

1 **On the prevalence of uninformative parameters in statistical models applying model**  
2 **selection in applied ecology**  
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## 13 **Abstract**

14 Research in applied ecology provides scientific evidence to guide conservation policy and  
15 management. Applied ecology is becoming increasingly quantitative and model selection via  
16 information criteria has become a common statistical modeling approach. Unfortunately,  
17 parameters that contain little to no useful information are commonly presented and interpreted as  
18 important in applied ecology. I review the concept of an uninformative parameter in model  
19 selection using information criteria and perform a literature review to measure the prevalence of  
20 uninformative parameters in model selection studies applying Akaike's Information Criterion  
21 (AIC) in 2014 in four of the top journals in applied ecology (*Biological Conservation*,  
22 *Conservation Biology*, *Ecological Applications*, *Journal of Applied Ecology*). Twenty-one  
23 percent of studies I reviewed applied AIC metrics. Many (31.5 %) of the studies applying AIC  
24 metrics in the four applied ecology journals I reviewed had or were very likely to have  
25 uninformative parameters in a model set. In addition, more than 40 % of studies reviewed had  
26 insufficient information to assess the presence or absence of uninformative parameters in a  
27 model set. Given the prevalence of studies likely to have uninformative parameters or with  
28 insufficient information to assess parameter status (71.5 %), I surmise that much of the policy  
29 recommendations based on applied ecology research may not be supported by the data analysis. I  
30 provide warning signals and a decision tree to help reduce the prevalence of uninformative  
31 parameters in studies applying model selection with information criteria. The four warning  
32 signals and decision tree should assist authors, reviewers, and editors to screen for uninformative  
33 parameters in studies applying model selection with information criteria. In the end, careful  
34 thinking at every step of the scientific process and greater reporting standards are required to  
35 detect uninformative parameters in studies adopting an information criteria approach.

36 **Keywords:** Akaike Information Criterion, biostatistics, conservation biology, information  
37 theory, model selection bias, pretending variables, transparency, variable importance

38

## 39 **Introduction**

40 Conservation biology emerged as a crisis discipline in the 1970s in response to evidence of  
41 widespread declines in biodiversity [1]. Along with the evolution of new technologies (e.g.  
42 Remote Sensing, Geographic Information Systems) and increasing availability of environmental  
43 (e.g. Land-use) and biodiversity (e.g. species occurrence records) data, the discipline has  
44 developed into a rigorous quantitative science [2-4]. These advances in methods and data allow  
45 applied ecologists to tackle complex problems at larger temporal and spatial scales than before.  
46 The application of quantitative analyses and the interpretation of these analyses in applied  
47 ecology is particularly important as research in this field often informs policy and management  
48 practices [5-7].

49         Around the same time as the field of conservation biology was emerging, Akaike [8] was  
50 paving the way for the broad application of the information criteria (IC) approach to statistics for  
51 evaluating data-based evidence for multiple working hypotheses [9,10]. Model selection using  
52 IC is now a common type of analysis in applied ecology (Fig 1). This statistical approach  
53 encourages *a priori* development of multiple working hypotheses and presents formal methods  
54 for weighing the evidence supporting the different hypotheses [see reviews in 10-12]. As with  
55 any quantitative method, there are many challenges and ways to misuse IC techniques and recent  
56 work has highlighted some important issues in the application of IC in ecology, evolution,  
57 wildlife management and conservation biology. For example, Galipaud et al. [13] show how  
58 model averaging using the sum of model weights can overestimate parameter importance and

59 Mac Nally et al. [14] present a plea for including absolute measures of model goodness-of-fit  
60 when possible as the top ranked model determined by IC may not be a “good” model. In  
61 addition, several researchers have called on the need for independent model validation [14-16]  
62 and better reporting of methods and results to facilitate critical evaluation of research  
63 conclusions (i.e. greater transparency [17]). Here, I focus on one issue in IC; uninformative  
64 parameters (*sensu* [18]). Uninformative parameters have received some attention in the literature  
65 [e.g. 11,18-20] but this issue is still prevalent in applied ecology.

66 *INSERT FIGURE 1 HERE*

67 **Figure 1.** Summary of the use of IC and prevalence of uninformative parameters in articles  
68 reviewed from four top applied ecology journals (*Biological Conservation, Conservation*  
69 *Biology, Ecological Applications, Journal of Applied Ecology*). Articles were classified into four  
70 different categories for the prevalence of uninformative parameters in model sets – see main text  
71 for description of categories. Note that many papers in these journals do not use statistical  
72 analyses (e.g. essays). UP = uninformative parameter.

