Comparison of Prognostic Accuracy of the quick Sepsis-Related Organ Failure Assessment between Short- & Long-term Mortality in Patients Presenting Outside of The Intensive Care Unit – A Systematic Review & Metaanalysis

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ABSTRACT

Objective

In year 2016, quick Sepsis-Related Organ Failure Assessment (qSOFA) was introduced as a better sepsis screening tool compared to systemic inflammatory response syndrome (SIRS). The purpose of this systematic review and meta-analysis is to evaluate the ability of the qSOFA in predicting short- and long-term mortality among patients outside the intensive care unit setting.

Method

Studies reporting on the qSOFA and mortality from MEDLINE (published between 1946 and 15^{th} December 2017) and SCOPUS (published before 15^{th} December 2017). Hand-checking of the references of relevant articles was carried out. Studies were included if they involved inclusion of patients presenting to the ED; usage of Sepsis-3 definition with suspected infection; usage of qSOFA score for mortality prognostication; and written in English. Study details, patient demographics, qSOFA scores, short-term (<30 days) and long-term (\geq 30 days) mortality were extracted. Two reviewers conducted all reviews and data extraction independently.

Results and Discussion

A total of 39 studies met the selection criteria for full text review and only 36 studies were included. Data on qSOFA scores and mortality rate were extracted from 36 studies from 15 countries. The pooled odds ratio was 5.5 and 4.7 for short-term and long-term mortality respectively. The overall pooled sensitivity and specificity for the qSOFA was 48% and 85% for short-term mortality and 32% and 92% for long-term mortality, respectively. Studies reporting on short-term mortality were heterogeneous (Tau=24%, I²=94%, P<0.001), while long-term mortality studies were homogenous (Tau=0%, I²<0.001, P=0.52). The factors contributing to heterogeneity may be wide age group, various clinical settings, variation in the timing of qSOFA scoring, and broad range of clinical diagnosis and criteria. There was no publication bias for short-term mortality analysis.

Conclusion

qSOFA score showed a poor sensitivity but moderate specificity for both short and long-term mortality prediction in patients with suspected infection. qSOFA score may be a cost-effective tool for sepsis prognostication outside of the ICU setting.

Key Words: emergency service hospital; mortality; outside the ICU; prognosis; qSOFA; sepsis

INTRODUCTION

Sepsis is a syndrome of uncertain pathophysiology that is recognized by a group of clinical signs and symptoms in patients with suspected infection ¹. Sepsis is a significant cause of mortality worldwide. In the last decade, an estimated 31.5 million sepsis patients have been treated globally per year, including 5.3 million deaths due to sepsis ². The diagnosis of sepsis is challenging, as a reliable test for its early confirmation is not available. Given the morbidity and mortality of sepsis, the ability to perform risk stratification in the early phase of a patient's illness may help physicians to effectively manage and improve their outcome.

The Third International Consensus Definitions for Sepsis and Septic Shock (Sepsis-3) defined sepsis as life-threatening organ dysfunction caused by a dysregulated host response to infection¹, known as severe sepsis in previous definitions ³. The formerly used Systemic Inflammatory Response Syndrome (SIRS) criteria for early identification of sepsis was considered impractical and inefficient ⁴. Subsequently, Sepsis-3 proposed using the quick Sepsis-Related Organ Failure Assessment (qSOFA) as a new risk stratification tool to distinguish patients who are likely to have sepsis under the new definition. The qSOFA criteria recommend screening patients for three clinical signs i.e. tachypnoea, altered mental status, and hypotension, rather than requiring blood tests. Ongoing efforts have been directed toward examining the ability of qSOFA to predict poor outcomes in patients with infection ⁵. However, several studies have suggested that qSOFA criteria lack accuracy for predicting mortality in patients outside the intensive care unit compared to other early scoring systems ^{6,7}. Since the qSOFA is a new scoring system, the clinical practicality of these criteria as a screening tool for infectious patients has not been fully evaluated.

The intention of this systematic review and meta-analysis was to evaluate qSOFA as a mortality predictor in patients presenting outside of the intensive care unit (ICU). We hypothesized that qSOFA can predict short- and long-term mortality in sepsis patients. The prognostic accuracy of qSOFA score for both short- (\leq 30 days) and long-term (>30 days) mortality was analysed.

