1	Proxy gene-by-environment Mendelian randomization study confirms a causal effect of maternal
2	smoking on offspring birthweight, but little evidence of long-term influences on offspring health
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19 Key Words: gene × environment, Mendelian randomization, proxy, maternal smoking, pregnancy

20 Abstract

21	Objective To validate a novel proxy gene-by-environment (G×E) Mendelian randomization (MR)
22	approach by replicating the previously established effect of maternal smoking heaviness in
23	pregnancy on offspring birthweight, and then use GxE MR to investigate the effect of smoking
24	heaviness in pregnancy on offspring health outcomes in later life and grandchild's birthweight.
25	Design A proxy G×E MR using participants' genotype (i.e. rs16969968 in CHRNA5) as a proxy for their
26	mother's genotype.
27	Setting UK Biobank.
28	Participants 289,684 white British men and women aged 40-69 in UK Biobank.
29	Main outcome measures Participants' birthweight and later life outcomes (height, body mass index,
30	lung function, asthma, blood pressure, age at menarche, years of education, fluid intelligence score,
31	depression/anxiety, happiness), and birthweight of female participants' first child.
32	Results In our proof of principle analysis, each additional smoking-increasing allele was associated
33	with a 0.018 (95% confidence interval (CI): -0.026, -0.009) kg lower birthweight in the "maternal
34	smoking during pregnancy" stratum, but no meaningful effect (-0.002kg; 95% CI: -0.008, 0.003) in
35	the "maternal non-smoking during pregnancy" stratum (interaction P-value=0.004). We found little
36	evidence of an effect of maternal smoking heaviness on participants' later life outcomes. We found
37	the differences in associations of rs16969968 with grandchild's birthweight between grandmothers
38	who did versus did not smoke were heterogeneous (interaction P-value=0.042) among female
39	participants who did (-0.020kg per allele; 95% CI: -0.044, 0.003) versus did not (0.007kg per allele;
40	95% CI: -0.005, 0.020) smoke in pregnancy.
41	Conclusions Our study demonstrated how offspring genotype can be used to proxy for mothers'
42	genotype in G×E MR. We confirmed the previously established causal effect of maternal smoking on

43 offspring birthweight but found little evidence of an effect on long-term health outcomes in the

- 44 offspring. For grandchild's birthweight, the effect of grandmother's smoking heaviness in pregnancy
- 45 may be modulated by maternal smoking status in pregnancy.
- 46 (word count: 300)

47 WHAT IS ALREADY KNOWN TO THIS TOPIC

- 48 Heavier maternal smoking in pregnancy causes lower offspring birthweight
- 49 Maternal smoking in pregnancy is also associated with offspring outcomes in later life and
- 50 grandchild's birthweight, but it is not known whether these associations are causal
- 51 Understanding the transgenerational causal effects of maternal smoking heaviness in pregnancy is
- 52 important to inform public health policies

53 WHAT THIS STUDY ADDS

- 54 The proxy gene-by-environment Mendelian randomization approach can be used to explore
- 55 maternal effects on offspring phenotypes when maternal genetic information is unavailable
- 56 The approach confirmed the causal effect of smoking on offspring birthweight.
- 57 Maternal smoking status in pregnancy modulates the effect of grandmother's smoking heaviness in
- 58 pregnancy on grandchild's birthweight, highlighting the importance of smoking cessation before
- 59 pregnancy in each generation

60 Introduction

61 The developmental origins of health and disease hypothesis proposes that early life experiences, 62 including those in utero, can have long-term health effects, and maternal pregnancy exposures are important to long-term health of offspring (1). Heavier maternal smoking in pregnancy is known to 63 be causally associated with lower offspring birthweight (2-6), but its other effects in offspring are 64 less clear. Multivariable regression in observational data showed that heavier maternal smoking 65 66 during pregnancy was associated with offspring being shorter (7) and more overweight/obese (8, 9), 67 and having higher blood pressure (10), but had mixed associations with age at menarche (11), respiratory (12), cognitive (13), and mental health (14). Heavier maternal smoking in pregnancy has 68 69 also been associated with higher grandchild's birthweight in certain subpopulations (15-17). It is unclear whether these associations reflect a causal effect of maternal smoking in pregnancy, as they 70 71 may be due to residual confounding. Some studies have assessed this using paternal smoking as a 72 'negative control' since an effect via uterine environment would be observed in mothers but not 73 fathers, such that similar-magnitude associations would indicate confounding via shared familial, 74 social, environmental and genetic factors (2, 5, 18). Negative control studies suggest little evidence 75 of a causal effect on offspring body mass index (BMI) (2, 5, 8), blood pressure (19, 20) and 76 depression (21).

77 Mendelian randomization (MR) provides an alternative way to explore this question by using single 78 nucleotide polymorphisms (SNPs) as instrumental variables (IVs) for an exposure of interest. MR is 79 less prone to confounding as germline genetic variants are randomly allocated at meiosis and not 80 influenced by subsequent socioeconomic and health behaviours (22, 23). MR has been applied in a 81 gene-by-environment (G×E) framework (24, 25), which requires variation in the strength of the 82 gene-exposure association across strata of another factor. If there is a causal effect of the IV on the 83 outcome via the exposure of interest, then we would expect the association of the IV with the 84 outcome to vary in proportion to the gene-exposure association. The rs1051730/rs16969968

85 (CHRNA5) SNPs, previously robustly associated with smoking heaviness amongst smokers (26), have 86 been widely used as IVs for smoking heaviness in GxE MR studies (3, 27-29). A causal effect of the 87 smoking heaviness IV on an outcome should be seen amongst ever but not amongst never smokers 88 if the effect is via smoking heaviness rather than other pathways (24, 25). G×E MR has also been 89 used to assess cross-generational causal effects. A smoking heaviness IV has been associated with 90 lower offspring birthweight amongst mothers who smoked in pregnancy but not amongst mothers 91 who did not smoke in pregnancy, suggesting the genetic instrument affects birthweight through 92 maternal smoking (3).

