

1 **On the relationship between high-order linkage disequilibrium and**
2 **epistasis**

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20 Keywords: *Arabidopsis thaliana*, High order linkage disequilibrium, Epistasis,

21 Molybdenum, Leaf

22

23 Data availability:

24 i) Genome wide re-sequencing data are available as part of *Arabidopsis thaliana* 1001

25 genomes project <http://1001genomes.org/data-center.html>, ii) 250 K SNP chip data

26 were available in the *Arabidopsis thaliana* Regmap panel

27 (http://bergelson.uchicago.edu/?page_id=790), iii) The Molybdenum level for 340

28 *Arabidopsis thaliana* are available in

29 <https://doi.org/10.1371/journal.pgen.1005648.s005>. iv) Corresponding genotypes are

30 extracted from a subset of the Regmap panel

31 (http://bergelson.uchicago.edu/?page_id=790).

32

33 Running title: High-order LD and statistical epistasis

34

ABSTRACT

35 A plausible explanation for statistical epistasis revealed in genome wide association
36 analyses is the presence of high order linkage disequilibrium (LD) between the
37 genotyped markers tested for interactions and unobserved functional polymorphisms.
38 Based on findings in experimental data, it has been suggested that high order LD
39 might be a common explanation for statistical epistasis inferred between local
40 polymorphisms in the same genomic region. Here, we empirically evaluate how
41 prevalent high order LD is between local, as well as distal, polymorphisms in the
42 genome. This could provide insights into whether we should account for this when
43 interpreting results from genome wide scans for statistical epistasis. An extensive and
44 strong genome wide high order LD was revealed between pairs of markers on the high
45 density 250k SNP-chip and individual markers revealed by whole genome sequencing
46 in the *A. thaliana* 1001-genomes collection. The high order LD was found to be more
47 prevalent in smaller populations, but present also in samples including several
48 hundred individuals. An empirical example illustrates that high order LD might be an
49 even greater challenge in cases when the genetic architecture is more complex than
50 the common assumption of bi-allelic loci. The example shows how significant
51 statistical epistasis is detected for a pair of markers in high order LD with a complex
52 multi allelic locus. Overall, our study illustrates the importance of considering also
53 other explanations than functional genetic interactions when genome wide statistical
54 epistasis is detected, in particular when the results are obtained in small populations of
55 inbred individuals.
56

57

INTRODUCTION

58 The genetic architecture of most biological traits is complex and involves multiple
59 genes, whose effects are often influenced by interactions with other genes and
60 environmental factors. To study the relative contributions by genes, environmental
61 factors and their interactions in segregating populations, statistical genetic approaches
62 are commonly used to partition the genetic variance to additive and dominance
63 variance of individual loci and epistatic interaction variance between them (Lynch
64 and Walsh 1998). In principle, the variance partitioning is performed by associating
65 the phenotypic variation for a trait in a population with linear combinations of the
66 genotypes within and/or across loci. How the genotypes are combined (parameterized)
67 in the model is determined by the genetic model used in the analysis. The classic
68 quantitative genetics models are parameterized to capture the genetic variance in a
69 hierarchical manner. First, a main additive allele-substitution is defined. Then, if
70 accounted for, dominance is modeled as a single-locus deviation from additivity and
71 genetic interactions as multi-locus deviations from single locus additivity and
72 dominance (Nelson *et al.* 2013). As a consequence of this, the genetic contributions of
73 individual and combinations of loci described as additive, dominance and epistatic
74 variances are unlikely to reflect the underlying biological mechanisms (Carlborg *et al.*
75 2006; Phillips 2008; Huang *et al.* 2012; Sackton and Hartl 2016; Forsberg *et al.* 2017).
76

77 Although the ultimate aim of a genetic association study is generally to detect
78 functional polymorphisms, most often genotypes are only scored for a reduced set of
79 polymorphisms (genetic markers). These reduced marker sets are selected with the
80 aim to tag as many of the unobserved functional polymorphisms as possible. The
81 statistical inferences of the underlying genetic architecture made from such reduced
82 sets of markers can, however, be problematic in some cases. For example, multiple
83 unobserved functional polymorphisms can lead to associations to individual markers
84 that do not properly represent the causal variants (Platt *et al.* 2010), and high order
85 linkage disequilibrium (LD) to single functional polymorphism can lead to indirect
86 statistical epistatic associations to pairs of markers (Wood *et al.* 2014). Here, we focus
87 on high order linkage disequilibrium defined as when two genotyped markers tag an
88 un-genotyped polymorphism (see Materials and Methods section). It is still unknown
89 how prevalent and strong such high order LD is in the genome, making it difficult to

90 estimate how many reported pairwise statistical epistatic interactions are due to such
91 LD. However, the study by *Wood et al* (*Wood et al.* 2014) presented results
92 suggesting that many of the significant statistical epistatic interactions detected
93 between pairs of local markers by *Hemani et al.* (*Hemani et al.* 2014) might be due to
94 high-order LD to unobserved, linked sequence polymorphisms in the same genomic
95 region. Many past and current studies of genetic interactions in, for example,
96 *Drosophila*, plant, animal and human populations (*Shimomura et al.* 2001; *Anholt et*
97 *al.* 2003; *Caicedo et al.* 2004; *Segrè et al.* 2004; *Carlborg et al.* 2006; *Hemani et al.*
98 2014) rely on genome-wide statistical analyses of pairwise interactions between
99 selected sets of markers as in (*Hemani et al.* 2014). With the increasing interest in,
100 and availability of, sufficiently large datasets for epistatic association analyses it is
101 therefore important to also evaluate the risk of making false inferences about loci
102 being involved in functional genetic interactions from findings of statistical epistasis,
103 when they instead are due to high order LD.

