

1 Spatial autocorrelation of the environment
2 influences the patterns and genetics of
3 local adaptation

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11 Abstract

12 Environmental heterogeneity can lead to spatially varying selection, which can, in turn, lead to local
13 adaptation. Population genetic models have shown that the pattern of environmental variation in space
14 can strongly influence the evolution of local adaptation. In particular, when environmental variation is
15 highly autocorrelated in space local adaptation will more readily evolve. Despite this long-held prediction,
16 the evolutionary genetic consequences of different patterns of environmental variation have not been
17 thoroughly explored. In this study, simulations are used to model local adaptation to different patterns of
18 environmental variation. The simulations confirm that local adaptation is expected to increase with the
19 degree of spatial autocorrelation in the selective environment, but also show that highly heterogeneous
20 environments are more likely to exhibit high variation in local adaptation, a result not previously
21 described. Spatial autocorrelation in the environment also influences the evolution and genetic
22 architecture of local adaptation, with different combinations of allele frequency and effect size arising
23 under different patterns of environmental variation. These differences influence the ability to
24 characterise the genetic basis of local adaptation in different environments. Finally, I analyse a large-scale
25 provenance trial conducted on lodgepole pine and find suggestive evidence that spatially autocorrelated
26 environmental variation leads to stronger local adaptation in natural populations of lodgepole pine.
27 Overall, this work emphasizes the profound importance that the spatial pattern of selection can have on
28 the evolution of local adaptation and how spatial autocorrelation should be considered when formulating
29 hypotheses in ecological and genetic studies.

30 Lay Summary

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32 Many species exhibit local adaptation to environmental variation across their ranges. Theoretical
33 population genetics predicts that the evolution of local adaptation and patterns of genetic variation
34 underlying it will be influenced by the spatial pattern of variation across a species' range. However, this
35 prediction has not been thoroughly explored for cases of complex heterogeneous landscapes. In this
36 paper, I analyse simulations and empirical data to characterise the effects that the spatial pattern of
37 environmental variation can have on the evolution of local adaptation and the genetics underlying it.
38 From these analyses, I show that the pattern of environmental variation influences the average level of
39 local adaptation, variation in local adaptation as well as the genetics underlying this important
40 phenomenon.

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42 Introduction

43 Local adaptation is an important phenomenon in the natural world. Along with phenotypic plasticity,
44 local adaptation can dictate the extent of environmental heterogeneity that a species can tolerate,
45 shape its geographic range (Kirkpatrick & Barton, 1997), and help predict how it will respond to changing
46 environments (Rellstab et al., 2021). In forest trees, for example, local adaptation is widely observed
47 (Leites & Benito Garzón, 2023) and is central to plans for adapting forestry practice in light of climate
48 change (O'Neill & Gómez-Pineda, 2021; Ying & Yanchuk, 2006). Local adaptation can be defined as a
49 kind of genotype-by-environment interaction for fitness, where individuals have higher chances of
50 survival and/or reproduction when they are reared at home as opposed to away, though several other
51 definitions are used in the literature (Blanquart et al., 2013; Kawecki & Ebert, 2004). Local adaptation is
52 a property of a particular population at a particular point in time rather than a property of a species as a
53 whole. For example, populations at range edges are often expected to be maladapted to their
54 conditions, while populations in the core of a range may be well adapted (Angert et al., 2020). Locally
55 adapted populations may harbour genetic variation that could help buffer susceptible ones against the
56 detrimental effects of climate change (Aitken & Whitlock, 2013), which are already wreaking havoc on
57 important species around the world (Hartmann et al., 2022). A deep understanding of local adaptation,
58 the agents that have given rise to it and the genetics that underpin this phenomenon is thus important
59 for our understanding of biodiversity and for species management and conservation in the
60 Anthropocene (Aitken & Whitlock, 2013; Exposito-Alonso, 2023; Wadgymer et al., 2022).

61
62 The ultimate cause of local adaptation is variation in the environment. Whether it is biotic (e.g.
63 disease/parasite prevalence or intraspecific competition) or abiotic (e.g. climate, geology or
64 photoperiod), variation in the environment may lead to spatially varying selection pressures where
65 phenotypic optima differ over a landscape. Such variation in selection across space has been well
66 documented (e.g. Siepielski et al., 2013) and there are, of course, myriad aspects of the environment
67 that could conceivably induce spatially varying selection, many of which would be highly inter-
68 correlated. However, while there is an infinite number of ways to describe the environment, most of
69 these may be functionally disconnected from a species' biology. A recent review by Wadgymer et al
70 (2022) highlighted a critical gap in our knowledge of local adaptation - that the aspects of environmental
71 variation that have given rise to local adaptation (what they term the 'agents of selection') are unknown
72 in most cases. For example, in the absence of experimental evidence many genetic studies have
73 assumed that various climatic measures recorded in databases such as WorldClim correspond to agents
74 of selection and search for the genetic basis of local adaptation using those data (Lasky et al., 2023).
75 However, that a particular aspect of environmental variation could conceivably induce spatially varying
76 selection is not a guarantee that it will have led to local adaptation.