93 It is usually difficult to investigate transgenerational associations due to a lack of data across the 94 generations of interest. Thus, previous work has sought to test transgenerational associations using 95 available traits as proxies for unmeasured traits of interest. A Norwegian cohort aimed to examine 96 whether women's smoking in adulthood was related to their mothers' smoking habits (that were not 97 recorded) and hence used maternal smoking-related mortality as a proxy (30). Recently, a case-98 control by proxy approach has been proposed (31). Participants' genotypes were used to proxy 99 unavailable parental genotypes, and their associations were tested against parental diagnosis of 100 Alzheimer's disease in UK Biobank (31), since Alzheimer's disease was much more prevalent in the 101 parents than the participants (aged between 40 and 69 at baseline in 2006-2010 (32)). Our study 102 aimed to demonstrate how an analogous approach can be used within a G×E MR framework to test maternal-offspring effects when maternal genotype is not available, using offspring genotype as a 103 104 proxy for the maternal genotype. First, we performed a proof of principle analysis to demonstrate 105 this approach, testing the previously established finding that maternal smoking in pregnancy leads to 106 lower offspring birthweight. Second, we tested for causal effects of maternal smoking on offspring 107 later life outcomes. Finally, we tested for a causal effect of grandmother's smoking on grandchild's 108 birthweight.

109 Methods

110 Study population

- 111 Our study was conducted using UK Biobank, a population-based cohort of more than 500,000 men
- and women in the UK. This study collected a large and diverse range of data from physical measures,
- 113 questionnaires and hospital episode statistics (32). Of 463,013 participants of European descent with
- genetic data passing initial quality control (i.e. genetic sex same as reported sex, XX or XY in sex
- chromosome and no outliers in heterozygosity and missing rates) (33), 289,684 participants (54%
- 116 women) of white British descent were eligible for inclusion in our analyses (Supplementary Figure 1).
- 117 We refer to the UK Biobank participants as generation one (G1), and their parents and offspring as
- 118 G0 and G2, respectively.

119 Genetic IV for maternal smoking

- 120 The rs16969968 SNP located in *CHRNA5* has been robustly associated with smoking heaviness (26).
- 121 Ideally, we would use the maternal rs16969968 as an IV for the heaviness of maternal smoking, but
- in UK Biobank parental genetic data are not available. Hence, we used rs16969968 of the UK Biobank
- 123 participants as a proxy for that of their mothers', coded as the number of smoking heaviness

increasing alleles.

125 Smoking phenotypes

We used participants' answers to the question "Did your mother smoke regularly around the time when you were born?" as a proxy for G0 smoking during pregnancy. Participants were also asked to report their smoking status (current/former/never). We derived a binary ever versus never measure of smoking status by combining current and former smokers. For female participants with at least one live birth, we derived a measure denoting whether they smoked during the pregnancy of their first child (see Supplementary Methods).

132 Outcomes in participants (G1)

133 We used baseline data measured at the UK Biobank initial assessment center. Anthropometric traits 134 included participants' birthweight (kg, self-reported), standing height (cm) and BMI (kg/m², 135 constructed from standing height and weight). To assess lung function, forced vital capacity (L) and 136 forced expiratory volume in 1-second (L) were measured by spirometry. Participants reported 137 whether they had had asthma via the question "Has a doctor ever told you that you have had any of 138 the following conditions?" (with an option of asthma) (34). Systolic and diastolic blood pressure 139 (mmHg) were measured twice using a digital monitor or a manual sphygmomanometer if the digital 140 monitor could not be employed, and we took the average of the two readings. Female participants 141 reported their age at menarche. We derived years of education based on qualifications achieved by 142 participants, as described previously (35). We included follow-up data of a subset of participants to 143 define intelligence and depression/anxiety. Fluid intelligence score was generated as an unweighted 144 sum of the number of correct answers given to 13 questions, and we used the earliest score if we 145 had data at multiple time points (36). We defined depression/anxiety cases as participants that 146 either answered "Yes" to "Have you ever seen a general practitioner (GP) for nerves, anxiety, tension 147 or depression?" or "Have you ever seen a psychiatrist for nerves, anxiety, tension or depression?". 148 or had hospital episode coded using ICD-10 (37). Happiness was assessed via a question – "In general how happy are you?", with six categories ranging from "extremely happy" to "extremely unhappy". 149 150 *Outcomes in participants' offspring (G2)*

- 151 The female participants with at least one live birth were asked to report their first child's
- birthweight. Male participants were not asked to report the birthweight of their offspring.

153 Statistical analyses

- Proof of principle analysis: testing the causal effect of maternal (G0) smoking heaviness in pregnancy
 on participants' (G1) birthweight
- 156 In this proof of principle analysis, we seek to replicate the finding, previously established using GxE

157 MR and many other methods (6), that heavier maternal smoking causes lower offspring birthweight. 158 We use our proxy GxE approach, where participants' (G1) genotype is used as a proxy for their 159 mothers' (G0) genotype. To assess whether rs16969968 affects participants' birthweight via G0 160 smoking in pregnancy, we stratified our G1 sample by G0 smoking status during pregnancy, and then 161 tested the associations of rs16969968 with birthweight in each stratum using multivariable linear 162 regression. Since birth precedes smoking initiation, participants' genotype cannot affect birthweight 163 through their own smoking heaviness, which means we do not need to consider smoking status of 164 participants (Figure 1A). We included participants' sex as a covariate to reduce variation in their 165 birthweight and the first ten principal components to control for population stratification. We 166 assumed an additive genetic effect and identified the strength of interaction between strata using 167 Cochran's Q test for heterogeneity.

168 Testing for causal effects of G0 smoking in pregnancy on G1 later life outcomes

169 We use the proxy GxE MR approach to test for causal effects of maternal (G0) smoking heaviness on 170 offspring (G1) height, BMI, lung function, asthma, blood pressure, age at menarche, education, 171 intelligence, depression/anxiety and happiness. In contrast to our proof of principle example where 172 participants smoking in adulthood cannot influence their birthweight, participants' rs16969968 173 could affect these outcomes via both maternal (G0) and participants' (G1) smoking heaviness (Figure 174 1B). To assess whether rs16969968 may affect these outcomes via maternal versus participants' 175 smoking, we stratified on both maternal and participants' smoking status. In each stratum, we 176 examined associations of rs16969968 with height, BMI, lung function, blood pressure, age at menarche, education and intelligence using linear regression, asthma and depression/anxiety using 177 178 logistic regression, and happiness using ordinal logistic regression. We included participants' age at 179 baseline, sex and the first ten genetic principal components as covariates.