73 An “uninformative parameter” (*sensu* [18]) also known as a “pretending variable” (*sensu*  
74 [9,11]), is a variable that has no relationship with the response, makes little to no improvement in  
75 the log-likelihood of a model (i.e. model fit) but can be included in a model ranked close to  
76 models with informative parameters based on IC. Interpreting uninformative parameters as  
77 important is a Type I error in statistics (i.e. false-positive [21]). If the interpretation of  
78 uninformative parameters as important is common, particularly in policy and management  
79 related fields such as conservation biology or medicine, then the policy recommendations of  
80 research may not be supported by the data analysis. What is more, poor data analysis and  
81 interpretation can lead to the natural selection of bad science (*sensu* [22]). My objectives are

82 twofold; i) review and operationalize the concept of an uninformative parameter in model  
83 selection using IC and ii) quantify the prevalence of uninformative parameters in model selection  
84 using IC in applied ecology. Based on the results, I end with recommendations for how to screen  
85 for uninformative parameters in model selection studies using IC.

## 86 **Identifying uninformative parameters**

87 In this section, I provide background on model selection using IC and formally present the  
88 concept of an uninformative parameter in this context. A clear definition of this concept is  
89 essential before presenting the methods and results of the quantitative review of uninformative  
90 parameters in model selection using IC in applied ecology.

91 Several IC exist for assessing the weight the evidence in support of different hypotheses  
92 formulated as competing statistical models [9] but I focus on the most commonly applied tool,  
93 Akaike's Information Criterion (AIC) and related variations (e.g. AIC<sub>c</sub>). AIC is defined as

$$94 \quad AIC = -2 \log L + 2K$$

95 where  $L$  is the likelihood of the model given the data and  $K$  is the number of estimated  
96 parameters in the model.  $K$  is included as a penalty for adding additional parameters to the  
97 model, therefore AIC prioritizes parsimonious models. It is customary practice in model  
98 selection with IC to rank competing models from lowest to highest AIC or more specifically to  
99 rank models based on the  $\Delta AIC$  which is the difference in AIC between a focal model and the  
100 model with the lowest AIC [9]. Top ranked models are models with a  $\Delta AIC = 0$ . Models with  
101  $\Delta AIC < 2$  are often considered equally supported or not differentiable from the top ranked model  
102 [9]. It is easy to see that two models (subscript 1 and 2) having identical  $\log L$  (i.e. same fit to the  
103 data) but differing only by 1 estimated parameter (i.e.  $K_1 - K_2 = 1$ ) will have a difference in AIC  
104  $= 2$ . Given identical  $\log L$ , the model with the additional parameter will have a *larger* AIC than

105 the model with one less parameter and therefore the model with one additional parameter will be  
106 ranked *below* the simpler model. Likewise, models with identical  $\log L$  and differing by 2  
107 estimated parameters will have a difference in  $AIC = 4$ , and so on.

108 Uninformative parameters occur when there are nested models or more specifically, more  
109 complex versions of simpler models, in a model set [11,18,23-25]. Importantly, if the  $\log L$  of a  
110 model has not improved with the addition of a parameter, it is likely that this additional  
111 parameter does not improve model fit and should be considered an uninformative parameter.  
112 However, if adding a parameter to a model improves the model fit, then the  $\log L$  will increase  
113 and the AIC will decrease (i.e. the model with the extra parameter will be ranked *above* the  
114 model with one less parameter). See Table 1 for an illustration of uninformative parameters from  
115 some recent empirical research. Next, I outline the warning signals for uninformative parameters  
116 and summarize these warning signs in a decision tree that can be used to formally identify  
117 uninformative parameters in IC analyses (Fig 2).

118

119 **Table 1** Here, I illustrate uninformative parameters from a real example derived from analyses in Yalcin and Leroux [26]. The  
120 objective of this study was to assess the relative and combined effects of land-use change and climate change on the colonization and  
121 extinction of species. We used a case study in Ontario, Canada where birds were surveyed in standardized grids during two time  
122 periods (1981-1985 and 2001-2005). Below I provide results for a subset of the colonization models of one of the study species, black-  
123 throated blue warbler (*Setophaga caerulescens*). In the colonization model, the black-throated blue warbler is observed as absent in a  
124 grid in the first time period and the response is warbler absence (0) or presence (1) in the second time period. We selected covariates  
125 based on *a priori* hypotheses. These covariates measured changes in land-use (% change in land-cover in each grid (%LCC), %  
126 change in land-cover in 20km buffers surrounding each grid (%LCCb) and change in Net Primary Productivity ( $\Delta$ NPP)) and climate  
127 (change in mean winter temperature ( $\Delta$ MWT), change in mean summer temperature ( $\Delta$ MST), and change in mean winter precipitation  
128 ( $\Delta$ MWP)) during the time period between bird sampling. All models include sampling effort (SE) in order to control for variable  
129 sampling effort across grids and between time periods. Yalcin and Leroux [26] fit generalized linear models with a binomial error  
130 structure and a logit link for local colonization models for the black-throated blue warbler. See for full details on data, methods, and  
131 hypotheses pertaining to each covariate used in these models. Table 1 provides a summary of AIC model selection results and  
132 parameter estimates (95% Confidence Interval) for a sub-set of the colonization models considered for this species. By following the  
133 decision tree in Fig 2, Yalcin and Leroux [26] identified the variable %LCCb is an uninformative parameter in models 2, 4, and 6.

