

1 Causal relevance of obesity on the leading causes of death in women
2 and men: A Mendelian randomization study

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4 Short title: Relevance of obesity to the leading causes of death in women and men

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49 **Abstract**

50 **Background**

51 Obesity traits are causally implicated with risk of cardiometabolic diseases. It remains unclear whether
52 there are similar causal effects of obesity traits on other non-communicable diseases. Also, it is largely
53 unexplored whether there are any sex-specific differences in the causal effects of obesity traits on
54 cardiometabolic diseases and other leading causes of death. We therefore tested associations of sex-
55 specific genetic risk scores (GRSs) for body mass index (BMI), waist-hip-ratio (WHR), and WHR
56 adjusted for BMI (WHRadjBMI) with leading causes of mortality, using a Mendelian randomization
57 (MR) framework.

58 **Methods and Findings**

59 We constructed sex-specific GRSs for BMI, WHR, and WHRadjBMI, including 565, 324, and 338
60 genetic variants, respectively. These GRSs were then used as instrumental variables to assess
61 associations between the obesity traits and leading causes of mortality using an MR design in up to
62 422,414 participants from the UK Biobank. We also investigated associations with potential mediators
63 and risk factors, including smoking, glycemic and blood pressure traits. Sex-differences were
64 subsequently assessed by Cochran's Q-test (P_{het}).

65 Up to 227,717 women and 194,697 men with mean (standard deviation) age 56.6 (7.9) and 57.0 (8.1)
66 years, body mass index 27.0 (5.1) and 27.9 (4.2) kg/m² and waist-hip-ratio 0.82 (0.07) and 0.94 (0.07),
67 respectively, were included. Mendelian randomization analysis showed that obesity causes coronary
68 artery disease, stroke (particularly ischemic), chronic obstructive pulmonary disease, lung cancer, type
69 2 and 1 diabetes mellitus, non-alcoholic fatty liver disease, chronic liver disease, and acute and chronic
70 renal failure. A 1 standard deviation higher body mass index led to higher risk of type 2 diabetes in
71 women (OR 3.81; 95% CI 3.42-4.25, $P=8.9\times 10^{-130}$) than in men (OR 2.78; 95% CI 2.57-3.02,
72 $P=1.0\times 10^{-133}$, $P_{het}=5.1\times 10^{-6}$). Waist-hip-ratio led to a higher risk of chronic obstructive pulmonary
73 disease ($P_{het}=5.5\times 10^{-6}$) and higher risk of chronic renal failure ($P_{het}=1.3\times 10^{-4}$) in men than women.

74 A limitation of MR studies is potential bias if the genetic variants are directly associated with
75 confounders (pleiotropy), but sensitivity analyses such as MR-Egger supported the main findings. Our
76 study was also limited to people of European descent and results may differ in people of other
77 ancestries.

78 **Conclusions**

79 Obesity traits have an etiological role in the majority of the leading global causes of death. Sex
80 differences exist in the effects of obesity traits on risk of type 2 diabetes, chronic obstructive
81 pulmonary disease, and renal failure, which may have implications on public health.

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95 **Abbreviations**

96 BMI, Body mass index; CAD: coronary artery disease; CLD, chronic liver disease; COPD, chronic
97 obstructive pulmonary disease; DBP, diastolic blood pressure; FG, fasting glucose; FI, fasting insulin;
98 GIANT, Genetic Investigation of ANthropometric Traits; GRS, genetic risk score; GWAS, genome-
99 wide association study; MAGIC, the Meta-Analyses of Glucose and Insulin-related traits Consortium;
100 MR, Mendelian randomization; NAFLD, non-alcoholic fatty liver disease; OR, odds ratio; T1D, type
101 1 diabetes; T2D, type 2 diabetes; SBP, systolic blood pressure; SD, standard deviation; SNP, single
102 nucleotide polymorphism; WHO, the World Health Organization; WHR, waist-hip-ratio;
103 WHRadjBMI, waist-hip-ratio adjusted for body mass index.

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122 **Introduction**

123 It is increasingly evident that obesity negatively impacts human health and the prevalence of obesity is
124 increasing world-wide (1). Obesity and central fat distribution, commonly measured by body mass
125 index (BMI; obesity usually defined as BMI >30 kg/m²) and waist-hip-ratio (WHR), respectively,
126 have been linked to cardiometabolic diseases and death in observational studies (2–5). However,
127 conventional observational studies can be affected by bias, confounding, and reverse causation, which
128 might lead to erroneous findings. Mendelian randomization (MR) offers an approach to circumvent
129 these issues by using single nucleotide polymorphisms (SNPs) that reliably associate with an exposure
130 as an instrument to test the causal relationship between an exposure and outcome (6). Owing to the
131 nature of genotypes and therefore genetic associations, MR estimates should be less affected by
132 confounding and reverse causation (S1 Supporting Information) (6). Previous studies have found
133 causal relationships between for example higher BMI and WHR adjusted for BMI (WHRadjBMI) and
134 type 2 diabetes (T2D) and coronary artery disease (CAD), mostly using a limited number of
135 previously known obesity-associated SNPs (7–12). However, previous studies have not thoroughly
136 investigated causal sex-specific relationships, nor have they comprehensively investigated the role that
137 obesity traits play in the leading causes of death beyond these cardiometabolic diseases.

138 Obesity traits differ between women and men—for example, regional obesity prevalence rates often
139 vary between the sexes (13,14), women have higher SNP-based heritability for WHR (15), and >90%
140 of WHRadjBMI-associated SNPs that show evidence of sexual dimorphism have larger effect sizes in
141 women than men (15). Observational studies have indicated that waist-related traits might be more
142 strongly associated with cardiometabolic outcomes in women, although previous studies are
143 inconclusive (16–20). Only a few studies have investigated sex differences in the effect of genetic risk
144 for obesity-related traits on disease risk (7,10,12). These studies have mostly been restricted to waist-
145 related traits and T2D and CAD, using a limited number of analyses and/or SNPs, but without finding
146 evidence of differences in disease risk between men and women (7,10,12).

147 A sex difference in the effect of obesity traits on major causes of death could signify that disease
148 burden arising from obesity may be differential in women and men, allowing prioritizing of public

149 health resources and potentially, sex-specific preventative strategies. We therefore investigated the
150 extent to which obesity traits causally impact the risk of the major global causes of death, and whether
151 relationships with disease are differential between women and men, exploiting recent advances in
152 discovery of obesity-associated SNPs (15).

153 **Methods**

154 **Data sources and study participants**

155 The UK Biobank is a prospective UK-based cohort study, with 488,377 genotyped individuals aged
156 40-69 when recruited (21). UK Biobank has a Research Tissue Bank approval (Research Ethics
157 Committee reference 16/NW/0274, this study's application ID 11867), and all participants gave
158 informed consent.

159 In the present study, genotype data for up to 422,414 individuals were included, after general genotype
160 and sample quality control procedures, and exclusion of people of non-European ancestry (S1
161 Supporting Information). Participant characteristics are in Table 1.

162 **Table 1. Characteristics of UK Biobank Participants included in the study.**

Characteristic	Men	Women
Individuals, N (%)	194,697 (46.1)	227,717 (53.9)
British, N (%) ^a	173,947 (89.3)	201,278 (88.4)
Age, mean (SD), years	57.0 (8.1)	56.6 (7.9)
UK BiLEVE array, N (%) ^b	23,187 (11.9)	22,755 (10.0)
Body mass index, mean (SD), kg/m ²	27.9 (4.2)	27.0 (5.1)
Waist circumference, mean (SD), cm	97.1 (11.3)	84.5 (12.5)
Hip circumference, mean (SD), cm	103.5 (7.6)	103.3 (10.3)
Waist-hip-ratio, mean (SD)	0.94 (0.07)	0.82 (0.07)
Systolic blood pressure, mean (SD), mmHg	144.8 (19.4)	138.0 (21.2)
Diastolic blood pressure, mean (SD), mmHg	86.6 (11.0)	82.3 (11.1)
Type 2 diabetes cases, N (%)	11,768 (6.0)	6,533 (2.9)
Coronary artery disease cases, N (%)	24,430 (12.5)	11,565 (5.1)
Breast cancer cases, N (%)	-	14,294 (6.3)
Chronic liver disease cases, N (%)	822 (0.4)	542 (0.2)
Colorectal cancer cases, N (%)	3,145 (1.6)	2,368 (1.0)
COPD cases, N (%)	7,890 (4.1)	6,789 (3.0)
Dementia cases, N (%)	580 (0.3)	448 (0.2)
Infertility cases, N (%)	85 (0.0)	1,588 (0.7)
Lung cancer cases, N (%)	1,473 (0.8)	1,244 (0.5)