Height and age at menarche manifest around the time of puberty such that participants' ownsmoking can only affect these if they started smoking before these outcomes are determined. We

182 conducted sensitivity analyses for these outcomes stratifying G1 participants according to whether
183 they were ever smokers before achieving their adulthood height (assuming age at 17 for men and 15
184 for women (38)) or their age at menarche.

185 Testing for causal effects of G0 smoking in pregnancy on grandchild's (G2) birthweight

- 186 To test for a causal effect of participants mothers' smoking on birthweight of participants' offspring,
- 187 we stratified G1 women based on their own and their mothers' smoking status during pregnancy, as
- rs16969968 could affect G2 birthweight through both G0 and G1 smoking heaviness (Figure 1C).
- 189 Within each stratum, we assessed associations of rs16969968 with G2 birthweight using linear
- 190 regression, adjusting for the first ten genetic principal components. We estimated the strength of
- 191 interaction between G0 smokers and G0 non-smokers within each G1 stratum. We also calculated a
- difference (39) in those associations between G0 smokers and G0 non-smokers within each G1
- 193 stratum, and estimated the strength of interaction between two differences to investigate whether

194 G1 smoking status modulates the effect of rs16969968 on G2 birthweight.

- 195 Our G × E MR may be vulnerable to collider bias (29, 40, 41), as we stratified on smoking status
- 196 (Supplementary Figure 2). Therefore, we tested associations of rs16969968 with G0 and G1 smoking
- 197 status and potential confounders available in UK Biobank (see Supplementary Methods). We also
- 198 tested observational associations of maternal (G0) smoking status with offspring (G1) smoking status
- and all outcomes for comparison with our MR results. Analyses were performed using R version
- 200 3.5.1 (R Foundation for Statistical Computing, Vienna, Austria).
- 201 Patient and public involvement

The current research was not informed by patient and public involvement because it used secondary
data. However, future research following on from our findings should be guided by patient and
public opinions.

205 Results

206	Characteristics of participants across sex are shown in Supplementary Table 1. Each additional
207	smoking-increasing allele of participants' rs16969968 was associated with a 1.02 (95% confidence
208	interval [CI]: 1.01, 1.03; P-value = 5×10^{-3}) higher odds of their mothers' smoking in pregnancy, a 0.98
209	(95% CI: 0.97, 0.99; P-value = 7×10^{-4}) lower odds of being an ever (versus never) smoker themselves,
210	and a 1.06 (95% CI: 1.04, 1.09; P-value = 3×10^{-7}) higher odds that female participants were a smoker
211	(versus non-smoker) in their own pregnancy. We found little evidence of an association between
212	rs16969968 and potential confounders, with small associations for participants' age and years of
213	education in some strata (Supplementary Table 2).
214	Our proof of principle analysis found that, amongst participants whose mothers smoked in
215	pregnancy, each additional smoking-increasing allele was associated with a 0.018kg lower
216	birthweight (95% CI: -0.026, -0.009) after adjustment for covariates (Figure 2). Amongst participants
217	whose mothers did not smoke in pregnancy, we found little evidence for an association of
218	rs16969968 with birthweight (-0.002kg [95% CI: -0.008, 0.003]), and we observed heterogeneity
219	between these associations (interaction P-value = 0.004).
220	Figure 3 showed estimates of rs16969968 on the 12 outcomes in the UK Biobank participants.
221	Overall, within each stratum, the estimates were broadly consistent between those whose mothers
222	smoked and those whose mothers did not, except for height among participants who never smoked
223	(all interaction P-values were in Supplementary Table 3). Each additional smoking-increasing allele
224	was associated with a 0.115cm lower height (95% CI: -0.200, -0.030) among never smokers whose
225	mothers smoked in pregnancy, but a 0.002cm lower height (95% CI: -0.057, 0.053) among never
226	smokers whose mothers did not smoke in pregnancy (interaction P-value = 0.029). However, this
227	difference was not observed amongst ever smokers (Figure 3A). We obtained largely consistent
228	results in sensitivity analyses (Supplementary Figure 3).

229 Figure 4 showed estimates of rs16969968 on grandchild's birthweight. Among female participants

230 who did not smoke in pregnancy, each additional smoking-increasing allele was associated with a 231 0.007kg higher grandchild's birthweight difference (95% Cl: -0.005, 0.020) between grandmothers 232 who did versus did not smoke in pregnancy. However, this difference was -0.020kg per allele (95% 233 CI: -0.044, 0.003) among female participants who smoked in pregnancy. These two differences were 234 heterogeneous (-0.028kg per allele [95% CI: -0.055, -0.001]; interaction P-value=0.042). 235 The directions of observational estimates were consistent with our MR estimates for both 236 participants' and their child's birthweight. Our observational analyses also found associations of 237 maternal smoking in pregnancy with offspring later life outcomes, where smoking in pregnancy was associated with lower height, higher BMI, poorer lung function, higher risk of asthma, earlier age at 238

239 menarche, higher blood pressure, and poorer cognitive and mental health (Supplementary Table 4).

240 Discussion

241 *Principle findings and comparison with the literature*

In this study, we have demonstrated how G×E MR can be used to test transgenerational causal
effects of maternal smoking heaviness in pregnancy using participants' genotype as a proxy for their
mothers' genotype. Our proof of principle analysis identified an effect of heavier maternal smoking
on lower offspring birthweight, consistent with previous studies (2-6). Our MR study also confirmed
previously established causal effects of participants' smoking on their own health, where heavier
smoking reduced BMI (27) and lead to impaired lung function (42), but found little evidence of an
effect on asthma risk (43) or blood pressure (28).

