic effects. By aggregating the estimated regression weights of autosomal single-
60 nucleotide-polymorphisms (SNPs) from genome-wide association studies (GWAS), polygenic
61 scores have been used to assist in several important clinical functions, including disease
62 prediction²⁰, risk stratification²¹, enriching clinical trials^{6,22}, and facilitating disease screening²³.
63 However, because the standard practice in GWAS treated sex as a confounding factor for
64 autosomal effects, the basis of polygenic scores, the estimated odds ratios, are devoid of sex-
65 dependent effects. Given the complexity of the moderating effects of sex on disease etiology⁹,
66 applying sex-agnostic polygenic scores may produce substantially biased risk quantifications.
67 Such scores could underestimate the genetic risk of AD for women, since *APOE* e4, one of the
68 most well-established risk factors for AD, has stronger effects on AD onset among women than

69 among men¹⁹. Given the heightened awareness of utilizing polygenic effects beyond *APOE* as
70 biomarkers for AD^{5,22,24–27}, understanding the sex-dependent polygenic effects for AD is
71 imperative for their application to clinical settings.

72 To investigate whether there are sex-dependent polygenic effects in addition to *APOE*,
73 we performed a sex *crossover* study (Methods and Figure 1) – we derived polygenic scores from
74 separate GWAS on men and women in the training cohorts (Alzheimer’s Disease Genetic
75 Consortium, ADGC, n = 17855; See Methods and Table 1), and then applied each of the sex-
76 dependent regression weights to both men and women in independent cohorts (National
77 Alzheimer’s Coordinate Center cohort, NACC, n = 6076; Religious Orders Study and Rush
78 Memory and Aging Project, ROSMAP, n = 599) to determine if there were a differential
79 performance in predicting AD. Importantly, the regression weights used as the basis for the
80 scores were based on Cox regressions, thereby capturing differences in clinical progression
81 between men and women as hazard functions.

82

83 **Methods**

84 Study Design

85 The crossover analysis is illustrated in Figure 1. First, we performed sex-stratified genome-wide
86 analyses on age-at-onset of AD, using imputed genotypes and phenotypic data from the
87 Alzheimer’s Disease Genetic Consortium (ADGC)^{28–30}. To ensure independence between the
88 training and validation cohorts, we performed an extensive check on potential sample overlap
89 and removed any overlapping individuals from the training data. The final training data included
90 7158 men and 10697 women (Table 1). Genome-wide Cox regression analyses were performed

91 on men and women separately to obtain sex-dependent weights. Detailed descriptions of the
92 analytical methods can be found in the following section and in the Supplemental Materials.

93 After obtaining the sex-dependent Cox regression weights for each autosomal SNP from
94 the ADGC data, we applied these weights to two independent cohorts (Table 1), generating men-
95 dependent polygenic hazard score (mPHS) and women-dependent polygenic hazard score
96 (wPHS) for every participant. Thus, we can compare whether sex-matched models (mPHS on
97 men and wPHS on women) has better predictive power than sex-mismatched models (wPHS on
98 men and mPHS on women), as a cross-over comparison (Figure 1).

99 The first independent cohort was obtained from the National Alzheimer’s Coordinate
100 Center (NACC). NACC recruits case series as a nationwide recruiting effort funded by National
101 Institute of Aging, involving clinical centers across United States. Given the longitudinal design
102 of NACC, we examined whether sex-matched PHS predicted dementia onset better than sex-
103 mismatched PHS. The cohort characteristics of NACC can be found in Table 1.

104 The second independent cohort was the Religious Orders Study and Rush Memory and
105 Aging Project (ROSMAP). ROS and MAP are two community-based cohort studies that enrolled
106 individuals without dementia, all of whom agreed to longitudinal follow-up and organ donation,
107 enabling us to examine the distribution of neuropathology among as a function of sex-specific
108 PHS. All participants signed an informed consent, Anatomic Gift Act, and repository consent
109 allowing their data to be shared. Both studies were approved by an Institutional Review Board of
110 Rush University. Details of the studies, generation of genomic data, and neuropathologic data
111 collection have been previously reported^{31,32}. We investigated whether sex-matched PHS has
112 stronger associations with neuropathology in the brain than sex-mismatched PHS. For those who

113 have both genotyping data and autopsy results were included in this analysis (n = 599). Detailed
114 characteristics of ROSMAP can be found in Table 1.