METHODS

Study Eligibility Criteria and Search Strategy

A systematic review and meta-analysis of the literature was conducted to identify relevant studies regarding the role of the qSOFA in mortality prognostication among patients with suspected infection presented outside of the intensive care unit after obtaining consent from UKM Research Ethical Committee (UKM PPI/111/8/JEP-2017-769). We used MEDLINE via Ovid Medline to conduct a comprehensive search of health science journals (published between 1946 and 15 December 2017) and SCOPUS (published before 15 December 2017); Hand-checking of the references of relevant articles was carried out. The search team comprised of three clinicians, a statistician and a scientist. The search strategy involved a combination of the following 2 sets of keywords 1) 'quick sequential organ failure assessment' OR 'quick SOFA' OR 'qSOFA' OR 'quick sepsis related organ failure assessment' and 2) 'mortalit*'. This meta-analysis was registered in PROSPERO (CRD42017079364). The search strategies were shown in S1 Table.

Identification and Selection of Studies

Study selection was performed based on their titles or abstracts, and only studies which appeared to fulfil the eligibility criteria were selected for full-text review. To be included, studies must fulfil the following criteria: inclusion of patients presenting to the ED; usage of Sepsis-3 definition with suspected infection; usage of qSOFA score for mortality prognostication; and written in English. Papers were excluded if they were: related to review articles; articles without complete texts; or animal studies.

Data Extraction and Study Appraisal

The selection of papers inclusion in this review was completed in four phases. First, an initial search of the selected databases was performed using the pre-specified keywords to identify relevant keywords and index terms. Second, a thorough search was conducted in which papers that failed to meet the inclusion criteria based solely on their titles and abstracts were excluded. In the third phase, the remaining papers from the second phase were extensively reviewed, and papers that did not meet our inclusion criteria were excluded. Finally, all relevant data from the included papers was subjected to meta-analysis to determine conclusions regarding the proposed hypothesis.

After the initial screening of titles and abstracts by two independent reviewers, who are clinicians, incomplete articles were removed. The remaining papers were screened again by the two reviewers. To minimize errors, both reviewers were trained under a consensus standard and then practiced using several articles for calibration. Any discrepancies were resolved through discussion with a third reviewer who is an Emergency Physician. We applied the QUADAS-2 (Quality Assessment of Diagnostic Accuracy Studies) criteria to assess the quality of all selected articles. The risk of bias of each included study was summarized (S2 and S3 Tables). Data extraction was conducted independently in a standardized manner with a data collection form. Study data including author, publication year, type of study conducted, brief description of the study population/sample and methods used in the study, index/reference time interval, and mortality outcome were extracted from the full text of each article and summarized in detail (S4 and S5 Tables). In-hospital mortality was categorized as short-term mortality. This analysis was reported according to Transparent Reporting of Systematic Reviews and Meta-Analyses (PRISMA) guideline. A flow diagram of study identification and articles selection for the meta-analyses can be found in Figure 1.

Statistical Analysis

The primary statistical analysis was performed using the Review Manager 5 (Version 5.3.5) software manufactured by Cochrane Community and the Comprehensive Meta-Analysis Software (CMA, Version 3) manufactured by Biostat. Based on this model, the pooled sensitivity, specificity and the odds ratio (OR) with the 95% CI were determined. Random effects model was used to report short- and long-term mortality individually with estimates of sensitivity, specificity and ORs. The I² statistic was calculated to determine the proportion of between-study variation caused by heterogeneity, with suggested thresholds for low (25%-49%), moderate (50%-74%), and high (75%) values. The publication bias of the included studies was assessed using effective sample-size funnel plot (OR values vs sample size of each study), Begg-adjusted rank correlation tests and the Egger regression asymmetry test for small study effects.

Data availability

The authors declare that all data supporting the findings of this study are available within the paper and its supplementary information files.

RESULTS

The search identified relevant studies from MEDLINE via Ovid Medline (1946 to 15 December 2017) and SCOPUS database (through 15 December 2017). The numbers of relevant records identified in MEDLINE and SCOPUS were 42 and 80, respectively, for a total of 122 references retrieved from the electronic database. Forty-two records were identified as duplicates and were removed from our selection. Of the 80 remaining references, 41 were excluded based on titles and abstracts: 22 did not meet the primary objective of our review, two did not meet our inclusion criteria, two studies were published in languages other than English, and 15 were other articles including review, consensus, perspective, commentary and editorial papers. The full texts of 39 studies were successfully retrieved. Three papers were excluded due to incomplete data in study (S6 and S7 Tables). The authors of the three studies fulfilled the inclusion criteria and were included. The characteristics of the included studies ⁵⁻⁴⁰ are summarized in Table 1.

The prognostic accuracy of the qSOFA was evaluated in different countries, with most studies conducted in the United States of America and Europe, followed by Asia countries, Africa, New Zealand and Australia. Among all the articles included, six of them were conducted in Asia countries. The cut-off values of the Glasgow Coma Scale (GCS) used for altered mentation in the qSOFA included GCS less than 15, 14 and 13 respectively, and the remaining studies only stated altered mentation. A total of 35 studies reported on the prognostic accuracy of the qSOFA for short-term mortality⁵⁻³⁹, while only 3 articles reported on long-term mortality ^{14,31,40}.