104

105 Here, we empirically explore the prevalence and strength of high order LD within and
106 between chromosomes in publically available high-density SNP and whole-genome
107 re-sequencing data from the model plant *Arabidopsis thaliana*. Two locus LDs are
108 calculated between the markers selected for the 250k *A. thaliana* SNP chip that have
109 been the basis for many GWAS analyses in the past, and the additional SNPs revealed
110 by whole genome sequencing using data from the 1001 genomes project (*Atwell et al.*
111 2010; *Cao et al.* 2011; *Horton et al.* 2012; *Schmitz et al.* 2013; *Alonso-Blanco et al.*
112 2016). Strong high order LD was found to be common both within and across
113 chromosomes between pairs of markers from the SNP-chip and the sequencing
114 polymorphisms and often the combined genotype of the marker pair tagged the
115 genotype of the sequencing markers better than any single marker on the SNP chip.
116 The risk of falsely inferring genetic interactions between markers on different
117 chromosomes in a two-locus interaction analysis might increase in situations when the
118 underlying genetic architecture is more complex, for example when a single locus
119 contains multiple functional alleles. This is illustrated using an empirical example
120 from a second public *A. thaliana* dataset (*Forsberg et al.* 2015). Overall, this study
121 provides new insights that deepen our understanding about the link between high
122 order LD and statistical epistasis to guide researchers when interpreting results

123 obtained from epistatic genetic association analyses.

124

125 MATERIALS AND METHODS

126

127 **Methods**

128 When an individual marker is in complete linkage disequilibrium ($r^2 = 1$) with a
129 functional polymorphism affecting a studied trait, a single-locus association test
130 between the marker and the trait will capture all the phenotypic variance contributed
131 by the functional polymorphism. A basic assumption in genetic association studies is
132 that at least one genotyped marker will be in sufficiently high LD with each functional
133 polymorphism to detect it in this way. In reality, however, not all functional
134 polymorphism will be in such perfect LD with a genotyped marker, and then there is a
135 risk that the joint genotype of two (or more) markers tags the genotype of the
136 functional polymorphism better than any single marker (high-order LD > single-
137 marker LD). This will, as discussed below, influence the significances of the trait-
138 marker associations detected in a genetic association analysis and the inferences made
139 about the genetic architecture of the trait.

140

141 *Quantifying high order linkage disequilibrium*

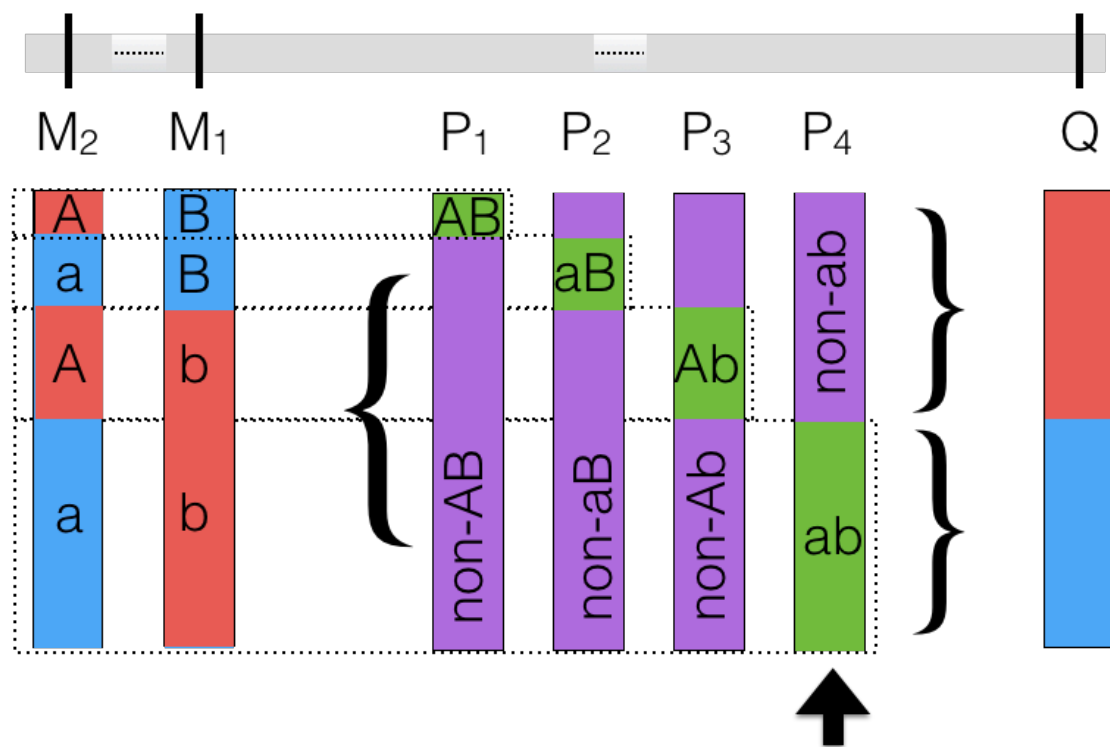
142 We calculate the high order LD between pairs of predictors (here genotyped SNP
143 markers) and single targets (here un-genotyped SNP polymorphisms) following (Hao
144 *et al.* 2007).

145

146 Consider a pair of bi-allelic predictor SNPs (M_1 and M_2 ; Figure 1). These markers can
147 together form four two-locus genotypes: AB, Ab, aB and ab (Figure 1). We now want
148 to know whether any two-locus predictor could tag the single locus target genotype
149 better than any of the individual predictor genotypes (i.e. evaluate whether
150 $\max(\text{second-order LD}) > \max(\text{single order LD})$). To calculate the high order LD
151 between the two predictors (M_1 and M_2) and the single target (Q), the two-locus
152 M_1M_2 genotype is used to create a multi-allelic pseudo marker (P) with four alleles
153 (Figure 1). In this way, a second-order LD (r^2) can be calculated for each of the
154 possible ways that M_1 and M_2 together can tag the genotype at Q (Figure 1).

155

156 The calculation of the second order LD therefore first involves creating the four
 157 possible bi-allelic pseudomarkers (P_1 , P_2 , P_3 & P_4 ; Figure 1) from the two locus
 158 M_1/M_2 genotypes. These are assigned the genotypes $P_1\{AB, \text{non-AB}\}$, $P_2\{Ab, \text{non-}$
 159 $Ab\}$, $P_3\{aB, \text{non-aB}\}$ and $P_4\{ab, \text{non-ab}\}$, respectively. The $LD-r^2$ is then computed
 160 between the target (Q) and the four bi-allelic pseudomarkers (P_1 , P_2 , P_3 & P_4). For
 161 each pair of predictors, the second order LD is then defined as the $LD-r^2$ for the
 162 pseudomarker with the highest $LD-r^2$ to the target. Pseudomarkers with higher $LD-r^2$
 163 to the target (Q) than 0.3 are kept for further analyses. The $LD-r^2$ values were
 164 computed using the software *LdCompare* (Hao *et al.* 2007).