Model	SE	$\Delta$ NPP	%LCC	%LCCb	$\Delta$ MST	$\Delta$ MWT	$\Delta$ MWP	K	log L	$\Delta$ AIC <sub>C</sub>	*Pseudo R <sup>2</sup>
1	0.30	0.06	0.04		0.25	-2.54		6	-276.14	0.00	0.28

	(0.24,0.39)	(0.03,0.09)	(0.02,0.06)		(0.10,0.40)	(-4.64,-0.46)					
2	0.30 (0.24,0.39)	0.06 (0.03,0.09)	0.04 (0.01,0.06)	<b>0.00</b> <b>(-0.05,0.05)</b>	0.25 (0.10,0.41)	-2.54 (-4.65,-0.47)		7	-276.14	2.00	0.28
3	0.30 (0.22,0.39)	0.05 (0.02,0.08)			0.25 (0.10,0.41)	-2.06 (-4.23,-0.09)	0.07 (0.02,0.12)	6	-278.18	4.07	0.27
4	0.30 (0.22,0.38)	0.05 (0.02,0.08)		<b>0.03</b> <b>(-0.02,0.07)</b>	0.29 (0.12,0.45)	-2.04 (-4.23,-0.12)	0.07 (0.01,0.12)	7	-277.63	4.97	0.27
5	0.30 (0.22,0.38)	0.06 (0.04,0.09)			0.32 (0.18,0.46)		0.08 (0.03,0.13)	5	-279.95	5.61	0.26
6	0.30 (0.22,0.38)	0.06 (0.04,0.09)		<b>0.03</b> <b>(-0.02,0.07)</b>	0.35 (0.20,0.50)		0.08 (0.03,0.13)	6	-279.34	6.39	0.26
Intercept								1	-344.09	125.91	0.00

134 \*McFadden's pseudo R<sup>2</sup>

135



136 *INSERT FIGURE 2 HERE*

137 **Figure 2.** Decision tree for identifying models with uninformative parameters in a model set  
138 based on warning signals (see main text). This decision tree was used to assess the prevalence of  
139 uninformative parameters in top applied ecology journals (see Quantitative review). Note that the  
140 particular cut-off for the first step will vary based on the IC used (see main text).

141 Here I focus on cases where one additional estimated parameter may be an uninformative  
142 parameter but the logic also applies for cases where a model contains two additional estimated  
143 parameters and both may be uninformative parameters. These warning signals should be  
144 assessed in sequence (i.e. they build on each other, Fig 2). An uninformative parameter may exist  
145 in a model set if:

- 146 1. there are two models that differ by one estimated parameter that are within  $AIC \leq 2$  of  
147 each other. Authors must screen *all* possible model pairs in a model set (i.e. not just top  
148 ranked models) as a parameter may not be uninformative in every model in which it  
149 appears given varying levels of multi-collinearity among covariates. Note that different  
150 IC metrics will yield slightly different cut-off points for detecting this first warning  
151 signal. For example, based on  $AIC_c$  ( $AIC_c = AIC + \frac{2K^2 + 2K}{n - K - 1}$ ) and a sample size ( $n$ ) of 30,  
152 two models with identical  $\log L$  and differing by only 1 parameter will have  $\Delta AIC_c =$   
153 2.21. Consequently, the particular cut-off for this first warning signal should be  
154 considered in light of the specific IC metric used.
- 155 2. the model with one additional parameter (as outlined in warning signal 1) is ranked *below*  
156 the model with one less parameter (i.e. less parsimonious model  $AIC >$  more  
157 parsimonious model  $AIC$ ). This suggests that the model with one additional parameter  
158 does not have a much better fit (i.e.  $\log L$ ) than the simpler model.

159 3. the models identified in warning signals 1 and 2 have virtually identical  $\log L$ . Nearly  
160 identical  $\log L$  suggests that the additional parameter is not contributing to improving  
161 model fit. This warning signal is subjective as there will be very few cases where the  $\log$   
162  $L$  of two different models are identical. Consequently, authors must decide what is a  
163 sufficient difference to demonstrate that the added parameter contains useful information  
164 about the data. A strength of model selection with IC is that it allows researchers to use  
165 all available information to draw inference [9]. If authors are too strict in the cut-off for  
166 what they consider useful information, then authors risk losing inferential power. To  
167 avoid committing a Type I error, it may be best to err on the side of caution and to lose  
168 some information than to mis-interpret uninformative parameters as useful information.  
169 Given  $\log L$  is a relative measure based on the data, there is no specific cut-off to  
170 determine if  $\log L$  are similar. In lieu of a specific cut-off, researchers should assess  
171 parameter estimates and confidence intervals as a final step to identify uninformative  
172 parameters (see warning signal 4 [18]).