115 For comparison purposes, we also examined the performance of polygenic risk scores
116 (PRS) in the same manner as described above, except using weights from logistic regressions
117 while controlling for age-at-ascertainment. This is intended to investigate the benefit of using
118 Cox regressions in contrast to the standard GWAS approach.

119

120 Estimating sex dependent hazards for autosomal SNPs

121 To obtain sex-dependent weights for each SNP, we fitted genome-wide Cox regression models
122 on men and women separately. This stratified approach was intended to capture sex-specific
123 effects from autosomal SNPs without explicitly modeling interaction terms. This stratified
124 approach also allows for differences in the shape of the baseline hazard function between men
125 and women. As noted in prior studies on sex-dependent genetic effects⁹, although the total
126 sample size for GWAS is thus reduced by half, stratified models are computationally simple and
127 avoid the need for additional assumptions on the nature of sex interactions. Furthermore, hazard
128 ratio estimation is facilitated by utilizing Martingale residuals under null³³:

$$\hat{\beta} = (x^T x)^{-1} x^T M_0$$

129 where x is the mean centered genotype dosage and M_0 is the Martingale residuals of the null
130 model. More detailed discussion about the hazard estimates from case-control study can be found
131 in Supplemental Materials.

132 For ADGC data, we used the age-at-onset as the time-to-event and the age-at-last-visit as
133 the censoring time for Cox regression while controlling for dosages of *APOE* e2 and e4, the first
134 five genetic principal components, and indicators of recruiting sites. After filtering (minor allele

135 frequencies greater than 1 percent, in Hardy-Weinberg equilibrium, missing rate less than 10
136 percent, located outside of *APOE* or major histocompatibility complex regions), 6,784,887
137 imputed SNPs were included in our analyses. The resulting men- and women-derived hazard
138 ratios were used to generate the corresponding sex-dependent PHS. For comparison purposes, we
139 also performed standard GWAS with logistic regressions for the same 6,784,887 SNPs. All
140 covariates are the same in the models except age-at-ascertainment is now treated as one of the
141 covariates. The estimated sex-dependent odds ratios were then used to generate the
142 corresponding polygenic risk scores (PRS). Because our focus was on polygenic effects over and
143 above the effects of *APOE*, we excluded any SNPs located within *APOE* region when we
144 calculated all polygenic scores.

145

146 *Deriving polygenic hazard scores and polygenic risk scores*

147 The polygenic scores are the product sum of GWAS obtained weights and genotypes of
148 individuals in the two test cohorts:

$$S_i = \sum_{j=1}^M G_{ij} \beta_j$$

149 for individual i , the score S_i is the product sum of genotypes G_{ij} and weights β_j for M SNPs. To
150 make PHS and PRS comparable, we used the identical pruning and clumping process to select
151 independent SNPs for generating the scores. The parameters include clumping within 250kb and
152 linkage disequilibrium greater than 0.1, resulting in 251,040 independent SNPs for generating
153 the scores. No p-value thresholds were imposed to avoid using different numbers of SNPs
154 between the PHS and the PRS. Men-derived scores used weights for SNPs based on the GWAS
155 of men in ADGC, and similarly women-derived scores only used weights from GWAS of

156 women in ADGC. Both men- and women-derived scores were then computed for each
157 participant in the validation cohorts using the same autosomal SNPs. Crossover analyses can thus
158 be used to compare the predictive performance of sex-matched vs. sex-mismatched scores in the
159 validation cohorts.

160

161 Statistical analysis

162 We implemented genome-wide Cox regression for efficiently estimating hazard ratios across
163 millions of SNPs. P-values of the Cox regressions were obtained using score tests³⁴. The logistic
164 regression GWAS were performed using PLINK. All genome-wide analyses were done using
165 ADGC data, separately for men and women. In order to provide an intuitive interpretation on the
166 obtained weights, we also calculated gene-based effect sizes using Pascal³⁵. Pascal obtained
167 gene-based p-values are based on a linkage-disequilibrium weighted average of effect sizes of
168 SNPs located within 50Kb regions of the gene body.

