

1 Leveraging Transcriptomics Data for Genomic Prediction Models 2 in Cassava

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12 Abstract

13 Background

14 *Genomic prediction models were, in principle, developed to include all the available marker
15 information; with this approach, these models have shown in various crops moderate to high
16 predictive accuracies. Previous studies in cassava have demonstrated that, even with relatively
17 small training populations and low-density GBS markers, prediction models are feasible for
18 genomic selection. In the present study, we prioritized SNPs in close proximity to genome regions
19 with biological importance for a given trait. We used a number of strategies to select variants
20 that were then included in single and multiple kernel GBLUP models. Specifically, our sources of
21 information were transcriptomics, GWAS, and immunity-related genes, with the ultimate goal to
22 increase predictive accuracies for Cassava Brown Streak Disease (CBSD) severity.*

23 Results

24 *We used single and multi-kernel GBLUP models with markers imputed to whole genome
25 sequence level to accommodate various sources of biological information; fitting more than one
26 kinship matrix allowed for differential weighting of the individual marker relationships. We
27 applied these GBLUP approaches to CBSD phenotypes (i.e., root infection and leaf severity three
28 and six months after planting) in a Ugandan Breeding Population (n = 955). Three means of
29 exploiting an established RNAseq experiment of CBSD-infected cassava plants were used.
30 Compared to the biology-agnostic GBLUP model, the accuracy of the informed multi-kernel
31 models increased the prediction accuracy only marginally (1.78% to 2.52%).*

32 Conclusions

33 *Our results show that markers imputed to whole genome sequence level do not provide enhanced
34 prediction accuracies compared to using standard GBS marker data in cassava. The use of
35 transcriptomics data and other sources of biological information resulted in prediction accuracies
36 that were nominally superior to those obtained from traditional prediction models.*

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38 Background

39 Genomic Selection (GS) [1] is a breeding method that exploits high-throughput genotyping
40 technologies, novel statistical methods and the availability of genomic information. It has been
41 used extensively in animal breeding and promises to impact plant breeding, particularly within
42 clonally propagated and perennial plant systems [2].

43 GS approaches tend to avoid marker selection, and instead, all the marker information is utilized
44 within the prediction models. Given such scenario where the number of predictors (p), is greater
45 than the number of available observations (n) traditional regression models achieve poor
46 predictive ability as a result of multicollinearity and overfitting among the predictors [2,3].
47 Several statistical methods have been explored to overcome these problems; shrinkage
48 methods, where the regression coefficients are shrunk towards zero, are widely used for
49 genomic predictions [4]. These methods include Genomic Best Linear Unbiased Predictions
50 (GBLUP) [5], Bayesian regression [1,6], Least Absolute Shrinkage and Selection Operator (LASSO)
51 [4] and ridge regression BLUP (rr-BLUP) [7]. Recently, machine learning methods have been
52 proposed for genome-enabled predictions as they are capable of dealing with the dimensionality
53 problem in a flexible manner [8,9]. Performance comparisons among these models have been
54 conducted in several plant species [10–13] showing that the best statistical approach depends
55 highly on the trait and the species that is being analyzed.

56 GS predictions rely on linkage disequilibrium (LD) between the markers and the Quantitative
57 Trait Loci (QTL). Given the dramatic drop in sequencing costs, full-genome sequence data was
58 proposed to be used in genomic predictions [14]. Simulation studies suggest that the use of
59 whole genome sequence data would result in increased accuracy of genomic predictions [14–
60 16] because the accuracy that can be achieved by the prediction model is no longer tied to the
61 LD-QTL relationship as the causal mutations are present in the dataset [15].

62 Whole-genome sequencing is still prohibitively expensive for most crop breeding programs as
63 the number of individuals evaluated can reach the tens of thousands. An efficient and cost-
64 effective approach is to impute the whole-genome sequence variants of the individuals using a
65 low-density genotyping platform and a previously sequenced reference population (reference
66 panel) [17]. This system is widely used in human genetics, where large-scale sequencing efforts,
67 like the 1000 Genome Project [18], provides standard reference panels for imputation.

68 In livestock and some crops, breeding populations are typically derived from a small group of
69 common ancestors within a few generations in the past. Thus, these populations tend to have a
70 small effective population size (N_e); this is a perfect scenario for performing whole genome

71 imputation (WGI) as low-density markers will be able to adequately trace the haplotypes
72 inherited from the ancestors [15] easing the imputation process.

73 Genomic prediction models tend to use unannotated anonymous markers, even when this is
74 currently slowly changing, most models do not take into consideration whether SNPs are close
75 to genic or regulatory regions. When imputing markers to whole sequence level, the number of
76 predictors utilized increases significantly and so does the $p \gg n$ problem; this might prevent the
77 model to put sufficient weight on the causal variants [19] thus affecting prediction accuracies.
78 The use of biological priors has been proposed to both alleviate this problem and reduce the
79 computational burden associated with models using millions of markers [20].

80 Over the last few years, several methods have been developed to incorporate biological or
81 functional information into Association Studies and Genomic Prediction. In cattle, for example,
82 Fortes et al. used an Associated Weighted Matrix (AWM) [21] to infer a set of genes related to
83 beef tenderness. They later demonstrated that making genomic predictions with only SNPs near
84 the inferred genes for beef tenderness resulted in prediction accuracies that were higher than
85 when the entire marker set was used [22]. Other methods have sought to exploit biological
86 information while avoiding marker selection. Su et al. [23] for example, tested a genomic BLUP
87 (GBLUP) model where the relationship matrix was weighted using prior Bayesian models or
88 GWAS summary statistics [23,24].

89 In contrast to the traditional GBLUP that assumes that all SNPs have the same effect-size
90 distribution, methods like GFBLUP [25] or MultiBLUP [26] add one or multiple genomic random
91 effects that quantify the importance of different marker sets respectively. These marker sets are
92 typically defined by some source of biological evidence (i.e., metabolic pathway, sequence
93 annotation, transcriptomics, evolutionary constraints).

94 A Bayesian method that has also been implemented to leverage biological information in
95 prediction efforts is BayesRC [27] which uses a mixture of normal distributions to model SNP
96 effects and include prior biological knowledge. BayesRC [28] allows the user to *a priori* allocate
97 the SNPs into classes where each class is believed to have a different probability of containing
98 causal variants for the trait. The aforementioned genomic feature modeling approaches
99 (GFBLUP, MultiBLUP, and BayesRC) were designed to improve prediction accuracies of complex
100 traits if the groups of markers selected are enriched for causal variants [28,29].

101 Transcriptomics studies have allowed researchers to investigate gene expression dynamics of
102 different organisms in different tissues, conditions or developmental stages [30]. It can be of aid
103 to discover genes and pathways that are involved in the regulation of complex traits, potentially

104 revealing genomic regions that would be enriched in variants affecting specific traits [25,31].
105 Transcriptomics studies have already been used effectively as a source of biological priors to
106 predict complex traits in cattle [20,25]. These studies showed that using informed models could
107 slightly improve prediction accuracies when making same breed predictions and that the
108 observed improvement was more evident with a greater genetic distance between the training
109 and validation population (across-breed predictions).

122 Cassava (*Manihot esculenta*) is a major staple crop in parts of sub-Saharan Africa and is the
123 primary source of calories for millions of people across the world [32]. Cassava Brown Streak
124 Disease (CBSD) is a viral disease that hampers the production of cassava and is considered a
125 serious threat to food security in Africa [33,34]. CBSD is caused by two distinct single-stranded
126 RNA viruses, Cassava Brown Streak Virus (CBSV) and Ugandan Cassava Brown Streak Virus
127 (UCBSV) [34–36]. Recently, transcriptomics data in cassava has been used to unravel the
128 transcriptional dynamics of cassava plants under infection by both UCBSV [37] and CBSVs [38].

129 In the present study, CBSD phenotypes (root infection and leaf severity three and six months
130 after planting) from a Ugandan Breeding Population (n=955) were analyzed using whole genome
131 imputation (WGI) data (~5 million SNPs) and biological information coming from transcriptomics
132 experiments [37,38], Genome-Wide Association Studies (GWAS) [39] and in-silico identification
133 of immunity-related genes [40,41]. Our main objective was first to assess the feasibility of
134 performing whole genome imputation in cassava and second to test if prediction accuracies can
135 be enhanced by using WGI together with biological priors using GBLUP-derived models.

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144 Methods

145 Plant material

146 Two diverse cassava populations were combined and used as a composite set for this study;
147 individuals in this composite data set represented the genetic diversity of the Ugandan cassava
148 gene pool. The first population (“Training”) was comprised of a panel of 414 cassava accessions
149 from the breeding program of the National Crops Resources Research Institute (NaCRRI) in
150 Namulonge, Uganda. This population was the first used to train genomic prediction models for
151 applied breeding at NaCRRI. The second population, (“GWAS”) was developed by Kayondo et al.
152 [39] and was comprised of 540 accessions. This population is derived from 49 parents from the
153 International Institute of Tropical Agriculture (IITA), The International Center for Tropical
154 Agriculture (CIAT) in Colombia and some landraces of East Africa. Briefly, the “Training” panel
155 was evaluated in two years (2012-2013), and three locations in an alpha-lattice design, and the
156 “GWAS” panel was evaluated in a single year (2015) at three locations using an augmented
157 randomized complete block design. For more information on both populations, please refer to
158 [39]. For a list of the accessions used, see Table S1.

159 Phenotyping Platform

160 The composite plant population was phenotyped for three separate traits: foliar CBSD severity
161 measured three (CBSD3) and six (CBSD6) months after planting and CBSD severity in the storage
162 roots (CBSDR) after a year. Briefly, CBSD severity was scored based on a 5-point scale with a
163 score of 1 implying an asymptomatic plant while a score of 5 would mean over 50% of leaf vein
164 clearing for foliar symptoms (CBSD3 and CBSD6) and 50% of root-core being covered by necrosis
165 for CBSDR. Please refer to Kayondo et al. [39] for further details.

166 Genotyping by sequencing and imputation

167 Genotyping-by-sequencing (GBS) libraries [42] were constructed as previously described [43].
168 Marker genotypes were called using the TASSEL 5.0 GBS discovery pipeline [44] after aligning
169 the reads to the *Manihot esculenta* Version 6 assembly. Genotype calls were stored in 18 Variant
170 Calling Format (VCF) files (one per cassava chromosome). The VCF files were filtered using
171 VCFtools [45]; individual marker calls were masked if the read depth was lower than 3x, cassava
172 genotypes with > 80% missing calls and SNP markers missing more than 60% were removed.
173 Insertions, deletions, and multi-allelic markers were also withdrawn from the dataset. Beagle
174 4.1 software [46] with default parameter settings was used for imputation. In total 173k SNPs
175 were called among 986 individuals. This dataset was further filtered by an Estimated Allelic r-
176 squared statistic (AR²) > 0.3 and a minimum Minor Allele Frequency (MAF) of 1%. The final set

177 herein referred to as the “GBS” dataset, included 41,530 SNP markers called among the 954
178 individuals.