77
78 Population genetic studies have revealed numerous factors that can influence the evolution of local
79 adaptation in a particular location, the most important being the strength of selection and rates of gene
80 flow. The strength of natural selection is of foremost importance because larger fitness consequences
81 for deviating from the optimal phenotype in a particular location can potentially lead to greater
82 evolutionary change (Falconer & MacKay, 1995). The rate of gene flow is important because migration
83 into a region experiencing idiosyncratic selection can overwhelm that selection, preventing regional trait
84 differences (i.e. local adaptation) from accumulating (Nagylaki, 1975; Slatkin, 1978; Wright, 1931;
85 Yeaman, 2015). In discrete space models, the evolution of local adaptation in a particular location
86 depends on the ratio of gene flow from dissimilar environments (m) to the strength of selection (s), m/s
87 (Slatkin, 1978; Wright, 1931; Yeaman, 2015). In models of continuous space, patterns of dispersal
88 relative to the strength of selection can be used to determine the minimum size a region experiencing
89 idiosyncratic selection needs to be for locally adaptive differences to accumulate, the so-called

90 “characteristic length” of a cline. Specifically, when dispersal is modelled as a diffusion process, the
91 characteristic length is the standard deviation of dispersal distances (σ) relative to the square root of the
92 strength of selection (i.e. σ/\sqrt{s}). Such characteristic lengths can be described for individual alleles
93 (Nagylaki, 1975; Slatkin, 1973) or for stabilising selection acting on polygenic traits (Barton, 1999;
94 Slatkin, 1978), see reviews by Felsenstein, (1976) and Lenormand (2002). Furthermore, the mean values
95 of polygenic traits will more closely track changes in phenotypic optima over space if those changes are
96 small (Barton, 1999; Slatkin, 1978). Because natural populations often inhabit large spatial ranges
97 encompassing complex patterns of environmental variation, relative rates of gene flow among regions
98 of high environmental similarity or dissimilarity will vary across the landscape, influencing the evolution
99 of local adaptation.

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101 That the spatial pattern of environmental heterogeneity will influence the evolution of local adaptation
102 has been recognised since at least the 1960s (Antonovics, 1971; Antonovics & Bradshaw, 1970; Forester
103 et al., 2016; Hadfield, 2016; Levins, 1966; Schiffers et al., 2014). Indeed, the overall level of local
104 adaptation a species exhibits can be strongly affected by the pattern of environmental variation over
105 space (Forester et al., 2016; Gilbert & Whitlock, 2017; Hadfield, 2016; Schiffers et al., 2014) and several
106 studies have framed this concept in terms of the spatial autocorrelation of the environment (Hadfield,
107 2016; Urban, 2011). Spatial autocorrelation describes the similarity of observations from nearby
108 locations and can be quantified, for example, using Moran’s I (Moran, 1950). Consider the maps of
109 environmental heterogeneity shown in Figure 1A. When the environment that gives rise to spatially
110 varying selection across a species’ range exhibits high spatial autocorrelation (e.g. the right-hand map in
111 Figure 1A), selection pressures may be similar over large areas and changes in environment over space
112 will tend to be gradual. On the other hand, when the environment exhibits weak autocorrelation (e.g.
113 the left-hand map in Figure 1A), regions experiencing idiosyncratic selection will be comparatively small
114 and selection pressures may change rapidly over space. Of course, other factors such as variation in
115 population density, the magnitude of environmental variation and the scale of dispersal will also affect
116 the outcomes of spatially varying selection. All else being equal, though, a species with restricted
117 migration will tend to evolve the strongest local adaptation if agents of selection exhibit high levels of
118 spatial autocorrelation.

119

120 It is likely that the spatial pattern of selection across a species’ range affects the genetic architecture
121 underlying local adaptation. In simple two-patch models, the relative balance of selection and migration
122 influences the number of alleles underlying local adaptation, their effect sizes and rates of allelic
123 turnover (Yeaman & Whitlock, 2011). In continuous, linearly varying landscapes, frequencies of alleles
124 contributing to locally adaptive traits are expected to vary in relation to their effect sizes and the
125 proximity of local trait means to phenotypic optima (Polechová & Barton, 2015). However, such models
126 may not fully predict the patterns of genetic variation expected in complex landscapes where the
127 relative balance of selection and migration vary across space. Previous studies have examined the
128 effects of landscape structure on the genetics of local adaptation in single locus models (Forester et al.,
129 2016) or in models of population expansion (Gilbert & Whitlock, 2017; Schiffers et al., 2014), but it is
130 unclear how the polygenic architecture of local adaptation will be influenced by the spatial pattern of
131 environmental variation.

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133 In this paper, I examine patterns and the genetic bases of local adaptation in complex landscapes.
134 Following previous studies, I cast spatial patterns of environmental variation across a species’ range in
135 terms of spatial autocorrelation. With population genetic simulations, I examine the patterns and
136 genetic architectures of local adaptation that evolve in environments that vary in their degree of spatial
137 autocorrelation. These simulations show ways that the pattern of environmental variation across a

138 species' range can influence the genetic variation underlying local adaptation. Finally, I analyse empirical
139 data from a large-scale experiment in lodgepole pine and find evidence suggesting a link between spatial
140 autocorrelation in climatic/environmental variation with the strength of local adaptation in a natural
141 system. Taken together, the results of this study highlight the importance of considering the spatial
142 pattern of environmental variation in studies of local adaptation.