169 In NACC, we used 1). Cox regression to examine the predictive power of polygenic
170 scores on AD age-at-onset, and 2). linear mixed effects model to examine the associations
171 between polygenic scores and rate of clinical progression, defined as changes in Cognitive
172 Dementia Rating – Sum of Boxes (CDR-SB). All models were controlled for *APOE* status
173 (dosages of e2 and e4) and education levels. The main analysis of NACC included 2628 men and
174 3448 women. We also examined whether the patterns of association remained constant if we
175 restricted analyses to neuropathologically-confirmed cases; 817 men and 706 women from
176 NACC had post-mortem neuropathological examinations. To ensure the consistency of the units,
177 all results are based on standardized polygenic scores, comparing changes in 1 standard deviation
178 (SD) of scores.

179 In ROSMAP, we analyzed the relationship between the neuropathological burden at
180 autopsy and sex-dependent polygenic scores. Four quantifications of neuropathology were
181 included, i.e., the percentage area occupied by β -amyloid, and the density of tau-positive
182 neurofibrillary tangles. Because those neuropathological measures were skewed, we performed a
183 square root transformation to normalize the neuropathology data. We also determined Braak
184 stage, and Consortium to Establish a Registry for Alzheimer's disease (CERAD) score. All
185 regression models controlled for *APOE* status (dosages of e2 and e4), age-at-death, and
186 education level. To ensure the consistency of the units, all results are based on standardized
187 polygenic scores and neuropathological data, comparing neuropathological variations in 1
188 standard deviation (SD) of scores.

189

190 Code availability

191 The code for the genome-wide Cox regressions will be available on GitHub

192

193 Data availability

194 The summary statistics for genome-wide hazard estimates and gene-based analyses will be found
195 in the Supplemental Materials.

196

197 Distribution of hazard weights

198 First, we performed genome-wide Cox regressions for AD age-at-onset on ADGC individuals
199 (men/women = 7158/10697). The models were controlled for first 5 genetic principal
200 components, *APOE* status, and recruiting sites (Methods). We noticed that there are different top
201 hits between men and women (Figure 2A. and Fure 2B.). Men had a GWAS-significant locus on

202 1q32.2, encompassing *CRI*, and women had a GWAS-significant locus on 2q14.3, encompassing
203 *BINI*. In addition to GWAS-significant loci, polygenic signals below the GWAS-significant
204 threshold are important for deriving polygenic scores. To provide an intuitive summary on the
205 sex-dependent polygenic effects, we performed gene-based analyses using Pascal³⁵. Figure 2C
206 illustrates the sex-dependent distributions from gene-based analyses. Gene clusters on 19q13.32
207 continue to show consistent effects between men and women, with trends for sex-specific genetic
208 effects. For example, the effect sizes of *BINI*, *MS4A6A*, *DNAJA2*, and *FERMT2* are larger
209 among women while *FAM193B*, *C2orf47*, *TYW5* have larger effect sizes among men.
210 Additionally, the tau-related gene, *MAPT*, shows stronger effects on men than on women.

211

212 Predicting clinical manifestations in NACC

213 By aggregating the hazard weights obtained from genome-wide Cox regressions of ADGC, we
214 derived women specific polygenic hazard scores (wPHS) and men specific polygenic hazard
215 scores (mPHS), using standard pruning and clumping process (Methods), for every individual in
216 the NACC cohort (men/women = 2628/3448), resulting in sex-matched model (men with mPHS
217 and women with wPHS) and sex-mismatched model (men with wPHS and women with mPHS).
218 To avoid the confounding of *APOE* due to imputations, we excluded any genetic variants located
219 at *APOE* region (Methods). For clinically determined AD onset, the sex-matched model
220 consistently performed more accurately than the sex-mismatched model (Figure 3A). After
221 controlling for *APOE* status, sex-matched PHS has a hazard ratio (HR) of 1.26 (95% CI: 1.26 –
222 1.32, $p < 1e-16$) and sex-mismatched PHS has a hazard ratio (HR) of 1.14 (95% CI: 1.09 – 1.19,
223 $p = 1e-10$). Sex-matched PHS performed significantly better than sex-mismatched PHS ($p =$
224 0.001). Subgroup analyses indicate that stronger predictive power in sex-matched models than