249 Our tests of effects of maternal smoking heaviness on offspring later life health outcomes were not 250 conclusive, given a lack of precision for many of our MR estimates. We found little evidence of an 251 effect on BMI, lung function, asthma, blood pressure, cognition, depression/anxiety or happiness. 252 These findings were consistent with negative control studies for BMI (2, 8), blood pressure (19, 20) 253 and depression/anxiety (21), although our estimation of interactions is not directly quantitatively 254 comparable to their estimation of effects of ever/never smoking or smoking heaviness categories in 255 observational studies. Our MR results found little evidence to support findings from our own and 256 previous observational studies indicating maternal smoking led to poorer lung function (44), higher 257 risk of asthma (45, 46), and lower happiness in offspring (47). This may be due to residual 258 confounding in observational associations, or because of low statistical power in MR. Previous 259 studies did not use the same cognition measurement approaches as used UK Biobank, making our 260 results for this outcome less comparable. We observed lower offspring adulthood height according 261 to maternal smoking in never smokers but not in ever smokers, which could be a chance finding 262 given we tested multiple outcomes.

We found little evidence of an effect of maternal smoking in pregnancy on offspring age at
menarche. However, we did find an effect of rs16969968 on age at menarche across strata of

smoking status (of both the participant and participants' mother) suggesting that rs16969968 may
have horizontal pleiotropic effects on age at menarche (e.g. via smoking outside of pregnancy).
Future MR studies could examine this (25).

268 Our observational results were consistent with previous observational studies (15-17) by showing a positive association of grandmother's (G0) smoking in pregnancy with grandchild's (G2) birthweight 269 270 after adjusting for mother's (G1) smoking in pregnancy. Although our G×E MR was vulnerable to 271 insufficient statistical power, we did find evidence that female G1 smoking in pregnancy modulates 272 the effect of G0 smoking heaviness in pregnancy on G2 birthweight, consistent with previous 273 observational findings (15-17). These results highlight the importance of both grandmother's and 274 maternal smoking in pregnancy for fetal growth, which could have implications for public health 275 interventions aiming to reduce the prevalence of low birthweight.

276 Strength of weakness of this study

277 We now discuss some limitations of this work. First, our proxy G×E MR used offspring genotype as a 278 proxy for maternal genotype and offspring rs16969968 contains 50% information from fathers. This 279 may cause regression dilution bias in each stratum, where the measurement error in the SNP biases 280 associations towards the null (48). However, we checked the extent that this might affect our 281 results, by comparing the associations of participant's rs16969968 with their own birthweight versus 282 their child's birthweight for smokers during pregnancy, and found little difference (-0.005kg (95% CI: 283 -0.020, 0.009)) between them. Second, we stratified on smoking status which rs16969968 was 284 weakly associated with. Stratification on colliders (between rs16969968 and outcomes) may bias our 285 MR estimates (see Supplementary Figure 2) (40, 41). Additionally, we used a highly selected sample 286 related to smoking (49) and had missing data in outcomes. These may also make our MR estimates 287 vulnerable to selection bias (50). However, previous evidence (29, 51) and our genetic associations 288 with measured confounders indicated that these selection effects may not be large enough to have a 289 considerable impact on our MR estimates. Third, rs16969968 predicts life-course smoking heaviness

and not just in pregnancy. Women who smoked in pregnancy may also smoke outside of pregnancy.
Therefore, the effect of maternal smoking might be via other pathways such as poor oocyte quality
for offspring birthweight, or postnatal maternal smoking (e.g. passive smoke exposure) for
adulthood outcomes among offspring (52).

294 Fourth, both participants' and their mothers' (G0) smoking status may be misclassified. Participants 295 were asked to report whether their mother smoked around the time of their birth and we used this 296 as our measure of G0 smoking in pregnancy. This means that G0 smokers might have smoked during 297 all their pregnancy, part of their pregnancy or started smoking shortly after giving birth. Effects of 298 smoking heaviness in pregnancy may vary according to the duration and pregnancy period during 299 which a woman smoked. For instance, previous work found that smoking in the first trimester was 300 not associated with lower birthweight in offspring suggesting that later stages may be more 301 important for fetal growth (3, 15). Similarly, participants reported their smoking status at baseline, 302 but this may not reflect their smoking status at an important time point for a given outcome. For, 303 instance, participants' height and age at menarche can only be affected by their own smoking 304 behaviour if they started smoking before achieving adult height or the onset of puberty. We 305 performed sensitivity analyses for height and age at menarche using estimates of participants 306 smoking status before these outcomes. For height, this assumed that men and women achieved 307 their adult height at 17 and 15 years old (38), respectively, as this information was not available in 308 UK Biobank. Fifth, we tested several hypotheses which increases the probability that our identified 309 associations may be due to chance. Finally, our study may lack statistical power due to small sample 310 sizes in strata and the low power of tests for interactions (53). We were unable to account for 311 grandchild's sex in our models assessing the impact of grandmother's smoking in pregnancy since 312 that is unavailable in UK Biobank, which may also reduce our statistical power. MR studies with 313 larger sample sizes and hence greater statistical power are needed to further investigate 314 transgenerational effects of smoking heaviness, together with studies in which both maternal and 315 offspring genotype are known.

316 Conclusion

317	G×E MR demonstrates how offspring genotype can be used to proxy for maternal genotype to
318	investigate causal effects of maternal smoking heaviness in pregnancy when maternal genotype is
319	unavailable. We demonstrated our proxy GxE approach by replicating the previously identified effect
320	of heavier smoking on lower offspring birthweight. We found little evidence of a causal effect of
321	maternal smoking heaviness on offspring's later life outcomes. Finally, we found evidence that the
322	effect of grandmother's smoking in pregnancy on grandchild's birthweight may be modulated by
323	mother's smoking status in pregnancy. Further studies with larger sample sizes are needed to
324	improve statistical power.

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328 Footnotes

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version of the manuscript, critically reviewed and revised the manuscript and approved the final

version of the manuscript as submitted. LACM contributed to the design of the study, critically

reviewed and revised the manuscript and approved the final version of the manuscript as submitted.