Table 1. Summary of Characteristics of Studies Included

Source	No. of Partici pants	Me an Age	Men, No. (%)	Main Inclusion Criteria	Outcome
hort-term mor	tality	, y			
April (8),	214	68	126	ED patients admitted to any ICU with suspected or proven	In-hospital
2017	211	00	(59%)	infection	mortality
Askim (9),	1535	61 ^{<i>a</i>}	813	New onset of suspected or confirmed infection according to	30-day mortality
2017			(53%)	the ESS47	
Brabrand	3824	65 ^a	2426	Patients presenting or discharged with suspected infection	In-hospital
(10), 2016			(63%)		mortality and/or
					ICU stay >3day
Chen (11),	1641	73 ^a	968	Patients with CAP or healthcare-associated pneumonia	28-day mortality
2016			(59%)		
Churpek (6),	30677	58	14561	Patients with suspicion of infection in wards or ED	28-day mortality
2017			(47%)		
Churpek (12),	53849	57	24719	Patients meeting suspicion criteria in ED or wards	In-hospital
2017			(46%)		mortality
DeGroot (13),	2280	61	1315	ED patients with suspected infection	In-hospital
2017	2502	NT 4	(58%)		mortality
Donnelly	2593	NA	NA	Admitted patients who meet SIRS criteria, SOFA and qSOFA	28-day mortality
(14), 2017 Finkalaztain	150	64 ^{<i>a</i>}	02	criteria	In hoonital
Finkelsztein (15), 2017	152	04	83 (55%)	Patients with suspicion of infection admitted to the medical ICU from emergency department or hospital wards	In-hospital mortality
(15), 2017 Forward (16),	161	70	(55%) 89	Non-ICU inpatients who triggered the hospital SK pathway	In-hospital
2017	101	70	89 (55%)	with acute deterioration and suspected or proven infection	mortality
Freund (17),	879	67 ^a	(<i>35</i> %) 465	Patients admitted to ED with clinical suspicion of infection	In-hospital
2017	0.7	57	(53%)	a water to an and to an and the chine of suspector of infection	mortality
Giamarellos	3436	NA	NA	Patients with signs of infection	28-day mortality
(18), 2017	-				j - · · · · · · · · · · · · · · · · · ·
Gonzalez	1071	84	544	Patients ≥75 years old clinically diagnosed with acute	30-day mortality
(19), 2017			(51%)	infection in ED	
Haydar (20),	199	71 ^a	109	ED patients treated for suspected sepsis	In-hospital
2017			(55%)		mortality
Henning (21),	7637	60	3799	ED patients admitted to the hospital with an infection-related	In-hospital
2016			(50%)	diagnosis	mortality
Huson (22),	329	34 ^{<i>a</i>}	125	Patients with suspected infection with ≥ 2 SIRS criteria	In-hospital
2017	450	254	(38%)		mortality
Huson (23),	458	35 ^a	243	patients admitted to the adult medical ward with suspected	In-hospital
2017	1205	c = a	(53%)	infection	mortality
Hwang (24)	1395	65 ^{<i>a</i>}	787 (56%)	Patients who received a diagnosis of severe sepsis or septic	28-day mortality
2017 Kim (25)	615	54	(56%) 204	shock during ED stay Patients with fever and chemotherapy-induced neutropenia	28-day mortality
Kim (25), 2017	015	54	204 (33%)	r aucins with rever and chemotherapy-mouced neuropenia	∠o-uay monality
Kim (26),	125	76	(33%) 78	Patients admitted to ED with discharge diagnosis of CAP	28-day mortality
2017	125	70	(62%)	r adonts admitted to ED with discharge diagnosis of CAI	20 day morality
Kolditz (27)	9327	63 ^{<i>a</i>}	5249	Patients with CAP	30-day mortality
2016			(56%)		.,
LeGuen (28),	182	72 ^a	88	Patients reviewed by the RRT	30-day mortality
2017			(48%)		•
Moskowitz	24164	64	12299	Patients with suspected infection presented to ED	In-hospital
(29), 2017			(51%)		mortality
Patidar (30),	124	57	NA	Cirrhotic patients hospitalized non-electively for infectious	30-day mortality
2017	102	<u> </u>	100	etiologies	2 0 1
Quinten (31),	193	60	108	Non-trauma patients in ED with suspected infection or sepsis	28-day mortality
2017	6074	~ ~ ~	(56%)	Detions with all is a line with a CAD	20 1 11:
Ranzani (32),	6874	66	4259	Patients with clinical diagnosis of CAP	30-day mortality
2017 Dothmon	2026	NT A	(62%)	Detion to admitted to be an ital with severi-	In head'tel
Rothman (33), 2017	3926	NA	NA	Patients admitted to hospital with sepsis	In-hospital
UND. ZULT				Patients with suspected infection	mortality In-hospital
Seymour (5),	66522	60	27446	Patients with suspected intection	In_hognital