179 *Imputation to whole-genome sequence data*

180 Beagle 4.1 [46] and Impute2 [47,48] were tested and compared for imputation accuracy, marker
181 density, and marker distribution. For both software’s, a Cassava Haplotype Map (HapMap) of
182 241 accessions was used as a reference panel. This reference panel represented cultivated,
183 hybrid and wild cassava relatives and contained 28 million SNP markers [49].

184 *Beagle Imputation*

185 Imputation using Beagle 4.1 was performed in two steps (Figure S1). During the “BEAGLE Stage
186 I” phase, a subset of the HapMap markers was used, including bi-allelic SNPs with MAF greater
187 than 1%. Additionally, a 10bp thinning filter was set up, meaning that only one marker per 10bp
188 was allowed. The resulting set included 716k markers with MAF > 1% and AR2 > 0.3. The BEAGLE
189 Stage I marker set was then used in the second round of full HapMap imputation. The second
190 marker dataset, “BEAGLE Stage II” had 2 million markers exposed to the same MAF and AR2
191 filters. The genetic positions of the HapMap markers were inferred using a smooth spline fit to
192 the 22,403-marker composite map published by the International Cassava Genetic Map
193 Consortium (ICGMC) [50]. The genetic positions were forced to be monotonically increasing,
194 which is a requirement for BEAGLE to run properly. Beagle 4.1 ran with default parameters. For
195 this manuscript, only the ‘BEAGLE Stage II’ markers were considered, and herein it will be
196 referred to as the “BEAGLE” dataset.

197 *Impute2 Imputation*

198 Imputation using IMPUTE2 was performed in a single step (Figure S2). The number of haplotypes
199 used as “custom” reference panel (-k_hap) was set to 400, the effective population size (Ne) to
200 1000, and the imputation window to 5Mb. The genetic positions of the HapMap were inferred
201 as described in the “Beagle Imputation” section of this manuscript. The IMPUTE2 software,
202 however, requires knowing the recombination rate between the current position and next
203 position on the map. This recombination rate was calculated using the following formula:

$$204 \quad RR = \frac{cM_{i+1} - cM_i}{Mb_{i+1} - Mb_i}$$

205 Where **cM** represents the genetic position of each marker “*i*” and **Mb** notes the physical position
206 in megabases. The accuracy of the imputation was assessed using internally-calculated
207 concordance tables. Briefly, IMPUTE2 masks the genotypes of one variant at a time from the
208 study data (GBS markers) and then imputes the masked genotypes with information from the
209 reference panel and the nearby variants. The percentage of concordance between the masked

210 and the imputed genotypes for each 5Mb imputed window were subsequently calculated
211 (Figure S3). Additionally, allele frequencies and imputation quality distributions were calculated
212 and depicted by the IMPUTE2 information measure statistic “info” [48] (Figure S4) and
213 imputation quality by allele frequencies (Figure S5).

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215 **Biological Information**

216 Three different sources of biological information related to CBSD resistance were used in this
217 study.

218 *Transcriptomics profiling*

219 RNAseq data were obtained from two experiments. The first experiment [37] focused on
220 profiling the transcriptome response across seven time-points after infection with UCBSV. Two
221 contrasting cassava genotypes were used: ‘Namikonga’ (CBSD resistant) and ‘Albert’ (CBSD
222 susceptible) (Figure S6). The 84 libraries (Table S2) were checked for read quality using FastQC
223 [51]. The Tuxedo Suite of programs [52,53] was then used to process the sequenced data. Reads
224 in FASTQ formats were aligned to the *M. esculenta* reference genome v6 [54] using TopHat
225 v2.1.1/Bowtie v2.2.8 [55]/[56]. A reference annotation of the cassava gene models (v6.1) from
226 the Phytozome database was provided (<https://phytozome.jgi.doe.gov>). This version of the gene
227 annotation contained a total of 33,033 transcripts. The minimum and maximum intron length
228 were set to 10 and 15,000bp respectively; the remaining parameters were set to default values.
229 Subsequently, the *Cuffdiff* program within Cufflinks version 2.2.1 [57] was used to identify
230 differentially expressed (DE) genes at each time-point among infected plants and controls. A
231 false discovery rate of 0.01 after the Benjamini-Hochberg correction for multiple testing was
232 used.

233 The second transcriptomics data was taken from Anjanappa et al. [38]. In this experiment, two
234 cassava genotypes, the resistant ‘KBH 2006/18’ and the susceptible ‘60444’, were challenged
235 against a mix of CBSV strains (CBSV – TAZ-DES-01 and UCBSV – TAZ-DES-02). RNAseq was
236 performed 28 days after infection; this time point was selected because it showed homogenous
237 virus titer levels across the biological replicates in the susceptible genotype. Raw reads were not
238 re-analyzed; a list of DE genes was extracted from the Anjanappa et al. manuscript (Table S3).

239 *Quantitative Trait Loci*

240 Kayondo et al. recently reported two major QTLs for CBSD foliage symptoms [39], one near the
241 end of chromosome 11 and another on chromosome 4 that collocates with a previously
242 reported, large introgression from wild cassava (Figure S7). Bi-parental QTL mapping has also

243 identified hits on chromosomes 4 and 11 for foliar symptoms [58] and chromosome 11 for root
244 necrosis [59]. Small effect QTLs related to CBSD symptoms on roots were also detected, but they
245 were not considered in this study.

246 *Immunity-related genes*

247 The most common disease resistance genes in plants are those belonging to the NBS-LRR family
248 [60]. This highly conserved gene family has already been identified and positioned in a previous
249 version of the cassava genome (Cassava Genome v5.0) [61]. In that study 228 NBS-LRR and 99
250 partial NBS-LRR genes were reported. Positions for each NBS-LRR genes were updated to fit in
251 the latest cassava genome assembly (<http://phytozome.gov>, Cassava Genome v6) using Blast+
252 [62] (Table S4). Additionally, immune-related genes listed by Soto et al. [41] were added to this
253 list (Table S4).

254 *Associating markers with genes*

255 Markers that appeared within the coding region of a gene (defined as 5'UTR to 3'UTR, including
256 introns) were considered to be “tagging” that gene. Bedtools [63,64] and in-house scripts
257 (available from the GitHub page of this manuscript) were used to associate SNP markers to genes
258 of interest.

259 *Co-expression Networks using WGCNA*

260 Weighted Co-expression Network Analysis (WGCNA) [65,66] was used to identify highly
261 correlated genes across different time-points based on their expression. Briefly, Fragments Per
262 Kilobase of exon per Million reads (FPKM) were log₂ transformed. Genes without variation
263 across the seven timepoints were filtered out using a Coefficient of variance ($CV = \sigma/\mu$) cutoff
264 of 0.9. Analyses were performed using the ‘WGCNA’ package in R programming software [67].
265 As previously described [66], ‘WGCNA’ calculates an expression Pearson’s correlation matrix for
266 the genes, this matrix is later raised to a power β (0.8 in this study) before continuing with the
267 clustering procedure. The ‘WGCNA’ *treecut* parameter was set to 0.85; the three parameters CV,
268 β and *treecut* values were selected based on the number and quality of the co-expression
269 modules identified. All other parameters were set to the package’s default values. To visualize
270 the general trend of each module, eigengenes were calculated as the first principal component
271 of the normalized expression values of all genes within a module and plotted as a heatmap
272 [68,69].

273 *Genomic Selection Models*

274 A two-step approach was used to evaluate genomic predictions in this study. This method was
275 used to increase computational efficiency and control for differences in experimental design

276 between different datasets. The first step involved accounting for trial-design variables using
277 linear mixed models to calculate de-regressed Best Linear Unbiased Predictions (BLUPs), and the
278 second step used the de-regressed BLUPs as phenotypes in the prediction model.

279 *Genotypic value estimation*

280 De-regressed BLUPs were calculated according to Garrick et al. [70]. The procedure has been
281 described previously [12,71] and for this composite population specifically in Kayondo et al. [39].
282 Briefly, a mixed model was fit with the population mean and location as fixed effects and clone
283 and breeding design variables (i.e., block, range) as random effects. BLUPs for clones represents
284 an estimate of the total genetic value (estimated genetic value, EGV). Clone effect BLUPs (EGVs)
285 were then extracted as the de-regressed BLUPs following:

$$286 \quad dBLUPs = \frac{BLUPs}{1 - \frac{PEV}{\sigma_{\mu}^2}}$$

287 Where σ_{μ}^2 is the genetic variance and **PEV** is the prediction error variance of the BLUPs. Solutions
288 for both σ_{μ}^2 and PEV were retrieved from the mixed models solved using the *lmer* function of
289 ‘lme4’ package [72] in R software.

290 *Prediction models*

291 We used three variations of the classic GBLUP to predict estimated breeding values (GEBV) for
292 CBSD related traits:

293 **GBLUP** was fit using a linear mixed model of the form:

$$294 \quad dBLUPs = 1_n\beta_0 + Zg + e, \quad \mathbf{g} \sim N(0, K\sigma_g^2) \quad , \quad \mathbf{e} \sim N(0, I\sigma_e^2)$$

295 Where the solution for **g** represents the GEBVs. Briefly, β_0 is the mean, vector **g** is the random
296 effect for the genetic markers, **Z** is a design matrix pointing observations to genotype identities,
297 and **e** are the residuals. We assume that **g** has a known covariance structure defined by the
298 genomic realized relationship matrix **K**. The genomic relationship matrix **K** was constructed using
299 SNP dosages and an Rcpp [73] implementation of the function *A.mat* in the R package ‘rrBLUP’
300 [74]. GBLUP predictions ran using the function *emmreml* in the ‘EMMREML’ R package [75].

301 **GFBLUP** [29,76] is a modification of the traditional GBLUP that includes an additional genetic
302 random effect; the linear mixed model followed the form:

$$303 \quad dBLUPs = 1_n\beta_0 + Zf + Zr + e, \quad \mathbf{f} \sim N(0, K_f\sigma_f^2) \quad , \quad \mathbf{r} \sim N(0, K_r\sigma_r^2) \quad , \quad \mathbf{e} \sim N(0, I\sigma_e^2)$$

304 where K_f and K_r were genomic relationship matrices built using the SNPs within and outside
305 the genomic feature. Specifically, K_f was calculated with markers thought to be enriched for
306 causal variants and K_r was calculated with the rest of the markers in the genome. The
307 relationships matrices were calculated as described before and the GFB LUP predictions were
308 conducted using the *emmremlMultiKernel* function in the 'EMMREML' R package [75].