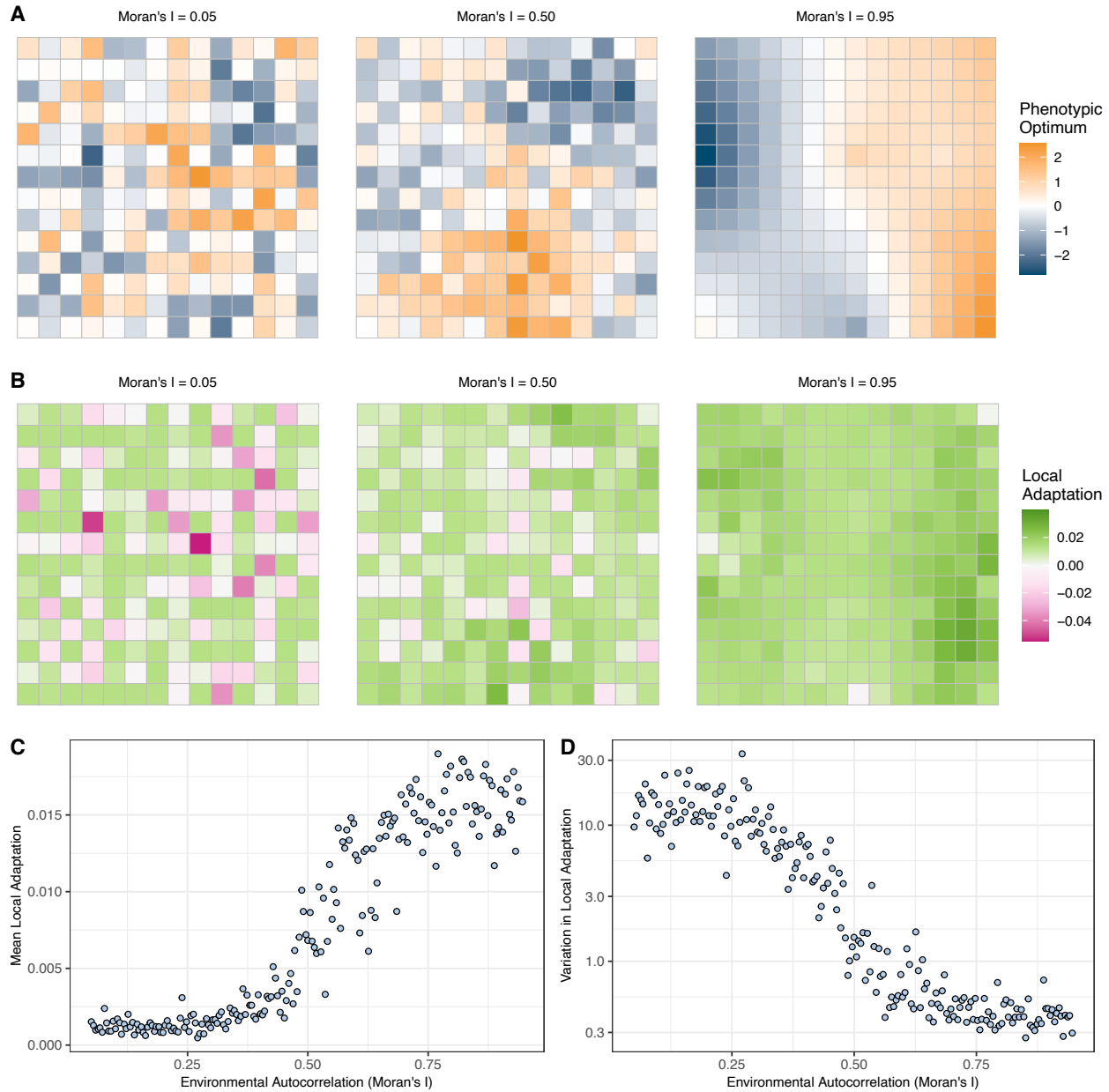
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144 Results and Discussion

145 Simulating local adaptation to spatially heterogeneous environments

146 To further understand the effects that the spatial pattern of environmental variation can have on
147 patterns of local adaptation, I constructed a simulation model of spatially varying selection. I used
148 forward-in-time population genetic simulations in SLiM (4.1; Haller & Messer, 2023) modelling a 2-
149 dimensional stepping-stone metapopulation of 196 demes (i.e. a 14x14 grid). Migration was restricted
150 to adjacent demes with rates of gene flow that resulted in pronounced population structure with clear
151 isolation-by-distance (Figure S1). Spatially varying selection was modelled as stabilising selection, where
152 each deme (d) had a particular phenotypic optimum (θ_d), i.e. an individual in deme d had higher fitness
153 if its phenotype was close to θ_d . The strength of stabilising selection was set such that an individual with
154 the optimal phenotype for the deme with the most negative optimum translocated into the deme with
155 the most positive optimum would experience a 50% reduction in fitness (strong selection) or a 25%
156 reduction in fitness (moderate selection). I used a quantitative trait model to study local adaptation,
157 because it is thought that the traits involved in local adaptation are generally polygenic (Savolainen et
158 al., 2013). Additionally, there is evidence that the genetic basis of local adaptation can involve alleles
159 that have spatially antagonistic fitness effects as well as conditionally neutral effects (Anderson et al.,
160 2013) and both kinds of effects can arise in a quantitative trait model given variation in genetic
161 backgrounds and environments. I constructed a set of 200 maps of normally distributed environmental
162 variation that varied in the degree of spatial autocorrelation (three examples are shown in Figure 1A). I
163 quantified spatial autocorrelation in the environment using Moran's I , which varied from 0.05 (weak
164 autocorrelation) to 0.95 (strong autocorrelation) in the maps I constructed. The maps of environmental
165 variation were used to specify phenotypic optima in individual simulations.

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167
168 **Figure 1** The pattern of environmental heterogeneity influences the outcomes of spatially varying
169 selection. A) Three examples of environmental heterogeneity with similar distributions of phenotypic
170 optima ranging from low autocorrelation on the left to high autocorrelation on the right. B) The pattern
171 of local adaptation that evolved on the landscapes shown in panel A. C) The average local adaptation
172 that arises as a function of Moran's I across 200 maps of environmental variation. C) The coefficient of
173 variation for local adaptation across the 200 maps. The simulation results shown are for cases with
174 mean $F_{ST} \sim 2\%$ and moderate stabilising selection.

175
176 The effects of environmental structure on patterns of local adaptation that evolved in simulations were
177 profound. I measured local adaptation in simulated populations by comparing an individual's fitness at
178 "home" versus "away" following the method outlined by (Blanquart et al., 2013). Using this method, the
179 observed local adaptation for each deme (LA) in the metapopulation was computed (e.g. Figure 1B). As
180 expected, the mean local adaptation across populations (\overline{LA}) increased with the degree of spatial

181 autocorrelation in the environment (Figure 1C), consistent with Hadfield (2016). This result held over
182 various levels of gene flow and strengths of stabilising selection (Figure S2). Varying the rate of gene
183 flow and/or the strength of selection had an effect on the level of local adaptation that arose in a given
184 case, but a pattern of increasing \overline{LA} with Moran's I was always observed (Figure S2). In natural
185 populations, it is likely that multiple aspects of environmental variation will induce spatially varying
186 selection. In cases where simulated populations had multiple traits subject to selection due to different
187 aspects of environmental variation, the trait corresponding to the more autocorrelated environment
188 exhibited greater local adaptation (Figure S3).