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Shetty (34), 2017	12555	50 ^{<i>a</i>}	6585 (52%)	Patients with suspected infection, suspected or confirmed sepsis	Mortality and/or prolonged ICU stay ≥72hours
Singer (35), 2016	22530	54	10589 (47%)	ED patients whom qSOFA score could be calculated according to simultaneous reporting of vital signs and a MEWS score	In-hospital mortality
Szakmany (36), 2017	380	74 ^a	180 (47%)	Patients with high degree of clinical suspicion of infection	30-day mortality
Tusgul (37), 2017	886	80	462 (52%)	Patients with suspected infection without alternative diagnosis, or microbiologically proven infection found in the ED workup	In-hospital mortality
Umemura (38), 2017	387	74 ^a	232 (60%)	ED patients admitted to ICU with diagnosis of severe sepsis	In-hospital mortality
Wang (39), 2016	477	73 ^a	295 (62%)	Patients treated at ED with clinically diagnosed infection	28-day mortality
Williams (7), 2017	8871	49 ^{<i>a</i>}	4453 (50%)	ED patients admitted with a diagnosis indicating presumed or potential infection	30-day mortality
Long-term mor	tality				
Donnelly (14), 2017	2593	NA	NA	Admitted patients who meet the SIRS criteria, SOFA and qSOFA criteria	1-year mortality
Quinten (31), 2017	193	60	108 (56%)	Non-trauma patients in ED with suspected infection or sepsis	6-month mortality
Rannikko (40), 2017	497	68 ^{<i>a</i>}	262 (53%)	Adult patients admitted to the ED who had blood culture- positive sepsis	90-day mortality

Abbreviations: ED, emergency department; ICU, intensive care unit; ESS47, Emergency Symptoms and Signs algorithm for infection; CAP, community acquired pneumonia; NA, not available; SIRS, systemic inflammatory response syndrome; SOFA, Sequential organ failure assessment; qSOFA, quick sequential organ failure assessment; SK, "Sepsis Kills"; RRT, Rapid Response Team; MEWS, Modified Early Warning System.

^a median

The qSOFA and short-term mortality

In this meta-analysis, 35 studies with 269,544 patients reported on the prognostic accuracy of the qSOFA and short-term mortality, including 27 retrospective studies ^{5-8,10,11,13-16,18,20,22,24-27,29,32-35,37-40}, while 8 studies were prospective studies ^{9,12,17,19,21,23,28,30,36}. Due to the heterogeneity of the inclusion criteria, a random-effects model was used to calculate the pooled sensitivity and specificity of the included studies. The forest plot for the sensitivity and specificity of the qSOFA predicting short-term mortality is shown in Figure 2. The pooled sensitivity was 48% and the specificity was 86%. The pooled odds ratio (OR) was 5.5 (95% CI: 5.32-5.72; Q =566.67; degree of freedom, df = 34; p<0.01), indicating that an elevated qSOFA score was associated with increased short-term mortality. The forest plot for the odds ratio is shown in Figure 3. We detected significant heterogeneity according to the heterogeneity tests (Tau = 24%; I² = 94%; p<0.001). Publication bias was not detected as shown in the funnel plot (S1 Figure). Egger's regression and Begg's test revealed no statistical significance with p=0.84 (2-tailed) and p=0.46 respectively, indicating no publication bias (S8 Table).

The qSOFA and long-term mortality

Only three studies with a total of 3,076 patients reported on the prognostic accuracy of the qSOFA and long-term mortality. Among these studies, two were retrospective ^{14,40} and one was prospective study ³¹. The forest plot for the sensitivity and specificity of the qSOFA for predicting long-term mortality is shown in Figure 2. The pooled sensitivity and specificity were calculated using a random-effects model, which yielded a pooled sensitivity of 32% and a pooled specificity of 92%. The three studies reported distinct mortality intervals, including 90-day mortality (sensitivity 56%, specificity 79%) ⁴⁰, 6-month mortality (sensitivity 33%, specificity 85%) ³¹ and 12-month mortality (sensitivity 21%, specificity 95%) ¹⁴. The forest plot for the odds ratio is shown in Figure 3. The pooled OR was 4.7 (95% CI: 3.54-6.12; Q = 1.31; df = 2), and the studies were homogenous (Tau = 0%; I²<.001; p=0.52). However, publication bias was not assessed due to the small number of studies included in the long-term mortality analysis.