309 **MULTIBLUP** [26] was also used. This method is similar to GFB LUP but allows for multiple genetic
310 random effects. As with GFB LUP method, predictions were conducted using the
311 *emmremlMultikernel* function implemented in the 'EMMREML' R package.

312 *Cross-validation*

313 The accuracy of genomic prediction was measured as the correlation between the total genetic
314 value (EGV, the random genetic effect from the first step regression model, not de-regressed)
315 and the GEBVs. We used 25 replications of a five-fold cross-validation scheme to obtain
316 unbiased estimates of the prediction accuracies. The process of cross-validation used in this
317 study was previously detailed by Wolfe et al. [13].

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332 Results

333 Describing the population

334 We used the GBS marker dataset (~40K SNPs) to describe the LD patterns, population structure,
335 and MAF distribution within a composite set of cassava varieties (Figure 1). After plotting the
336 mean LD score (As in GCTA-LDS, [77]) of each variant, we noted a high level of LD heterogeneity
337 across the entire cassava genome. Major LD peaks were not observed in centromeric regions,
338 as would be expected with the common fall in recombination rate. Some high LD clusters were
339 observed, however, near to the telomeres (Figure 1a). High LD across chromosome 4 and at the
340 end of chromosome 1 were consistent with two relatively recent introgressions from a wild
341 cassava relative [54]. The unique LD pattern in these two chromosomes was evident after
342 plotting a regular LD decay plot (Figure 1b). Principal component analysis (PCA) on the dosage
343 marker matrix (Figure 1c) indicated that there is little genetic differentiation between the two
344 populations merged for composite analysis in this study. Moreover, the percentage of variance
345 explained by the first two PCs was only 8.95%. The allele frequency distribution was also similar
346 between the two populations (Figure 1d).

347 Imputation to whole genome sequence

348 We compared two different methods to impute the GBS dataset to a whole-genome sequence.
349 BEAGLE and IMPUTE2 methods have been challenged before regarding imputation accuracy and
350 computational time, the results of which suggest that both approaches are sufficiently robust
351 [78]. To select genetic markers that would “tag” candidate genes, we focused on the number
352 and distribution of higher quality imputed SNPs ($AR^2/info > 0.3$, $MAF > 0.01$) across the cassava
353 genome. Using IMPUTE2 resulted in high-quality markers, more tagged genes (Figure 2a), and
354 better marker distribution (Figure 2b, Figure S8) than BEAGLE. The total number of predicted
355 genes in the current cassava assembly was 33,033. We tagged 32% of them using GBS markers,
356 70% using the BEAGLE imputed dataset, and 91% when using IMPUTE2. Other quality control
357 tests were performed on the IMPUTE2 dataset, including imputation accuracies per
358 chromosomal segments, distribution of allele frequencies, and “info” quality scores (Figure S3-
359 S5).

360 Impact of Imputation level on Genomic Prediction accuracies

361 Prediction accuracies of a regular GBLUP model for three CBSD-related traits are shown in Figure
362 3. Specific conclusions regarding the impact of different imputation levels on prediction
363 accuracies are not possible, as there is not a common trend among the three traits. We did
364 note, however, that there was not a significant increase in prediction accuracy using different

365 imputation levels. Moreover, when evaluating Cassava Brown Streak Disease severity six months
366 after planting (CBSD6), the accuracy using only GBS data was consistently higher than any of the
367 imputation methods tested. We also compared the prediction accuracies using one subset of
368 markers from IMPUTE2 that matched the position of the GBS markers (Impute2GBS) and
369 another subset using only SNPs imputed with the highest reliability ($AR^2/info > 0.9$, Impute290,
370 $n = 371,524$). Again, the prediction accuracies resulting from these subsets were nearly identical
371 to those obtained using the full GBS and IMPUTE2 dataset (Figure 3).

372 [Accounting for known QTLs](#)

373 Kayondo et al. [39] previously conducted a Genome Wide Association Study (GWAS) and
374 identified two big effect QTLs for foliar CBSD severity using the same cassava population
375 presented in this manuscript. The first identified QTL was very wide and located in the middle
376 of chromosome 4. This QTL appeared to co-locate with a recent introgression from a wild
377 cassava relative. The second QTL was located at the end of chromosome 11 (Fig. S7).

378 This study sought to evaluate the relative importance of these QTLs for genomic prediction
379 accuracy. We first ran a genomic prediction GBLUP model which included two genomic random
380 effects: the first built with markers from chromosome 4 and the second built with markers from
381 chromosome 11. We compared the partial and total accuracies of this model with another two-
382 kernel GBLUP model built with two random chromosomes, excluding chromosomes 4 or 11
383 (Figure 4). A clear difference in prediction accuracy was observed when chromosomes
384 containing QTLs (blue) and random chromosomes (white) were compared. Since these QTLs
385 were detected on foliar symptoms, we observed that the influence of chromosome 4 and 11 is
386 higher in predictions of foliar phenotypes (CBSD6) than in necrosis on roots (CBSDR).
387 Additionally, when we compared the total accuracy of the model including only the
388 chromosomes with identified QTLs, we observed the prediction accuracy for CBSD6 was very
389 close to the model calculated using all 18 cassava chromosomes. We then fit a model with three
390 kernels (i.e. chromosome 4, 11 and the rest of the genome) to investigate if there was any
391 additional variance beyond the chromosomes containing the important QTL (Figure S9). The
392 total prediction accuracy increased slightly for each measured trait, but it did not reach the
393 accuracy level obtained when all markers were used in a single kernel model. This result suggests
394 that marker partitioning is performed at the cost of prediction accuracy.

395

396 Using Transcriptomics data

397 Amuge et al. [37] profiled the response of two contrasting cassava genotypes to infection with
398 UCBSV. RNA samples were collected across seven time points after inoculation by grafting with
399 UCBSV and deep sequenced using the Illumina platform (Figure S6). Relative virus titer was
400 quantified from the RNAseq libraries as the number of reads mapping to either CBSV or UCBSV
401 genomes (Figure S10). Additionally, reads mapped to either of these genomes were de-novo
402 assembled using Trinity [79] as a means of confirming the virus infecting the plant was only
403 UCBSV and not CBSV (Figure S11). As previously demonstrated by Amuge et al., the
404 transcriptional response of the two genotypes evaluated was radically different after UCBSV
405 infection. While the tolerant cassava variety ('Namikonga') showed a strong response across
406 most of the seven timepoints, the susceptible variety ('Albert') showed no transcriptional
407 response between 24 hours and 8 days after infection (Figure 5, Table S5). Under the assumption
408 that tagging and prioritizing SNPs close to genes contributing to the plant-virus interaction would
409 increase prediction accuracies, we proceeded to explore different means of exploiting this
410 dataset to locate these genes of interest.

411 *Differentially expressed genes*

412 The most direct way to use the transcriptome dataset was to apply a GFBLUP procedure using
413 the SNPs inside each Differentially Expressed (DE) gene as genomic features. We ran this analysis
414 for two traits (CBSD6, CBSDR) and compared prediction accuracies between each GFBLUP model
415 and the regular GBLUP model using the whole genome sequence imputed dataset (WGI) (Figure
416 6). In total, we ran eleven different GFBLUP models, including one comprised of DE genes across
417 all time points (DE-all). While there were differences in the mean prediction accuracies between
418 the models, none of them were significant.

419 *Genes having a significant interaction between genotype and inoculation status*

420 An alternative means of selecting genes of importance across all DE genes was to consider only
421 those genes with a significant interaction with Genotype-by-Inoculation status (herein referred
422 to as *GxI* genes). To accomplish this, a mixed model was fit for each gene:

$$423 \quad E \sim \text{reps} + G * I * T + e$$

424 Where **E** is expression in FPKM, **reps** encompasses the three replicates as a random effect and
425 **G * I * T** describes the three-way, fixed effect interaction among inoculation status (**I**, infected
426 or control), Genotype (**G**, susceptible or resistant) and the different time points (**T**). The p-values
427 for each **G * I** interaction were extracted and corrected for multiple testing using a 5% FDR. Out
428 of the total set of 33,033 genes in the cassava genome, 1,392 showed a significant *GxI*

429 interaction at 5% FDR and 292 at 1% FDR (Table S6). The genomic distribution of these genes
430 appeared to be uniform (Figure 7a). When using GFBLUP, we noted that partitioning SNPs into
431 two kernels based on whether they tagged *Gxl* genes (at both 0.05 and 0.01 FDR thresholds) was
432 not advantageous for prediction accuracies (Figure S12).

433 Based on previous results demonstrating the importance of large-effect QTLs on chromosomes
434 4 and 11, we partitioned the *Gxl* SNPs into three kernels: chromosome 4, chromosome 11 and
435 the rest of the genome. In this model, only SNPs inside the significant *Gxl* genes (5% FDR) were
436 considered. This was in contrast to the GFBLUP approach, where a kernel with information from
437 the rest of the genome was fit. Thus, the number of SNPs used was much lower than the GFBLUP
438 approach. The prediction accuracies using this three-kernel model were similar to those using
439 the WGI dataset, despite using less than 2% of the SNPs (Figure 7b). To test that the *Gxl*
440 associated SNPs were relevant for prediction, we also ran a model using a different random set
441 of SNPs during each of each of the 25 rounds of cross-validation. These random SNPs were in
442 approximately linkage equilibrium with the *Gxl*-associated SNPs. The *Gxl*-associated SNPs
443 showed significantly better prediction accuracies than when random SNPs were used (Figure
444 7b). Given the apparently good results using the three-kernel method, we fit the same model
445 with an extra kernel to account for the rest of the genome and while we expected an additional
446 boost in prediction accuracies, we did not observe an increase (Figure S13). Whether the rest of
447 the genome SNPs has spurious associations that decrease prediction accuracies or if there is an
448 implicit “cost” for partitioning the genome in a multiBLUP model, are hypotheses that were not
449 tested in this manuscript.

450 *Co-expression modules*

451 We used Weighted Gene Correlation Network Analysis (WGCNA) [65,66] to identify correlated
452 genes based on their expression patterns across the different timepoints. WGCNA allows the
453 identification of modules of genes that are more correlated within each other than they are to
454 genes outside the module [65]. This unsupervised method was used to identify modules of co-
455 expressed genes and test if any of these modules were more important or enriched in causal
456 variants, the result of which would increase prediction accuracies for any of the CBSD related
457 traits under a GFBLUP framework.