173 4. the additional parameter identified from warning signals 1-3 has a parameter estimate  
174 near zero with a confidence interval overlapping 0 [11,18,20,27]. A parameter estimate  
175 near zero suggests that there is no relationship between this variable and the response  
176 variable. Arnold [18] and Galipaud et al. [20] provide specific guidance on confidence  
177 interval interpretations for identifying uninformative parameters.

178 By sequentially searching for the above warning signals, authors can identify all possible  
179 uninformative parameters in a model set (Fig 2). In order for readers of scientific papers to  
180 independently assess these warning signals, it follows that authors must provide all information  
181 required to interpret model selection with IC analyses.

182           While some recent research has demonstrated issues with uninformative parameters  
183 usually as part of broader studies [11,18,20,25,27], none have documented the prevalence of  
184 uninformative parameters in applied ecology and focused on solutions. Next, I provide a  
185 quantitative review of the prevalence of uninformative parameters in four of the top journals in  
186 applied ecology.

## 187 **Methods**

188 I reviewed all 2014 articles in four of the top journals in applied ecology; *Biological*  
189 *Conservation*, *Conservation Biology*, *Ecological Applications*, and *Journal of Applied Ecology*  
190 for evidence of uninformative parameters. Specifically, I downloaded every article for each  
191 journal and I searched for the terms AIC or Akaike Information Criterion. I retained all articles  
192 with the term AIC in it. Following this first pass, I removed all articles that did not apply AIC in  
193 their analysis (i.e. they just mention AIC in the text).

194           I determined the presence or absence of uninformative parameters by systematically  
195 searching for the four warning signals in the order listed in the previous section and outlined in  
196 the decision tree (Fig 2). For warning signal 1, I only focused on pairs of models that differ by  
197  $AIC \sim 2$  and one estimated parameter. I used  $AIC \sim 2$  as a cut-off as different articles used  
198 different AIC metrics (e.g. AIC,  $AIC_c$ ,  $qAIC$ ). I did not focus on cases where two models differ  
199 by 2 or more parameters (i.e. differ by  $AIC \sim 4$ ) – so my assessment of the prevalence of  
200 uninformative parameters is a minimum or conservative estimate. In many cases, authors did not  
201 provide sufficient information to fully determine if a model set included a model with an  
202 uninformative parameter. For example, AIC tables or estimates of model coefficients were often  
203 absent and when AIC tables were provided, key information such as the number of estimated  
204 parameters ( $K$ ) or  $\log L$  were often omitted. Consequently, I identified four possible

205 uninformative parameter outcomes for each article in the study; i) articles with uninformative  
206 parameters, ii) articles with no uninformative parameters, iii) articles *very likely* to have  
207 uninformative parameters, iv) articles with insufficient information to identify uninformative  
208 parameters. These possible outcomes can be interpreted as follows. An article was classed as  
209 outcome i) if it had all four warning signals and outcome ii) if it did not have one of the warning  
210 signals. I assumed that the occurrence of one model with one uninformative parameter was  
211 sufficient to classify an article as having uninformative parameters. In most cases where there  
212 was one model with confirmed or very likely uninformative parameters, there were many models  
213 with uninformative parameters in the model set. I do not, however, report on the number of  
214 uninformative parameters per article. An article was classified as outcome iii) if it had the first  
215 three warning signals and as outcome iv) if there was insufficient information to assess any of  
216 the warning signals.

217         The article classification followed a two-step process. In the first step, two reviewers with  
218 experience in model selection with IC (lead author and A. Tanner (MSc working with lead  
219 author)) independently placed each article into one of the four outcomes listed above. In step  
220 two, the lead author reviewed the independent responses and flagged any articles with  
221 disagreement between reviewers ( $n = 16$  or 9 % of studies). Then the lead author re-read and re-  
222 assigned each article that had initial disagreement between reviewers. I extracted the following  
223 information from each article: basic article information (authors, title, journal, issue, pages), IC  
224 used (i.e. AIC, AIC<sub>c</sub>, qAIC), the presence or absence of  $\Delta$ AIC, parameter estimates, model  
225 averaging, and stepwise IC and the uninformative parameter ranking (i.e. yes, no, very likely,  
226 insufficient information). All data are available online [28].

## 227 **Results**

228 The literature review revealed 329, 187, 163, and 182 articles published in 2014 in *Biological*  
 229 *Conservation*, *Conservation Biology*, *Ecological Applications*, and *Journal of Applied Ecology*,  
 230 respectively (Table 2). From this total, there were 87 (26 %), 22 (12 %), 33 (20 %), 39 (21 %)  
 231 articles from *Biological Conservation*, *Conservation Biology*, *Ecological Applications*, and  
 232 *Journal of Applied Ecology*, respectively that used AIC metrics in their analysis (Table 2, Fig 1).  
 233 While only 21 % of articles (n = 181 / 861) in these journals apply AIC, many papers in these  
 234 journals do not use statistical analyses (e.g. essays).