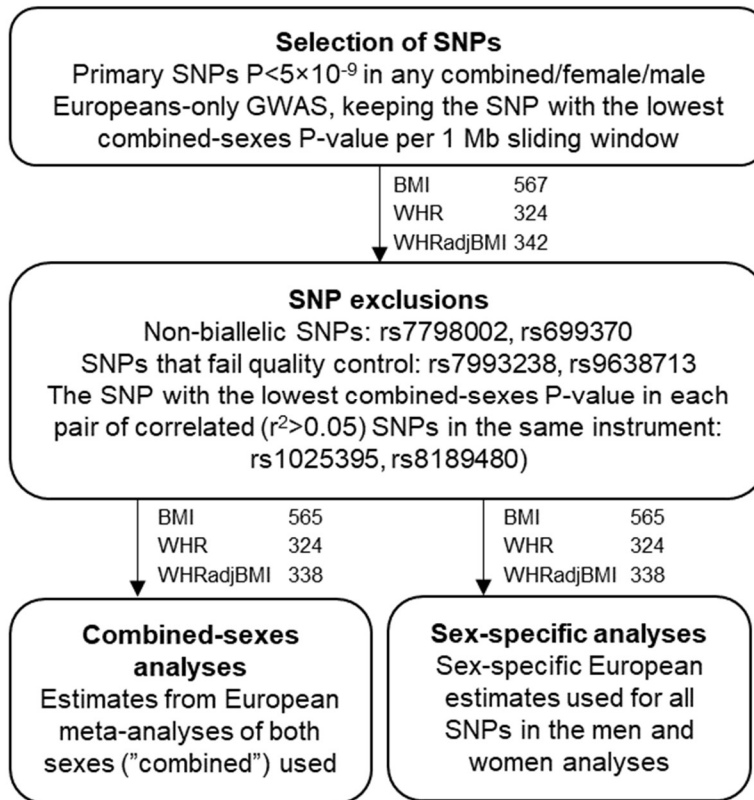
Characteristic	Men	Women
NAFLD cases, N (%)	912 (0.5)	778 (0.3)
Renal failure cases, N (%)	5,704 (2.9)	3,902 (1.7)
Renal failure, acute, cases, N (%)	3,045 (1.6)	1,643 (0.7)
Renal failure, chronic, cases, N (%)	2,581 (1.3)	2,019 (0.9)
Stroke cases, N (%)	6,329 (3.3)	4,437 (1.9)
Stroke, hemorrhagic, cases, N (%)	929 (0.5)	972 (0.4)
Stroke, ischemic, cases, N (%)	2,167 (1.1)	1,177 (0.5)
Type 1 diabetes cases, N (%)	824 (0.4)	675 (0.3)

163 COPD, chronic obstructive pulmonary disease; NAFLD, non-alcoholic fatty liver disease; SD, standard deviation.
164 ^aParticipants were denoted as “British” if they were in the British ancestry subset as defined by the UK Biobank (21) (based
165 on self-report of British ancestry and similar ancestry according to principal components analysis)
166 ^bUK BiLEVE array is the number of participants genotyped on that array as opposed to the UK Biobank Axiom array
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168 Instruments

169 We evaluated several approaches to construct sex-specific genetic risk scores (GRSs) for BMI, WHR,
170 and WHRadjBMI (S1 Supporting Information, Fig A-B in S1 Supporting Information). The approach
171 with the highest ranges of trait variance explained and F-statistics for the relevant obesity trait, and
172 with no demonstrable heterogeneity between men and women, was selected as the main model. In this
173 model, GRSs were constructed by including the primary (“index”) genome-wide significant ($P < 5 \times 10^{-9}$)
174 SNPs in the men, women, or combined-sexes analyses in the largest genome-wide association study
175 (GWAS) available with sex-specific European summary statistics, a meta-analysis of the Genetic
176 Investigation of ANthropometric Traits (GIANT) (22,23) and the UK Biobank (Fig 1, S1 Supporting
177 Information) (15). Primary SNPs were identified in the original GWAS (15) by proximal and joint
178 conditional analysis using GCTA in associated loci. Associated loci included all SNPs (associated
179 with the GWAS obesity trait $P < 0.05$) ± 5 Mb around a top SNP ($P < 5 \times 10^{-9}$) and that were in linkage
180 disequilibrium (LD; $r^2 > 0.05$) with the top SNP; overlapping loci were merged (15). We then kept the
181 SNP with the lowest combined-sexes P-value within each 1 Mb sliding window to limit correlation
182 between SNPs discovered in different sex-strata in each obesity trait. We excluded non-biallelic SNPs
183 (N=2), SNPs that failed quality control (N=2), and one SNP per pair with long-distance linkage
184 disequilibrium ($r^2 > 0.05$, N=2) (S1 Supporting Information). For the combined-sexes analyses, SNPs
185 were weighted using estimates from the combined-sexes European meta-analyzed GWASs. For the
186 men- and women-only analyses, SNPs were weighted by their sex-specific European estimates. All

187 SNPs were orientated so that the effect allele corresponded to a higher level of the investigated obesity
188 trait.



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190 **Fig 1. SNP- and weight selection flowchart with number of SNPs for each obesity trait.**

191 SNPs were selected by including the primary ("index") variants for each associated (with SNPs $P < 5 \times 10^{-9}$) locus
192 (assessed for a minimum of ± 5 Mb around the top SNP and including all SNPs in linkage disequilibrium
193 $R^2 > 0.05$ and $P < 0.05$, and with primary variants as determined through joint and conditional testing using GCTA
194 in the original study (15)), in any of the men, women, and combined-sexes genome-wide association studies for
195 each obesity trait (15). To further ascertain independence for SNPs selected from different sex-stratified genome-
196 wide association studies and to have the same set of SNPs for all sex-strata, only the SNP with the lowest
197 combined-sexes P-value within each 1 Mb-window was kept. SNPs that were non-biallelic ($N=2$) or that failed
198 quality control ($N=2$) were removed, as was one SNP in each pair with long-distance linkage disequilibrium
199 ($N=2$, using a linkage disequilibrium threshold of $r^2 < 0.05$ and removing the SNP with the highest combined-
200 sexes P-value). All SNPs were then weighted by their sex-specific European estimates for the men- and women
201 analyses, and by the combined-sexes European estimates for the combined-sexes analyses, using estimates from
202 the original genome-wide association study (15). BMI, body mass index; SNP, single nucleotide polymorphism;
203 WHR, waist-hip-ratio; WHRadjBMI, waist-hip-ratio adjusted for body mass index.

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205 **Outcomes**

206 We investigated associations between three obesity traits (BMI, WHR, and WHRadjBMI) with all

207 non-communicable diseases on the World Health Organization's (WHO) list of leading mortality

208 causes world-wide and in high-income countries (24); CAD, stroke (including ischemic, hemorrhagic,
209 and of any cause), chronic obstructive pulmonary disease (COPD), dementia, lung cancer, T2D and
210 type 1 diabetes (T1D), colorectal cancer, renal failure (including acute, chronic and of any cause) and
211 breast cancer in women (Table A in S1 Supporting Information). In addition, we included infertility,
212 non-alcoholic fatty liver disease (NAFLD) and chronic liver disease (CLD) as they have previously
213 been linked to obesity and represent important and increasing burdens of disease (25–31). For T2D
214 and T1D, we drew case definitions from a validated algorithm for prevalent T2D and T1D (using
215 “probable” and “possible” cases) and those the algorithm denoted as “diabetes unlikely” were used as
216 controls (32). For CAD, we used the same case and control definitions as a large GWAS (33). Case
217 and control criteria for the other disease outcomes were defined using self-report data, data from an
218 interview with a trained nurse, and hospital health outcome codes in discussion between two licensed
219 medical practitioners (Table A in S1 Supporting Information). For CAD, acute renal failure, chronic
220 renal failure, stroke of any cause, ischemic stroke and hemorrhagic stroke, exclusions for certain codes
221 were also made in the control groups after defining the case groups.

222 To assess potential mediation, we also investigated associations between the obesity traits and the
223 potential cardiometabolic risk factors systolic blood pressure (SBP), diastolic blood pressure (DBP),
224 fasting glucose (FG), fasting insulin (FI), and smoking status.

225 Baseline measurements were used for all continuous traits, including BMI, WHR, WHRadjBMI, SBP
226 and DBP. For SBP and DBP, the mean of the up to two baseline measurements were used. Fifteen
227 mmHg to SBP and 10 mmHg to DBP were added if blood pressure lowering medications were used
228 (defined as self-reported use of such in data-fields 6153 and 6177), as in previous blood pressure
229 GWASs and as suggested in simulation studies (34,35). These anthropometric and blood pressure
230 measurements were then standardized by rank inverse normal transformation of the residuals after
231 regression of the trait on baseline age, age², assessment centre, and, if applicable, sex. This was done
232 separately in the men and women only analyses, but jointly in the combined analyses, after any sample
233 quality exclusions (S1 Supporting Information). WHRadjBMI was generated in a similar manner, but
234 with adjustment for BMI as well, as in the original GWAS (15).