458 Of the 33,033 total genes in the reference cassava genome, 5,574 passed an ad-hoc Coefficient
459 of Variance filter ($CV = 0.9$) and were used in downstream analysis. From the remaining 5,574
460 genes, 2,789 were assigned to 16 modules containing between 43 and 991 genes (Table S7). A
461 total of 2,785 genes could not be assigned to any module (Grey module). Eigengenes for each

462 module were calculated and plotted in a heatmap depicting modules as rows and the
463 timepoints, genotypes, and inoculation status as columns (Figure 7a). While some modules are
464 noisy with a broad co-expression pattern across different timepoints and conditions, some of
465 them are correlated at only one or two conditions (yellow, etan, and green). Other modules are
466 dependent on time after infection, regardless of genotype or inoculation status (turquoise).
467 Interestingly, two modules (black and cyan) grouped genes with ‘Namikonga’ and ‘Albert’
468 specific expression across all timepoints (Figure 7a).

469 We then used the identified modules to fit a GFBLUP model for each module. The accuracies
470 obtained are shown in Figure 7b. For CBSD severity six months after planting (CBSD6) and
471 severity on roots (CBSDR), none of the GFBLUP models provided a significant advantage in
472 prediction accuracy over the traditional GBLUP (WGI). For CBSD severity three months after
473 planting (CBSD3), however, one of GFBLUP module model (red, 154 genes, 3,558 SNPs) obtained
474 a prediction accuracy higher than WGI. Using WGCNA as a proxy to identify genomic features
475 helped to marginally improve the genomic prediction accuracy for only one of the traits tested.

476 *Other biological data*

477 As a final step in this analysis, we incorporated all the available biological information, including
478 large-effect QTL peaks, *GxI* genes, and previously identified immunity-related genes. The
479 immunity-related genes included NBS-LRR genes[40], immunity-related genes as annotated by
480 Soto et al. [41], and DE genes proposed to have a major role in the resistance response against
481 joint UCBSV and CBSV infection in a single-point transcriptomics study (Table S3) [38].

482 Multi-kernel GBLUP models were fit with SNPs tagging each biological information category;
483 chromosome 11 large-effect QTL, chromosome 4 large-effect QTL, *GxI* significant genes, and
484 immunity related genes (Fig 8). A small increase in prediction accuracy for each of the traits was
485 obtained through various combinations of the information above. For CBSD3, a three-kernel
486 model with the chromosome 11 large-effect QTL, tagged *GxI* genes, and genes present in the
487 red WGCNA module increased accuracy by 1.7% (Fig 8a). For CBSD6, a four-kernel model using
488 QTLs from both chromosome 11 and chromosome 4, tagged *GxI* genes, and the immunity-
489 related genes resulted in a 2.52% increase in prediction accuracy (Fig 8b). Finally, a three-kernel
490 model considering only the chromosome 11 large-effect QTL, the immunity related genes, and
491 the tagged *GxI* genes resulted in a prediction accuracy increase of 2.52% for roots phenotyped
492 one year after planting (Fig 8c).

493 Discussion

494 In this study, we explored the improvement of genomic prediction in cassava through the
495 integration of transcriptomics data, the genetic architecture of CBSD, biological priors, and
496 whole sequence variants. Our results provide insight on how incorporating biological
497 information into prediction models can impact genomic prediction within this important staple
498 crop. Also, we explored models which can be extended to its use on other sources of biological
499 data such as regulatory elements, evolutionary conserved regions, chromatin accessibility
500 assays, and eQTLs.

501 *SNP imputation to Whole-genome sequence*

502 Compared to the prediction accuracies obtained using GBS markers, imputed sequence data
503 produced no advantage when applied to CBSD related traits. This behavior has been noted in
504 other animal empirical studies, where marginal [80] or absent increases in prediction accuracy
505 and reliability were observed [19,81–83]. Simulation studies, however, have reported significant
506 gains in prediction accuracy under some circumstances (i.e., low MAF of the causal variants)
507 [14–16]. As reviewed before [19], several reasons may account for this lack of increase in
508 prediction accuracy when using imputed sequence data. Problems with the imputation method
509 itself, small reference panels, and causal variants with low MAF may result in difficulties
510 imputing sequence data. Additionally, many markers could result in models failing to put
511 sufficient weight on the causal variants (i.e. a severe “ $p \gg n$ ” problem).

512 In our study, an imputation reference panel of only 240 individuals was used to impute a dataset
513 of 955 highly related individuals from NACRRI (Namulonge, Uganda). Additionally, the cassava
514 genome has at least two major and recent introgressions from wild relatives [54] on
515 chromosomes 1 and 4. Since wild cassava individuals are underrepresented in the reference
516 panel [49] introgressed regions showed a significant drop in imputation accuracies (Fig S3).
517 Moreover, the overall imputation accuracy in this dataset was significantly lower than when a
518 larger and more diverse target panel was used. While these factors have affected the prediction
519 accuracies, the purpose of using imputed sequence data in this study was to tag the maximum
520 number of genes rather than just increase predictive accuracies by imputing to sequence level.
521 That is, imputation was performed as a means of ensuring relevant genes could be tagged and
522 used as additional information in the model.

523

524 *Genetic Architecture of CBSD*

525 Genetic architecture of a trait is an important consideration when implementing different
526 genomic prediction models. Genetic architecture can vary drastically from trait to trait but also
527 from species to species. For example, in maize, most agronomic traits are controlled by many
528 small effect loci. This is in contrast to rice, where many agronomic traits, including grain yield,
529 have large effect QTLs [84].

530 Resistance to CBSD in cassava was historically considered to be a quantitative trait under the
531 control of several contributing loci. However, large-effect QTLs were recently detected using
532 association studies in a diverse population [39] and by traditional bi-parental QTL mapping
533 [58,59]. In the present study, we showed that when genomic predictions were performed using
534 only markers belonging to chromosomes containing the large-effect QTLs (i.e. chromosomes 4
535 and 11), nearly the same prediction accuracies were obtained as when markers across the
536 genome were used (Fig 4a). Since these QTLs were originally detected in leaves, it was no
537 surprise that the prediction accuracies were not as high when the same models were used to
538 predict CBSD severity on roots (Fig 4b). These data suggest an absence of correlation between
539 root and shoot symptoms in cassava plants affected by CBSD. This phenomenon has been
540 previously described; infected plants may show severe shoot symptoms and mild root necrosis
541 or vice versa [85]. Moreover, the severity of symptoms has been demonstrated not to be
542 correlated with virus titer, especially for resistant or tolerant varieties [85].

543 Previous research has tackled the problem of incorporating genotype-phenotype associations
544 to boost genomic prediction by either adding significant markers as fixed effects [86,87] or by
545 weighting the Genomic Relationship Matrix (GRM) with marker association information [88,89].
546 While we did not focus on any of these methods, tracking known QTLs allowed us to utilize
547 better the information obtained from the transcriptomics experiment.

548 *On using Transcriptomics to Aid Genomic Prediction*

549 Transcriptomics data has been used before as a source of biological priors for genomic
550 prediction in cattle [25,28]. Like in the present study, Fang et al. [25] used transcriptomic regions
551 responsive to Intra Mammary Infection (IMI) to fit a GFBLUP model that included a separate
552 genomic effect of SNPs within DE genes. Similarly, MacLeod et al. used a novel Bayesian method
553 (BayesRC), that allowed the incorporation of biological information by defining classes of
554 variants likely to be enriched for causal mutations [28]. Both studies showed a minimal increase
555 in prediction accuracies for within-breed predictions and a true benefit was observed only with
556 across-breed predictions.

557 In this study, we analyzed existing transcriptomic data using three different approaches to
558 explore multiple hypotheses related to the introgression of transcriptomics into genomic
559 prediction models. The first approach exploited DE genes specific to each measured disease
560 timepoint and cassava genotype (i.e DE genes six hours after infection in Namikonga) to fit a
561 series of GFBLUP models. This approach explored whether any timepoint-genotype combination
562 would be more enriched for causal variants and, thus, more useful for improving prediction
563 accuracies. No increase in prediction accuracy was observed. This result was expected as we did
564 not expect the response of individual genotypes to be representative of the entire population.
565 Further, there were a total of 9,379 DE genes found in at least at one time point; this is close to
566 one-third of the entire predicted gene set in the cassava reference genome.

567 To narrow the number of DE genes, we then hypothesized that genes exhibiting a significant
568 statistical interaction between inoculation status (Control vs. Infected) and genotype
569 ('Namikonga' vs. 'Albert') might be more relevant for CBSD related traits. Only 1,391 genes were
570 significant to $G \times I$ ($q < 0.05$), and, while the multi-kernel GBLUP models performed better than
571 when selecting the same number of random genes, the prediction accuracy remained the same
572 as the full GBLUP model.

573 Finally, we used WGCNA to infer modules of co-expressed genes within the RNAseq dataset.
574 This method has been used in several organisms to identify biologically meaningful gene
575 modules, and it has helped to generate useful insights into how genes interact under certain
576 conditions [66,69,90–92]. We assumed that modules consisting of highly interconnected genes
577 would be enriched in causal variants and promote an increase in prediction accuracy under a
578 GFBLUP framework. Only one module for one trait (red, CBSD3), however, showed a marginal
579 increase in prediction accuracy

580 There are many reasons why we think the approaches using transcriptomics did not result in
581 larger increases in prediction accuracy. First, the RNAseq data came from only two cassava
582 varieties, and its transcriptome response may not be representative of the composite set used
583 in this study. Secondly, samples were collected during the early (i.e., <54 days) response of the
584 plant to the infection. In contrast, the phenotypes were collected in the field three, six, and
585 twelve months after planting. Thirdly, the plants were infected with only UCBSV (as confirmed
586 by de-novo assembly of the viral reads, Fig S11), while under field conditions it is common to
587 observe co-infection of CBSV and UCBSV [93]. Anjanappa et al. [38] previously showed that the
588 response of cassava to a combined CBSV and UCBSV infection was significantly stronger in the
589 susceptible variety than in the resistant variety. These results are in contrast to the current

590 study, where ‘Namikonga’ showed a stronger response when only infected by UCBSV. As such,
591 we can infer that the transcription response of cassava plants infected only with UCBSV may not
592 be representative of infected plants in the field. Fourth, Increasing the accuracy of predictions
593 using closely related individuals with long-range LD might not be an easy task in future breeding
594 efforts. Rather, genomic prediction methods that incorporate biological priors may be more
595 beneficial in across-breed prediction, where the LD structure is disrupted [28,76,82]. Specifically,
596 Fang et al. found only a small increase (3.2% to 3.9%) in prediction accuracies by using GFBLUP
597 and transcriptomics data when predicting milk traits within Holstein cows; the same study
598 observed a 164% gain in prediction accuracy when the prediction was performed across-breeds.