235 **Table 2** Summary statistics (number and percentage of articles) of uninformative parameter  
 236 assessment for four top journals in applied ecology. Articles were classified into four different  
 237 categories for the prevalence of uninformative parameters in model sets – see main text for  
 238 description of categories. The number of articles and percent of articles reported are compared to  
 239 the subset of articles with AIC per journal, except in the final row which reports the totals across  
 240 all journals. UP = uninformative parameter.

Journal (Total # in 2014)	Total # (%) with AIC	Number of articles (%) with			
		UP	very likely UP	no UP	insufficient information
Biological Conservation (329)	87(26)	7(8)	20(23)	25(29)	35(40)
Conservation Biology (187)	22(12)	1(5)	7(32)	5(23)	9(41)
Ecological Applications (163)	33(20)	0(0)	8(24)	8(24)	17(51)
J. of Applied Ecology (182)	39(21)	3(8)	11(28)	12(31)	13(33)
Total (861)	181(21)	11(6)	46(25)	50(28)	74(41)

241  
 242

243           Across all journals there was at least one model with an uninformative parameter in an  
244 article's model set in 6 % of cases and no model with an uninformative parameter in an article's  
245 model set in 28 % of cases. Only 4 % of articles self-identified uninformative parameters and  
246 removed them from their model set. *Biological Conservation* and *Journal of Applied Ecology*  
247 had the highest percentage of articles adopting an AIC approach where the presence or absence  
248 of uninformative parameters could be confirmed (Table 2). This statistic goes hand in hand with  
249 the fact that these two journals had the lowest percentage of articles with insufficient information  
250 to assess uninformative parameters, albeit these percentages were still high (*Biological*  
251 *Conservation* = 40 %, *Journal of Applied Ecology* = 33 %). *Ecological Applications* had no  
252 confirmed cases of models with uninformative parameters but it also had the highest percentage  
253 of articles with insufficient information to identify uninformative parameters (51 %, Table 2).  
254 Note that in many cases, there is no possibility for uninformative parameters as a model set may  
255 be very simple with a null model (i.e. intercept only) and one additional model with one fixed  
256 effect or a set of non-nested models (i.e. models with no overlapping parameters). For example,  
257 Barnes et al. [29]'s model set to investigate the response of dung beetle communities to land-use  
258 management in Afromontane rainforests in Nigeria included four non-nested models and  
259 therefore there is no possibility for uninformative parameters in their model set. Consequently,  
260 the percentage of studies with no uninformative parameters should be higher than the percentage  
261 of studies with uninformative parameters.

262           In 23 to 32 % (grand mean 25 %) of articles across the four journals there was evidence  
263 that uninformative parameters were very likely based on the information presented in the article  
264 (i.e. warning signals 1-3 were confirmed, Fig 1). Altogether, nearly 1/3 (31.5 %) of all articles

265 considered had or were very likely to have models with an uninformative parameter in the model  
266 set (Table 2, Fig 1).

## 267 **Discussion**

268 Applied ecologists are increasingly being called on to support evidence-based  
269 environmental and natural resource management. The evidence we provide, therefore, must be  
270 based on sound empirical design, statistical analyses, and interpretations of these analyses [5]. In  
271 this study, I conducted a quantitative review of the prevalence of uninformative parameters in  
272 model selection using IC in applied ecology. My review revealed two main findings with  
273 potential impacts on the field of applied ecology; i) many articles applying model selection with  
274 IC in this study had or were very likely to have at least one model in a model set with one  
275 uninformative parameter (Table 2, Fig 1) and ii) many articles had insufficient information to  
276 identify uninformative parameters in their model set. These two issues stand to reduce the  
277 validity of inference drawn from statistical analyses applying model selection using IC in applied  
278 ecology.

279 In many of the articles reviewed herein, uninformative parameters were reported as  
280 important and often interpreted as such. For example, *Biological Conservation* [30 – author  
281 names withheld] report the following results for two competing models (i.e.  $y \sim \text{time}$ ;  $y \sim \text{time} +$   
282 weather) of florican (*Sypheotides indicus*) detection in semiarid grasslands in India: “The time  
283 model had smallest  $AIC_c$  value, more precise effect ( $\beta = 0.62_{\text{Mean}} \pm 0.31_{\text{SE}}$ ) and parsimony than  
284 the time and weather model ( $\Delta AIC_c = 1.54$ ,  $\beta = 0.56 \pm 0.31$  [time],  $0.28 \pm 0.34$  [weather]). Time  
285 had stronger influence ( $AIC_c\text{-wt} = 0.61$ ) than weather ( $AIC_c\text{-wt} = 0.31$ ) on display frequency...”.  
286 The two models differ by one parameter, have almost identical log  $L$  (i.e. differ by 0.69) and the  
287 parameter estimate for weather overlaps zero. In this case, weather is an uninformative parameter

288 and weather should be removed from the model set and presented as having little to no support  
289 (i.e. not interpreted as important). In contrast to this example many of the papers that did have  
290 uninformative parameters did not interpret these parameters as important. For example, Rudolphi  
291 et al. [31] have many uninformative parameters in their model sets to investigate the impacts of  
292 logging on bryophytes and lichens. However, they restrict their interpretation to parameters with  
293 95 % confidence estimates that do not overlap zero.