165
 166 **Figure 1.** Illustration of how the pseudomarkers (P_1 , P_2 , P_3 , P_4) used in the estimation of the second
 167 order linkage disequilibrium between a pair of linked or unlinked markers (predictors; M_1 and M_2)
 168 and a third linked or unlinked functional polymorphism (target; Q) are created. The pseudomarkers
 169 together represent the possible bi-allelic formulations of the two-locus M_1M_2 genotypes. The maximum
 170 pairwise $LD-r^2$ between the target and the four pseudomarkers (P_4) defines the second order LD
 171 between the predictors (M_1 , M_2) and the target (Q).

172

173

174 *Statistical epistasis emerging from high order linkage disequilibrium*

175 In a genetic association study in an inbred or haploid population, two-locus epistasis

176 is typically modelled as:

177

$$178 \quad Y = a_1\beta_1 + a_2\beta_2 + a_1a_2\beta_{12} + e \quad [1]$$

179

180 Here, a_1 and a_2 are indicator variables for the genotypes at two genotyped markers, M_1
181 and M_2 , taking values 1/-1 for the two alternative homozygous genotypes AA vs aa
182 and BB vs bb, respectively. a_1a_2 is an indicator variable for the interaction between a_1
183 and a_2 taking value 1 for the two-locus genotypes AABB and aabb and -1 for AAbb
184 and aaBB. β_1 , β_2 and β_{12} are the corresponding estimates for the marginal (additive)
185 effects and the additive-by-additive interaction between the loci.

186

187 The aim of a statistical epistatic analysis is to include an interaction term in the model
188 [1] to estimate the deviations of the two-locus genotype-values (AABB, AAbb, aaBB
189 and aabb) from the predictions obtained by the marginal (additive) effects (Alvarez-
190 Castro and Carlborg 2007). However, a non-zero estimate of the interaction term in
191 model [1] does not, as noted e.g. by Wood et al. (Wood *et al.* 2014) necessarily have
192 to result from a genetic interaction. It could, for example, instead emerge from a
193 second-order LD between two markers and a single functional polymorphism. Here,
194 refer back to Figure 1. Now assume that a trait is determined by a single functional
195 locus (Q). Two markers, M_1 and M_2 , are genotyped but neither of these markers
196 individually tag the causal genotype (blue) at Q well. However, the causal (blue)
197 allele at Q is tagged perfectly by one of the two-locus M_1M_2 genotypes (ab; Figure 1),
198 while the other three M_1M_2 two-locus genotypes (aB, Ab and AA; Figure 1) are only
199 present together with the no-effect (red) allele at locus Q. When fitting model [1] to
200 the genotypes of marker M_1 and M_2 , the estimate for the interaction term (β_{12}) will be
201 non-zero, illustrating how statistical epistasis can emerge from the second-order LD
202 between M_1 and M_2 and Q. This example illustrates a scenario similar to what was
203 empirically observed in (Wood *et al.* 2014), where physically linked markers in low
204 LD with each other tagged haplotypes that were in high order LD with a
205 polymorphism that was unobserved in the original study.

206

207 *Classifying identified high order linkage disequilibrium triplets depending on the*
208 *distance between the loci*

209 Here, we evaluate the prevalence and strength of high order LD between pairs of
210 markers selected for genotyping on a 250k SNP chip (predictors) and a third locus
211 revealed by whole genome sequencing (targets) using publicly available datasets in *A.*
212 *thaliana* (Cao *et al.* 2011; Alonso-Blanco *et al.* 2016). Three types of high order LD
213 are defined based on the locations of the predictors relative to the target. If both
214 predictors are located within 1Mb of the target it is classified as cis-cis. If only one
215 predictor is closer than 1Mb it is classified as cis-trans. If none is closer than 1Mb it is
216 classified as trans-trans. The choice of a 1Mb threshold to define cis vs trans
217 predictors is arbitrary, but we consider it useful for evaluating how common high
218 order LD is between predictors near (local/cis) and far (global/trans) from the target.

219

220 **Material**

221

222 *The genome wide prevalence of high order linkage disequilibrium in publically*
223 *available Arabidopsis thaliana datasets*

224 The *A. thaliana* 1001-genomes project has released complete genome sequences for
225 hundreds of wild collected accessions (<http://www.1001genomes.org>). Here, we used
226 whole-genome SNP data on 728 accessions scored by whole genome re-sequencing
227 (Cao *et al.* 2011; Alonso-Blanco *et al.* 2016). The predictors used in our analysis was
228 a subset of the SNPs selected for the 250k *A. thaliana* SNP chip (Horton *et al.* 2012)
229 ($n = 200,352$ in total; $MAF > 0.05$) and the targets a subset of the SNPs revealed
230 using whole-genome re-sequencing ($n = 1,641,240$ in total; $MAF > 0.05$) (Table 1).
231 Although the results from the analyses of this data will be specific to this species and
232 dataset, it is assumed that the relationships between targets and predictors will be a
233 realistic representation of what to be expected also in other populations. This is
234 because the selection of markers for the high-density 250k SNP chip, was done for the
235 purpose of genetic association studies following similar procedures as used also in
236 other species and populations.

237

238 The reason for only studying a subset of the possible targets and predictors is that it
239 was not computationally feasible to exhaustively evaluate the high order LD between
240 all possible pairs of predictors selected for the 250k SNP chip and all the targets
241 revealed by genome sequencing. Instead, the second order LD was exhaustively

242 calculated for all targets and predictors i) within a randomly selected 6 Mb window on
243 chromosome 2 as well as ii) between three randomly selected windows from different
244 chromosomes (Table 1). Computations were performed for the entire population (n =
245 728 individuals) and two smaller random samples of n = 100 and n = 50 individuals.
246 The results for the populations with n = 100 and n = 728 are reported in the main
247 manuscript and the results for n = 50 is reported in the Supplementary material.