599 Cassava Brown Streak Disease is currently present only in East and Southern Africa. Thus the
600 Western African material cannot be evaluated for resistance to this disease because of the
601 dangers of propagating the disease. In this scenario, a genomic selection model might be trained
602 in the eastern African population(s) to predict resistance to CBSD in western germplasm. While
603 these populations are not as divergent as cattle breeds, we expect that the LD structure between
604 these two populations would be weaker and thus favor a model that uses prior biological
605 information.

606 **Conclusions**

607 The Genomic Prediction approach using prior biological information and markers imputed to
608 whole-genome sequence achieved only a marginal increase in the accuracy of prediction for
609 CBSD related traits. We believe that additional functional genomics research together with
610 bigger reference panels that would improve imputation accuracies and a more precise
611 phenotyping platform are necessary to unlock the potential of biology-assisted prediction
612 models. Moreover, we think that this kind of novel approaches would provide insights into the
613 genetic mechanisms underlying quantitative traits.

614

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618

619 Abbreviations

620 **GS:** Genomic Selection **GWAS:** Genome-Wide Association Studies **CBSD:** Cassava Brown Streak
621 Disease; **CBSV:** Cassava Brown Streak Virus **UCBSV:** Ugandan Cassava Brown Streak Virus **GBS:**
622 Genotyping-By-Sequencing **BLUP:** Best Linear Unbiased Prediction **GEBV:** Genomic Estimated
623 Breeding Values **LD:** Linkage Disequilibrium **SNP:** Single Nucleotide Polymorphism **DE:**
624 Differentially Expressed; **EGV** Estimated Genetic Value

625 Declarations

626 Ethics approval and consent to participate

627 Not applicable.

628

629 Consent for publication

630 Not applicable.

631

632 Availability of data and material

633 Sequences for every gene presented in this article are available in the Phytozome v10.1
634 repository, <http://phytozome.jgi.doe.gov> (*Manihot esculenta* v6.1). Scripts used in this
635 manuscript are available at, https://github.com/tc-mustang/CBSD_Transcriptomics. Dosage
636 matrices and Variant Call Format (VCF) files can be accessed upon request through a secure FTP
637 server. The transcriptome data from Amuge et al. [37] is available in the SRA BioProject ID
638 PRJNA360340.

639 Competing Interests

640 The authors declare that they have no competing interests.

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645 Authors’ contributions

646 The study was conceived and designed by RL, DPC and JLJ. AO and IK were in charged of data
647 collection. MF and TA advised on CBSD and performed the transcriptomics study. The data
648 analysis was performed by RL. The manuscript was written by RL and DPC. JLJ critically revised
649 the manuscript with important scientific and statistical content All authors read and approved
650 the final manuscript.

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654

655 References

656 1. Meuwissen TH, Hayes BJ, Goddard ME. Prediction of total genetic value using genome-wide
657 dense marker maps. *Genetics*. 2001;157:1819–29.

658 2. Desta ZA, Ortiz R. Genomic selection: genome-wide prediction in plant improvement. *Trends*
659 *Plant Sci*. Elsevier Ltd; 2014;19:592–601.

660 3. Jannink J-L, Lorenz AJ, Iwata H. Genomic selection in plant breeding: from theory to practice.
661 *Brief. Funct. Genomics*. 2010;9:166–77.

662 4. de Los Campos G, Hickey JM, Pong-Wong R, Daetwyler HD, Calus MPL. Whole-genome
663 regression and prediction methods applied to plant and animal breeding. *Genetics*. *Genetics*
664 *Society of America*; 2013;193:327–45.

665 5. Kolbehdari D, Schaeffer LR, Robinson JAB. Estimation of genome-wide haplotype effects in
666 half-sib designs. *J. Anim. Breed. Genet*. 2007;124:356–61.

667 6. de los Campos G, Naya H, Gianola D, Crossa J, Legarra A, Manfredi E, et al. Predicting
668 quantitative traits with regression models for dense molecular markers and pedigree.
669 *Genetics*. 2009;182:375–85.

670 7. Hoerl AE, Kennard RW. Ridge Regression: Biased Estimation for Nonorthogonal Problems.
671 Taylor & Francis Group; 2012;

672 8. Long N, Gianola D, Rosa GJM, Weigel KA, Avendaño S. Machine learning classification
673 procedure for selecting SNPs in genomic selection: application to early mortality in broilers. *J.*
674 *Anim. Breed. Genet*. 2007;124:377–89.

675 9. Machine learning methods and predictive ability metrics for genome-wide prediction of
676 complex traits. *Livest. Sci. Elsevier*; 2014;166:217–31.

677 10. Heslot N, Yang H-P, Sorrells ME, Jannink J-L. Genomic Selection in Plant Breeding: A
678 Comparison of Models. *Crop Sci. The Crop Science Society of America, Inc.*; 2012;52:146.

- 679 11. Ornella L, Singh S, Perez P, Burgueño J, Singh R, Tapia E, et al. Genomic Prediction of
680 Genetic Values for Resistance to Wheat Rusts. *Plant Genome J.* 2012;5:136.
- 681 12. Rutkoski JE, Heffner EL, Sorrells ME. Genomic selection for durable stem rust resistance in
682 wheat. *Euphytica.* 2010;179:161–73.
- 683 13. Wolfe MD, Del Carpio DP, Alabi O, Ezenwaka LC, Ikeogu UN, Kayondo IS, et al. Prospects for
684 Genomic Selection in Cassava Breeding. *Plant Genome. Crop Science Society of America;*
685 2017;0:0.
- 686 14. Meuwissen T, Goddard M. Accurate Prediction of Genetic Values for Complex Traits by
687 Whole-Genome Resequencing. *Genetics.* 2010;185:623–31.
- 688 15. Druet T, Macleod IM, Hayes BJ. Toward genomic prediction from whole-genome sequence
689 data: impact of sequencing design on genotype imputation and accuracy of predictions.
690 *Heredity (Edinb).* Nature Publishing Group; 2014;112:39–47.
- 691 16. Clark SA, Hickey JM, van der Werf JH. Different models of genetic variation and their effect
692 on genomic evaluation. *Genet. Sel. Evol.* 2011;43:18.
- 693 17. Yan G, Qiao R, Zhang F, Xin W, Xiao S, Huang T, et al. Imputation-Based Whole-Genome
694 Sequence Association Study Rediscovered the Missing QTL for Lumbar Number in Sutan Pigs.
695 *Sci. Rep.* 2017;7:615.
- 696 18. Auton A, Abecasis GR, Altshuler DM, Durbin RM, Abecasis GR, Bentley DR, et al. A global
697 reference for human genetic variation. *Nature.* 2015;526:68–74.
- 698 19. Calus MPL, Bouwman AC, Schrooten C, Veerkamp RF. Efficient genomic prediction based
699 on whole-genome sequence data using split-and-merge Bayesian variable selection. *Genet.*
700 *Sel. Evol. BioMed Central;* 2016;48:49.
- 701 20. MacLeod IM, Bowman PJ, Vander Jagt CJ, Haile-Mariam M, Kemper KE, Chamberlain AJ, et
702 al. Exploiting biological priors and sequence variants enhances QTL discovery and genomic
703 prediction of complex traits. *BMC Genomics.* 2016;17:144.
- 704 21. Fortes MRS, Reverter A, Zhang Y, Collis E, Nagaraj SH, Jonsson NN, et al. Association weight
705 matrix for the genetic dissection of puberty in beef cattle. *Proc. Natl. Acad. Sci. U. S. A.*
706 2010;107:13642–7.

- 707 22. Snelling WM, Cushman RA, Keele JW, Maltecca C, Thomas MG, Fortes MRS, et al. Breeding
708 and Genetics Symposium: networks and pathways to guide genomic selection. *J. Anim. Sci.*
709 2013;91:537–52.
- 710 23. Su G, Christensen OF, Janss L, Lund MS. Comparison of genomic predictions using genomic
711 relationship matrices built with different weighting factors to account for locus-specific
712 variances. *J. Dairy Sci.* 2014;97:6547–59.
- 713 24. de Los Campos G, Vazquez AI, Fernando R, Klimentidis YC, Sorensen D. Prediction of
714 complex human traits using the genomic best linear unbiased predictor. *PLoS Genet. Public*
715 *Library of Science*; 2013;9:e1003608.
- 716 25. Fang L, Sahana G, Ma P, Su G, Yu Y, Zhang S, et al. Exploring the genetic architecture and
717 improving genomic prediction accuracy for mastitis and milk production traits in dairy cattle by
718 mapping variants to hepatic transcriptomic regions responsive to intra-mammary infection.
719 *Genet. Sel. Evol.* 2017;49:44.
- 720 26. Speed D, Balding DJ. MultiBLUP: improved SNP-based prediction for complex traits.
721 *Genome Res. Cold Spring Harbor Laboratory Press*; 2014;24:1550–7.
- 722 27. Kemper KE, Reich CM, Bowman PJ, Vander Jagt CJ, Chamberlain AJ, Mason BA, et al.
723 Improved precision of QTL mapping using a nonlinear Bayesian method in a multi-breed
724 population leads to greater accuracy of across-breed genomic predictions. *Genet. Sel. Evol.*
725 2015;47:29.
- 726 28. MacLeod IM, Bowman PJ, Vander Jagt CJ, Haile-Mariam M, Kemper KE, Chamberlain AJ, et
727 al. Exploiting biological priors and sequence variants enhances QTL discovery and genomic
728 prediction of complex traits. *BMC Genomics. BioMed Central*; 2016;17:144.
- 729 29. Edwards SM, Sørensen IF, Sarup P, Mackay TFC, Sørensen P. Genomic Prediction for
730 Quantitative Traits Is Improved by Mapping Variants to Gene Ontology Categories in
731 *Drosophila melanogaster*. *Genetics.* 2016;203.
- 732 30. Lowe R, Shirley N, Bleackley M, Dolan S, Shafee T. Transcriptomics technologies. *PLoS*
733 *Comput. Biol. Public Library of Science*; 2017;13:e1005457.
- 734 31. Costa V, Aprile M, Esposito R, Ciccodicola A. RNA-Seq and human complex diseases: recent
735 accomplishments and future perspectives. *Eur. J. Hum. Genet.* 2013;21:134–42.