DISCUSSION

This systematic review and meta-analysis revealed that most of the included studies suggested that a qSOFA score of ≥ 2 was able to predict short and long-term mortality. A total of 35 studies were reviewed, and the quality of the studies varied. Most of the studies had good quality according to QUADAS-2. Seven studies showed evidence of bias (S2 and S3 Tables). These seven studies had excluded many missing data and missing data analysis were not mentioned.

Our analysis revealed that qSOFA score exhibited fair sensitivity and specificity in predicting mortality. The pooled specificity of qSOFA in this study was higher compared to the SIRS (66%)⁴¹. According to our analysis, qSOFA can predict sepsis mortality, with the odds of 5.5 for the short-term mortality and 4.7 for the long-term mortality.

All 36 papers reporting on short-term mortality showed clinical, methodological and statistical heterogeneity. Factors that may have contributed to the high heterogeneity include the mean age (ranging from 54 to 84 years old), various clinical settings, variation in the timing of qSOFA scoring, and broad range of clinical diagnosis and criteria. This heterogeneity contributed to a lower pooled sensitivity of the qSOFA that may not represent the actual accuracy of the qSOFA. However, this finding was expected as the study population were diverse and multiple confounding factors were present. All studies showed positive direction in the forest plot reflecting a high pooled odds ratio. The funnel plot revealed no publication bias for the studies investigating qSOFA in predicting short-term mortality. Recently, three new publications reported on qSOFA short-term mortality prediction with similar findings⁴²⁻⁴⁴. Nevertheless, these studies did not perform further analysis on qSOFA long-term mortality prediction nor compared its prognostic accuracy with short-term mortality.

We found three studies which reported on qSOFA prognostication for long-term mortality. These studies showed clinical and methodological heterogeneity, but they were statistically homogenous. The performance of the qSOFA in long-term mortality prediction was more specific but less sensitive compared to its performance in short-term mortality. In studies reporting on longterm mortality, we found that the qSOFA was more specific but less sensitive. Since age is a confounding factor for mortality, we postulate that aging may contribute to this observation. In addition, longer mortality periods may result in the inclusion of deaths from extraneous factors. Further studies will be important to provide insight into this intriguing finding.

The qSOFA score consists of 3 criteria, and ranges from 0-3 points, with 1 point each for systolic hypotension (<100 mmHg), tachypnoea (>22/min), and altered mentation⁵. The three criteria of the qSOFA are much related to the pathophysiology of sepsis leading to organ injury. In sepsis, profound changes occur in the endothelium, causing widespread tissue oedema due to increased leukocyte adhesion, vasodilation, a shift to a pro-coagulant state and loss of barrier function, resulting in low blood pressure. Likewise, endothelial changes in severe sepsis causes increased permeability of lung capillaries, leading to accumulation of exudate fluid in the interstitial spaces of the lung, which floods into the alveoli and induces alveolar epithelial barrier dysfunction. These changes lead to acute respiratory distress and tachypnoea due to perfusion-ventilation mismatch, arterial hypoxemia and reduced lung compliance. Furthermore, systemic endothelial dysfunction affects the blood-brain barrier. Inflammatory cytokines and cells entered the brain, causing perivascular oedema, leukoencephalopathy, oxidative stress, and global neurotransmitter alteration, resulting in altered mental status⁴⁵. These complex changes in sepsis contributes to the high specificity of the qSOFA in detecting sepsis.

Sepsis was redefined in 2016 and the qSOFA was introduced to replace the SIRS as a better criteria for sepsis. The advantage of the qSOFA is that it can be repeatedly performed over time without laboratory investigations, which can be time-consuming ²⁵. Since sepsis can deteriorate in a short period of time, a simple screening tool for early detection is warranted. The SIRS criteria introduced in previous sepsis definitions^{3,46} as a screening tool for sepsis was found to be overly sensitive relative to its specificity ⁶. A screening tool with high sensitivity and poor specificity can result in an excessive number of false positives, leading to unnecessary diagnostic or therapeutic procedures. Over-diagnosing patients poses a significant economic impact and further increases patients' medical burden. In addition to qSOFA scoring, several publications have suggested lactate level could be of valuable biomarker when added to the original qSOFA score and may improve its prognostic value^{21,29,34}. These studies provide insight into modification of the qSOFA which may improve its sensitivity and efficacy in detecting patients with sepsis. Efforts to modify the qSOFA could consider combining the present scoring criteria with other sepsis biomarkers such as C-reactive protein (CRP), procalcitonin (PCT), and serum secretory phospholipase A2-IIa (spla2-IIa) ⁴⁶⁻⁴⁹.