225 sex-mismatched models is evident for both men and women (Supplemental Figure 1A). When
226 we limited our analysis to those with neuropathological disease confirmation ($n = 1523$), the
227 crossover effects were consistent (HR: 1.21, $p = 2e-9$, Figure 3B, Supplemental Figure 1B) and
228 retaining significant difference between sex-matched and mismatched models ($p = 0.008$). Figure
229 3C shows the performance of polygenic scores in predicting clinical progressions as CDR-SB
230 changes during longitudinal follow-up in NACC. Sex-matched PHS was predictive of annual
231 changes of CDR-SB ($\beta: 0.057$, 95% CI: 0.049 – 0.064, $p < 1e-16$) and performed better than sex-
232 mismatched PHS ($\beta: 0.043$, 95% CI: 0.035 – 0.050). The difference between sex-matched PHS
233 and sex-mismatched PHS was statistically significant ($p = 0.006$). In contrast, PRS from logistic
234 regressions showed no significant associations regardless of which sex-dependent PRS were
235 applied (Figure 3A-C).

236

237 *Predicting neuropathology in ROSMAP*

238 Figure 4 demonstrates the association strengths across four different types of neuropathology.
239 After controlling for age at death, education levels, and *APOE* status, sex-matched models have
240 significantly stronger associations than sex-mismatched models for all neuropathological
241 measures (p values for differences between sex-matched and sex-mismatched PHS as $5e-5$, $4e-7$,
242 0.007 , and $5e-4$ for amyloid deposition, CERAD score, tau associated neurofibrillary tangles,
243 and Braak score, respectively). None of the sex-mismatched models reached statistical
244 significance in predicting neuropathology based on polygenic components. Table 2 summarizes
245 the variance explained for subgroup analyses on each neuropathology. Compared to sex-
246 mismatched models, wPHS applied to women increased the variance explained 6 percent, 5
247 percent, 3 percent, and 6 percent for amyloid deposition, CERAD score, neurofibrillary tangles,

248 and Braak score, respectively. Applying mPHS to men would increase 1 percent, 3 percent, 3
249 percent, and 4 percent for amyloid deposition, CERAD score, neurofibrillary tangles, and Braak
250 score, respectively. In general, variance explained attributable to the polygenic components for
251 sex matched models can reach up to 89 percent of variance explained by *APOE* only. In contrast,
252 sex-matched PRS had no significant associations with any neuropathologies except CERAD
253 score, whereas no evident sex differences after controlling for *APOE* (Figure 4 and Supplemental
254 Figure 2).

255

256 Understanding the sex-dependent polygenic architecture of AD

257 By modeling the disease courses as time-to-clinical-onset, the polygenic hazard approach
258 revealed sex-dependent autosomal effects on AD after controlling for *APOE*. Sex-matched PHS
259 showed better prediction of both clinical age-at-onset and neuropathological manifestations than
260 sex-mismatched PHS, implying that genetic risk factors differ between men and women. These
261 findings have implications not only for the etiology of AD, but also a new approach to examine
262 sex differences in genetic risks.

263 Many of the genes highlighted by our analyses have been implicated in AD in prior
264 reports^{28–30,36}. Yet, our survival analyses revealed a complex landscape of sex-dependency
265 across the genome. Loci such as *BINI*, *MS4A6A*, *DNAJA2*, and *FERMT2* contribute higher risk
266 to women than to men. Previous GWAS have identified *BINI* and *MS4A6A* as risk loci for AD³⁶,
267 but our results indicate that their effects may be sex dependent, especially for pathologica aging
268 processes. Experimental studies have found that *FERMT2* is associated with amyloid deposition
269³⁷ whereas *DNAJA2* interacts with protein tau aggregation³⁸. When aggregating those

270 differences as PHS, the sex-dependency of the genetic effects emerged, indicating there are
271 divergent pathological pathways between men and women.