189
190 The spatial pattern of environmental variation did not just affect the average level of local adaptation,
191 though, it also had a large influence on the variation in local adaptation across the landscape (Figure 1B).
192 The coefficient of variation in local adaptation [$CV(LA)$] across the landscape decreased rapidly with
193 increasing autocorrelation (Figure 1D, S2B). When environmental variation was weakly autocorrelated,
194 the $CV(LA)$ was as much as 30x higher than for more highly autocorrelated environments (Figure 1D).
195 Variation in the degree of local adaptation across a species range is understudied in the population
196 genetics literature but has important implications (see below).

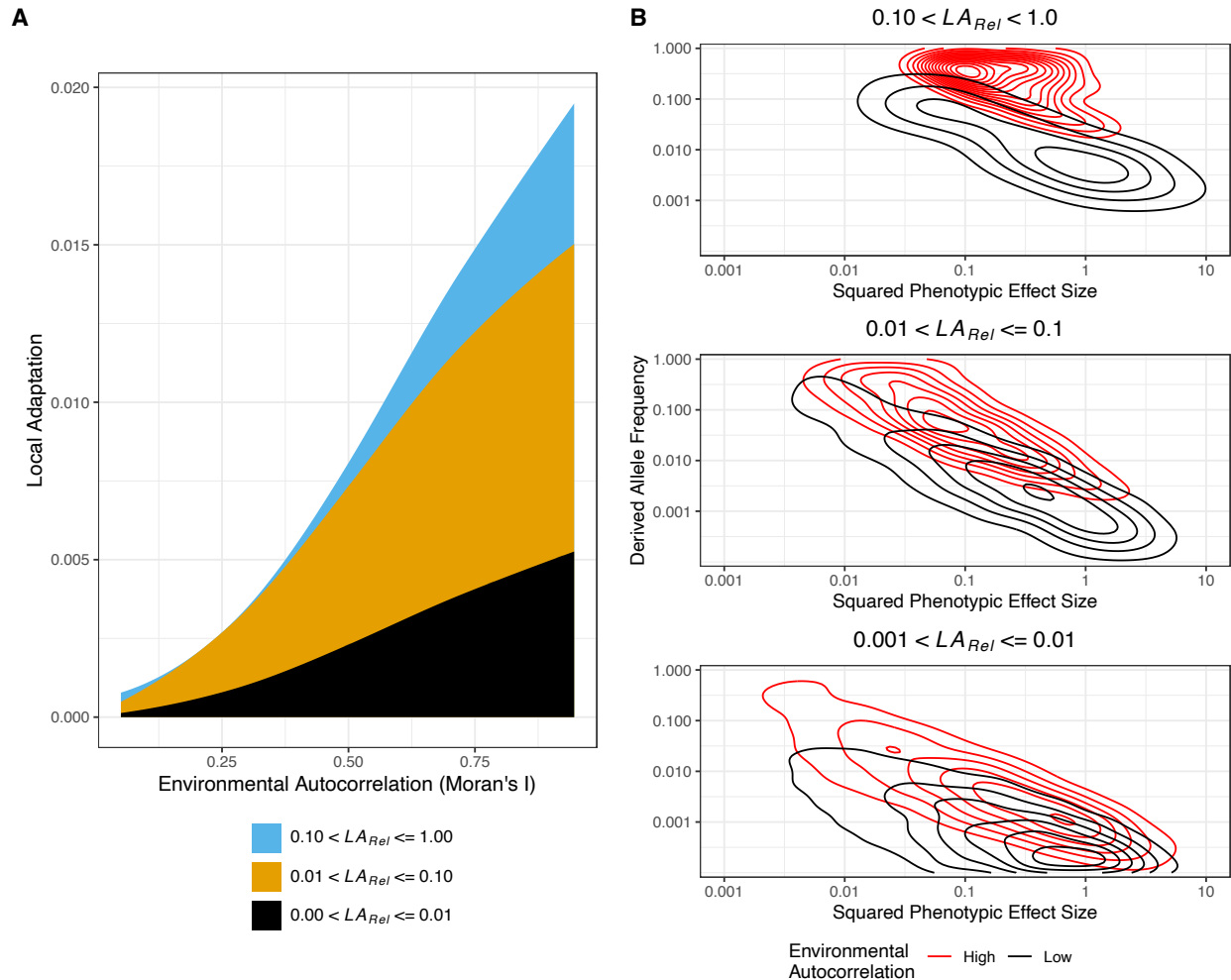
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198 At a finer scale, environmental variation in the immediate vicinity of a particular deme predicted its level
199 of local adaptation and genetic variation, as predicted by theory (Barton, 1999; Guillaume & Whitlock,
200 2007; Slatkin, 1978). Demes that were surrounded by populations with highly similar phenotypic optima
201 evolved greater local adaptation than demes bordering more dissimilar environments (Figure S4A). This
202 was particularly evident when the overall landscape was weakly autocorrelated (Figure S4A),
203 presumably because in highly autocorrelated landscapes most demes are surrounded by similar
204 environments. Furthermore, additive genetic variance (V_A) for the trait under selection was highest in
205 demes surrounded by dissimilar environments (Figure S4B), suggesting that gene flow among locally
206 divergent populations has an effect of increasing genetic variability. Such a positive correlation between
207 V_A and local environmental heterogeneity has been reported in lodgepole pine (Yeaman & Jarvis, 2006).
208 The magnitude of a species' response to selection on a trait is expected to be proportional to V_A
209 (Falconer & MacKay, 1995), thus the pattern of environmental variation that local adaptation evolves
210 under may influence how a species responds to changing environments.