Although the qSOFA exhibited high specificity and low sensitivity in most of the studies included in our meta-analysis, seven papers showed contradictory results. The studies reported that qSOFA was highly sensitive but had poor specificity. On further analysis, four of the studies had sample populations comprised of patients who were directly admitted from the ED to the ICU ^{8,15,24,38}, and two other studies included high numbers of Human Immunodeficiency Virus carriers

^{22,23}. The remaining paper had a distinct study population including elderly and disabled patients, in whom assessment of altered mental status was regarded as challenging ²⁰. The population included in these studies were more specific and likely to present to the ED with greater illness severity. Due to the specificity of these study populations, patients in these studies tended to be screened as positive as reflected by the identification of more true-positive patients compared to the other studies' populations, resulting in heightened sensitivity of the qSOFA.

LIMITATIONS

In this meta-analysis, we successfully retrieved all full-texts and a standardized tool was used to examine the quality of the included papers. One limitation of our analysis was the inclusion of too few articles reporting on long-term mortality. Second, we discovered that the study populations were substantially diverse, as some included specific groups of patients with infection. However, all of the included patients fulfilled our inclusion criterion of patients with suspected infection. Since random sampling was not performed in most of the studies included, a sampling bias is likely. Some studies had combined outcomes of mortality and/or ICU admission, thus complicating precise categorization of outcomes^{10,34}. We classified in-hospital mortality as short-term mortality. Since in-hospital mortality may be longer than 30 days, this assumption may lead to a misclassification bias and mask the true predictive ability of the qSOFA. Most of the included studies were retrospective studies, posing a certain disadvantage as these studies relied on available medical records. Therefore, missing records or data may have influenced the results and the predictive accuracy of qSOFA in the current analysis. In addition, most of the studies were single-centered with variability across methods and study designs, which contributed to heterogeneity. Multiple confounders were likely to coexist, which may jeopardize the validity of these studies. Future research should consider prospective randomization in sampling method to minimize sampling bias. More studies exploring the qSOFA for long-term mortality prediction should be conducted in the near future.

CONCLUSION

This meta-analysis revealed that the qSOFA score had a poor sensitivity but moderate specificity for both short and long-term mortality prediction in patients with suspected infection. Compared to short-term mortality, the qSOFA had superior predictive ability for long-term mortality. Further research on modification of qSOFA may improve its sensitivity in detecting patients with sepsis for

prompt intervention. In general, qSOFA score may be a cost-effective tool for sepsis

prognostication outside of the ICU setting.

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Supplementary Information is available in the online version of the paper

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Table and Figure's Title and Legend

Table 1. Summary of Characteristics of Studies Included.

Figure 1. Identification and Selection of Articles for Meta-analysis. Flow chart shows process of article selection and exclusion throughout the study.

Figure 2. Sensitivity and Specificity of quick Sepsis-Related Organ Failure Assessment

(**qSOFA**) in Predicting Short-term and Long-term Mortality. Studies included into the metaanalysis and their corresponding sensitivity and specificity of quick Sepsis-Related Organ Failure Assessment (qSOFA) values in predicting short- and long-term mortality is shown using a forest plot.

Figure 3. Odds Ratio of quick Sepsis-Related Organ Failure Assessment (qSOFA) in

Predicting Short-term and Long-term Mortality. Odds of each study is shown in the forest plot. All studies found odds ratio of > 1 for quick Sepsis-Related Organ Failure Assessment (qSOFA) in predicting short- and long-term mortality.

Supplemental Digital Content Tables and Figure Legend

S1 Table. Table showing search strategy for both MEDLINE and SCOPUS including the keywords used and results from the database.

S2 Table. A summary of QUADAS-2 results for included studies from MEDLINE assessing the risk of bias and applicability concerns. Three studies from MEDLINE showed evidence of bias.

S3 Table. A summary of QUADAS-2 results for included studies from SCOPUS assessing the risk of bias and applicability concerns. Four studies from SCOPUS showed evidence of bias.

S4 Table. A detailed summary of type of study, subjects included, methodology and outcome of selected studies from MEDLINE

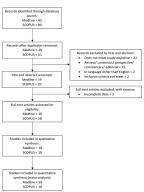
S5 Table. A detailed summary of type of study, subjects included, methodology and outcome of selected studies from SCOPUS.

S6 Table. Table showing list of articles excluded by title and abstract, including reason of exclusion. A total of 41 studies were excluded by title and abstract.

S7 Table. Table showing list of included and excluded full-text articles, and reason of exclusion. A total of 36 studies were included after retrieving full-text articles and three studies were excluded due to incomplete data.

S1 Figure. Funnel plot showing publication bias for short-term mortality.

S8 Table. Egger and Begg's test were done for small-study effects, and no publication bias was detected.