333 GDS conceptualized the study, contributed to the design of the study, critically reviewed and revised

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335 The corresponding author attests that all listed authors meet authorship criteria and that no others

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346 **Ethical approval:** The UK Biobank received ethical approval from the research ethics committee (REC

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- 348 **Data sharing:** The data reported in this paper are available by applying directly to the UK Biobank. All
- 349 code used to produce the results can be accessed at <u>https://github.com/MRCIEU/MR-maternal-</u>
- 350 <u>smoking</u>. Git tag v0.1 corresponds to the version presented here.
- 351 **Transparency:** The lead author (the manuscript's guarantor) affirms that this manuscript is an
- 352 honest, accurate, and transparent account of the study being reported; that no important aspects of
- 353 the study have been omitted; and that any discrepancies from the study as originally planned (and, if
- 354 relevant, registered) have been explained.
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359 References

- 1.Sharp GC, Lawlor DA, Richardson SS. It's the mother!: How assumptions about the causal primacy
 of maternal effects influence research on the developmental origins of health and disease. *Soc Sci*
- 362 Med 2018;213:20-7. doi: 10.1016/j.socscimed.2018.07.035
- 363 2.Davey Smith G. Negative control exposures in epidemiologic studies. *Epidemiology* 2012;23:350-1;
 364 author reply 1-2. doi: 10.1097/EDE.0b013e318245912c
- 365 3.Tyrrell J, Huikari V, Christie JT, *et al.* Genetic variation in the 15q25 nicotinic acetylcholine receptor 366 gene cluster (CHRNA5-CHRNA3-CHRNB4) interacts with maternal self-reported smoking status
- during pregnancy to influence birth weight. *Hum Mol Genet* 2012;21:5344-58. doi:
- 368 10.1093/hmg/dds372
- 4.Rice F, Harold GT, Boivin J, Hay DF, van den Bree M, Thapar A. Disentangling prenatal and inherited
 influences in humans with an experimental design. *Proc Natl Acad Sci U S A* 2009;106:2464-7. doi:
 10.1073/pnas.0808798106
- 372 5.Davey Smith G. Assessing intrauterine influences on offspring health outcomes: can
- epidemiological studies yield robust findings? *Basic Clin Pharmacol Toxicol* 2008;102:245-56. doi:
 10.1111/j.1742-7843.2007.00191.x
- 6.Krieger N, Davey Smith G. The tale wagged by the DAG: broadening the scope of causal inference
 and explanation for epidemiology. *Int J Epidemiol* 2016;45:1787-808. doi: 10.1093/ije/dyw114
- 7.Howe LD, Matijasevich A, Tilling K, *et al.* Maternal smoking during pregnancy and offspring
 trajectories of height and adiposity: comparing maternal and paternal associations. *Int J Epidemiol*2012;41:722-32. doi: 10.1093/ije/dys025
- 8.Riedel C, Schonberger K, Yang S, *et al.* Parental smoking and childhood obesity: higher effect
 estimates for maternal smoking in pregnancy compared with paternal smoking--a meta-analysis. *Int J Epidemiol* 2014;43:1593-606. doi: 10.1093/ije/dyu150
- 9.Albers L, Sobotzki C, Kuss O, et al. Maternal smoking during pregnancy and offspring overweight: is
 there a dose-response relationship? An individual patient data meta-analysis. Int J Obes (Lond)
 2018;42:1249-64. doi: 10.1038/s41366-018-0050-0
- 10.Brion MJ, Leary SD, Lawlor DA, Davey Smith G, Ness AR. Modifiable maternal exposures and
 offspring blood pressure: a review of epidemiological studies of maternal age, diet, and smoking.
 Pediatr Res 2008;63:593-8. doi: 10.1203/PDR.0b013e31816fdbd3
- 11.Chen Y, Liu Q, Li W, Deng X, Yang B, Huang X. Association of prenatal and childhood environment
 smoking exposure with puberty timing: a systematic review and meta-analysis. *Environ Health Prev Med* 2018;23:33. doi: 10.1186/s12199-018-0722-3
- 12.Jayes L, Haslam PL, Gratziou CG, *et al.* SmokeHaz: Systematic Reviews and Meta-analyses of the
 Effects of Smoking on Respiratory Health. *Chest* 2016;150:164-79. doi: 10.1016/j.chest.2016.03.060
- 13.Clifford A, Lang L, Chen R. Effects of maternal cigarette smoking during pregnancy on cognitive
- parameters of children and young adults: a literature review. *Neurotoxicol Teratol* 2012;34:560-70.
 doi: 10.1016/j.ntt.2012.09.004

- 14.Tiesler CM, Heinrich J. Prenatal nicotine exposure and child behavioural problems. *Eur Child Adolesc Psychiatry* 2014;23:913-29. doi: 10.1007/s00787-014-0615-y
- 15.Ding M, Yuan C, Gaskins AJ, et al. Smoking during pregnancy in relation to grandchild birth weight
 and BMI trajectories. *PLoS One* 2017;12:e0179368. doi: 10.1371/journal.pone.0179368
- 401 16.Miller LL, Pembrey M, Davey Smith G, Northstone K, Golding J. Is the growth of the fetus of a non402 smoking mother influenced by the smoking of either grandmother while pregnant? *PLoS One*403 2014;9:e86781. doi: 10.1371/journal.pone.0086781
- 404 17.Hypponen E, Davey Smith G, Power C. Effects of grandmothers' smoking in pregnancy on birth
 405 weight: intergenerational cohort study. *BMJ* 2003;327:898. doi: 10.1136/bmj.327.7420.898

18.Macdonald-Wallis C, Tobias JH, Davey Smith G, Lawlor DA. Parental smoking during pregnancy
and offspring bone mass at age 10 years: findings from a prospective birth cohort. *Osteoporos Int*2011;22:1809-19. doi: 10.1007/s00198-010-1415-y