294 The quantitative review revealed that more than 40 % of all articles had insufficient  
295 information to identify uninformative parameters (Table 2, Fig 1). This lack of transparency in  
296 reporting of methods and results has been highlighted previously [e.g. 10,17,32,33]. The missing  
297 information ranged from not reporting the number of parameters or log  $L$  per model, to not  
298 reporting parameter estimates, and in many cases not presenting any AIC table.

299 Based on my findings, I present the following recommendations for reducing erroneous  
300 interpretation of uninformative parameters from model selection studies in applied ecology.  
301 First, once authors have identified all uninformative parameters in a model set, I recommend that  
302 all models with uninformative parameters be removed from the model set and that the model  
303 removal be noted in the results section (see discussion of full reporting below; [11,18]). In some  
304 cases, the top model may include an uninformative parameter uncovered elsewhere in the IC  
305 table and in such cases, the original top model should be removed from the model set. Models  
306 with interaction terms (i.e.  $X_1 * X_2$ ) where a component (e.g.  $X_1$ ) of the interaction is an  
307 uninformative parameter in the model set should be retained because a parameter may be  
308 informative (i.e. improve model fit) once it is in interaction with another parameter. However, if  
309 an interaction term is an uninformative parameter, then all models with the full interaction term  
310 should be removed from the model set. The type of variable (i.e. continuous or categorical) will



311 influence the approach to removing models with uninformative parameters. Continuous variables  
312 and categorical variables with two levels usually have one estimated parameter and advice for  
313 removal of uninformative parameters above can be followed. Categorical variables with more  
314 than 2 levels will have  $n - 1$  estimated parameters where  $n$  is the number of levels. It is possible  
315 that one level of a multi-level categorical variable is uninformative but others are informative. In  
316 these cases, authors should retain the categorical variable but interpret the results for every level  
317 making a clear distinction between the informative and uninformative levels.

318         Second, a solution to detecting and removing uninformative parameters from analyses is  
319 to report sufficient information to assess the warning signals of uninformative parameters (see  
320 Fig 2, [11,18]). Proper reporting of quantitative analyses should be a default in scientific  
321 research. Transparency will allow peer review to help identify uninformative parameters at  
322 various stages of the review process. At minimum, papers using model selection with IC must  
323 report AIC tables with  $K$ ,  $\log L$ ,  $\Delta AIC$ , absolute measure of goodness-of-fit (see [14]) and  
324 parameter estimates with some measure of confidence intervals for all models [9,10].  
325 Abbreviated AIC tables (i.e. models with  $\Delta AIC < 8$ ) may occur in the main text as per Burnham  
326 et al. [10] but the AIC table for the full model set prior to removal of models with uninformative  
327 parameters should be placed in supplement. Graphical presentations of modeled relationships  
328 also may be useful for understanding relationships [34,35] and detecting uninformative  
329 parameters.

330         As described in Arnold [18], authors must not sacrifice full reporting when removing  
331 models with uninformative parameters. Specifically, authors should present all models in the  
332 methods and report the presence of uninformative parameters and subsequent model removal in  
333 the results. If done correctly, readers should be able to identify all models considered by authors

334 and the particular parameters that were uninformative. Examples for clear reporting of all models  
335 considered and removal of uninformative parameters can be seen in Devries et al. [36], Fondell  
336 et al. [37], Beauchesne et al. [38] and Fitzherbert et al. [39].

337 Third, some IC techniques are more prone to uninformative parameters than others and  
338 steering away from such approaches can help reduce the occurrence of uninformative  
339 parameters. Cade [40] and Galipaud et al. [13,20] convincingly demonstrate the perils of model  
340 averaging by summed IC weights (but see [41]). Most articles considered in the quantitative  
341 review which used model averaging by summed AIC weights were very likely to have  
342 uninformative parameters. For example, *Conservation Biology* [42– author names withheld]  
343 present summed AIC weights for several models with uninformative parameters for the effects of  
344 land-use (i.e. mining vs agriculture) on West African rainforest bird richness (see their Figs 3  
345 and 4).