248

249 **Table 1.** *Regions and SNPs selected for evaluation of second order LD.*

	Window 1	Window 2	Targets ¹	Predictors ²	Filtered targets ³
Region 1	Chr2: 8-14Mb	-	70,712	6,053	-
Regionpair 1	Chr1: 10-12Mb	Chr3: 10-12Mb	29,133	6,245	20,239
Regionpair 2	Chr2: 10-12Mb	Chr4: 10-12Mb	23,751	5,302	15,887
Regionpair 3	Chr2: 10-12Mb	Chr3: 10-12Mb	23,751	5,212	15,884
Genome			1,486,942	154,298	1,229,012

250 ¹Total number of polymorphic SNPs in the evaluated windows/genome in the population revealed via
251 whole-genome re-sequencing (Alonso-Blanco *et al.* 2016). ²Total number of polymorphic SNPs in the
252 two windows/genome included on the 250k AT SNP-chip (Horton *et al.* 2012); ³Number of target SNPs
253 in the two windows/genome with $LD-r^2 < 0.6$ to any individual predictor.

254

255 The predictor pairs in the evaluated windows in the genome with high order $LD-r^2 >$
256 0.6 to a target were classified as cis-cis/cis-trans/trans-trans. To extrapolate these
257 findings to the genome level, the proportions of all evaluated predictor pairs that
258 displayed these patterns were calculated and then multiplied with the total number of
259 possible cis-cis/cis-trans/trans-trans pairs in the genome (Table S1).

260

261 *Analyzing a public A. thaliana dataset for two locus statistical epistasis*

262 A publicly available dataset including 340 *Arabidopsis thaliana* accessions were used
263 for a genome wide association analysis. In short, the plants were grown in a controlled
264 environment with 6 biological replicate plants per accession. Analyses by Inductively
265 Coupled Mass Spectroscopy (ICP-MS) provided estimates of leaf molybdenum
266 concentration as described in (Baxter *et al.* 2010; Forsberg *et al.* 2015). The
267 accessions were genotyped for 141,385 SNP markers with $MAF > 0.15$ (Atwell *et al.*
268 2010; Baxter *et al.* 2010; Shen *et al.* 2012; Forsberg *et al.* 2015). A more thorough
269 description of the dataset can be found in (Baxter *et al.* 2010; Forsberg *et al.* 2015). In
270 an earlier study of this dataset (Forsberg *et al.* 2015), it was revealed that a large

271 fraction of the genetic variance for this trait was explained by a single linkage block
272 containing several low-frequency, large effect structural variants that were poorly
273 tagged by the genotyped SNPs. This linkage block was originally identified due to its
274 large marginal, variance heterogeneity effect in the population (Shen *et al.* 2012). It is
275 known that statistical epistasis and genetic variance heterogeneity can emerge from
276 similar genetic architectures (Forsberg *et al.* 2015), and this population was therefore
277 selected for further evaluations of whether high order LD between the genotyped
278 SNPs and these hidden polymorphisms could lead to statistical epistasis in a two locus
279 association analysis. We performed an exhaustive, two-dimensional genome scan for
280 pairwise statistical epistasis between the genotyped markers and the level of
281 molybdenum in the leaf using the software *plink* (Purcell *et al.* 2007) without control
282 for population structure. Thereafter, each pair of loci that passed the genome wide
283 significance threshold in the initial scan was fitted in a two-locus epistatic genetic
284 model [1] using *hglm* function in *hglm* package (Rönnegård *et al.* 2010) to correct for
285 the possible effects of population structure via the genomic kinship matrix as in
286 (Forsberg *et al.* 2015). The significance threshold used to infer significant interacting
287 pairs ($p < 3.2 \times 10^{-10}$) was defined as a Bonferroni corrected nominal 5% significance
288 threshold. The correction was done for an estimated number of independent
289 association tests assumed to equal the number of independent LD blocks in the *A.*
290 *thaliana* genome as described in (Lachowiec *et al.* 2015).

291

292 *Data availability*

293 Genome wide re-sequencing data are available as part of the *Arabidopsis thaliana*
294 1001 genomes project <http://1001genomes.org/data-center.html>. The 250 K SNP chip
295 data are available as part of the genotype data for the *Arabidopsis thaliana* Regmap
296 panel (http://bergelson.uchicago.edu/?page_id=790). The Molybdenum levels for the
297 340 *Arabidopsis thaliana* accessions are available in
298 <https://doi.org/10.1371/journal.pgen.1005648.s005>

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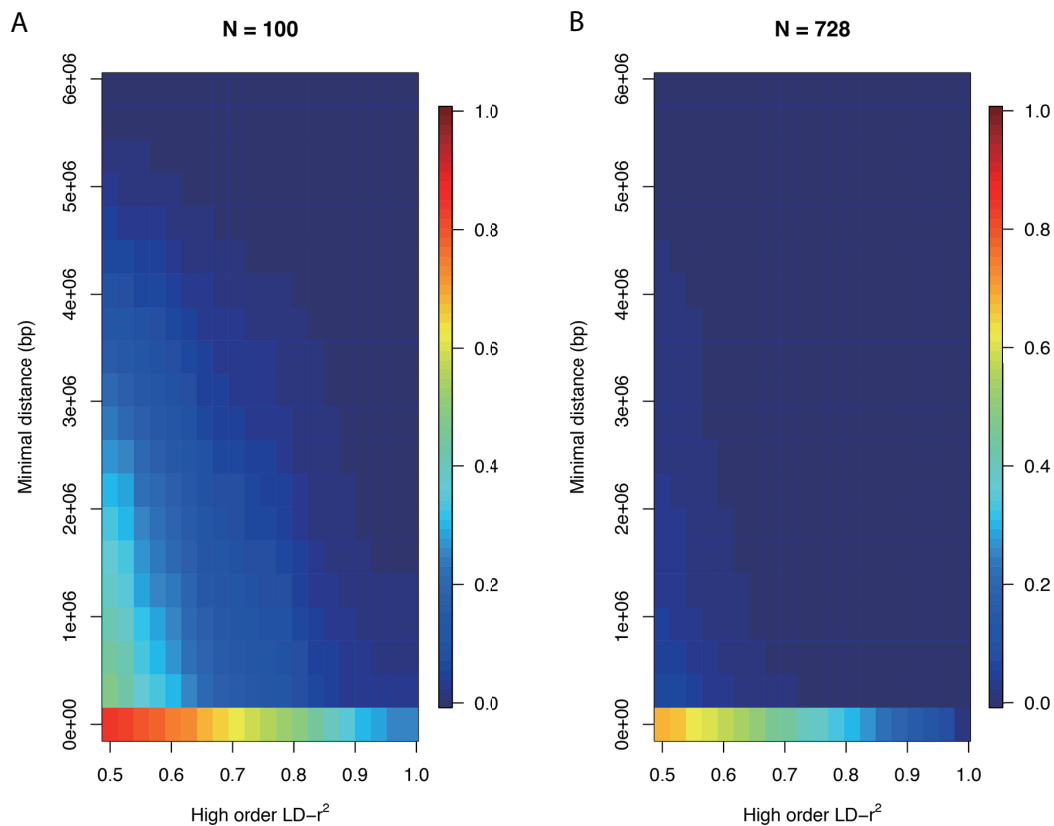
301 **RESULTS**