- 409 19.Brion MJ, Leary SD, Davey Smith G, Ness AR. Similar associations of parental prenatal smoking
 410 suggest child blood pressure is not influenced by intrauterine effects. *Hypertension* 2007;49:1422-8.
 411 doi: 10.1161/hypertensionaha.106.085316
- 20.Leary SD, Brion MJ, Lawlor DA, Smith GD, Ness AR. Lack of emergence of associations between
 selected maternal exposures and offspring blood pressure at age 15 years. *J Epidemiol Community Health* 2013;67:320-6. doi: 10.1136/jech-2012-201784
- 21. Taylor AE, Carslake D, de Mola CL, *et al.* Maternal Smoking in Pregnancy and Offspring
 Depression: a cross cohort and negative control study. *Sci Rep* 2017;7:12579. doi: 10.1038/s41598017-11836-3
- 22.Davey Smith G, Lawlor DA, Harbord R, Timpson N, Day I, Ebrahim S. Clustered environments and
 randomized genes: a fundamental distinction between conventional and genetic epidemiology. *PLoS Med* 2007;4:e352. doi: 10.1371/journal.pmed.0040352
- 421 23.Davey Smith G, Ebrahim S. 'Mendelian randomization': can genetic epidemiology contribute to 422 understanding environmental determinants of disease? *Int J Epidemiol* 2003;32:1-22. doi:
- 423 24.Davey Smith G. Use of genetic markers and gene-diet interactions for interrogating population424 level causal influences of diet on health. *Genes Nutr* 2011;6:27-43. doi: 10.1007/s12263-010-0181-y
- 425 25.Spiller W, Slichter D, Bowden J, Davey Smith G. Detecting and correcting for bias in Mendelian
 426 randomization analyses using Gene-by-Environment interactions. *Int J Epidemiol* 2018. doi:
 427 10.1093/ije/dyy204
- 428 26.Munafo MR, Timofeeva MN, Morris RW, *et al.* Association between genetic variants on
 429 chromosome 15q25 locus and objective measures of tobacco exposure. *J Natl Cancer Inst*430 2012;104:740-8. doi: 10.1093/jnci/djs191

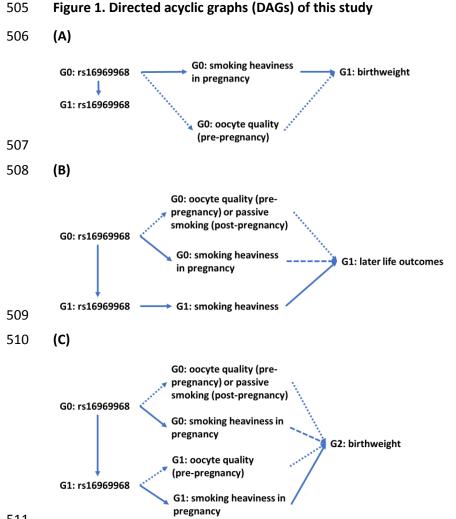
431 27.Freathy RM, Kazeem GR, Morris RW, *et al.* Genetic variation at CHRNA5-CHRNA3-CHRNB4
432 interacts with smoking status to influence body mass index. *Int J Epidemiol* 2011;40:1617-28. doi:
433 10.1002/lia/dt.m077

- 434 28.Linneberg A, Jacobsen RK, Skaaby T, et al. Effect of Smoking on Blood Pressure and Resting Heart
- 435 Rate: A Mendelian Randomization Meta-Analysis in the CARTA Consortium. *Circ Cardiovasc Genet*
- 436 2015;8:832-41. doi: 10.1161/circgenetics.115.001225
- 29.Millard LAC, Munafo M, Tilling K, Wootton RE, Davey Smith G. MR-pheWAS with stratification and
 interaction: Searching for the causal effects of smoking heaviness identified an effect on facial aging. *bioRxiv* 2018. doi: 10.1101/441907
- 30.Kvalvik LG, Skjaerven R, Klungsoyr K, Vollset SE, Haug K. Can 'early programming' be partly
 explained by smoking? Results from a prospective, population-based cohort study. *Paediatr Perinat*
- 442 *Epidemiol* 2015;29:50-9. doi: 10.1111/ppe.12164
- 31.Liu JZ, Erlich Y, Pickrell JK. Case-control association mapping by proxy using family history of
 disease. *Nat Genet* 2017;49:325-31. doi: 10.1038/ng.3766
- 32.Sudlow C, Gallacher J, Allen N, *et al.* UK biobank: an open access resource for identifying the
 causes of a wide range of complex diseases of middle and old age. *PLoS Med* 2015;12:e1001779. doi:
 10.1371/journal.pmed.1001779
- 33.Mitchell R, Hemani G, Dudding T, Paternoster L. UK Biobank Genetic Data: MRC-IEU Quality
 Control, Version 1. 06 Nov 2017. <u>https://data.bris.ac.uk/data/dataset/3074krb6t2frj29yh2b03x3wxj</u>.

34.Zhu Z, Lee PH, Chaffin MD, *et al.* A genome-wide cross-trait analysis from UK Biobank highlights
the shared genetic architecture of asthma and allergic diseases. *Nat Genet* 2018;50:857-64. doi:
10.1038/s41588-018-0121-0

- 35.Okbay A, Beauchamp JP, Fontana MA, *et al.* Genome-wide association study identifies 74 loci
 associated with educational attainment. *Nature* 2016;533:539-42. doi: 10.1038/nature17671
- 36.Hill WD, Marioni RE, Maghzian O, *et al.* A combined analysis of genetically correlated traits
 identifies 187 loci and a role for neurogenesis and myelination in intelligence. *Mol Psychiatry*2019;24:169-81. doi: 10.1038/s41380-017-0001-5
- 37.Howard DM, Adams MJ, Shirali M, *et al.* Genome-wide association study of depression
 phenotypes in UK Biobank identifies variants in excitatory synaptic pathways. *Nat Commun*2018;9:1470. doi: 10.1038/s41467-018-03819-3
- 38.Tanner JM, Whitehouse RH, Takaishi M. Standards from birth to maturity for height, weight,
 height velocity, and weight velocity: British children, 1965. I. Arch Dis Child 1966;41:454-71. doi:
- 39.Knezevic A. StatNews # 73: Overlapping Confidence Intervals and Statistical Significance. Oct
 2008. <u>https://www.cscu.cornell.edu/news/statnews/stnews73.pdf</u>.
- 465 40.Munafo MR, Tilling K, Taylor AE, Evans DM, Davey Smith G. Collider scope: when selection bias
 466 can substantially influence observed associations. *Int J Epidemiol* 2018;47:226-35. doi:
 467 10.1093/ije/dyx206
- 468 41.Taylor AE, Davies NM, Ware JJ, VanderWeele T, Smith GD, Munafo MR. Mendelian randomization
 469 in health research: using appropriate genetic variants and avoiding biased estimates. *Econ Hum Biol*470 2014;13:99-106. doi: 10.1016/j.ehb.2013.12.002