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212 Environmental structure and the genetic architecture of local adaptation

213 The results so far demonstrate that the structure of the environment can have a clear impact on the
214 patterns of local adaptation that evolve, but does it influence the genetic basis of that adaptation?
215 Under a model of spatially varying stabilising selection, each polymorphism that affects the phenotypes
216 under selection will influence local adaptation, but the extent of this will depend on its effect size,
217 where it is present and its allele frequencies. For each polymorphism in a simulation, I quantified the
218 contribution it makes to mean local adaptation (\overline{LA}) as follows. I shuffled the presence/absence of a
219 particular polymorphism across the landscape, effectively erasing its contribution to local adaptation.
220 Average local adaptation was then recalculated without the contribution of the focal polymorphism
221 (\overline{LA}_i). The relative contribution of the focal polymorphism to mean local adaptation across the
222 landscape was then calculated as $LA_{Rel,i} = (1 - \overline{LA}_i/\overline{LA})$. For example, a polymorphism with $LA_{Rel} \approx$
223 1.0 would be the basis of all local adaptation, while one with $LA_{Rel} \approx 0.0$ would have no effect. Note,
224 LA_{Rel} is not strictly a proportion (see Methods for details).

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Figure 2 The genetic architecture of local adaptation is influenced by the structure of the environment.

A) The proportion of total local adaptation explained by alleles that individually explain different amounts of local adaptation (LA_{Rel}) varies as a function of environmental autocorrelation. Lines represent LOESS curves fit with a span parameter of 1.5. B) The mean allele frequencies compared to the squared effect sizes of polymorphisms that underly local adaptation differ depending on the pattern of the environment. The contour lines indicate regions with high densities of points. High autocorrelation refers to data from maps with the 50 highest values of Moran's I. Low autocorrelation refers to data from maps with the 50 lowest values of Moran's I. Results in both panels come from simulations with $F_{ST} = 0.02$ and moderate stabilising selection.

The distribution of locally adaptive effects varied in relation to the pattern of environmental variation (Figure 2A, S5). In environments exhibiting weak autocorrelation, polymorphisms that individually made a large contribution to local adaptation across the species' range ($LA_{Rel} > 0.10$) were largely absent and polymorphisms that made intermediate ($0.01 < LA_{Rel} < 0.10$) and small contributions ($LA_{Rel} < 0.01$) explained most of the local adaptation that evolved (Figure 2A). In environments that were more highly autocorrelated, polymorphisms with $LA_{Rel} > 0.10$ made a substantial contribution to local adaptation alongside those with intermediate and small effects, particularly under strong stabilising selection (Figure S5). These general patterns held over different levels of gene flow (Figure S5).

246 Patterns of genetic variation underlying locally adaptive polymorphisms varied depending on the degree
247 of spatial autocorrelation in the environment. In general, locally adaptive polymorphisms with similar
248 LA_{Rel} tended to have smaller phenotypic effects but larger allele frequencies in highly versus weakly
249 autocorrelated environments (Figure 2B). In highly autocorrelated environments, alleles may readily
250 spread among populations facing similar environmental challenges. However, when alleles with large
251 phenotypic effects spread among neighbouring demes, they may cause individuals to overshoot their
252 respective phenotypic optima, so the alleles that are maintained may tend to have smaller phenotypic
253 effects. In weakly autocorrelated environments, on the other hand, genes flowing from one location to
254 another have a much greater chance of encountering highly divergent environments, preventing locally
255 adaptive alleles from spreading across wide regions. In such cases, phenotypic effects may need to be
256 large for locally adaptive mutations to withstand the swamping effects of gene flow. These general
257 patterns were observed with both strong and moderate selection (Figure 2B, S6A) as well as over
258 varying levels of gene flow (Figure S6A). Indeed, the patterns of allele frequency versus phenotypic
259 effect still held when looking at absolute effects on local adaptation, though they were much less
260 pronounced (Figure S6B).

261
262 Characterising the genetic basis of local adaptation is important, and researchers generally attempt to
263 do so using one of two strategies; by comparing phenotypic variation for traits important for local
264 adaptation to genetic variation (i.e. genome-wide association studies, GWAS) or environmental variation
265 to genetic variation (i.e. genotype-environment association analysis, GEA analysis)(reviewed in Lasky et
266 al. 2023). Since patterns of genetic variation underlying local adaptation can differ depending on the
267 pattern of environmental variation, statistical power to identify the genetic basis of local adaptation will
268 likely vary for different aspects of the environment. To demonstrate this, I performed GWAS on
269 phenotypes for 1,000 randomly chosen individuals from the simulations and corrected for population
270 structure using the kinship matrix. Figure S7 shows that the $-\log_{10}(p\text{-values})$ for alleles that contribute
271 similar levels of local adaptation tend to be smaller (i.e. there is less power) for high versus low
272 autocorrelation environments. Some of this difference may partially be due to the population structure
273 correction procedure (see below), but it demonstrates that the pattern of environmental variation that
274 gave rise to local adaptation can affect the ability to study the genetics of that adaptation.

275
276 There are numerous factors that may interact with the pattern of spatially varying selection to shape the
277 genetics of local adaptation that I did not explore here. The degree of genetic redundancy in relevant
278 traits, distribution of phenotypic effect sizes, mutation rates and patterns of dispersal can all influence
279 the genetics of local adaptation (e.g. Láruson et al., 2020; Yeaman, 2013; Yeaman & Whitlock, 2011).
280 Follow up studies looking at how such factors influence the genetic basis of local adaptation in
281 differently structured environments are needed. However, the results from the simulations should
282 provide researchers seeking to characterise the genetic basis of local adaptation with useful intuition.