302 This study aims to answer the following questions by analyzing two public *A.*
303 *thaliana* datasets: How common can we expect high order LD to be between pairs of

304 SNPs selected for genotyping and hidden sequence variants in the genome? Is high
305 order LD primarily observed between predictors tightly linked to a target functional
306 polymorphism (in cis) as in (Wood *et al.* 2014), or is it also observed for predictors
307 unlinked to the target (in trans)? How dependent is the prevalence of high order LD
308 and cis vs trans predictors on the population size? We also present an empirical
309 example where high order LD exists between a cis-trans predictor pair with
310 significant statistical epistasis and a locus displaying a strong genetic variance
311 heterogeneity due to independent contributions by multiple linked polymorphisms
312 (Forsberg *et al.* 2015). This illustrates how complex inheritance patterns of individual
313 loci, something usually not explored in GWAS data, further complicates the
314 interpretation of detected statistical epistatic signals.

315

316 *The population size affects the prevalence and location of predictors in high order LD*
317 The high order LD- r^2 values for all pairs of predictors and individual targets in a 6Mb
318 window on Chromosome 2 (Table 1) is shown for populations with $n = 100$ and $n =$
319 728 individuals in Figure 2. The strongest second order LD- r^2 was observed where at
320 least one predictor is located near the target (y-axis). When the sample size was
321 smaller ($n = 100$; Figure 2A), strong second order LD- r^2 was rather common also
322 when both predictors were located far from the target. For example, 20% of the
323 targets had a high order LD- $r^2 > 0.65$ with a predictor pair where at least one of the
324 predictors was located more than 1Mb away from it. Even though the prevalence of
325 strong high order LD- r^2 decreases when the sample size increases, it is still common
326 in the large population ($n = 728$; Figure 2B), with the highest prevalence when at least
327 one of the predictors is located close to the target.

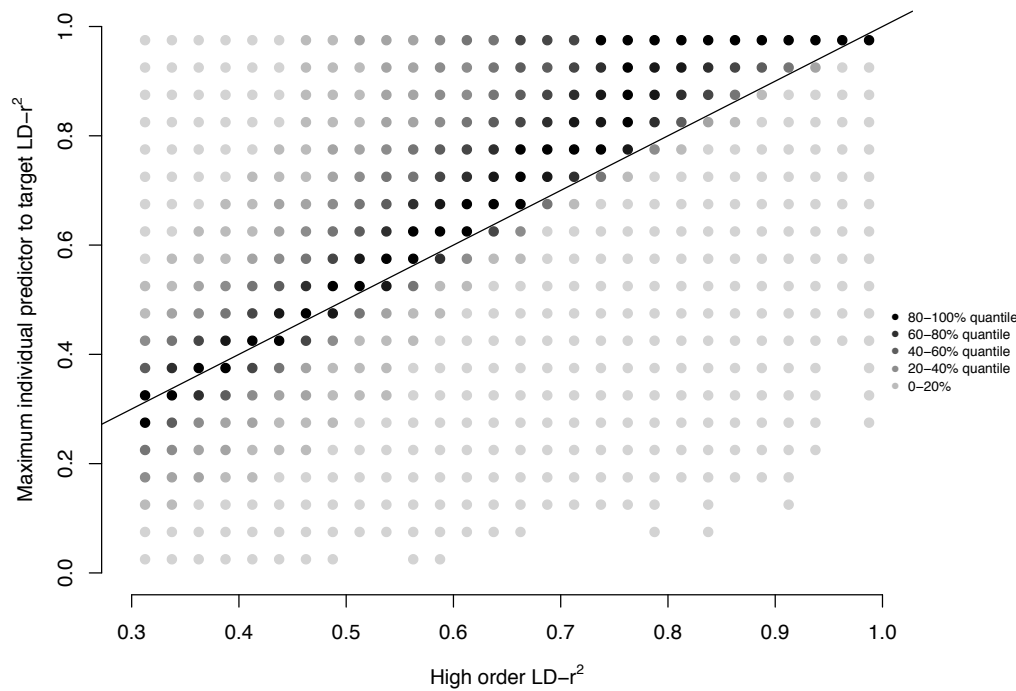


328

329 **Figure 2.** Illustration of how the prevalence of high order LD- r^2 to the targets in a 6Mb window on *A.*
330 *thaliana* chromosome 2 (8 – 14Mb) depends on distance of the predictors from the target. The color
331 gradient illustrates the proportion of predictor pairs that reach a particular LD- r^2 (x-axis) depending
332 on the distance between the nearest predictor and the target (y-axis). Results are presented for
333 populations with $n = 100$ (A) and $n = 728$ (B) individuals.

334

335 Strong high-order LD- r^2 between a predictor pair and a target is mostly observed
336 when at least one of the predictors is in strong individual LD- r^2 with the target.
337 However, as illustrated in Figure 3, many cases also exist where the high order LD- r^2
338 is strong while the LD- r^2 to the individual predictors is weak.



339

340 **Figure 3.** Strong second order LD- r^2 exists also when the individual predictor to target LD- r^2 is weak.

341 The intensity of each dot illustrates the number of cases with a particular high order LD- r^2 / maximum

342 individual predictor to target LD- r^2 combination. Dots below the line are cases where the high order

343 LD- r^2 stronger than any individual predictor to target LD- r^2 ($n=728$).

344

345 *Estimating the genome wide prevalence of strong high order linkage disequilibrium*

346 Figure 2 illustrates that high-order LD- r^2 exists where one or both predictors are

347 located close to the target as well as when one or both predictors are located further

348 away in the evaluated 6Mb window. The genome-wide prevalence of high order LD-

349 r^2 for the three different classes of predictor pairs, cis-cis/cis-trans/trans-trans (as

350 defined above) were next explored in three pairs of distant 2Mb windows in the

351 genome (Table 1) to provide data to estimate their genome-wide prevalence. Here,

352 only cases when individual predictors in the windows had lower individual LD- r^2 than

353 0.6 to the targets were considered.