- 736 32. Fauquet C, Fargette D, Munihor C. African Cassava Mosaic Virus : Etiology , Epidemiology ,
737 and Control. 1990;74.
- 738 33. Monger WA, Alicai T, Ndunguru J, Kinyua ZM, Potts M, Reeder RH, et al. The complete
739 genome sequence of the Tanzanian strain of Cassava brown streak virus and comparison with
740 the Ugandan strain sequence. Arch. Virol. Springer Vienna; 2010;155:429–33.
- 741 34. Ndunguru J, Sseruwagi P, Tairo F, Stomeo F, Maina S, Djinkeng A, et al. Analyses of Twelve
742 New Whole Genome Sequences of Cassava Brown Streak Viruses and Ugandan Cassava Brown
743 Streak Viruses from East Africa: Diversity, Supercomputing and Evidence for Further
744 Speciation. Melcher U, editor. PLoS One. Public Library of Science; 2015;10:e0139321.
- 745 35. Maruthi MN, Hillocks RJ, Mtunda K, Raya MD, Muhanna M, Kiozia H, et al. Transmission of
746 Cassava brown streak virus by Bemisia tabaci (Gennadius). J. Phytopathol. Blackwell Verlag
747 GmbH; 2005;153:307–12.
- 748 36. Mware B, Narla R, Amata R, Olubayo F, Songa J, Kyamanyua S, et al. Journal of General and
749 Molecular Virology. J. Gen. Mol. Virol. Academic Journals; 2009.
- 750 37. Amuge T, Berger DK, Katari MS, Myburg AA, Goldman SL, Ferguson ME. A time series
751 transcriptome analysis of cassava (Manihot esculenta Crantz) varieties challenged with
752 Ugandan cassava brown streak virus. Sci. Rep. 2017;7:9747.
- 753 38. Anjanappa RB, Mehta D, Okoniewski MJ, Szabelska A, Gruissem W, Vanderschuren H.
754 Molecular insights into cassava brown streak virus susceptibility and resistance by profiling of
755 the early host response. Mol. Plant Pathol. 2017;
- 756 39. Kayondo SI, Pino Del Carpio D, Lozano R, Ozimati A, Wolfe MD, Baguma Y, et al. Genome-
757 wide association mapping and genomic prediction unravels CBSB resistance in a Manihot
758 esculenta breeding population. bioRxiv. 2017;
- 759 40. Lozano R, Hamblin MT, Prochnik S, Jannink J-L. Identification and distribution of the NBS-
760 LRR gene family in the Cassava genome. BMC Genomics. 2015;16:360.
- 761 41. Soto JC, Ortiz JF, Perlaza-Jiménez L, Vásquez AX, Lopez-Lavalle LAB, Mathew B, et al. A
762 genetic map of cassava (Manihot esculenta Crantz) with integrated physical mapping of
763 immunity-related genes. BMC Genomics. ???; 2015;16:190.
- 764 42. Elshire RJ, Glaubitz JC, Sun Q, Poland JA, Kawamoto K, Buckler ES, et al. A robust, simple

- 765 genotyping-by-sequencing (GBS) approach for high diversity species. PLoS One. Public Library
766 of Science; 2011;6:e19379.
- 767 43. Hamblin MT, Rabbi IY. The Effects of Restriction-Enzyme Choice on Properties of
768 Genotyping-by-Sequencing Libraries: A Study in Cassava (). Crop Sci. The Crop Science Society
769 of America, Inc.; 2014;54:2603.
- 770 44. Glaubitz JC, Casstevens TM, Lu F, Harriman J, Elshire RJ, Sun Q, et al. TASSEL-GBS: a high
771 capacity genotyping by sequencing analysis pipeline. PLoS One. Public Library of Science;
772 2014;9:e90346.
- 773 45. Danecek P, Auton A, Abecasis G, Albers CA, Banks E, DePristo MA, et al. The variant call
774 format and VCFtools. Bioinformatics. 2011;27:2156–8.
- 775 46. Browning BL, Browning SR. Genotype Imputation with Millions of Reference Samples. Am.
776 J. Hum. Genet. 2016;98:116–26.
- 777 47. Howie B, Marchini J, Stephens M. Genotype Imputation with Thousands of Genomes. G3
778 Genes, Genomes, Genet. 2011;1.
- 779 48. Howie BN, Donnelly P, Marchini J. A Flexible and Accurate Genotype Imputation Method
780 for the Next Generation of Genome-Wide Association Studies. Schork NJ, editor. PLoS Genet.
781 Public Library of Science; 2009;5:e1000529.
- 782 49. Ramu P, Esuma W, Kawuki R, Rabbi IY, Egesi C, Bredeson J V, et al. Cassava haplotype map
783 highlights fixation of deleterious mutations during clonal propagation. Nat. Genet.
784 2017;49:959–63.
- 785 50. International Cassava Genetic Map Consortium (ICGMC). High-resolution linkage map and
786 chromosome-scale genome assembly for cassava (*Manihot esculenta* Crantz) from 10
787 populations. G3 (Bethesda). 2014;5:133–44.
- 788 51. Simon Andrews. FastQC A Quality Control tool for High Throughput Sequence Data
789 [Internet]. 2010 [cited 2017 May 24]. Available from:
790 <http://www.bioinformatics.babraham.ac.uk/projects/fastqc/>
- 791 52. Trapnell C, Williams BA, Pertea G, Mortazavi A, Kwan G, van Baren MJ, et al. Transcript
792 assembly and quantification by RNA-Seq reveals unannotated transcripts and isoform
793 switching during cell differentiation. Nat. Biotechnol. 2010;28:511–5.

- 794 53. Trapnell C, Pachter L, Salzberg SL. TopHat: discovering splice junctions with RNA-Seq.
795 *Bioinformatics*. 2009;25:1105–11.
- 796 54. Bredeson J V, Lyons JB, Prochnik SE, Wu GA, Ha CM, Edsinger-Gonzales E, et al. Sequencing
797 wild and cultivated cassava and related species reveals extensive interspecific hybridization
798 and genetic diversity. *Nat. Biotechnol*. 2016;34:562–70.
- 799 55. Kim D, Pertea G, Trapnell C, Pimentel H, Kelley R, Salzberg SL. TopHat2: accurate alignment
800 of transcriptomes in the presence of insertions, deletions and gene fusions. *Genome Biol*.
801 2013;14:R36.
- 802 56. Langmead B, Salzberg SL. Fast gapped-read alignment with Bowtie 2. *Nat. Methods*.
803 2012;9:357–9.
- 804 57. Trapnell C, Hendrickson DG, Sauvageau M, Goff L, Rinn JL, Pachter L. Differential analysis of
805 gene regulation at transcript resolution with RNA-seq. *Nat. Biotechnol*. 2012;31:46–53.
- 806 58. Nzuki I, Katari MS, Bredeson J V, Masumba E, Kapinga F, Salum K, et al. QTL Mapping for
807 Pest and Disease Resistance in Cassava and Coincidence of Some QTL with Introgression
808 Regions Derived from *Manihot glaziovii*. *Front. Plant Sci. Frontiers Media SA*; 2017;8:1168.
- 809 59. Masumba EA, Kapinga F, Mkamilo G, Salum K, Kulembeka H, Rounsley S, et al. QTL
810 associated with resistance to cassava brown streak and cassava mosaic diseases in a bi-
811 parental cross of two Tanzanian farmer varieties, Namikonga and Albert. *Theor. Appl. Genet.*
812 *Springer Berlin Heidelberg*; 2017;130:2069–90.
- 813 60. Meyers BC, Dickerman a W, Michelmore RW, Sivaramakrishnan S, Sobral BW, Young ND.
814 Plant disease resistance genes encode members of an ancient and diverse protein family
815 within the nucleotide-binding superfamily. *Plant J*. 1999;20:317–32.
- 816 61. Lozano R, Hamblin MT, Prochnik S, Jannink J-L. Identification and distribution of the NBS-
817 LRR gene family in the Cassava genome. *BMC Genomics*. 2015;16:360.
- 818 62. Camacho C, Coulouris G, Avagyan V, Ma N, Papadopoulos J, Bealer K, et al. BLAST+:
819 architecture and applications. *BMC Bioinformatics*. 2009;10:421.
- 820 63. Quinlan AR, Quinlan, R. A. BEDTools: The Swiss-Army Tool for Genome Feature Analysis.
821 *Curr. Protoc. Bioinforma*. Hoboken, NJ, USA: John Wiley & Sons, Inc.; 2014. p. 11.12.1-
822 11.12.34.

- 823 64. Quinlan AR, Hall IM. BEDTools: a flexible suite of utilities for comparing genomic features.
824 Bioinformatics. Oxford University Press; 2010;26:841–2.
- 825 65. Zhang B, Horvath S. A general framework for weighted gene co-expression network
826 analysis. Stat. Appl. Genet. Mol. Biol. 2005;4:Article17.
- 827 66. Childs KL, Davidson RM, Buell CR, Coghill P, Sammut S. Gene Coexpression Network
828 Analysis as a Source of Functional Annotation for Rice Genes. El-Sayed NM, editor. PLoS One.
829 Public Library of Science; 2011;6:e22196.
- 830 67. Langfelder P, Horvath S. WGCNA: an R package for weighted correlation network analysis.
831 BMC Bioinformatics. 2008;9:559.
- 832 68. Langfelder P, Horvath S. Eigengene networks for studying the relationships between co-
833 expression modules. BMC Syst. Biol. 2007;1:54.
- 834 69. Massa AN, Childs KL, Lin H, Bryan GJ, Giuliano G, Buell CR. The Transcriptome of the
835 Reference Potato Genome Solanum tuberosum Group Phureja Clone DM1-3 516R44. Zhang J,
836 editor. PLoS One. Public Library of Science; 2011;6:e26801.
- 837 70. Garrick DJ, Taylor JF, Fernando RL. Deregressing estimated breeding values and weighting
838 information for genomic regression analyses. Genet. Sel. Evol. 2009;41:55.
- 839 71. Wolfe MD, Rabbi IY, Egesi C, Hamblin M, Kawuki R, Kulakow P, et al. Genome-Wide
840 Association and Prediction Reveals Genetic Architecture of Cassava Mosaic Disease Resistance
841 and Prospects for Rapid Genetic Improvement. Plant Genome. 2016;9:0.
- 842 72. Bates D, Mächler M, Bolker B, Walker S. Fitting Linear Mixed-Effects Models Using **lme4**. J.
843 Stat. Softw. 2015;67:1–48.
- 844 73. Eddelbuettel D, François R. **Rcpp** : Seamless R and C++ Integration. J. Stat. Softw.
845 2011;40:1–18.
- 846 74. Endelman JB. Ridge Regression and Other Kernels for Genomic Selection with R Package
847 rrBLUP. Plant Genome J. Crop Science Society of America; 2011;4:250.
- 848 75. Akdemir D, Okeke UG. EMMREML: Fitting Mixed Models with Known Covariance
849 Structures. <https://cran.r-project.org/package=EMMREML>. 2015;R package version 3.1.