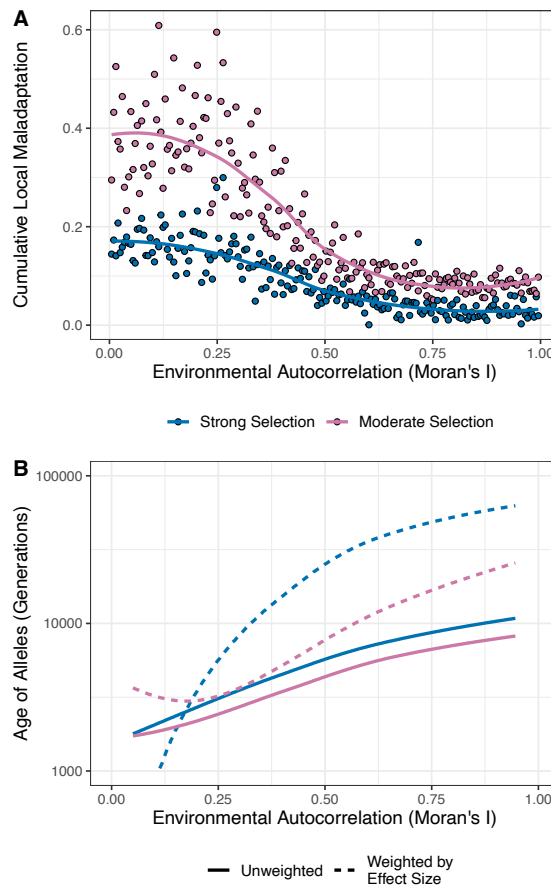
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284 **The evolution of local adaptation: maladaptation and allelic turnover**

285 In heterogeneous environments certain polymorphisms may have a net effect of reducing local
286 adaptation across a species' range. All populations will harbour such locally maladaptive alleles, because
287 any new mutation that increases distance between an individual's phenotype and the local optimum will
288 reduce local adaptation even if only by a small amount. By summing the effects of all polymorphisms
289 with $LA_{Rel} < 0$ across a simulation, I obtained a measure of the cumulative local maladaptation across a
290 meta-population. Note that the cumulative local maladaptation is analogous to "migration load". All
291 simulations exhibited some degree of maladaptation regardless of the level of autocorrelation, but the
292 cumulative effects of maladaptive alleles were always higher under weakly versus highly autocorrelated

293 environments (Figure 3A, S8B). Increasing the rate of gene flow increased the degree of maladaptation
294 and increasing the strength of selection decreased it (Figure S8B).

295

296 Regardless of the pattern of environmental variation in a simulation, levels of local adaptation had been
297 maintained at a steady state for many generations before they were sampled (Figure S9). However, the
298 average age of alleles underlying local adaptation increased with increasing autocorrelation in the
299 environment (Figure 3B). Furthermore, weighing the average allele age within a simulation by effect
300 size, the increase in allele age with spatial autocorrelation was even more pronounced (Figure 3B). Thus
301 large effect alleles, in particular, are maintained for longer times in more highly autocorrelated
302 landscapes. Taken together, these results demonstrate that the rate of allelic turnover is greater for
303 more weakly autocorrelated environments.



304

305 **Figure 3** Species-wide maladaptation and the age of locally adaptive alleles are influenced by the pattern
306 of environmental variation. A) Cumulative local maladaptation, the summed effects of all polymorphism
307 that have a net negative effect on local adaptation across a simulated species' range, decreases with
308 increasing autocorrelation. Points represent individual simulations and lines represent LOESS curves fit
309 with a span parameter of 1.5. B) Alleles underlying local adaptation tend to be older when the
310 environment is more highly autocorrelated. Lines represent LOESS curves fit to the data treating all
311 polymorphisms equally, or by giving higher weight to polymorphisms with greater effect size.

312

313 [Local adaptation, environmental autocorrelation, provenance trials and lodgepole pine](#)

314 The simulation results clearly show how the pattern of environmental variation across a species range
315 may influence the evolution of local adaptation and the genetics underlying it. In natural populations, if

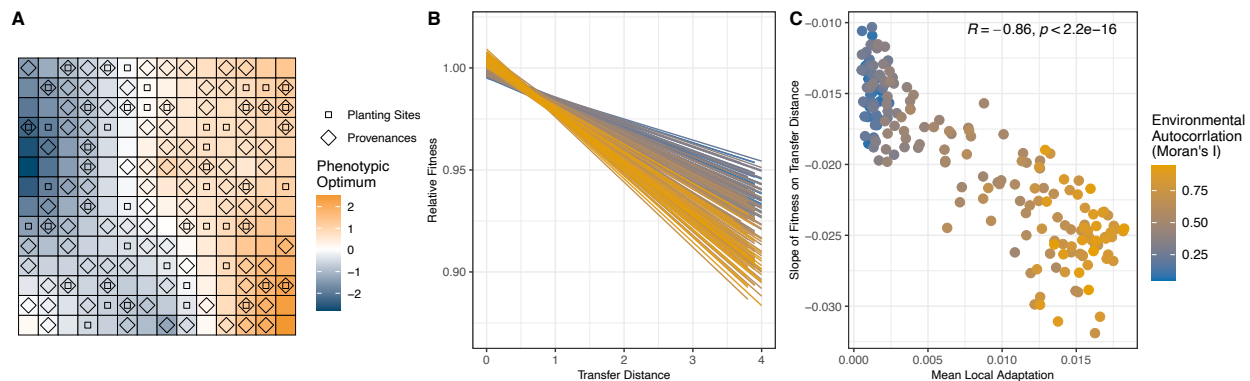
316 the relative strength of selection acting on different aspects of environmental variation is unknown, we
317 should perhaps predict that local adaptation will be strongest when the environment is highly
318 autocorrelated. Despite such strong predictions, though, empirical evidence that patterns of local
319 adaptation coincide with spatially autocorrelated features of the environment is lacking (Siepielski et al.,
320 2013). Spatial autocorrelation in patterns of biotic interactions can explain a large proportion of
321 variation in trait differentiation among populations in several species (Urban, 2011), but such variation is
322 not necessarily locally adapted. Local adaptation has been demonstrated in many forest tree species
323 using provenance trials (Leites & Benito Garzón, 2023), but not to test the prediction that the pattern of
324 environmental variation influences local adaptation.