- 471 42.Colak Y, Afzal S, Lange P, Nordestgaard BG. Smoking, Systemic Inflammation, and Airflow
- Limitation: A Mendelian Randomization Analysis of 98 085 Individuals from the General Population.
- 473 *Nicotine Tob Res* 2018. doi: 10.1093/ntr/nty077
- 474 43.Skaaby T, Taylor AE, Jacobsen RK, *et al.* Investigating the causal effect of smoking on hay fever
 475 and asthma: a Mendelian randomization meta-analysis in the CARTA consortium. *Sci Rep*476 2017;7:2224. doi: 10.1038/s41598-017-01977-w
- 44.Magnus MC, Henderson J, Tilling K, Howe LD, Fraser A. Independent and combined associations
 of maternal and own smoking with adult lung function and COPD. *Int J Epidemiol* 2018;47:1855-64.
- 479 doi: 10.1093/ije/dyy221
- 480 45.Silvestri M, Franchi S, Pistorio A, Petecchia L, Rusconi F. Smoke exposure, wheezing, and asthma
 481 development: a systematic review and meta-analysis in unselected birth cohorts. *Pediatr Pulmonol*482 2015;50:353-62. doi: 10.1002/ppul.23037
- 483 46.Accordini S, Calciano L, Johannessen A, *et al.* A three-generation study on the association of
 484 tobacco smoking with asthma. *Int J Epidemiol* 2018;47:1106-17. doi: 10.1093/ije/dyy031
- 47.Menezes AM, Murray J, Laszlo M, *et al.* Happiness and depression in adolescence after maternal
 smoking during pregnancy: birth cohort study. *PLoS One* 2013;8:e80370. doi:
- 487 10.1371/journal.pone.0080370
- 488 48.Hutcheon JA, Chiolero A, Hanley JA. Random measurement error and regression dilution bias.
 489 *Bmj* 2010;340:c2289. doi: 10.1136/bmj.c2289
- 490 49.Swanson JM. The UK Biobank and selection bias. *Lancet* 2012;380:110. doi: 10.1016/s0140-491 6736(12)61179-9
- 492 50.Hughes RA, Davies NM, Davey Smith G, Tilling K. Selection bias in instrumental variable analyses.
 493 *BioRxiv* 2018. doi: 10.1101/192237
- 494 51.Gkatzionis A, Burgess S. Contextualizing selection bias in Mendelian randomization: how bad is it
 495 likely to be? *Int J Epidemiol* 2018. doi: 10.1093/ije/dyy202
- 496 52.Lawlor D, Richmond R, Warrington N, *et al.* Using Mendelian randomization to determine causal
 497 effects of maternal pregnancy (intrauterine) exposures on offspring outcomes: Sources of bias and
- 498 methods for assessing them. *Wellcome Open Res* 2017;2:11. doi:
- 499 10.12688/wellcomeopenres.10567.1
- 53.Marshall SW. Power for tests of interaction: effect of raising the Type I error rate. *Epidemiol Perspect Innov* 2007;4:4. doi: 10.1186/1742-5573-4-4
- 502 54.Pearl J, Glymour M, Jewell NP. Causal inference in statistics A Primer: John Wiley & Sons Ltd,503 2016.

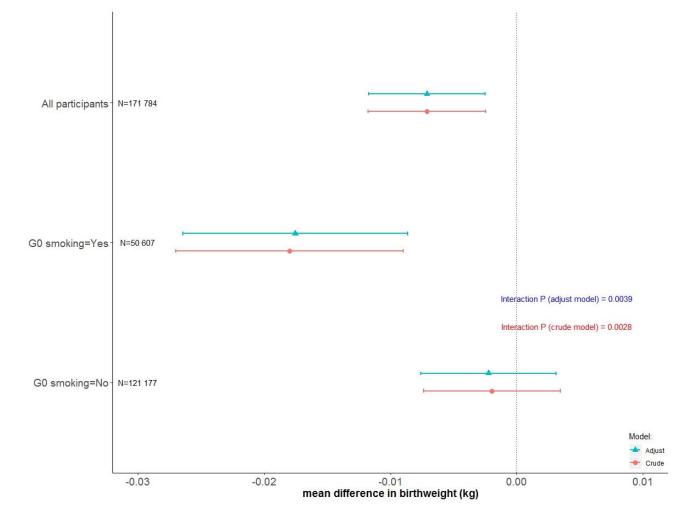




512 Generation (G)0: UK Biobank participants' mother; G1: UK Biobank participants themselves; G2: First 513 offspring of UK Biobank participants.

- A) Assessing the effect of G0 smoking heaviness on G1 birthweight: We used G1 rs16969968 as a
- 515 proxy for G0 rs16969968 and stratified on G0 smoking status in pregnancy. There is no backdoor
- 516 path (54) via G1 smoking heaviness since G1 cannot smoke before they were born. Maternal
- 517 smoking outside of pregnancy might influence the outcome (52), e.g. via oocyte quality, causing an
- 518 alternate path between rs16969968 and G1 birthweight (shown as>).
- B) Assessing the effect of G0 smoking on G1 later life outcomes: Besides the paths described in (A),
- 520 there is a backdoor path from G1 rs16969968 via G1 life-course smoking heaviness to the outcomes.
- 521 To estimate the effect of G0 smoking heaviness in pregnancy (shown as ---->), we need to block
- 522 this backdoor path by further stratifying on G1 smoking status.
- 523 C) Assessing the effect of G0 smoking on G2 birthweight: Besides the paths described in (A), there is
- a backdoor path from G1 rs16969968 via G1 smoking heaviness in pregnancy to the outcomes. To
- estimate the effect of G0 smoking heaviness in pregnancy (shown as ---->), we need to block this
- 526 backdoor path by further stratifying on G1 smoking status in pregnancy. G1 pre-pregnancy smoking
- 527 might influence G2 birthweight (shown as>).
- 528 See further DAGs in the Supplementary Figure 2 illustrating potential sources of bias due to 529 conditioning on a collider.