Study	19	FP.	FN	TN	Sensitivity (95N CI)	Specificity (95% O)	Senaltivity (35% O)	Specificity (\$5% C0
Short term mortality (S80 days)								
April 2017	25	127	4	- 48	0.90 [0.76, 0.97]	0.27 [0.21, 0.35]		
Askim 2017	8	59	60	1435	0.12 [0.05, 0.22]	0.96 (0.95, 0.97)		
Brabrand 2016	47	122	100	3555	0.32 0.25, 0.40	0.97 [0.95, 0.97]		
Chen 2016	291	273	256	821	0.53 [0.49, 0.57]	0.75 [0.72, 0.78]		
Churpek 2016	1133	10556	\$16	15432	0.69 [0.66, 0.71]	0.63 [0.62, 0.64]		
Churpek 2017	319	4166	1453	47911	0.18 [0.16, 0.20]	0.92 [0.92, 0.92]		
DeGroot 2017	45	278	97	1152	0.3210.25.0.401	0.87 (0.85, 0.86)		
Donnelly 2017	62	166	71	2220	0.47 [0.38, 0.55]	0.93 [0.92, 0.94]	-	
Finkelsztein 2007	26	71	2	\$2	0.90 [0.72, 0.94]	0.42 [0.33, 0.52]		-
Forward 2017	17	73		63	0.6810.46.0.851	0.4610.38.0.551		-
Freund 2017	52	166	22	622	0.7010.52, 0.001	0.79 (0.76, 0.62)		
Giamarellos 2017	525	755	340	1513	0.61 [0.57, 0.64]	0.71 (0.69, 0.72)		
Goezalez 2007	20	63	\$2	926	0.28 [0.18, 0.46]	0.94 [0.92, 0.95]		
Havdar 2017	20	26	2	01	0.9110.71.0.991	0.4610.38.0.531		
Henning 2016	173	1041	160	6263	0.52 10.46, 0.571	0.86 (0.85, 0.87)		
Happe 2007	13	76	2	235	0.87 [0.60, 0.96]	0.75 (0.70, 0.80)		
Hasen 2017	76	112	50	240	0.72 0.62 0.80	0.65 [0.63, 0.73]		
Hwang 2017	144	\$73	67	611	0.66 [0.62, 0.74]	0.52 [0.49, 0.54]	-	
Kim 2017	4	15	16	\$77	0.2010.06.0.441	0.97 (0.95, 0.96)	_	
Kim 2017	7	12	6	100	0.54 (0.25, 0.81)	0.89 (0.82, 0.94)		
Kolditz 2006	81	563	159	8485	0.29 [0.24, 0.35]	0.94 [0.93, 0.94]		
LeGues 2017	21	46	11	304	0.66 [0.47, 0.81]	0.69 [0.61, 0.77]	_	-
Mookpwitz 2017	457	3025	726	19943	0.35 10.35, 0.421	0.87 (0.85, 0.87)		
Petider 2017	8	5	17	94	0.32 [0.15, 0.54]	0.95 (0.89, 0.98)		
Oxinter 2017	5	29	2	157	0.71 0.29, 0.94]	0.84 [0.78, 0.89]		
Ranzani 2017	234	994	243	4549	0.49 [0.44, 0.54]	0.82 [0.81, 0.83]		
Rothman 2017	287	245	192	3202	0.6010.55.0.641	0.9010.89.0.911		
Seymour 2016	1040	10079	846	54557	0.55 [0.53, 0.57]	0.84 [0.84, 0.85]		
Shetty 2007	511	1258	561	10225	0.48 [0.45, 0.51]	0.89 [0.88, 0.90]		
Singer 2016	304	629	247	21550	0.30 [0.25, 0.35]	0.97 [0.97, 0.97]		
Szakmany 2007	18	22	60	269	0.23 [0.54, 0.34]	0.89 [0.85, 0.92]	-	
Tuani 2017	15	134	10	727	0.6010.33.0.791	0.6410.82.0.871	-	
Umernura 2017	84	240	5	58	0.94 [0.87, 0.98]	0.19 [0.15, 0.24]		
Wang 2016	81	108	50	238	0.62 [0.53, 0.70]	0.69 [0.64, 0.74]		
Williams 2017	164	741	163	7903	0.50 [0.45, 0.56]	0.91 [0.91, 0.92]		
Long-term mortality (> 30 days)								
Openally 2017	51	115	187	2003	0.21 (0.15, 0.27)	0.35 (0.94, 0.96)	-	
Quinter 2017	2	25	18	141	0.33 (0.17, 0.54)	0.85 (0.79, 0.90)		
Rannikko 2017	55	91	43	310	0.56 [0.46, 0.66]	0.80 [0.75, 0.84]		
							0 03 04 06 09 1	0.