235 Sex-specific summary-level data for plasma FG (in mmol/L, untransformed, corrected to plasma levels
236 using a correction factor of 1.13 if measured in whole blood in the original GWAS) and serum FI (in
237 pmol/L, ln-transformed) were kindly provided by the Meta-Analyses of Glucose and Insulin-related
238 traits Consortium (MAGIC) investigators and can be downloaded from
239 <https://www.magicinvestigators.org/downloads/> (36). SNPs in chromosome:position format were
240 converted to rsIDs using the file All_20150605.vcf.gz from the National Center for Biotechnology
241 Information (NCBI) (37) (available at
242 ftp://ftp.ncbi.nih.gov/snp/organisms/archive/human_9606_b144_GRCh37p13/VCF/). All SNPs were
243 then updated to dbSNP build 151 using the file RsMergeArch.bcp.gz, also from the NCBI (37)
244 (available at ftp://ftp.ncbi.nlm.nih.gov/snp/organisms/human_9606/database/organism_data/).

245 Smoking status was defined as self-report of being a current or previous smoker or having smoked or
246 currently smoking (most days or occasionally; any code 1 or 2 in any of the data fields 1239, 1249,
247 and 20116).

248 **Statistical analyses**

249 The GRSs were first assessed if they were robustly associated with their respective obesity traits by
250 computing trait variance explained and the F-statistics (S1 Supporting Information, Table B in S1
251 Supporting Information).

252 We explored the associations of sex-specific GRSs with outcomes in the UK Biobank (21). For
253 disease outcomes and smoking status, logistic regression was used while for continuous traits
254 (including evaluation of the GRSs in their respective obesity traits and the blood pressure traits) linear
255 regression was used. Associations of sex-specific GRSs with outcome traits that surpassed our P-value
256 thresholds were taken forward for MR to more formally quantify the effect of the obesity trait on the
257 outcome.

258 Individual-level MR was performed using the Wald method, with the instrumental variable estimate
259 being the ratio between the outcome and risk factor regressed separately on each GRS (38). Standard
260 errors were adjusted to take the uncertainty in both regressions into account by using the first two

261 terms of the delta method (39–41). MR regressions of the risk factors on the GRSs was performed in
262 controls only for the binary outcomes.

263 Adjustments were made for baseline age, age², array type, assessment centre, 10 principal
264 components, and sex if applicable, for all traits when in clinical units, and array and 10 principal
265 components if rank inverse normal transformed (where adjustment for age, age², assessment centre,
266 and if applicable sex had already been performed in the rank inverse normal transformation of the
267 residuals).

268 Two-sample MR was performed for the effect of the obesity traits on FG and FI, including the inverse-
269 variance weighted (IVW), MR-Egger, and weighted median methods (42–45).

270 For the obesity trait-risk factor analyses, the P-value threshold was set at <0.003 (=0.05/15) for the
271 regressions and the MRs, for the total of 15 obesity trait-risk factor combinations investigated in the
272 study. We conducted MRs for all obesity traits with smoking status for completeness, since we
273 performed analyses adjusting for smoking status as a sensitivity analysis. We also performed
274 summary-level MRs for the potential risk factors FG and FI directly, as we only had summary-level
275 data for these traits. For the obesity trait-disease analyses, the P-value thresholds for both the
276 regressions and the MRs were set at <0.001 (=0.05/51) for 51 obesity trait-disease combinations
277 investigated in the study. If a combined-sexes regression analysis identified evidence against the null
278 hypothesis it was taken forward for MR; if a regression analysis identified evidence against the null
279 hypothesis in either men or women, MR was performed in both sexes so sexual heterogeneity could be
280 assessed. Sexual heterogeneity between male and female estimates from the linear and logistic
281 regressions and the MRs was assessed using P-values from Cochran's Q test (46). To facilitate
282 comparisons between the obesity traits and sex-strata, estimates were computed per 1 standard
283 deviation (SD) higher obesity trait.

284 **Sensitivity analyses**

285 We performed several sensitivity analyses to ascertain robustness; we performed (a) analyses adjusting
286 for smoking status and (b) analyses restricted to those of genetically confirmed British ancestry only

287 (S1 Supporting Information). We also (c) evaluated the robustness of the MR findings by comparing
288 different weighting strategies, including use of unweighted and externally weighted (using weights
289 from the GIANT 2015 studies (22,23)) GRSs, and (d) investigated for pleiotropy and performed more
290 pleiotropy-robust sensitivity analyses (44,45) (S1 Supporting Information). We also (e) performed
291 logistic regressions using the same number of cases and controls in men and women for the disease
292 outcomes and (f) conducted analyses using stricter T2D and T1D case definitions (S1 Supporting
293 Information).

294 **Software**

295 The genotype data was handled PLINK v2.00aLM and PLINK v1.90b3 (47) (S1 Supporting
296 Information). Further data handling was performed in Python 3.5.2 (48) using the packages “pandas”
297 (49) and “numpy” (50), R version 3.4.3 (51) and the package “dplyr” (52), bash version 4.1.2(2) (53)
298 and awk (54). Statistical analyses and plots were performed using R version 3.4.3 (51) and packages
299 “ggplot2” (55), “mada” (56), “dplyr” (52), “gridExtra” (57), “lattice” (58), “grid” (51), “grDevices”
300 (51), “ggpubr” (59), and “MendelianRandomization” (42).

301 **Results**

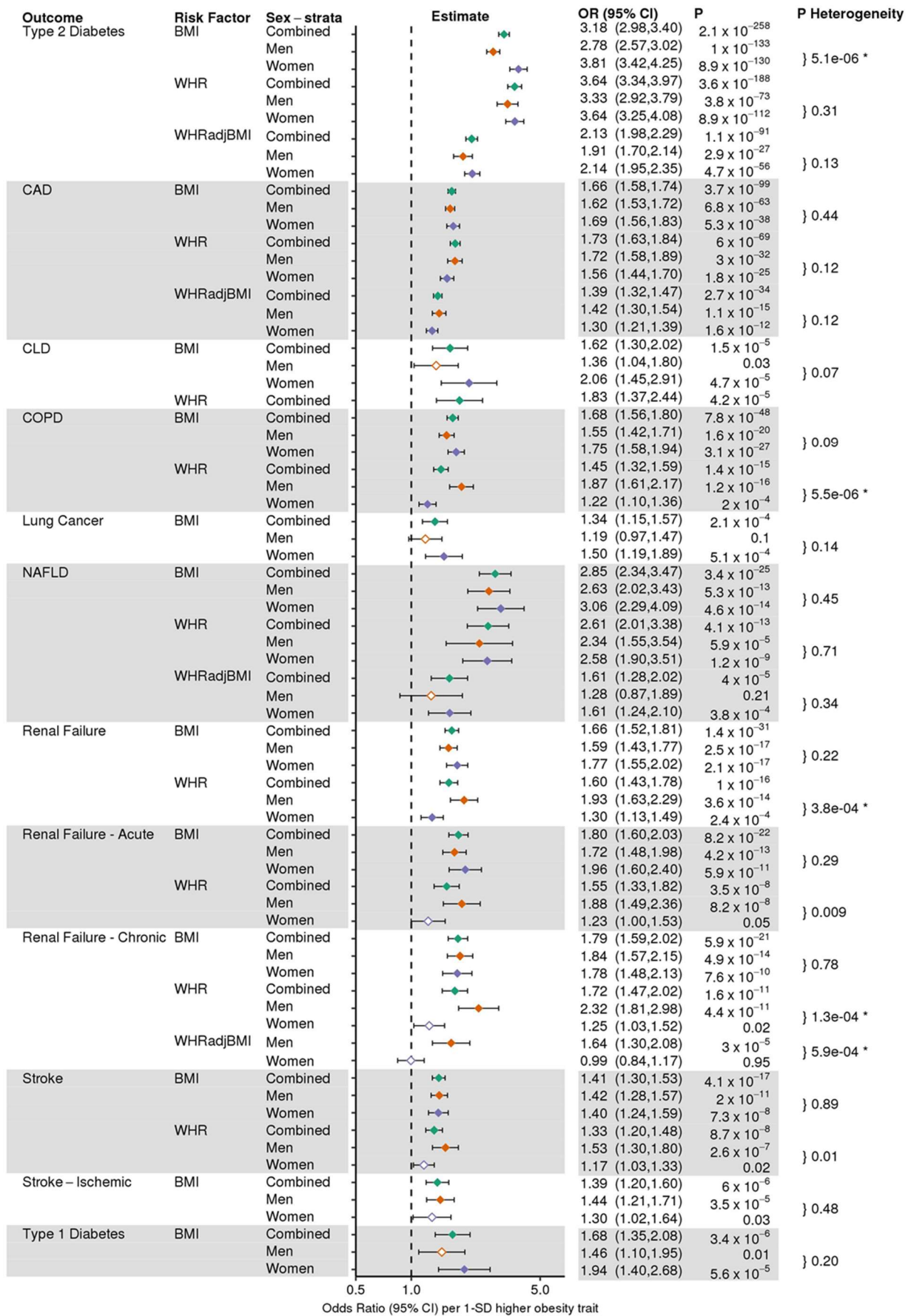
302 **Evaluation of genetic risk scores**

303 The GRSs included 565 SNPs for BMI, 324 for WHR and 338 for WHRadjBMI. Trait variance
304 explained varied between 2.5-7.1% and the F-statistic between 4,941-26,311, depending on trait and
305 sex-stratum (Table B in S1 Supporting Information). After having assessed the associations between
306 GRSs and risk factors and disease outcomes using regression analyses, associations that surpassed
307 correction for multiple testing were taken forward for MR (Table C-E and Fig C in S1 Supporting
308 Information).