- 850 76. Fang L, Sahana G, Ma P, Su G, Yu Y, Zhang S, et al. Exploring the genetic architecture and
851 improving genomic prediction accuracy for mastitis and milk production traits in dairy cattle by
852 mapping variants to hepatic transcriptomic regions responsive to intra-mammary infection.
853 *Bioinformatics*. *BioMed Central*; 2017;49:44.
- 854 77. Bulik-Sullivan BK, Loh P-R, Finucane HK, Ripke S, Yang J, Consortium SWG of the PG, et al.
855 LD Score regression distinguishes confounding from polygenicity in genome-wide association
856 studies. *Nat Genet*. Nature Publishing Group; 2015;advance on:1–7.
- 857 78. Ma P, Brøndum RF, Zhang Q, Lund MS, Su G. Comparison of different methods for imputing
858 genome-wide marker genotypes in Swedish and Finnish Red Cattle. *J. Dairy Sci*. 2013;96:4666–
859 77.
- 860 79. Haas BJ, Papanicolaou A, Yassour M, Grabherr M, Blood PD, Bowden J, et al. De novo
861 transcript sequence reconstruction from RNA-seq using the Trinity platform for reference
862 generation and analysis. *Nat. Protoc*. 2013;8:1494–512.
- 863 80. Heidaritabar M, Calus MPL, Megens H-J, Vereijken A, Groenen MAM, Bastiaansen JWM.
864 Accuracy of genomic prediction using imputed whole-genome sequence data in white layers. *J*.
865 *Anim. Breed. Genet*. 2016;133:167–79.
- 866 81. van Binsbergen R, Calus MPL, Bink MCAM, van Eeuwijk FA, Schrooten C, Veerkamp RF.
867 Genomic prediction using imputed whole-genome sequence data in Holstein Friesian cattle.
868 *Genet. Sel. Evol. BioMed Central*; 2015;47:71.
- 869 82. Veerkamp RF, Bouwman AC, Schrooten C, Calus MPL. Genomic prediction using
870 preselected DNA variants from a GWAS with whole-genome sequence data in Holstein-Friesian
871 cattle. *Genet. Sel. Evol. BioMed Central*; 2016;48:95.
- 872 83. Ni G, Cavero D, Fangmann A, Erbe M, Simianer H. Whole-genome sequence-based genomic
873 prediction in laying chickens with different genomic relationship matrices to account for
874 genetic architecture. *Genet. Sel. Evol. BioMed Central*; 2017;49:8.
- 875 84. Spindel J, Begum H, Akdemir D, Virk P, Collard B, Redoña E, et al. Genomic Selection and
876 Association Mapping in Rice (*Oryza sativa*): Effect of Trait Genetic Architecture, Training
877 Population Composition, Marker Number and Statistical Model on Accuracy of Rice Genomic
878 Selection in Elite, Tropical Rice Breeding Lines. Mauricio R, editor. *PLOS Genet*. International
879 Rice Research Institute; 2015;11:e1004982.

- 880 85. Kaweesi T, Kawuki R, Kyaligonza V, Baguma Y, Tusiime G, Ferguson ME. Field evaluation of
881 selected cassava genotypes for cassava brown streak disease based on symptom expression
882 and virus load. *Virologica Sinica*. 2014;11:216.
- 883 86. Bian Y, Holland JB. Enhancing genomic prediction with genome-wide association studies in
884 multiparental maize populations. *Heredity (Edinb)*. Nature Publishing Group; 2017;118:585–
885 93.
- 886 87. Spindel JE, Begum H, Akdemir D, Collard B, Redoña E, Jannink J-L, et al. Genome-wide
887 prediction models that incorporate de novo GWAS are a powerful new tool for tropical rice
888 improvement. *Heredity (Edinb)*. Nature Publishing Group; 2016;116:395–408.
- 889 88. Fragomeni BO, Lourenco DAL, Masuda Y, Legarra A, Misztal I. Incorporation of causative
890 quantitative trait nucleotides in single-step GBLUP. *Genet. Sel. Evol.* 2017;49:59.
- 891 89. Lee J, Cheng H, Garrick D, Golden B, Dekkers J, Park K, et al. Comparison of alternative
892 approaches to single-trait genomic prediction using genotyped and non-genotyped Hanwoo
893 beef cattle. *Genet. Sel. Evol. BioMed Central*; 2017;49:2.
- 894 90. Botía JA, Vandrovцова J, Forabosco P, Guelfi S, D'Sa K, Hardy J, et al. An additional k-means
895 clustering step improves the biological features of WGCNA gene co-expression networks. *BMC*
896 *Syst. Biol.* 2017;11:47.
- 897 91. Forabosco P, Ramasamy A, Trabzuni D, Walker R, Smith C, Bras J, et al. Insights into TREM2
898 biology by network analysis of human brain gene expression data. *Neurobiol. Aging*.
899 2013;34:2699–714.
- 900 92. Ballouz S, Verleyen W, Gillis J. Guidance for RNA-seq co-expression network construction
901 and analysis: safety in numbers. *Bioinformatics*. 2015;31:2123–30.
- 902 93. Ogwok E, Alicai T, Rey MEC, Beyene G, Taylor NJ. Distribution and accumulation of cassava
903 brown streak viruses within infected cassava (*Manihot esculenta*) plants. *Plant Pathol.*
904 2015;64:1235–46.
- 905
- 906
- 907

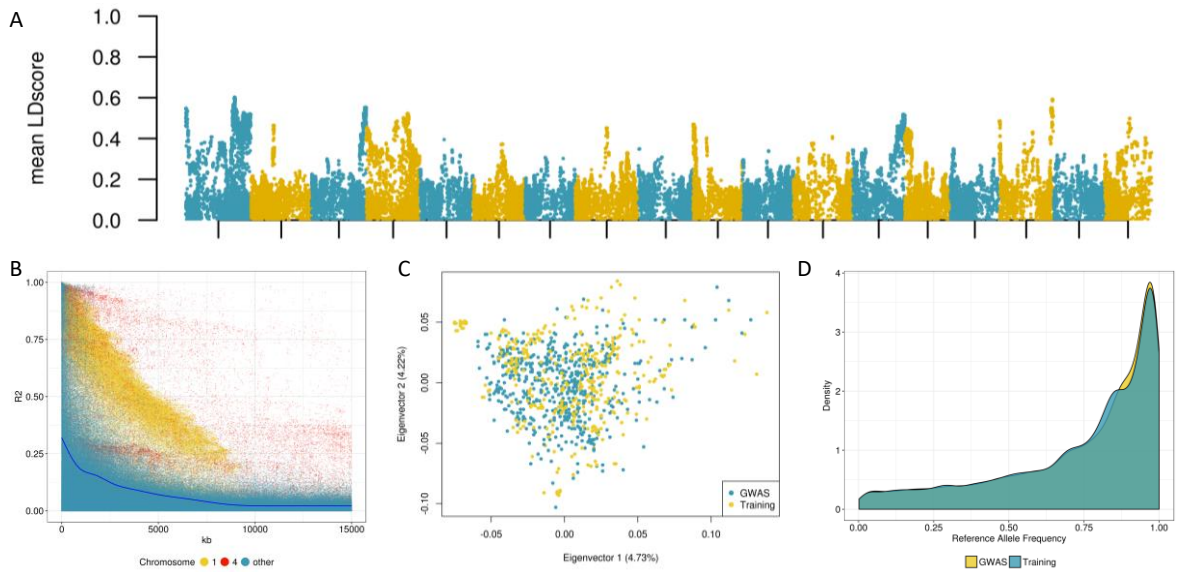


Fig 1.- Describing the Breeding Population. **a** Local LD patterns across the 18 cassava chromosomes as depicted by the mean LDscore of each marker. **b** LD decay plot, A random subset of all the r^2 values of SNPs closer than 15Mb were plotted. Chromosomes 1 and 4 were plotted separately to highlight the distortion in their LD patterns due to the introgressions. **c** Principal component analysis using the SNP marker matrix, the two breeding populations that were merged in this study are shown in different colors. **d** Distribution of the reference allele frequencies between the two breeding populations.

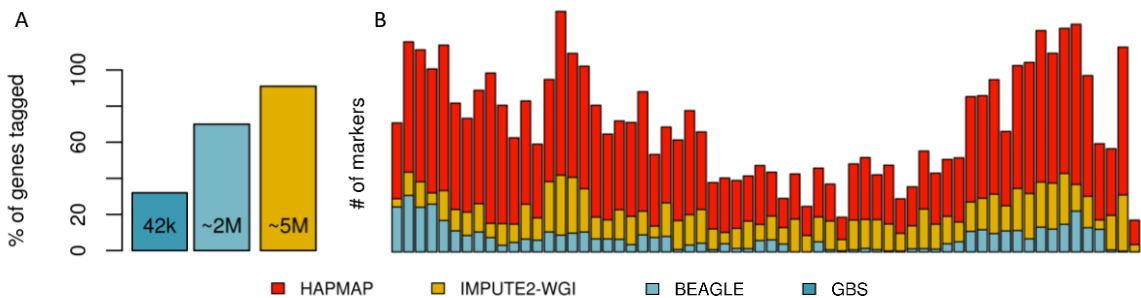


Fig 2.- Imputation to whole-genome sequence. **a** Percentage of genes “tagged” using different SNP marker sets, the numbers inside the plots represents the number of markers. All markers considered had a MAF higher than 1% and an imputation quality value AR2/info higher than 0.3 **b** Marker distribution across chromosome 12, each bar represents a bin of 0.5Mb. The red colored bars represents the “true” distribution of variability as reported in the cassava HAPMAP, in orange, the distribution of the IMPUTE2 dataset (~5M markers) and in blue the Beagle dataset (~2M markers).