354

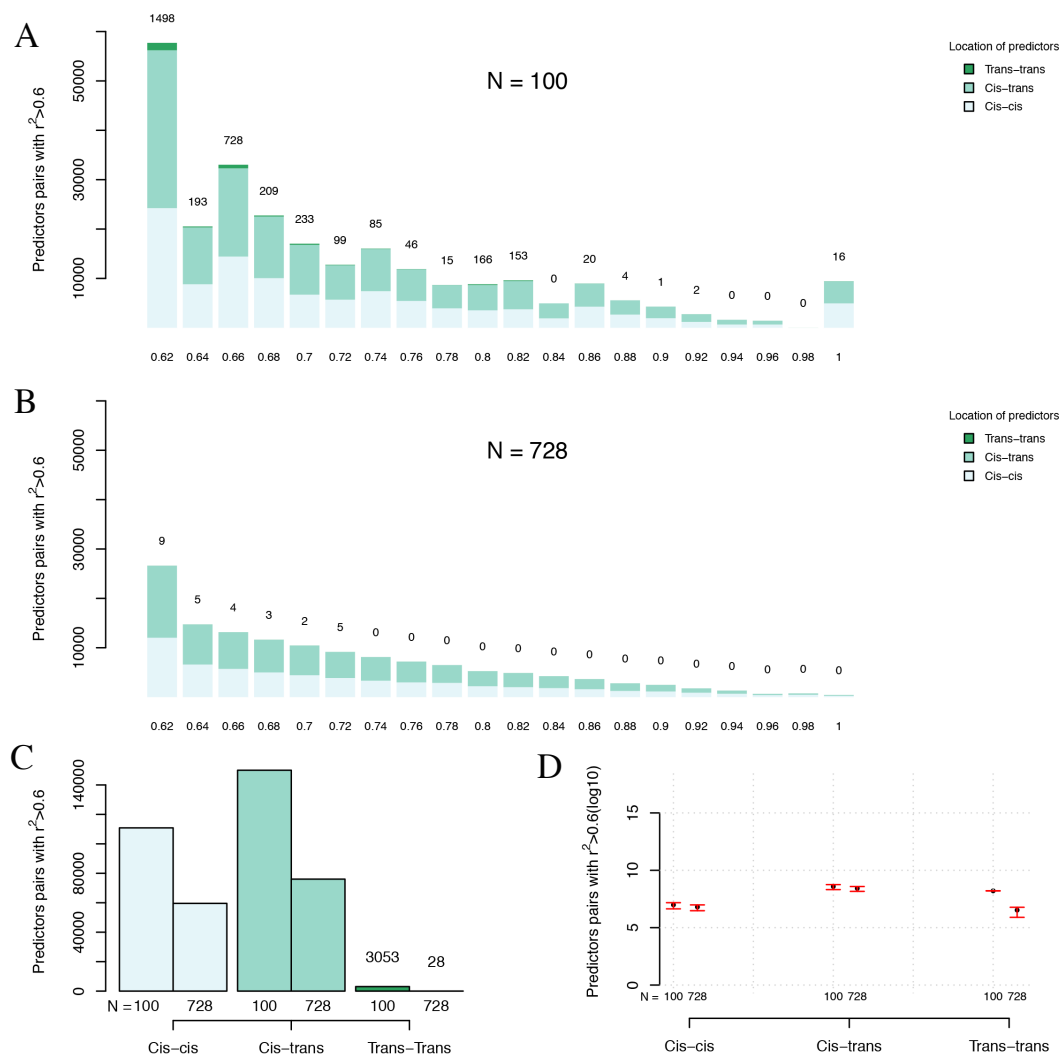
355 Overall, the fraction of predictor pairs that display higher second-order LD (LD- $r^2 >$

356 0.6) is low. In the smaller population ($n = 100$), less than 1 out of 10^6 evaluated

357 predictor pairs and in the larger population ($n = 768$) less than 1 out of 10^7 (Table S1).

358 However, since the total number of evaluated pairs was very large (around 10^{11}),

359 many cases were still detected. Regardless of population size, cis-cis and cis-trans
 360 pairs dominated (42/44% for $n = 100$, and 56/58% for $n = 728$; Figure 4A-C; Table
 361 S1). Trans-trans pairs existed, but were much less common ($\sim 1\%$ for $n = 100$, $<0.01\%$
 362 for $n = 728$, respectively, Figure 4A-C; Table S1). When extrapolating these results to
 363 a genome wide scale, this picture, however, changes dramatically (Figure 4D). Trans-
 364 trans and cis-trans predictor pairs are now much more common than cis-cis pairs due
 365 to their much higher genome-wide prevalence (35/18-fold for $n = 100$ and 35/0.3 for
 366 $n = 728$ more common; Figure 4D, Table S1). This result illustrates that it is a
 367 considerable risk to disregard high-order LD as a possible explanation for statistical
 368 epistatic interactions even at larger sample-sizes.
 369



370
 371 **Figure 4.** Number of predictor pairs of different classes in strong high order LD- r^2 (>0.6) to targets
 372 detected in the evaluated windows and estimated genome wide. The distribution of LD- r^2 values > 0.6

373 for the cis-cis, cis-trans,trans-trans predictor pairs for (A; $n = 100$) and (B; $n = 728$) The total number
374 of predictor pairs with high order LD- r^2 above 0.6 in the three classes are summarized in (C) and used
375 to estimate the total expected number of predictor pairs in the entire genome (D; error bars show the
376 estimation error estimated from the results obtained for the three window (Materials and Methods).