530 Figure 2. The associations of rs16969968 of UK Biobank participants with their own birthweight by their mothers' smoking status during pregnancy

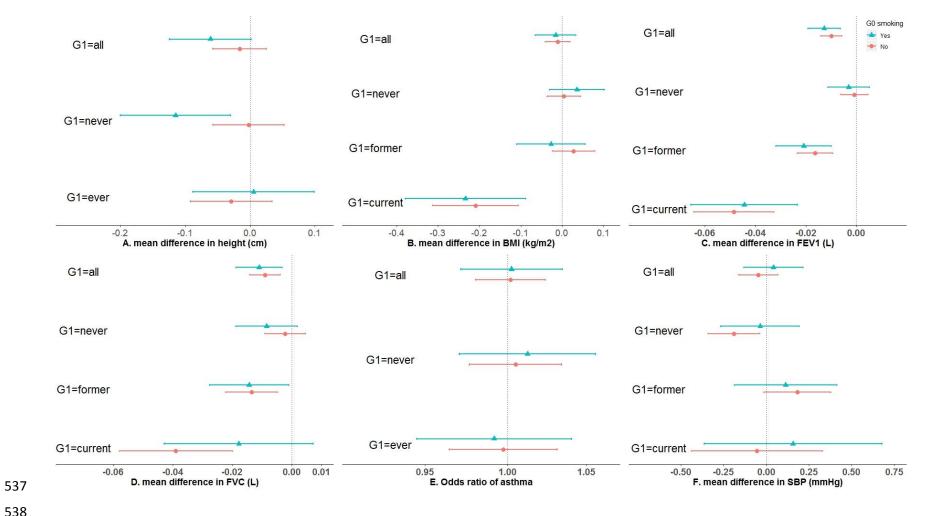


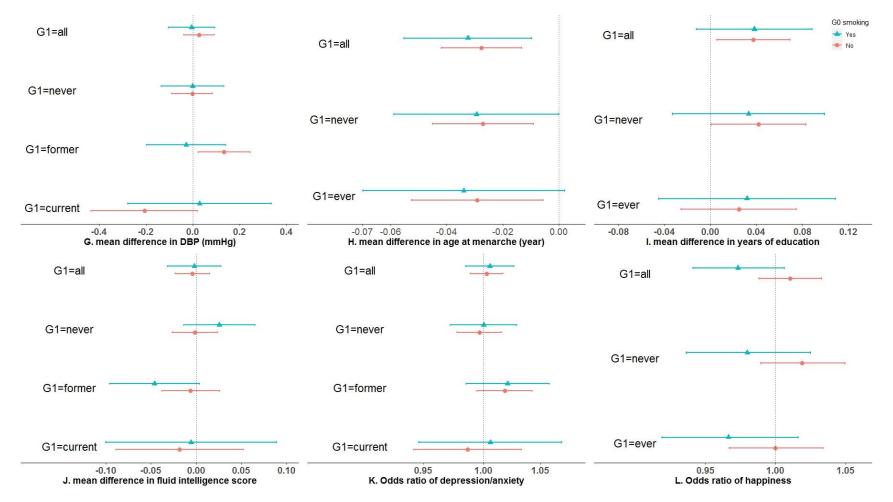
532 Generation (G)0: UK Biobank participants' mother; G1: UK Biobank participants themselves. Estimates are the mean difference of G1 birthweight per each

533 smoking-heaviness increasing allele of rs16969968. Associations are adjusted for sex of participants and the first ten principal components. The number of

⁵³⁴ participants was listed for each analysis.

Figure 3. The associations of rs16969968 with 12 outcomes in UK Biobank participants by their mothers' smoking status during pregnancy and their own smoking status



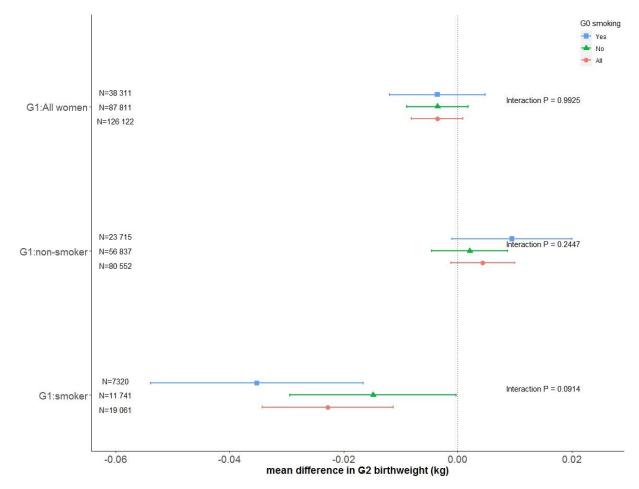


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Generation (G)0: UK Biobank participants' mother; G1: UK Biobank participants themselves. Estimates are the mean difference (or change in odds) of G1
 outcome per each smoking-heaviness increasing allele of rs16969968. We adjusted for age and sex of participants for outcomes except for menarche, and
 the first ten principal components for all 12 outcomes. We combined G1 current and former smokers into ever smokers for height, menarche, education,
 asthma and happiness to enlarge sample sizes given smoking cessation may not have a rapid impact on them.
 Abbreviations: BMI, body mass index, DBP, diastolic blood pressure; FEV₁, forced expiratory volume in 1-second; FVC, forced vital capacity; SBP, systolic

545 blood pressure.

546 Figure 4. The associations of rs16969968 of UK Biobank women participants with their first child's birthweight by their mothers' and their own smoking 547 status during pregnancy, after adjusting for the first ten genetic principal components



549 Generation (G)0: UK Biobank participants' mother; G1: UK Biobank participants themselves; G2: First offspring of UK Biobank participants. Estimates are the

550 mean difference of G2 birthweight per each smoking-heaviness increasing allele of rs16969968. Interactions are tested between G0 smokers (blue line) and 551 non-smokers (green line) with their P-values presented. All women in G1 included G1 smokers, G1 non-smokers and G1 women whose smoking status in

552 pregnancy was missing.