346 Stepwise AIC runs counter to the original intention of model selection with IC [9-12,21].  
347 Stepwise AIC does not encourage the creation of *a priori* hypotheses and models but is rather  
348 usually applied to all possible models. Stepwise AIC was common in the studies reviewed with  
349 14% of articles using some form of stepwise AIC in their analysis. The process of fitting all  
350 possible models without *a priori* reason is flawed [9-12,21] and will often inflate the occurrence  
351 of uninformative parameters relative to an *a priori* model selection approach [27]. Note that  
352 uninformative parameters may still occur in a model set based on *a priori* selection of variables.  
353 However, trying all possible models will almost surely lead to more uninformative parameters.  
354 Stepwise AIC also does not allow one to assess model selection uncertainty [27] which is a  
355 critical component of multiple hypothesis testing. While stepwise AIC has critical flaws, the end  
356 result likely does not include uninformative parameters as the stepwise process ends with one top

357 model and models with additional variables but higher AIC would have been thrown out during  
358 the stepwise process. That said, stepwise AIC should only be used when paired with *a priori*  
359 selection of variables and models.

360 Common advice to reduce uninformative parameters is to remove more complex or  
361 nested versions of simpler models in a model set [12,24,25]. This approach is not new to  
362 statistics [24] and it is commonly used in a Bayesian framework [43]. The articles in the data set  
363 that used this approach [e.g. 38,39,44] did not have uninformative parameters. Authors should  
364 think critically about nested models and only use the more complex versions of nested models if  
365 they represent *a priori* hypotheses for the phenomenon of interest.

## 366 **Conclusion**

367 I provide quantitative evidence of the prevalence of uninformative parameters in IC studies in  
368 applied ecology and recommendations on how to diagnose and remove these uninformative  
369 parameters. My review focused on the most widely used IC metric; AIC, but uninformative  
370 parameters should be considered when applying other IC metrics (e.g. Bayesian Information  
371 Criterion, Deviance Information Criterion). Model selection with IC is a powerful tool to assess  
372 the evidence supporting multiple working hypotheses but only if the tool is applied correctly.  
373 Given the close connection of applied ecology to conservation policy and management, careful  
374 thinking at every step of the process from the individual researchers (i.e. study design, statistical  
375 analysis, interpretation of results), reviewers (i.e. interpretation of results, transparency in  
376 reporting), and editors is required for valid inferences to be made. Additional vigilance can be  
377 facilitated by improving the reporting standards for statistical analyses [35,45] and by screening  
378 the statistical analyses of submitted articles. In the end, researchers must be critical of results and

379 seek statistical advice when in doubt - biodiversity and the reputation of the field of applied  
380 ecology depends on it.

381

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388

389 **Data accessibility:** Data available from figshare digital repository

390 <https://doi.org/10.6084/m9.figshare.6002582.v1>

391

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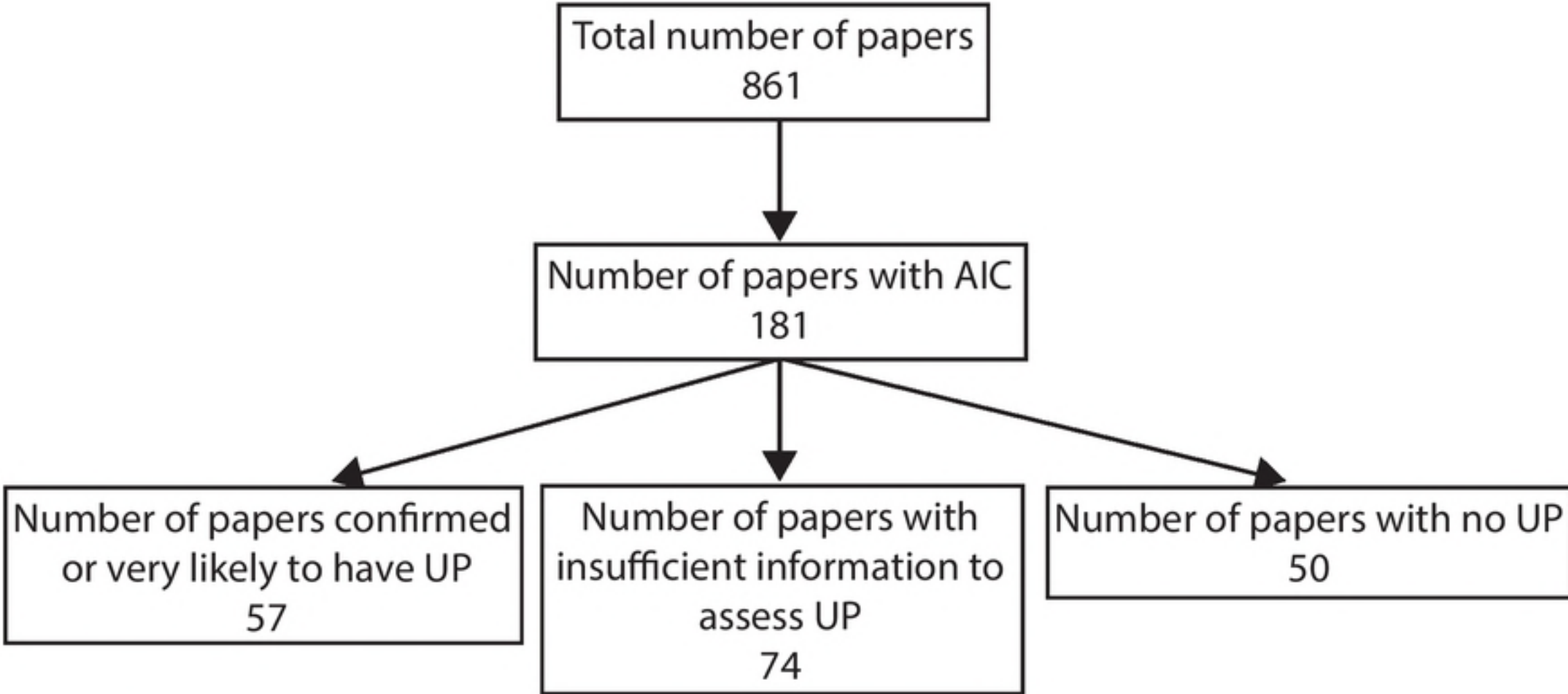