309 Several instruments were positively associated with smoking status and with higher estimates in men
310 than in women for both BMI as well as WHR (BMI: $P_{\text{het}}=4.7\times 10^{-4}$; WHR: $P_{\text{het}}=1.3\times 10^{-13}$;
311 WHRadjBMI $P_{\text{het}}=0.007$) (Table D in S1 Supporting Information). We therefore ran the individual-
312 level MRs adjusting for smoking status to assess potential mediation.

313 **Mendelian randomization of obesity with disease outcomes: all individuals**

314 Obesity traits were causally implicated with diseases that represent the major causes of death (Fig 2
315 and 3). All measures of obesity were strongly causally related to risk of CAD (odds ratio (OR) ranging
316 from 1.39 for WHRadjBMI to 1.73 for WHR in the combined analyses per 1-SD higher obesity trait).
317 For stroke, both BMI and WHR conferred higher risk (ORs 1.41 and 1.33, respectively). Strong effects
318 were seen for all obesity traits with T2D (OR range 2.13 to 3.64) and BMI also associated with risk of
319 T1D (OR 1.68). Obesity traits increased the risk of kidney disease, including both acute (ORs 1.55 for
320 WHR and 1.80 for BMI) and chronic (ORs 1.72 for WHR and 1.79 for BMI) renal failure. Measures
321 of obesity also causally impacted on risks of COPD (OR 1.68 for BMI and 1.45 for WHR) and lung
322 cancer (BMI OR 1.34). Adjusting for smoking status resulted in reduced magnitudes of effects for
323 COPD and lung cancer traits, suggesting potential mediation (Table F in S1 Supporting Information).
324 In addition to these endpoints, strong effects were seen for risk of NAFLD (OR range 1.61-2.85) and
325 CLD (ORs 1.62 for BMI and 1.83 for WHR).



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Fig 2. Causal effects of obesity traits on disease outcomes, overall and stratified by sex.

The obesity trait-disease combinations brought forward for Mendelian randomization, with estimates given in odds ratio (95% CI) per 1-SD higher obesity trait. Filled diamonds indicate that the P-value for the obesity trait

330 to disease endpoint surpasses our threshold for multiple testing; empty diamonds indicate that the P-value does
331 not surpass this threshold (Bonferroni-adjusted P-value-threshold set at <0.001 ($=0.05/51$) for 51 obesity trait-
332 disease outcome combinations in the study). * denotes that the P-value for heterogeneity (from Cochran's Q test)
333 surpasses our threshold for multiple testing; P_{het} -threshold set at <0.001 ($=0.05/48$) for 48 male-female
334 comparisons in the study (fewer since breast cancer analyses were performed in women only). \blacklozenge , combined-
335 sexes estimates; \blacklozenge , male estimates; \blacklozenge , female estimates; BMI, body mass index; CAD, coronary artery disease;
336 COPD, chronic obstructive pulmonary disease; NAFLD, non-alcoholic fatty liver disease; SD standard
337 deviation; WHR, waist-hip-ratio; WHRadjBMI, waist-hip-ratio adjusted for body mass index.
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388 **Fig 3. Overview of the sex-specific effect magnitudes and strengths of association of obesity traits on leading causes of death.**
 389 Leading causes of death defined as non-communicable diseases on the WHO top 10 lists of causes of death, globally and in high-income countries, with additional separate

390 analyses for subclasses of stroke, diabetes, and renal disease. No obesity trait (BMI, WHR, or WHRadjBMI)
391 genetic risk score associated with dementia, colorectal cancer, breast cancer (investigated in women only) or
392 hemorrhagic stroke – these are not shown on the plot. (A) Total number of deaths globally, in 1,000 deaths, as
393 estimated by the WHO for 2016 (60), stratified by sex. For diabetes, estimates for annual number of deaths are
394 for type 1 and type 2 diabetes combined. (B) Obesity trait-disease combinations taken forward for Mendelian
395 randomization showed with circles. Mendelian randomization associations with P-values surpassing our
396 threshold in yellow to red fill depending on P-value ($-\log_{10}$ P-value), white fill indicates a P-value not
397 surpassing our threshold. The size of the circles corresponds to the magnitude of the odds ratio estimate for the
398 Mendelian randomization estimate. Estimates and P-values from the MR analyses of the obesity traits with the
399 disease outcomes using the sex-specific estimates approach. BMI, body mass index; P, P-value; WHR, waist-
400 hip-ratio; WHRadjBMI, waist-hip-ratio adjusted for body mass index; WHO, World Health Organization.

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402 Sensitivity analyses, including restricting analyses to those of genetically confirmed British ancestry
403 only, use of different weighting strategies, analyses using more pleiotropy-robust methods, using the
404 same number of cases and controls in men and women, and use of more stringent diabetes case
405 definitions supported the main findings (S1 Supporting Information, Tables G,H and Fig D-F in S1
406 Supporting Information).

407 **Mendelian randomization of obesity with disease outcomes: sex-stratified analyses**

408 Five obesity trait-disease associations differed between women and men (Fig 2). The risk of T2D from
409 1-SD higher BMI was higher in women (OR 3.81; 95% CI 3.42-4.25, $P=8.9 \times 10^{-130}$) than men (OR
410 2.78; 95% CI 2.57-3.02, $P=1.0 \times 10^{-133}$), with strong evidence for sexual heterogeneity ($P_{\text{het}}=5.1 \times 10^{-6}$,
411 P_{het} -threshold set at <0.001 ($=0.05/48$) for 48 male-female disease estimate comparisons, since breast
412 cancer was investigated in women only). This sexual heterogeneity could also be observed in
413 sensitivity analyses where the number of cases in women and men was similar ($P_{\text{het}}=4.4 \times 10^{-5}$) (Table
414 H in S1 Supporting Information).

415 WHR increased risk of COPD to a greater extent in men (OR 1.87; 95% CI 1.61-2.17, $P=1.2 \times 10^{-16}$)
416 than in women (OR 1.22; 95% CI 1.10-1.36, $P=2.0 \times 10^{-4}$, $P_{\text{het}}=5.5 \times 10^{-6}$), per 1-SD higher WHR. While
417 the association of WHR with smoking was greater in men than in women (Table I in S1 Supporting
418 Information) and estimates of WHR with COPD for both men and women attenuated after adjustment
419 for smoking status, the association of WHR and COPD remained higher in men after adjusting for
420 smoking ($P_{\text{het}}=1.2 \times 10^{-4}$; Table F in S1 Supporting Information).

421 There was also evidence of WHR leading to a higher risk on renal failure in men than in women. Men
422 had a higher risk of chronic renal failure per 1-SD higher WHR, with the risk in men being OR 2.32

423 (95% CI 1.81-2.98, $P=4.4\times 10^{-11}$) and in women OR 1.25 (95% CI 1.03-1.52, $P=0.02$, $P_{\text{het}}=1.3\times 10^{-4}$),
424 with similar sex differences seen for WHRadjBMI. Men also had a higher risk of acute renal failure
425 (men: OR 1.88; 95% CI 1.49-2.36, $P=8.2\times 10^{-8}$; women: OR 1.23; 95% CI 1.00-1.53, $P=0.05$, per 1-SD
426 higher WHR, $P_{\text{het}}=0.009$), although the P_{het} -value did not pass our P_{het} -threshold.

427 Sensitivity analyses using different GRS weighting strategies strongly supported sex-differences in the
428 effect of BMI on T2D and WHR on chronic renal failure and COPD, but only weakly supported a sex-
429 difference in the effect of WHR on renal failure of any cause (S1 Supporting Information, Fig D,E in
430 S1 Supporting Information).

431 **Potential mechanisms**

432 To identify potential mediators, we assessed the relationship of obesity traits with blood pressure
433 (SBP, DBP), glycemic traits (FG, FI), and smoking status (Tables I-M in S1 Supporting Information).
434 All obesity traits causally impacted risk on SBP, DBP, FG and FI. The increase in DBP arising from
435 elevated BMI was greater in women than men ($P_{\text{het}}=3.5\times 10^{-5}$, P_{het} -threshold set at <0.003 ($=0.05/15$)
436 for 15 obesity trait-risk factor combinations). BMI and WHR both associated with higher risk of being
437 a smoker, with the magnitudes of effect being larger in men than women (BMI $P_{\text{het}}=0.002$; WHR
438 $P_{\text{het}}=3.7\times 10^{-14}$). WHRadjBMI was only associated with smoking status in men.