272 In addition to the pathogenesis of AD, these crossover analyses also highlight an
273 important aspect for modeling genetic risks – time. AD is an insidious, progressive disease.
274 When the genetic effects on disease risks are differentially expressed across time, the mean
275 liability model cannot readily capture differences in the underlying genetic risks^{39,40}. In our
276 analyses, PRS had limited predictive accuracy on both AD onset and neuropathology, regardless
277 of sex-dependencies. This strongly suggests that explicit modeling of time of clinical disease
278 onset using survival analyses is needed to reveal sex-dependent effects in polygenic signals.
279 Considering one of the key differences between men and women with respect to AD is the
280 temporal disease course, and hence the underlying hazard function, sex-dependent polygenic
281 effects may largely modulate the temporal disease course for AD.

282 Sex differences are ubiquitous in human biology and disease manifestations, yet are
283 rarely reported in terms of genetic risks⁹. Our results indicate that by explicitly modeling age-
284 dependent hazards in sex-stratified analyses, we can reveal these sex-dependent effects. In
285 addition to providing insight about sex-differences in AD pathophysiology, we also hope this
286 study will encourage improvements in GWAS study design to consider sex differences regarding
287 time of disease onset.

288

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292 ADGC and NACC can be found in the Supplemental Information.

293

294 **Author Contributions**

295 C.C.F., S.J.B., R.D., and A.M.D. conceived and designed the study. C.C.F., S.J.B., and R.D.

296 acquired, analyzed, and interpreted the data. C.C.F., S.J.B., R.D., and A.M.D. drafted the

297 manuscript. W.K.T., C.H.C., L.K.M., C.H.T., W.K., D.A.B., L.A.F., R.M., G.D.S., and O.A.A.

298 critically revised the manuscript for important intellectual content.

299

300 **Competing Interests**

301 C.C.F. is under employment of Multimodal Imaging Service, dba Healthlytix, in addition to his

302 research appointment at the University of California, San Diego. A.M.D. is a founder of and

303 holds equity interest in CorTechs Labs and serves on its scientific advisory board. He is also a

304 member of the Scientific Advisory Board of Healthlytix and receives research funding from

305 General Electric Healthcare (GEHC). The terms of these arrangements have been reviewed and

306 approved by the University of California, San Diego in accordance with its conflict of interest

307 policies.

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399

400 **Tables**

401 **Table 1. Characteristics of training samples and independent validating cohorts**

402

403 **Table 2. Variance explained of neuropathological indices for subgroup analysis in**

404 **ROSMAP**

405

406 **Figure Legends**

407 **Figure 1. Flow chart of our sex crossover analysis.**

408

409 **Figure 2. Effect size distributions of obtained hazard weights from sex stratified**

410 **genomewide Cox regressions.** A. Manhattan plot from genomewide Cox regression from men

411 in ADGC. B. Manhattan plot from genomewide Cox regression from women in ADGC. C.

412 Results from gene-based analysis. The diagonal dashed line represents the equivalent effect sizes

413 given the sample size differences. We listed top 10 rank genes in terms of $-\log_{10}(p)$ from the

414 Pascal. Genes in both top 10 rank list of men and women are colored in red. Genes in only top 10

415 rank list of women are colored in green and of men are colored in blue.

416

417 **Figure 3. Predictive performance of polygenic components in NACC.** Weights from Cox

418 regressions of training data were applied to all participants in NACC, yielding both mPHS and

419 wPHS for all. The hazard ratios of comparing 1 standard deviation differences in PHS, after

420 controlling *APOE* and education levels, are shown. A. Prediction of clinically defined AD. B.

421 Prediction in neuropathologically confirmed AD cases, C. Prediction in CDR-SB changes

422

423 **Figure 4. Associations with neuropathology in ROSMAP.** Sex dependent polygenic scores
424 were obtained for all participants in ROSMAP. The coloring schemes are consistent with Figure
425 3. All models controlled for age at death, education levels, and *APOE* status. A. Associations
426 with amyloid deposition, B. Associations with CERAD score, C. Associations with
427 neurofibrillary tangles, D. Associations with Braak score.
428

429 **Table 1. Characteristics of training samples and independent validating cohorts**

430

	Training samples		Independent testing cohorts			
	<i>ADGC*</i>		<i>NACC</i>		<i>ROSMAP</i>	
	Men	Women	Men	Women	Men	Women
Total N	7158	10697	2628	3448	220	379
Age - years (SD)	75.4 (7.7)	75.9 (8.2)	78.6 (9.4)	79.1 (9.8)	86.4 (6.3)	89.4 (6.2)
AD cases/events	42.7%	47.6%	52.3%	41.5%	37.7%	43.8%
APOE ε4 carriers	40.9%	43.3%	40.9%	37.9%	29.5%	28.4%

431

432 * Excluded any overlapping samples with NIA ADCs and ROSMAP.