377

378 *Linking high order LD and statistical epistasis in a two locus epistatic association*
379 *analysis in A. thaliana*

380 A publicly available dataset including 340 *Arabidopsis thaliana* accessions were used
381 for a genome wide association analysis for leaf molybdenum concentration This
382 dataset was earlier used by (Forsberg *et al.* 2015) to dissect a locus with a highly
383 significant variance heterogeneity association for leaf molybdenum concentration
384 (Shen *et al.* 2012) to the contributions of four independent associations in an extended
385 LD block on chromosome 2. Several of these associations were found to structural
386 variants that were poorly tagged by the SNP markers (Forsberg *et al.* 2015). Our
387 pairwise genome wide scans for pairs of epistatic loci identified 396 significant SNP
388 pairs. For 290 pairs both markers were located in the narrow region on chromosome 2
389 that was earlier dissected in detail (Forsberg *et al.* 2015). All these are examples of
390 cis-cis predictor pairs. The remaining 106 pairs contained one predictor in the
391 chromosome 2 region and another one elsewhere in the genome, being examples of
392 cis-trans predictor pairs.

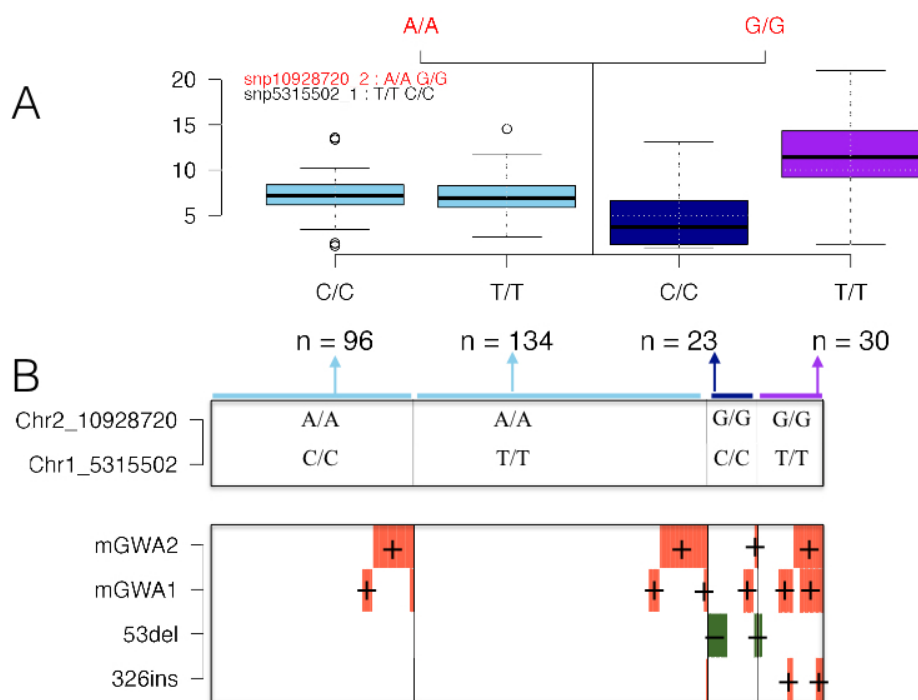
393

394 The strongest pairwise epistasis was detected for a cis-trans predictor pair (Figure 5A).
395 The accessions with the AA genotype at the predictor located in trans to the
396 chromosome 2 region (chromosome 1:5,315,502 bp) all have an intermediate
397 molybdenum level in the leaf (Figure 5A). The accessions with the GG allele at the
398 trans predictor have different levels of molybdenum in their leaves depending on
399 whether they carry the CC or TT genotype at the cis predictor in on chromosome 2
400 (10,928,720 bp). These differences explain the significant statistical epistasis detected
401 when fitting the two-locus epistatic model [1] to this data.

402

403 This statistical interaction could be due to a true genetic interaction. An alternative
404 explanation is however presented in Figure 5 There, the overlap between the two
405 locus genotypes for the cis-trans predictor pair (Figure 5A) and the alleles at the four

406 loci earlier reported to be associated with leaf molybdenum levels in this region
 407 (Forsberg *et al.* 2015) are illustrated. The multi-locus genotypes of the predictor pair
 408 tags different combinations of minor alleles at the four loci that were found to either
 409 increase (mGWA1, mGWA2, 326ins) or decrease (53del) leaf molybdenum levels in
 410 the accessions (Forsberg *et al.* 2015). The statistical epistatic interaction was detected
 411 due to the difference in molybdenum levels between accessions carrying the GGCC
 412 genotype (low molybdenum) and GGTT (high molybdenum). Figure 5B shows that
 413 the accessions with the GGCC genotype have the lowest frequency of the
 414 molybdenum increasing allele mGWA2 and the highest frequency of the molybdenum
 415 decreasing allele 53del. The accessions with the GGTT genotype instead have the
 416 highest frequencies of the molybdenum increasing alleles at mGWA2, mGWA1 and
 417 326ins. The genotypes AACC and AATT, with intermediate molybdenum levels, both
 418 have intermediate frequencies of the mGWA1 and mGWA1 increasing alleles and
 419 lack the 53del and 326ins alleles. A more parsimonious interpretation of these results
 420 is thus that the statistical epistasis at the predictor pair is due to the high order LD
 421 between them and the genotypes at the four loci located in the region on chromosome
 422 2.



423
 424 **Figure 5.** An illustration of how the high order LD between four polymorphisms affecting the level of
 425 molybdenum in the *A. thaliana* leaf (Forsberg *et al.* 2015), likely explains the significant statistical

426 *epistasis detected for a cis-trans predictor pair. (A) Boxplots illustrating the phenotypic distribution in*
427 *the four genotype classes defined by the cis-trans predictor pair with the strongest significant epistatic*
428 *interaction to the level of molybdenum in the *A. thaliana* leaf. (B) Illustration of the connection*
429 *between the two-locus genotypes of the predictor pair and the minor alleles at the four linked loci*
430 *associated with this trait on chromosome 2 (Forsberg et al. 2015). The top box in (B) illustrates the*
431 *two-locus genotype for the predictor pair, with the width of each sub-box indicating the number of*
432 *individuals in each genotype class in the population. In the bottom box in (B), each individual is*
433 *represented as a column, where green (molybdenum decreasing) and orange (molybdenum increasing)*
434 *colors indicates that the individual carry the minor alleles at the four loci identified in (Forsberg et al.*
435 *2015). *mGWA1* and *mGWA2* are SNP markers associated with the trait and *53del* and *326* are*
436 *structural polymorphisms (Forsberg et al. 2015).*