439 **Discussion**

440 Our study demonstrates that obesity is causally implicated in the etiology of two thirds of the leading
441 causes of death from non-communicable diseases (globally and in high-income countries) (24).

442 Furthermore, we identify that for some diseases, obesity conveys altered magnitudes of risk in men
443 and women. Such sexual dimorphism could be observed in the effects of BMI on T2D and waist-
444 related traits on COPD and renal failure. These findings have potential implications for public health
445 policy.

446 Obesity traits were causally related to higher risk of T2D, in alignment with previous studies (7–
447 12,20,61). We could not detect a sex difference in risk of T2D from higher WHR or WHRadjBMI.

448 Even though some observational studies have suggested that WHR may be a stronger predictor of T2D

449 risk in women than in men (19,20), studies investigating the effect on T2D risk from genetic
450 predisposition to higher WHRadjBMI have not found evidence of sexual heterogeneity (7,10,12). In
451 contrast, we found that BMI conferred a higher T2D risk in women than in men. Whereas men tend to
452 be diagnosed with T2D at lower BMI than women (62), there may be a stronger association between
453 increase of BMI and T2D risk in women than in men (16,19,61,63–66). Whether this reflects a
454 stronger causal effect of BMI on T2D risk in women has hitherto been unknown. We found no
455 evidence for sexual heterogeneity of the causal effect of BMI on potential glycaemic trait risk
456 mediators (FG and FI). There have been indications of higher BMI being observationally associated
457 with lower insulin sensitivity in men than in women, but this observed sex-difference may not reflect a
458 causal pathway or we are not capturing it by our glycaemic measurements (67–69). We also found
459 evidence of BMI causally increasing risk of T1D. Previous observational (70) and MR (71) studies
460 have implicated childhood BMI in risk of T1D. As SNPs associated with adult BMI have also been
461 found to affect childhood BMI (71,72), our results may well reflect the consequences of childhood
462 BMI on T1D rather than adult BMI. The results were robust to use of a stricter T1D case definition,
463 minimizing risk of erroneous finding due to misclassification of diabetes type.

464 Higher BMI, WHR and WHRadjBMI increased risk of CAD in both sexes, as shown previously (4,7–
465 9,11,12,16,18). Our obesity trait-CAD analyses did not show evidence for sexual heterogeneity.
466 Observational studies have indicated that waist-related traits may be more strongly associated with
467 cardiovascular disease in women and men, but have not been conclusive (16,18,73). However, a recent
468 study (12) investigated the effect of higher WHRadjBMI, lower gluteofemoral fat distribution, and
469 higher abdominal fat distribution, proxied by genetic variants, on CAD and T2D risk and found no
470 evidence that relationships differed between men and women, similar to our findings. BMI and WHR
471 have previously been observationally associated with risk of stroke (74–76) and a previous MR study
472 found a causal effect of BMI on ischemic stroke (77). However, some studies have found WHR to be
473 an epidemiological risk factor for stroke in men only (74,75). Our results confirm BMI as a causal risk
474 factor for overall stroke in both men and women. In women, the effects of WHR were directionally

475 consistent with harm, but the estimates were imprecise, probably reflecting insufficient power in the
476 sex-stratified analysis.

477 Our results also indicate that higher BMI and WHR increase risk of COPD and higher BMI the risk of
478 lung cancer; a likely common mechanism is through smoking. BMI has previously been implicated in
479 COPD, but is not an established epidemiological nor causative risk factor (8,78–80). Obesity may
480 directly contribute to COPD as its diagnosis is partly based on spirometry values, and obesity is
481 associated with lower lung function (80,81). Higher BMI also increased risk of lung cancer in our
482 study, similar to a previous MR study (82). Observational studies tend to identify associations between
483 smoking and lower body weight, but whereas smoking lowers body weight, higher BMI is associated
484 with increased smoking (82–85). We found associations between particularly BMI and WHR with
485 smoking propensity. To assess mediation, we therefore conducted analyses adjusting for smoking
486 status. This attenuated the associations between the obesity markers and risk of COPD and lung
487 cancer, suggesting that smoking status may be on the causal pathway between obesity, COPD and
488 lung cancer. This diminution does not discredit the validity of the MR analyses unadjusted for
489 smoking provided that the obesity instruments only affect smoking propensity through altered obesity
490 (86). Rather, they suggest that higher BMI impacts on disease beyond the immediate physiological
491 effects of obesity: by altering human behavior (i.e. increased smoking, likely motivated as a weight
492 loss strategy (87,88)) and this increased propensity to smoking has additional, far-reaching, deleterious
493 effects on human health, as evidenced by the higher risks of serious lung disease. Higher WHR was
494 associated with higher effects on both COPD and being a smoker in men than in women. Whereas the
495 sex difference in the effect of WHR on COPD persisted after adjustment for smoking status, we
496 cannot rule out that WHR has a higher effect on COPD in men than women through its effect on
497 smoking propensity, but that our smoking phenotype does not fully capture the life-long effects of
498 smoking in men and women.

499 Our results also provide further evidence for a role of obesity traits in both acute and chronic renal
500 failure using an MR design — previous MR studies assessing these relationships have not been
501 conclusive (7,8,89–91). Obesity may affect chronic renal disease through a number of mechanisms,

502 including structural changes in the kidney and through higher risks of mediating diseases, such as T2D
503 and renal cell carcinoma (91–95). We found central fat distribution (as measured by WHR and
504 WHRadjBMI) to have higher effects on chronic renal failure in men than in women, with evidence of
505 sexual heterogeneity. The reason for this sex difference is unclear — a recent MR study found both
506 BMI and WHR to increase risk of renal cell carcinoma but with no difference in risk between men and
507 women (95).

508 Obesity traits associated with increased risk of NAFLD and CLD (important and emerging causes of
509 chronic disease and mortality (27–30)), with the effect on CLD possibly mediated by NAFLD, since
510 CLD may be caused by NAFLD (28). A previous MR study found BMI to increase hepatic
511 triglyceride content (96). Our study confirms a role of both general obesity and central fat distribution
512 in NAFLD and CLD using an MR design. This strengthens evidence of a causal effect and emphasizes
513 the risk of increased CLD burden if the obesity prevalence continues to increase (1,27–30).

514 **Strengths and limitations**

515 Genetic instruments should only affect the outcome through the risk factor of interest and not through
516 any confounders (97,98). We performed sensitivity analyses (MR-Egger, weighted-median based
517 methods) more robust to such bias, which supported the main findings (44,45).

518 If instruments are weakly associated with their respective traits, it can introduce bias in MR studies
519 (99). We therefore only used instruments strongly associated with their respective risk factor, and
520 performed sensitivity analyses using a variety of SNP-selection and weighting approaches, including
521 unweighted and externally weighted scores, which also supported the main results (41,99,100).

522 Recent studies have also indicated that there may be slight population stratification in both GIANT
523 and UKBB, although such bias is likely to be minor (101,102). Our study was restricted to individuals
524 of Europeans ancestry; limiting our analyses to those of British ancestry only yielded near-identical
525 results. Associations between the obesity traits and outcomes may differ in other ancestries.

526 Finally, it is possible that our genetic instrument for WHRadjBMI might show features of collider bias
527 whereby SNPs included in the GRS associate with both higher WHR and lower BMI leading to

528 potentially spurious findings (103). We note that a recent GWAS (15) evaluated the potential for
529 collider bias in the WHRadjBMI GWAS and found limited evidence for such, although the GRS was
530 associated with higher WHR and lower BMI. The directional consistency of associations between
531 WHR and WHRadjBMI and disease endpoints in our analysis suggests that collider bias is unlikely to
532 represent a major source of error in this study.

533 **Conclusion**

534 Global prevalence of obesity is increasing (1). Our results implicate major obesity traits (BMI, WHR,
535 and WHRadjBMI) in the etiology of the leading causes of death globally, including CAD, stroke, type
536 2 and 1 diabetes, COPD, lung cancer and renal failure, as well as NAFLD and CLD. The risk increase
537 from obesity traits differs between men and women for T2D, renal failure and COPD. This
538 emphasizes the importance of improved preventative measures and treatment of obesity-related
539 disorders and implies that women and men may experience different disease sequelae from obesity,
540 with potential implications for provision of health services and health policy.