433 † NP – neuropathology. NP samples means number of samples with post-mortem neuropathology examinations.

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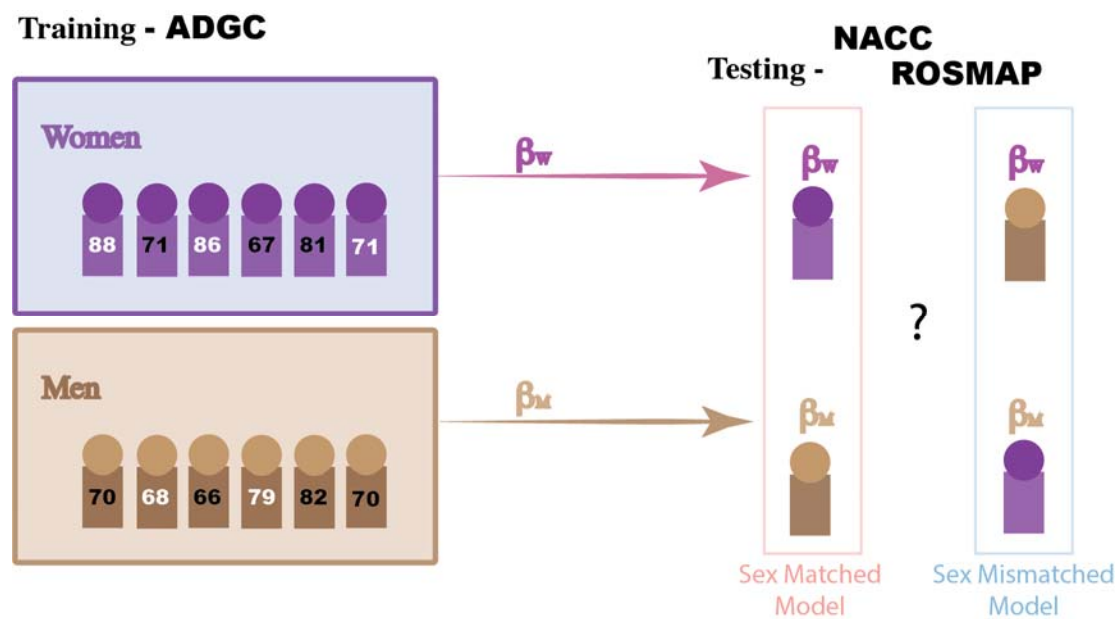
436 **Table 2. Variance explained of neuropathological indices for crossover models in ROSMAP**

Pathology	Testing Subjects	Covariates only	plus E2 + E4	Sex-matched PHS	PHS / APOE *	
	Amyloid	Women	2%	12%	17%	55%
Amyloid Related Pathology	Amyloid	Men	5%	12%	13%	11%
	CERAD	Women	1%	11%	16%	51%
	CERAD	Men	3%	9%	12%	59%
	Tangles	Women	2%	15%	19%	24%
Tau related pathology	Tangles	Men	4%	10%	13%	54%
	Braak	Women	5%	11%	17%	89%
	Braak	Men	8%	15%	19%	59%