Figure 1

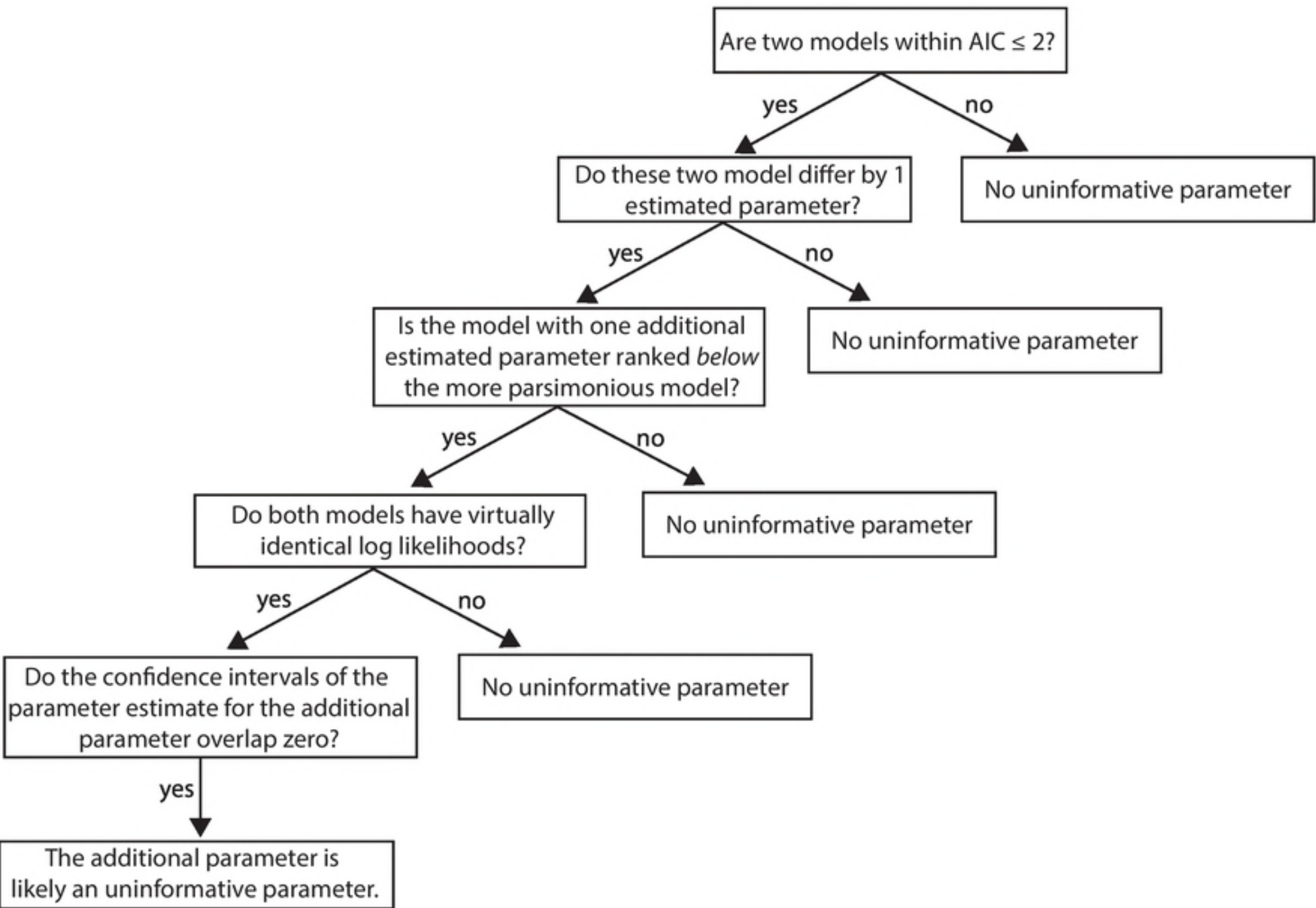


Figure 2