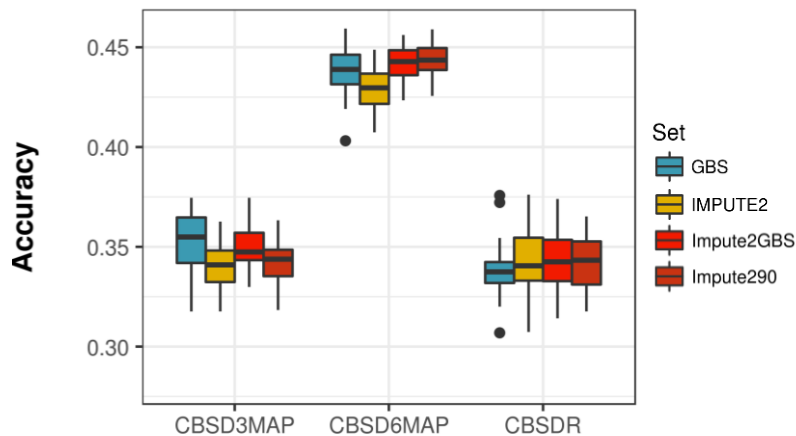


Fig 3.- Impact of Imputation level on Genomic Prediction Accuracies. Comparing prediction accuracies for three traits; CBSD severity on leaves 3 months after planting (CBSD3MAP), 6 months after planting (CBSD6MAP) and CBSD severity on roots one year after planting (CBSDR) when using GBS (42k SNPs), the whole-genome sequence imputed datasets using IMPUTE2 (~5M) and also prediction accuracies for a subset of the IMPUTE2 markers matching the position of the GBS set (Impute2GBS) and only marker with an “info” imputation quality score higher than 0.9 (Impute290)

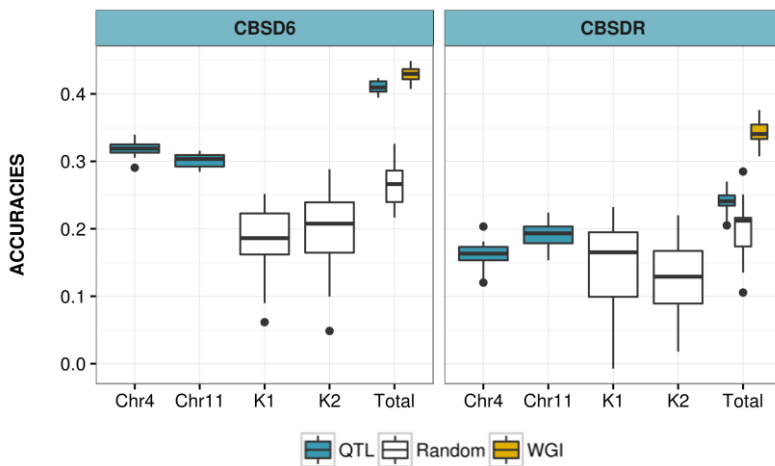


Fig 4.- Accounting for the effect of previously reported QTLs. Comparing the maximum accuracy using whole-genome imputation (yellow) with two kernel GBLUP model using chromosome 4 and 11 only (blue) and random chromosomes excluding 11 and 4 (white) in each cross-validation iteration. Partial accuracies are shown under Chr4 , Chr11, K1 and K2. Full model prediction accuracies are shown in “Total”. *CBSD6MAP: Foliar symptoms, CBSDR: Root symptoms.

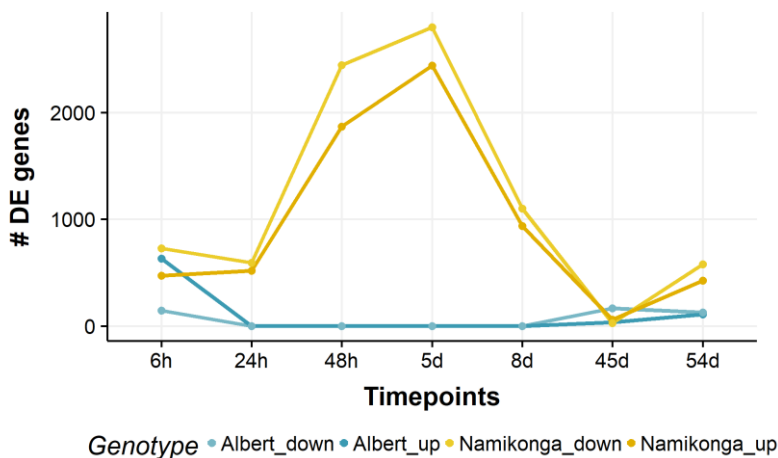


Fig 5.- Transcriptional response to Infection with UCBSV. The test for differentially expressed genes was conducted at each timepoint between the infected and control plants using Cuffdiff. Genes considered to be differentially expressed had a q-value < 0.01 (Benjamini-Hochberg correction for multiple testing). *h = hours after infection, d = days after infection

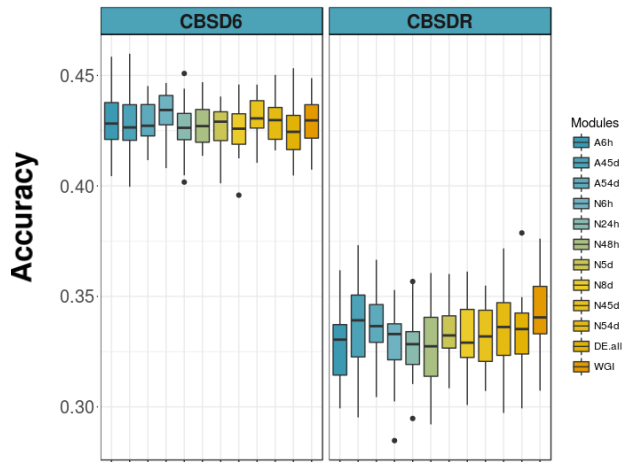


Fig 6.- Using DE genes for Genomic Prediction. GFBLUP models (Two-kernel GBLUP) were fitted. For each model the Genomic feature kernel comprised SNPs inside the genes that were DE at each time point for each genotype. Three models for the susceptible DE genes (A6h, A45d and A54d), seven for the tolerant (N6h – N54d) and one for the combined DE genes (DE.all) were performed. Boxplot were the result of 25 replications of 5-fold cross-validation.
 *A = Albert, N = Namikonga, h = hours after infection, d = days after infection

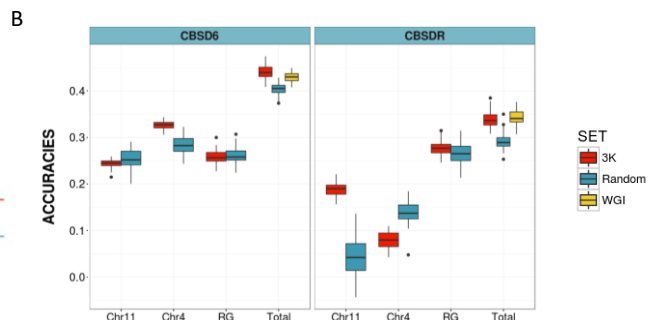
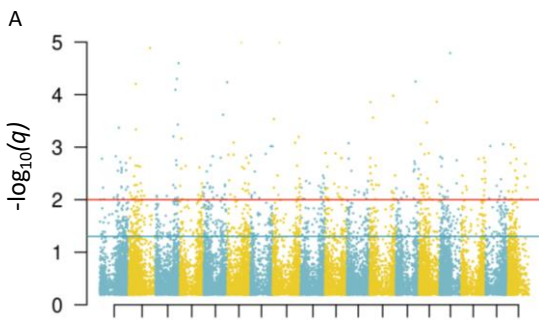


Fig 7.- Filtering DE genes. **a** A linear mix model was used to calculate the genes that showed a significant interaction between inoculation status and genotype. In the manhattan plot $-\log_{10}$ “q-values” (FDR corrected p-values) for the $G \times I$ interaction term was plotted for the 18 cassava chromosomes. The blue line is the threshold for 5% FDR and the red one for 1% FDR **b** Genomic predictions using three kernel GBLUP models. In red, the partial prediction accuracies (Chr11, Chr4 and RG) and total accuracy using only markers associated with significant $G \times I$ genes are compared with a three kernel model of random SNPs in blue and the regular single kernel GBLUP prediction using all the markers in yellow.

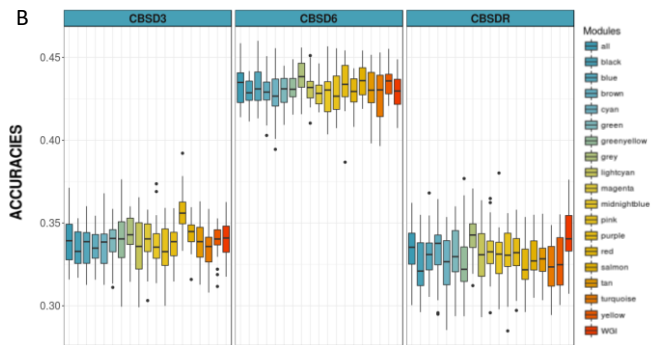
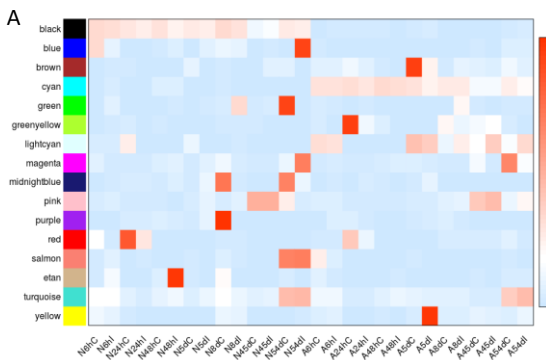


Fig 8.- Co-expression network analysis. **a** Heatmap of eigengenes representing each co-expression module as obtained by WGCNA. All timepoints for both genotypes including controls were included and presented as columns. The 16 identified co-expression modules are presented in each row. The eigengene values are a relative measure of expression levels of the genes in the module. **b** GFBLUP predictions using the modules information. As in figure 6 the genes in each module were used to build a GFBLUP model, one kernel using SNPs within each module genes and the other covering the rest of the genome. Total prediction accuracies were plotted.

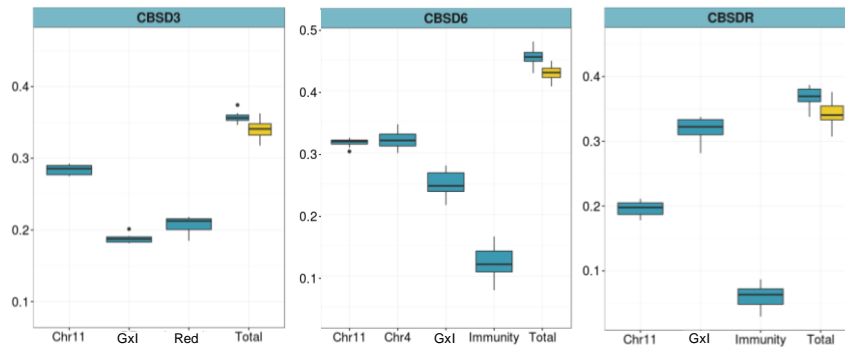


Fig 9.- Combining sources of evidence. Four and three kernel GBLUP model including markers surrounding previously reported QTLs (chr4 and chr11), Gxl genes found in this study, immunity related genes and the red WGCNA module (blue). Partial and Total accuracies are compared with the regular GBLUP model (yellow). A nominal increase in prediction accuracy of 1.7%, 2.52% and 2.5% was found for CBSD3, CBSD6 and CBSDR respectively.