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555 **References**

- 556 1. GBD 2015 Obesity Collaborators, Afshin A, Forouzanfar MH, Reitsma MB, Sur P, Estep K, et
557 al. Health Effects of Overweight and Obesity in 195 Countries over 25 Years. *N Engl J Med*.
558 2017 Jul 6;377(1):13–27.
- 559 2. Seidell JC, Oosterlee A, Thijssen MA, Burema J, Deurenberg P, Hautvast JG, et al. Assessment
560 of intra-abdominal and subcutaneous abdominal fat: relation between anthropometry and
561 computed tomography. *Am J Clin Nutr*. 1987 Jan;45(1):7–13.
- 562 3. Prospective Studies Collaboration. Body-mass index and cause-specific mortality in 900 000
563 adults: collaborative analyses of 57 prospective studies. *Lancet*. 2009 Mar 28;373(9669):1083–
564 96.
- 565 4. Taylor AE, Ebrahim S, Ben-Shlomo Y, Martin RM, Whincup PH, Yarnell JW, et al.
566 Comparison of the associations of body mass index and measures of central adiposity and fat
567 mass with coronary heart disease, diabetes, and all-cause mortality: a study using data from 4
568 UK cohorts. *Am J Clin Nutr*. 2010 Mar 1;91(3):547–56.
- 569 5. WHO. Obesity: Preventing and Managing the Global Epidemic. Report of a WHO
570 consultation. Geneva: World Health Organization; 2000.
- 571 6. Davies NM, Holmes M V, Davey Smith G. Reading Mendelian randomisation studies: a guide,
572 glossary, and checklist for clinicians. *BMJ*. 2018 Jul 12;362:k601.
- 573 7. Emdin CA, Khera A V, Natarajan P, Klarin D, Zekavat SM, Hsiao AJ, et al. Genetic
574 Association of Waist-to-Hip Ratio With Cardiometabolic Traits, Type 2 Diabetes, and
575 Coronary Heart Disease. *JAMA*. 2017 Feb 14;317(6):626–34.
- 576 8. Millard LAC, Davies NM, Tilling K, Gaunt TR, Smith GD. Searching for the causal effects of
577 BMI in over 300 000 individuals, using Mendelian randomization. *bioRxiv*. 2017 Dec
578 19;236182.
- 579 9. Lyall DM, Celis-Morales C, Ward J, Iliodromiti S, Anderson JJ, Gill JMR, et al. Association of
580 Body Mass Index With Cardiometabolic Disease in the UK Biobank. *JAMA Cardiol*. 2017
581 Aug 1;2(8):882.
- 582 10. Huang T, Qi Q, Zheng Y, Ley SH, Manson JE, Hu FB, et al. Genetic Predisposition to Central
583 Obesity and Risk of Type 2 Diabetes: Two Independent Cohort Studies. *Diabetes Care*. 2015
584 Jul;38(7):1306–11.
- 585 11. Dale CE, Fatemifar G, Palmer TM, White J, Prieto-Merino D, Zabaneh D, et al. Causal
586 Associations of Adiposity and Body Fat Distribution With Coronary Heart Disease, Stroke
587 Subtypes, and Type 2 Diabetes Mellitus: A Mendelian Randomization Analysis. *Circulation*.
588 2017 Jun 13;135(24):2373–88.
- 589 12. Lotta LA, Wittemans LBL, Zuber V, Stewart ID, Sharp SJ, Luan J, et al. Association of
590 Genetic Variants Related to Gluteofemoral vs Abdominal Fat Distribution With Type 2
591 Diabetes, Coronary Disease, and Cardiovascular Risk Factors. *JAMA*. 2018 Dec
592 25;320(24):2553.
- 593 13. Finucane MM, Stevens GA, Cowan MJ, Danaei G, Lin JK, Paciorek CJ, et al. National,
594 regional, and global trends in body-mass index since 1980: systematic analysis of health
595 examination surveys and epidemiological studies with 960 country-years and 9·1 million
596 participants. *Lancet*. 2011 Feb 12;377(9765):557–67.
- 597 14. Ng M, Fleming T, Robinson M, Thomson B, Graetz N, Margono C, et al. Global, regional, and
598 national prevalence of overweight and obesity in children and adults during 1980-2013: a
599 systematic analysis for the Global Burden of Disease Study 2013. *Lancet (London, England)*.
600 2014 Aug 30;384(9945):766–81.

- 601 15. Pulit SL, Stoneman C, Morris AP, Wood AR, Glastonbury CA, Tyrrell J, et al. Meta-analysis
602 of genome-wide association studies for body fat distribution in 694,649 individuals of
603 European ancestry. *Hum Mol Genet.* 2018 Sep 14;
- 604 16. Rost S, Freuer D, Peters A, Thorand B, Holle R, Linseisen J, et al. New indexes of body fat
605 distribution and sex-specific risk of total and cause-specific mortality: a prospective cohort
606 study. *BMC Public Health.* 2018 Dec 2;18(1):427.
- 607 17. Lind L, Ärnlöv J, Lampa E. The Interplay Between Fat Mass and Fat Distribution as
608 Determinants of the Metabolic Syndrome Is Sex-Dependent. *Metab Syndr Relat Disord.* 2017
609 Sep 1;15(7):337–43.
- 610 18. Dagenais GR, Yi Q, Mann JFE, Bosch J, Pogue J, Yusuf S. Prognostic impact of body weight
611 and abdominal obesity in women and men with cardiovascular disease. *Am Heart J.* 2005 Jan
612 1;149(1):54–60.
- 613 19. Meisinger C, Döring A, Thorand B, Heier M, Löwel H. Body fat distribution and risk of type 2
614 diabetes in the general population: are there differences between men and women? The
615 MONICA/KORA Augsburg Cohort Study. *Am J Clin Nutr.* 2006 Dec 1;84(3):483–9.
- 616 20. Wannamethee SG, Papacosta O, Whincup PH, Carson C, Thomas MC, Lawlor DA, et al.
617 Assessing prediction of diabetes in older adults using different adiposity measures: a 7 year
618 prospective study in 6,923 older men and women. *Diabetologia.* 2010 May 10;53(5):890–8.
- 619 21. Bycroft C, Freeman C, Petkova D, Band G, Elliott LT, Sharp K, et al. Genome-wide genetic
620 data on ~500,000 UK Biobank participants. *bioRxiv.* 2017 Jul 20;166298.
- 621 22. Locke AE, Kahali B, Berndt SI, Justice AE, Pers TH, Day FR, et al. Genetic studies of body
622 mass index yield new insights for obesity biology. *Nature.* 2015 Feb 12;518(7538):197–206.
- 623 23. Shungin D, Winkler TW, Croteau-Chonka DC, Ferreira T, Locke AE, Mägi R, et al. New
624 genetic loci link adipose and insulin biology to body fat distribution. *Nature.* 2015 Feb
625 12;518(7538):187–96.
- 626 24. World Health Organization. The top 10 causes of death [Internet]. 2018 [cited 2018 Jul 7].
627 Available from: <http://www.who.int/news-room/fact-sheets/detail/the-top-10-causes-of-death>
- 628 25. Campbell JM, Lane M, Owens JA, Bakos HW. Paternal obesity negatively affects male fertility
629 and assisted reproduction outcomes: a systematic review and meta-analysis. *Reprod Biomed*
630 *Online.* 2015 Nov 1;31(5):593–604.
- 631 26. van der Steeg JW, Steures P, Eijkemans MJC, Habbema JDF, Hompes PGA, Burggraaff JM, et
632 al. Obesity affects spontaneous pregnancy chances in subfertile, ovulatory women. *Hum*
633 *Reprod.* 2007 Dec 14;23(2):324–8.
- 634 27. Williams CD, Stengel J, Asike MI, Torres DM, Shaw J, Contreras M, et al. Prevalence of
635 Nonalcoholic Fatty Liver Disease and Nonalcoholic Steatohepatitis Among a Largely Middle-
636 Aged Population Utilizing Ultrasound and Liver Biopsy: A Prospective Study.
637 *Gastroenterology.* 2011 Jan 1;140(1):124–31.
- 638 28. Chalasani N, Younossi Z, Lavine JE, Diehl AM, Brunt EM, Cusi K, et al. The diagnosis and
639 management of non-alcoholic fatty liver disease: Practice Guideline by the American
640 Association for the Study of Liver Diseases, American College of Gastroenterology, and the
641 American Gastroenterological Association. *Hepatology.* 2012 Jun 1;55(6):2005–23.
- 642 29. Vernon G, Baranova A, Younossi ZM. Systematic review: the epidemiology and natural
643 history of non-alcoholic fatty liver disease and non-alcoholic steatohepatitis in adults. *Aliment*
644 *Pharmacol Ther.* 2011 Aug 1;34(3):274–85.
- 645 30. Estes C, Razavi H, Loomba R, Younossi Z, Sanyal AJ. Modeling the epidemic of nonalcoholic