	Dea	4	Surv	hard		Odds Ratio	Odds Ratio		
Study or Subgroup	Events		Funds.		Weinter	M-H, Fonders, 95% Cl	M-H, Bandom, 95% CI		
1.1 Shat form mortality									
April 2017	36	- 39	122	175	1.6%	331012880			
Askim 2017		69	59	1457	2.2%	21011.45.6.561			
Brakkand 2016	47	147	122	3877	3.1%	13,70 (8,27, 20,24)			
Chen 2216	291	547	272	1094	2.5%	2.4212.75.4.251	-		
Charpek 2016	319	1772	4166	52077	3.6%	2.5212.23.2.881	-		
Churpek 2217	1122	1549	10596	29229	2.6%	3,9213,43,4,251			
de0xxxt 2017	46	143	278	2137	3.1%	317(218,4.68)			
Denselly 2017	62	133	166	2395	3.1%	11.69 [8.02, 17.00]			
Finkelsztein 2017	26	22	71	123	1.2%	6.2511.82.22.101			
Forward 2817	17	- 25	73	139	1.9%	1.83(0.74, 4.64)			
Freund 2017	52	74	168	005	2.0%	9.10/5.27.15.411			
Giarsareños 2017	628	893	766	2568	3.5%	3,73(3,18, 4,38)	-		
Operaties 2017	20	72	63	222	2.7%	5.71 (3.21, 10.16)			
Haydar 2017	20	22	- 96	177	1.0%	8.44 (1.81, 37.18)			
Henning 2016	172	222	1041	7304	2.4%	6.51 (5.20, 8.14)	-		
Husen 2017	13	15	78	314	1.0%	19.87 (4.34, 88.07)			
Husen 2017a	76	105	112	252	2.9%	5.4212.25.8.781			
Hwang 2017	144	211	673	1184	3.3%	2.29 [1.88, 3.13]			
Kim 2017	- 4	20	18	595	1.4%	8.01 (2.43, 26.28)			
kim 2017a	÷.	13	12	112	1.3%	9,72 (2,89, 33,73)			
Kolditz 2016	81	280	561	9347	24%	6161489.8.081	-		
LeOsen 2017	21	32	46	150	21%	4.32(1.92, 9.68)			
Meskowitz 2017	467	1193	3028	22871	3.6%	4 24 13 25 4 281	-		
Patidar 2017	0	25	5	22	1.2%	0.05/2.50.20.201			
Quinten 2017	ŝ	- 7	- 29	189	0.9%	13 53 12 58 73 131			
Renzeni 2017	224	477	998	5547	2.5%	4.2913.62.5.321	-		
Reteran 2017	297	479	345	3447	3.6%	13.44 [10.85, 16.65]	-		
Sermour 2016	1040	1895	10079	64535	2.6%	6.6515.05.7.201			
Shefty 2817	611	1872	1258	11483	3.6%	7.4015.43, 8.461	-		
Singer 2017	104	251	629	22179	2.4%	14.43 [11.32, 18.38]	-		
Szakimane 2017	18	78	33	332	2.6%	2.45(1.23, 4.63)			
Tunpal 2017	15	- 25	124	991	2.1%	814(2.58.18.50)			
Universities 2017	- 24		240	295	1.0%	4.05 (1.58, 10.46)			
Wang 2016	81	121	108	346	2.0%	3 57 12 35 5 431			
Williamo 2017	164	327	741	0544	3.4%	10.59 (8.42, 13.33)	-		
Subtrinal (95% CD		12738		256806	92.8%	5.6014.63, 6.768	+		
Total events	6141		37079						
Heterogeneity: Tau* a	0.24: CN	a 663.	77. at = 3	6 /P < 0.0	00015/P	194%			
Test for overall effect	Z=17.02	P + 0.1	100011						
1.1.2 Long-term mort	alty								
Dennelly 2017	61	238	115	2148	3.2%	4.82(3.38, 6.93)			
Ouinten 2017	9	27	25	166	1.9%	2,9211.14, 6.981			
Rannibbo 2017	55		81	222	2.9%	5.0213.15.8.021			
Subfatal (95% Ct)		363		2713	8.8%	4.66 [3.54, 6.12]	•		
Total events	115		221						
Heterogeneits: Tau? = 0.00; Ch? = 1.21; df = 2 (P = 0.52); P = 0%									
Testfor overall effect Z = 11.04 (P × 0.00001)									
Tetal (95% CI)		13 90 1		259519	100.0%	5.40 [4.50, 6.56]	•		
Total events	6256		37300						
Heleropenety: Tau" = 0.22; Chi" = 555.43; df = 37 (P = 0.00001); P = 92%									
Testforoverall effect 2 = 18.55 /P < 0.00001) Favoura (dead) Favoura (dead)									
Test for subgroup diff	eneces:	Chiffe 1	17, df = 1	(P = 0.2)	0,7114	9%	· · · · · · · · · · · · · · · · · · ·		