437

438

DISCUSSION

439 High order linkage disequilibrium between combinations of genotyped markers, and
440 unobserved functional polymorphisms, can result in significant statistical epistasis in
441 genome wide association analyses. This was earlier illustrated empirically for linked
442 pairs of genotyped predictor SNPs and ungenotyped target polymorphisms in humans
443 by Wood *et al.* (Wood *et al.* 2014). Here, we present a new example from *A. thaliana*
444 where significant statistical epistasis between pairs of predictors is due to the effects
445 at a single loci and that only one of the statistically interacting loci was located near
446 the target. By exploring the prevalence of second order LD in the genome of the
447 1001-genomes *A. thaliana* collection, we find that although the total amount of high
448 order linkage disequilibrium decreases with increasing population sizes, it is still
449 highly prevalent both within and across chromosomes even in relatively large
450 populations (n = 728). It is was found to be most common when one predictor is in
451 high LD to (and located physically near) the target, but many cases exist where the
452 LD to the individual predictors is very weak but the high order LD is strong. The
453 choice of target and predictor SNPs in this study is arbitrary and therefore it is
454 difficult to assess how representative they are for the prevalence of high order LD in
455 other populations. However, they do suggest that strong high order LD can be
456 prevalent also in larger populations, indicating that statistical epistasis observed in
457 studies based on reduced representation genotyping (such as SNP-chips) need to be
458 interpreted with caution.

459

460 The most prevalent type of high order LD on a genome wide basis is that of cis-trans
461 predictor pairs, but also cis-cis pairs are common regardless of population size. The
462 prevalence of trans-trans pairs is high in smaller populations but decreases rapidly as
463 the population size increases. A possible biological explanation for the observation
464 that cis-cis and cis-trans high order LD pairs is relatively prevalent also at larger
465 population sizes would be that the number of, and variation in, the trans located
466 predictors is sufficiently large on a genome-wide basis to complement any
467 imperfection in the tagging of the functional polymorphism by the cis located
468 predictor. Whereas trans-trans high order LD will always result in falsely associated
469 loci, cis-trans and cis-cis high order LD presents an opportunity to identify true
470 functional loci for the trait. The problem in a real data analysis is that statistical
471 epistasis between a pair of predictors can emerge from true interactions or high order
472 LD within and across chromosomes. However, as the sample sizes increase the risk of
473 detecting pairs of predictors where none is located close to the true functional
474 polymorphism decreases. Before concluding that the detected association is due to
475 two interacting loci, further analyses of the associated pair are however recommended.
476

477 Whole-genome sequencing provides unprecedented opportunities to genotype most
478 segregating single nucleotide polymorphisms in the genome. Despite this, it is
479 unlikely that these will be able to tag all functional polymorphisms, such as larger
480 structural variants or multi-allelic functional loci due to tandem repeats. Hence, even
481 though the scenario of reduced representation genotyping with SNP-chips or similar
482 will soon be a technology of the past, association analyses will still be challenged by
483 the need to tag hidden polymorphisms with imperfect markers as illustrated in our
484 analyses of the complex locus affecting molybdenum levels in the *A. thaliana* leaf. In
485 fact, it is not unlikely that the problem with high order LD between SNP predictors
486 and hidden, complex functional loci will remain a major challenge in the future as the
487 increased number of markers generated by sequencing also increases the chance of
488 finding combinations of cis-cis or cis-trans predictors that tag these functional
489 polymorphisms better than any single marker. To evaluate the extent of this problem
490 one will, however, need a more comprehensive dataset than the one studied here
491 including a more complete scoring of all types of non-SNP polymorphisms in the

492 genome with potential effect on traits of interest.

493

494 The prevalence of high order LD is likely to be more of a concern in populations of
495 inbred or haploid individuals. These include, for example, inbred lines derived from
496 bi- and multi-parental crosses of plants and animals, as well as populations of wild
497 collected inbred plants (Churchill *et al.* 2004; Valdar *et al.* 2006; Kover *et al.* 2009;
498 Cao *et al.* 2011; Mackay *et al.* 2012). As heterozygotes are not present in these
499 populations, the number of multi locus genotype classes is smaller than in outbred
500 populations, making them attractive for studies of genetic interactions. As a common
501 approach to detect interactions in such populations is to identify pairs of loci
502 displaying significant statistical epistasis, such results need to be interpreted with
503 caution, as the analyzed populations are generally small. If one, or more, of the
504 functional polymorphisms in the genome are unknown and poorly tagged by the
505 genotyped markers, there is a risk that statistical interactions arise from high-order LD
506 between the genotyped markers and the hidden functional polymorphisms. Hence,
507 even though these populations increase the power to map loci displaying statistical
508 epistasis, there is also a risk of falsely concluding that the underlying genetic
509 architecture involves genetic interactions.

510

511

CONCLUSIONS

512 Statistical epistasis detected in genome wide association analyses can result from high
513 order LD between genotyped markers and unobserved functional polymorphisms.
514 This study revealed extensive and strong genome wide high order LD between pairs
515 of markers on a high density 250k SNP-chip and individual markers revealed by
516 whole genome sequencing in the *A. thaliana* 1001-genomes collection. The high
517 prevalence of strong high order LD in this dataset suggests that epistatic variance
518 detected between pairs of markers in association analyses, especially in small inbred
519 populations genotyped for reduced representation sets of markers, need to be
520 interpreted with caution. An empirical example is presented where a pair of markers
521 with significant statistical epistasis in a genome wide association analysis is in high
522 order LD with a complex multi allelic locus with large effects on the analyzed trait.
523 As complex functional loci such as this are unlikely to be captured by individual bi-
524 allelic SNP markers, even if millions of them are scored by whole genome sequencing,

525 it is important to evaluate also other explanations of statistical epistasis than
526 underlying genetic interactions in particular when small populations of inbred
527 individuals are studied.

528

529

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532

533

AUTHOR CONTRIBUTIONS

534 ÖC and SF initiated the study. ÖC, SF and YZ designed the project and the statistical
535 analyses; SF and YZ wrote analysis scripts and performed the data analyses. ÖC and
536 YZ summarized the results and wrote the initial version of the manuscript. All authors
537 contributed to the writing of the final version of the manuscript.

538

539

DISCLOSURE DECLARATION

540 The authors declare no competing interest.

541

542

SUPPLEMENTARY MATERIAL

543 Supplementary material is provided in Supplementary Figure 1 and Supplementary
544 Table 1.

545

546

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