- 646 fatty liver disease demonstrates an exponential increase in burden of disease. *Hepatology*.
647 2018;67(1):123–33.
- 648 31. Macaluso M, Wright-Schnapp TJ, Chandra A, Johnson R, Satterwhite CL, Pulver A, et al. A
649 public health focus on infertility prevention, detection, and management. *Fertil Steril*. 2010 Jan
650 1;93(1):16.e1-16.e10.
- 651 32. Eastwood S V, Mathur R, Atkinson M, Brophy S, Sudlow C, Flaig R, et al. Algorithms for the
652 Capture and Adjudication of Prevalent and Incident Diabetes in UK Biobank. Herder C, editor.
653 *PLoS One*. 2016 Sep 15;11(9):e0162388.
- 654 33. Nelson CP, Goel A, Butterworth AS, Kanoni S, Webb TR, Marouli E, et al. Association
655 analyses based on false discovery rate implicate new loci for coronary artery disease. *Nat*
656 *Genet*. 2017 Jul 17;49(9):1385–91.
- 657 34. Tobin MD, Sheehan NA, Scurrah KJ, Burton PR. Adjusting for treatment effects in studies of
658 quantitative traits: antihypertensive therapy and systolic blood pressure. *Stat Med*. 2005 Oct
659 15;24(19):2911–35.
- 660 35. International Consortium for Blood Pressure Genome-Wide Association Studies, Ehret GB,
661 Munroe PB, Rice KM, Bochud M, Johnson AD, et al. Genetic variants in novel pathways
662 influence blood pressure and cardiovascular disease risk. *Nature*. 2011 Sep 11;478(7367):103–
663 9.
- 664 36. Lagou V, Mägi R, Hottenga J-JJ. Fasting glucose and insulin variability: sex-dimorphic genetic
665 effects and novel loci. *Prep*. 2018;
- 666 37. Sherry ST, Ward MH, Kholodov M, Baker J, Phan L, Smigielski EM, et al. dbSNP: the NCBI
667 database of genetic variation. *Nucleic Acids Res*. 2001 Jan 1;29(1):308–11.
- 668 38. Wald A. The Fitting of Straight Lines if Both Variables are Subject to Error. *Ann Math Stat*.
669 1940 Sep;11(3):284–300.
- 670 39. Bautista LE, Smeeth L, Hingorani AD, Casas JP. Estimation of Bias in Nongenetic
671 Observational Studies Using “Mendelian Triangulation.” *Ann Epidemiol*. 2006 Sep
672 1;16(9):675–80.
- 673 40. Thomas DC, Lawlor DA, Thompson JR. Re: Estimation of Bias in Nongenetic Observational
674 Studies Using “Mendelian Triangulation” by Bautista et al. *Ann Epidemiol*. 2007
675 Jul;17(7):511–3.
- 676 41. Burgess S, Thompson SG. *Mendelian Randomization - Methods for Using Genetic Variants in*
677 *Causal Estimation*. 1st ed. Boca Raton, FL, USA: CRC Press, Taylor & Francis Group,
678 Chapman and Hall; 2005. 6, 20, 45-52, 67, 124 p.
- 679 42. Yavorska OO, Burgess S. MendelianRandomization: an R package for performing Mendelian
680 randomization analyses using summarized data. *Int J Epidemiol*. 2017 Dec 1;46(6):1734–9.
- 681 43. Burgess S, Butterworth A, Thompson SG. Mendelian randomization analysis with multiple
682 genetic variants using summarized data. *Genet Epidemiol*. 2013 Nov;37(7):658–65.
- 683 44. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments:
684 effect estimation and bias detection through Egger regression. *Int J Epidemiol*. 2015 Apr
685 1;44(2):512–25.
- 686 45. Bowden J, Davey Smith G, Haycock PC, Burgess S. Consistent Estimation in Mendelian
687 Randomization with Some Invalid Instruments Using a Weighted Median Estimator. *Genet*
688 *Epidemiol*. 2016 May;40(4):304–14.
- 689 46. Cochran W. The combination of estimates from different experiments. *Biometrics*. 1954;10:
690 101-29.

- 691 47. Purcell S, Neale B, Todd-Brown K, Thomas L, Ferreira MAR, Bender D, et al. PLINK: A Tool
692 Set for Whole-Genome Association and Population-Based Linkage Analyses. *Am J Hum*
693 *Genet.* 2007 Sep;81(3):559–75.
- 694 48. Python Software Fondation. Python, version 3.5.2 [Internet]. Available from:
695 <https://www.python.org/>
- 696 49. McKinney W. Data Structures for Statistical Computing in Python. *Proc 9th Python Sci Conf.*
697 2010;51–6.
- 698 50. Oliphant TE. *A guide to NumPy.* USA: Trelgol Publishing; 2006.
- 699 51. R Core Team. R: A language and environment for statistical computing. R Foundation for
700 Statistical Computing, Vienna, Austria [Internet]. 2017. Available from: [https://www.r-](https://www.r-project.org/)
701 [project.org/](https://www.r-project.org/)
- 702 52. Wickham H, Francois R, Henry L, Müller K. dplyr: A Grammar of Data Manipulation. R
703 package version 0.7.4. 2017;
- 704 53. Free Software Foundation. bash 4.1.2(2) [Internet]. 2007. Available from:
705 <https://www.gnu.org/software/bash/>
- 706 54. Free Software Foundation. GNU AWK 3.1.7 [Internet]. 1989. Available from:
707 <https://www.gnu.org/software/gawk/manual/gawk.html>
- 708 55. Wickham H. *ggplot2: Elegant Graphics for Data Analysis.* New York: Springer-Verlag; 2009.
- 709 56. Doebler P. mada: Meta-Analysis of Diagnostic Accuracy. R package version 0.5.8. [Internet].
710 2017. Available from: <https://cran.r-project.org/package=mada%0A>
- 711 57. Auguie B. gridExtra: Miscellaneous Functions for “Grid” Graphics [Internet]. R package
712 version 2.3. 2017. Available from: <https://cran.r-project.org/package=gridExtra>
- 713 58. Sarkar D. *Lattice: Multivariate Data Visualization with R.* New York: Springer; 2008.
- 714 59. Kassambara A. ggpubr: “ggplot2” Based Publication Ready Plots. R package version 0.1.7.
715 2018.
- 716 60. World Health Organization. *Global Health Estimates 2016: Deaths by Cause, Age, Sex, by*
717 *Country and by Region, 2000-2016.* Geneva; 2018.
- 718 61. Vazquez G, Duval S, Jacobs DR, Silventoinen K. Comparison of Body Mass Index, Waist
719 Circumference, and Waist/Hip Ratio in Predicting Incident Diabetes: A Meta-Analysis.
720 *Epidemiol Rev.* 2007 May 2;29(1):115–28.
- 721 62. Logue J, Walker JJ, Colhoun HM, Leese GP, Lindsay RS, McKnight JA, et al. Do men develop
722 type 2 diabetes at lower body mass indices than women? *Diabetologia.* 2011 Dec
723 30;54(12):3003–6.
- 724 63. Bray GA. Medical Consequences of Obesity. *J Clin Endocrinol Metab.* 2004 Jun 1;89(6):2583–
725 9.
- 726 64. Kautzky-Willer A, Harreiter J, Pacini G. Sex and Gender Differences in Risk, Pathophysiology
727 and Complications of Type 2 Diabetes Mellitus. *Endocr Rev.* 2016;37(3):278–316.
- 728 65. Colditz GA, Willett WC, Rotnitzky A, Manson JE. Weight gain as a risk factor for clinical
729 diabetes mellitus in women. *Ann Intern Med.* 1995 Apr 1;122(7):481–6.
- 730 66. Chan JM, Rimm EB, Colditz GA, Stampfer MJ, Willett WC. Obesity, fat distribution, and
731 weight gain as risk factors for clinical diabetes in men. *Diabetes Care.* 1994 Sep;17(9):961–9.
- 732 67. Sierra-Johnson J, Johnson BD, Bailey KR, Turner ST. Relationships between insulin sensitivity