325

326 Provenance trials involve planting multiple populations of a species in numerous common gardens to
327 assess how “transfer distance”, the distance between home and the common garden, affects
328 productivity (reviewed in Wadgymar et al., 2022). However, the structure of provenance trials is such
329 that the methods I used to quantify local adaptation in the simulations above are not necessarily
330 applicable. To demonstrate how provenance trial data could be used to quantify local adaptation, I
331 conducted *in silico* provenance trials on the simulations (e.g. Figure 4A). As expected, the slope of fitness
332 on transfer distance in provenance trials is increasingly negative with increasing autocorrelation (Figure
333 4B, S10) and strongly negatively correlated with mean local adaptation (\overline{LA}) in simulations (Figure 4C,
334 S10). Provenance trials may, thus, contain information that is useful for assessing whether the pattern of
335 environmental across a landscape is important in shaping local adaptation.

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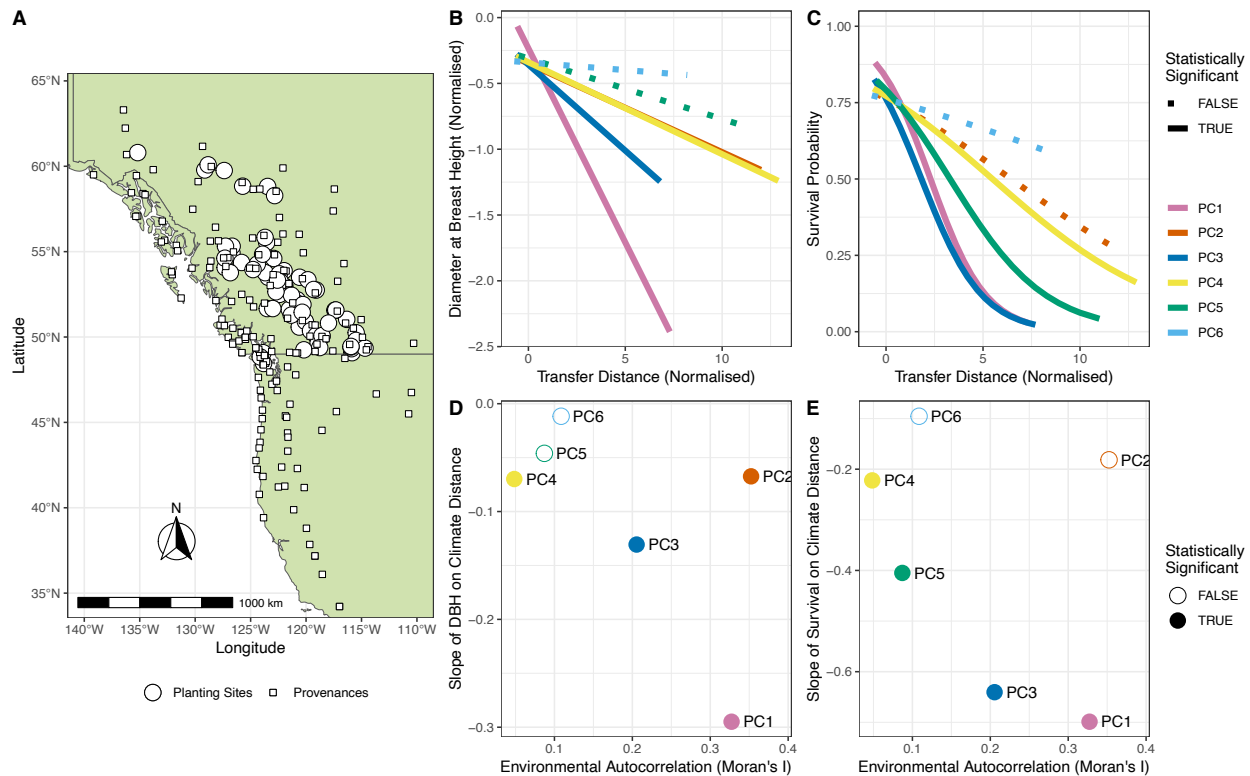
Figure 4 Comparing the results from a simulated provenance trial to measures of local adaptation. A) A map of a provenance trial conducted on a simulated population showing the locations of planting sites and provenances. B) Linear regressions of relative fitness on environmental transfer distance for landscapes with differing levels of environmental autocorrelation. C) The slope of relative fitness on transfer distance compared to the mean local adaptation (\overline{LA}) across simulated meta-populations. Spearman's ρ and its p -value are shown inset in the panel C.

The Illingworth trial is an exceptionally large provenance trial established by the Ministry of Forestry in British Columbia, Canada in the 1970s to establish seed-transfer guidelines for the lodgepole pine (*Pinus contorta*) (Illingworth, 1978). Seeds were collected from 140 provenances from Northwestern North America and seedlings were planted in a set of 62 sites distributed across British Columbia (Figure 5A). Phenotypic data has been recorded for around 60,000 individual trees since the Illingworth trial began and previous studies have used this data to demonstrate clear local adaptation in lodgepole pine (Mahony et al., 2020; Wang et al., 2006). Given the geographic breadth of the Illingworth trial (Figure 5A), it represents a suitable dataset to test the prediction that spatial patterns of environmental variation influence the evolution of local adaptation.