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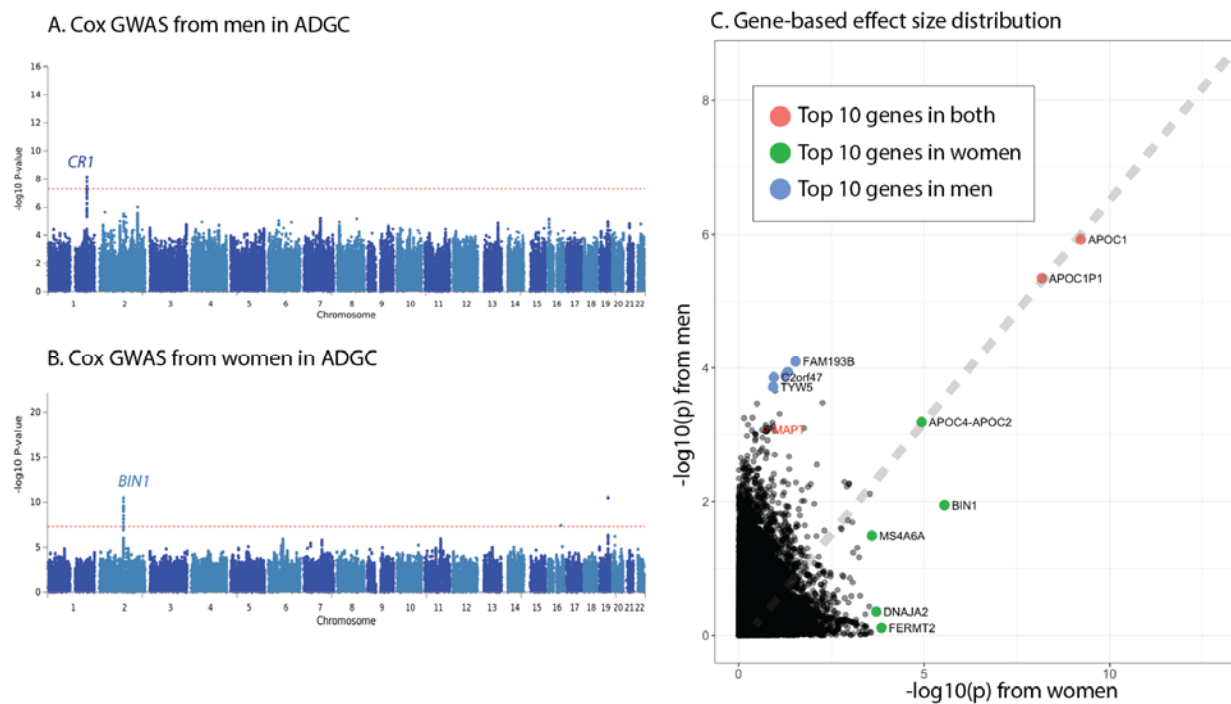
438 * Amount of variance explained attributable to polygenic component over the amount attributable to *APOE* dosages

Figure 1.



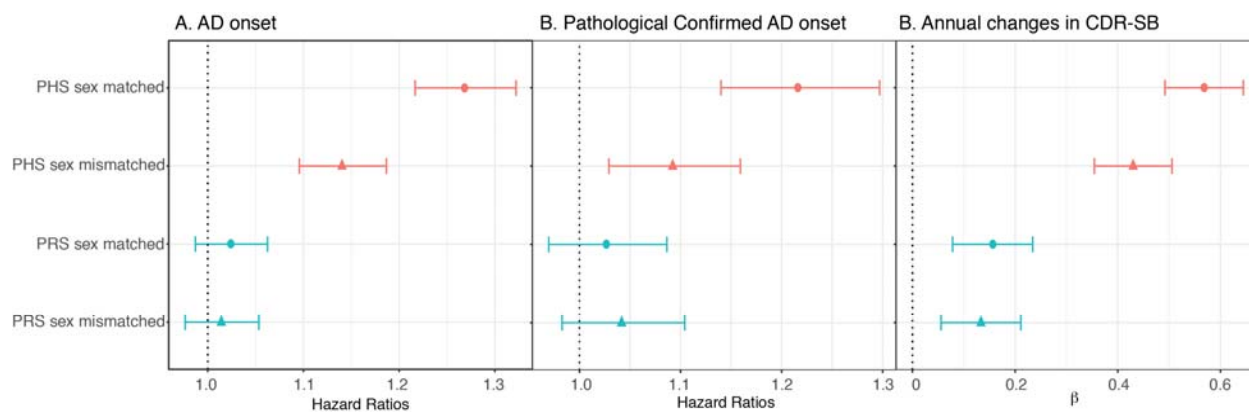
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Figure 2.



440

Figure 3.



441

Figure 4.