- 733 and measures of body fat in asymptomatic men and women. *Obes Res.* 2004 Dec;12(12):2070–
734 7.
- 735 68. Masharani U, Goldfine ID, Youngren JF. Influence of gender on the relationship between
736 insulin sensitivity, adiposity, and plasma lipids in lean nondiabetic subjects. *Metabolism.* 2009
737 Nov;58(11):1602–8.
- 738 69. Quon MJ. Limitations of the Fasting Glucose to Insulin Ratio as an Index of Insulin Sensitivity.
739 *J Clin Endocrinol Metab.* 2001 Oct 1;86(10):4615–7.
- 740 70. Verbeeten KC, Elks CE, Daneman D, Ong KK. Association between childhood obesity and
741 subsequent Type 1 diabetes: a systematic review and meta-analysis. *Diabet Med.* 2011 Jan
742 1;28(1):10–8.
- 743 71. Censin JC, Nowak C, Cooper N, Bergsten P, Todd JA, Fall T. Childhood adiposity and risk of
744 type 1 diabetes: A Mendelian randomization study. *Langenberg C*, editor. *PLoS Med.* 2017
745 Aug 1;14(8):e1002362.
- 746 72. Felix JF, Bradfield JP, Monnereau C, van der Valk RJP, Stergiakouli E, Chesi A, et al.
747 Genome-wide association analysis identifies three new susceptibility loci for childhood body
748 mass index. *Hum Mol Genet.* 2016 Jan 15;25(2):389–403.
- 749 73. Li C, Engström G, Hedblad B, Calling S, Berglund G, Janzon L. Sex differences in the
750 relationships between BMI, WHR and incidence of cardiovascular disease: a population-based
751 cohort study. *Int J Obes.* 2006 Dec 11;30(12):1775–81.
- 752 74. Hu G, Tuomilehto J, Silventoinen K, Sarti C, Männistö S, Jousilahti P. Body Mass Index,
753 Waist Circumference, and Waist-Hip Ratio on the Risk of Total and Type-Specific Stroke.
754 *Arch Intern Med.* 2007 Jul 9;167(13):1420.
- 755 75. Abete I, Arriola L, Etxezarreta N, Mozo I, Moreno-Iribas C, Amiano P, et al. Association
756 between different obesity measures and the risk of stroke in the EPIC Spanish cohort. *Eur J*
757 *Nutr.* 2015 Apr 6;54(3):365–75.
- 758 76. The Global Burden of Metabolic Risk Factors for Chronic Diseases Collaboration (BMI
759 Mediated Effects). Metabolic mediators of the effects of body-mass index, overweight, and
760 obesity on coronary heart disease and stroke: a pooled analysis of 97 prospective cohorts with
761 1·8 million participants. *Lancet.* 2014 Mar 15;383(9921):970–83.
- 762 77. Hagg S, Fall T, Ploner A, Magi R, Fischer K, Draisma HH, et al. Adiposity as a cause of
763 cardiovascular disease: a Mendelian randomization study. *Int J Epidemiol.* 2015 Apr
764 1;44(2):578–86.
- 765 78. Vozoris NT, O'Donnell DE. Prevalence, risk factors, activity limitation and health care
766 utilization of an obese, population-based sample with chronic obstructive pulmonary disease.
767 *Can Respir J.* 2012;19(3):e18-24.
- 768 79. Hanson C, Rutten EP, Wouters EFM, Rennard S. Influence of diet and obesity on COPD
769 development and outcomes. *Int J Chron Obstruct Pulmon Dis.* 2014;9:723–33.
- 770 80. Vestbo J, Hurd SS, Agustí AG, Jones PW, Vogelmeier C, Anzueto A, et al. Global Strategy for
771 the Diagnosis, Management, and Prevention of Chronic Obstructive Pulmonary Disease. *Am J*
772 *Respir Crit Care Med.* 2013 Feb 15;187(4):347–65.
- 773 81. Bedell GN, Wilson WR, Seebom PM. Pulmonary function in obese persons. *J Clin Invest.*
774 1958 Jul;37(7):1049–60.
- 775 82. Carreras-Torres R, Johansson M, Haycock PC, Wade KH, Relton CL, Martin RM, et al.
776 Obesity, metabolic factors and risk of different histological types of lung cancer: A Mendelian
777 randomization study. *Hu C*, editor. *PLoS One.* 2017 Jun 8;12(6):e0177875.

- 778 83. Morris RW, Taylor AE, Fluharty ME, Bjørngaard JH, Åsvold BO, Elvestad Gabrielsen M, et
779 al. Heavier smoking may lead to a relative increase in waist circumference: evidence for a
780 causal relationship from a Mendelian randomisation meta-analysis. The CARTA consortium.
781 *BMJ Open*. 2015 Aug 11;5(8):e008808.
- 782 84. Audrain-McGovern J, Benowitz NL. Cigarette smoking, nicotine, and body weight. *Clin*
783 *Pharmacol Ther*. 2011 Jul;90(1):164–8.
- 784 85. Rásky E, Stronegger WJ, Freidl W. The relationship between body weight and patterns of
785 smoking in women and men. *Int J Epidemiol*. 1996 Dec;25(6):1208–12.
- 786 86. Burgess S, Scott RA, Timpson NJ, Davey Smith G, Thompson SG, EPIC- InterAct
787 Consortium. Using published data in Mendelian randomization: a blueprint for efficient
788 identification of causal risk factors. *Eur J Epidemiol*. 2015 Jul 15;30(7):543–52.
- 789 87. Chiolero A, Faeh D, Paccaud F, Cornuz J. Consequences of smoking for body weight, body fat
790 distribution, and insulin resistance. *Am J Clin Nutr*. 2008 Apr 1;87(4):801–9.
- 791 88. Fulkerson JA, French SA. Cigarette smoking for weight loss or control among adolescents:
792 gender and racial/ethnic differences. *J Adolesc Health*. 2003 Apr;32(4):306–13.
- 793 89. Geng T, Smith CE, Li C, Huang T. Childhood BMI and Adult Type 2 Diabetes, Coronary
794 Artery Diseases, Chronic Kidney Disease, and Cardiometabolic Traits: A Mendelian
795 Randomization Analysis. *Diabetes Care*. 2018 May 1;41(5):1089–96.
- 796 90. Todd JN, Dahlström EH, Salem RM, Sandholm N, Forsblom C, FinnDiane Study Group the
797 FS, et al. Genetic Evidence for a Causal Role of Obesity in Diabetic Kidney Disease. *Diabetes*.
798 2015 Dec;64(12):4238–46.
- 799 91. van Zuydam NR, Ahlqvist E, Sandholm N, Deshmukh H, Rayner NW, Abdalla M, et al. A
800 Genome-Wide Association Study of Diabetic Kidney Disease in Subjects With Type 2
801 Diabetes. *Diabetes*. 2018 Jul 1;67(7):1414–27.
- 802 92. Tsuboi N, Utsunomiya Y, Kanzaki G, Koike K, Ikegami M, Kawamura T, et al. Low
803 glomerular density with glomerulomegaly in obesity-related glomerulopathy. *Clin J Am Soc*
804 *Nephrol*. 2012 May 1;7(5):735–41.
- 805 93. Kovesdy CP, Furth SL, Zoccali C, World Kidney Day Steering Committee on behalf of the
806 WKDS. Obesity and Kidney Disease: Hidden Consequences of the Epidemic. *Can J kidney*
807 *Heal Dis*. 2017;4:2054358117698669.
- 808 94. Wang F, Xu Y. Body mass index and risk of renal cell cancer: a dose-response meta-analysis
809 of published cohort studies. *Int J cancer*. 2014 Oct 1;135(7):1673–86.
- 810 95. Johansson M, Carreras-Torres R, Scelo G, Purdue MP, Mariosa D, Muller DC, et al. The
811 influence of obesity-related factors in the etiology of renal cell carcinoma—A mendelian
812 randomization study. Minelli C, editor. *PLOS Med*. 2019 Jan 3;16(1):e1002724.
- 813 96. Stender S, Kozlitina J, Nordestgaard BG, Tybjærg-Hansen A, Hobbs HH, Cohen JC. Adiposity
814 amplifies the genetic risk of fatty liver disease conferred by multiple loci. *Nat Genet*. 2017
815 Jun;49(6):842–7.
- 816 97. Greenland S. An introduction to instrumental variables for epidemiologists. *Int J Epidemiol*.
817 2000 Aug;29(4):722–9.
- 818 98. Martens EP, Pestman WR, de Boer A, Belitser S V, Klungel OH. Instrumental variables:
819 application and limitations. *Epidemiology*. 2006 May;17(3):260–7.
- 820 99. Burgess S, Davies NM, Thompson SG. Bias due to participant overlap in two-sample
821 Mendelian randomization. *Genet Epidemiol*. 2016 Nov;40(7):597–608.

- 822 100. Burgess S, Thompson SG. Use of allele scores as instrumental variables for Mendelian
823 randomization. *Int J Epidemiol*. 2013 Aug;42(4):1134–44.
- 824 101. Berg JJ, Harpak A, Sinnott-Armstrong N, Joergensen AM, Mostafavi H, Field Y, et al.
825 Reduced signal for polygenic adaptation of height in UK Biobank. *bioRxiv*. 2018 Jun
826 27;354951.
- 827 102. Haworth S, Mitchell R, Corbin L, Wade KH, Dudding T, Budu-Aggrey A, et al. Common
828 genetic variants and health outcomes appear geographically structured in the UK Biobank
829 sample: Old concerns returning and their implications. *bioRxiv*. 2018 Apr 11;294876.
- 830 103. Davey Smith G, Paternoster L, Relton C. When Will Mendelian Randomization Become
831 Relevant for Clinical Practice and Public Health? *JAMA*. 2017 Feb 14;317(6):589.
- 832
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855 **Supporting Information**

856 **S1 Supporting Information.**