I analyzed data from the Illingworth trial using a mixed-modelling approach. Different aspects of climatic variation across the sites in the Illingworth trial are highly intercorrelated (Figure S11A), so I used principal components analysis to identify independent axes of climatic/environmental variation in the dataset. I restricted the analysis to the first 6 principal components (PCs), as each one explained at least 1% of the variation in the data and combined they explained 95% of the variation (Figure S11B). I then regressed tree diameter at breast height and survival measured after 20 years on transfer distance between planting site and provenances in PC-space (Figure 5B-C) (see Methods for details). For both diameter at breast height and survival, there were significant negative relationships predicting phenotypic variation from climatic PCs (Figure 5B-C). If diameter at breast height and/or survival are considered proxied for fitness, then the results shown in Figures 5B-C indicates local adaptation along several dimensions of climatic variability.

Patterns in the Illingworth trial data suggest that local adaptation in lodgepole pine is strongest when climatic/environmental variation is highly spatially autocorrelated (Figure 5C-D). For both phenotypes, PC1 was the strongest climatic predictor (Figure 5B-C). PC1 captures climatic/environmental differences between coastal and inland locations (Figure S12). Spatial autocorrelation for PC1 was also among the highest in the dataset (Figure 5D-E). For survival, there was a negative relationship between the regression slope and Moran's I (Figure 5E), which is qualitatively similar to the analysis of simulated provenance trials (Figure 4B). However, I did not conduct a formal statistical test of the relationship

375 between spatial autocorrelation and local adaptation because only 4 PCs exhibited statistically
376 significant evidence for local adaptation so such an analysis would be underpowered. While this means
377 the results are merely suggestive rather than concrete, they are in line with the prediction that spatial
378 autocorrelation in the abiotic environment predicts the strength of local adaptation in natural
379 populations.



380
381 **Figure 5** Analysis of local adaptation in lodgepole pine from the Illingworth provenance trial. Panel A)
382 shows the map of provenances and planting sites in the Illingworth trials across the Northwest of North
383 America. B) Fitted relationships between tree diameter at breast height (DBH) and transfer distance for
384 8 principal components. C) Survival probability as a function of transfer function for 8 environmental
385 principal components. D) Linear regression coefficients for the relationships shown in B compared to
386 degree of autocorrelation in the environment. E) Logistic regression coefficients for the relationships
387 shown in C compared to the degree of spatial autocorrelation in the environment. Statistical significance
388 was assessed at $\alpha = 0.05$ after correcting for multiple comparisons using the Dunn-Šidak method.

389

390 Population structure and the genetic basis of local adaptation

391 Knowing the genetic basis of local adaptation in natural systems would give us a better understanding of
392 evolution, but may also be informative for conservation and management (Grummer et al., 2022). Many
393 studies have used methods that associate environmental/phenotypic variation with genotypes or allele
394 frequencies to characterise the genetic basis of local adaptation (Lasky et al., 2023). A large proportion
395 of species exhibit a specific pattern of population structure termed “isolation-by-distance” (IBD) where
396 genetic distance is positively correlated with geographic distance or measures of resistance based on
397 features of the landscape (Jenkins et al., 2010). IBD can arise with restricted migration, when rates of
398 gene flow are highest among parts of a species’ range that are in close proximity, though it can also
399 reflect demographic histories such as past population expansion (Slatkin, 1993). Greater levels of local
400 adaptation may evolve across a species’ range if the spatial pattern of environmental variation aligns

401 with the opportunity for migration (Figure 1). Thus, stronger local adaptation is expected to arise when
402 selection is co-linear or co-autocorrelated with patterns of population structure.

403

404 The relationship between population structure and the structure of the environment likely impacts our
405 ability to study the genetic basis of local adaptation. It is well established that a pattern of IBD can
406 confound the search for genes involved in local adaptation (Meirmans, 2012). Indeed characterising the
407 genetic basis of local adaptation when the agents of selection are co-linear or co-autocorrelated with
408 patterns of gene flow, termed “isolation by environment”, is particularly challenging (Wang & Bradburd,
409 2014). Many studies have analyzed genotype-environment associations (GEA) to characterize the
410 genetic basis of local adaptation. Such association methods often treat population structure as a
411 nuisance variable and various approaches are taken to correct for it. This is done for the statistical
412 necessity of establishing a suitable null model (Meirmans, 2012). For example, latent factor mixed
413 models (LFMMs) are widely used to conduct GEA analyses that correct for population structure (Caye et
414 al., 2019; Frichot et al., 2013). However, Lotterhos (2023) recently found that the sensitivity of the
415 LFMM method declined with increasing correlation between the environment and major axes of
416 population structure. This all suggests that characterising the genetic basis of local adaptation is
417 particularly difficult in the exact cases where local adaptation is expected to be strongest, when
418 selection pressures and population structure are highly co-autocorrelated over space. Careful sampling
419 strategies may alter the power of association methods (Lotterhos & Whitlock, 2015; Meirmans, 2015;
420 Wang & Bradburd, 2014), but such strategies require *a priori* hypotheses about locally adapted traits,
421 the agents of selection (e.g. Kreiner et al., 2022) and/or the genes involved (e.g. Fournier-Level et al.,
422 2011).

423

424 Implications for conservation management

425 A result from the simulations that was particularly striking is the heterogeneity in local adaptation that
426 can arise under different patterns of environmental variation. The only difference between the
427 simulations shown in Figure 1 is the pattern of environmental variation, yet the average level of local
428 adaptation increased by a factor of roughly 5x and the coefficient of variation in local adaptation
429 decreased by 30x when comparing the cases with the highest and lowest spatial autocorrelation in the
430 environment. High heterogeneity in local adaptation across a species’ range may impact conservation
431 interventions and population genetic analyses.

432

433 Any practical conservation intervention that uses average patterns of local adaptation to project
434 performance under changing climates should carefully consider heterogeneity in local adaptation. In
435 forestry, for example, seed-transfer guidelines based on average transfer functions from provenance
436 trials may give a misleading picture of performance for some provenances if there is high heterogeneity
437 in local adaptation. The applicability of a single transfer function would vary depending on how
438 heterogeneous local adaptation is among the populations in question.

439
440 In recent years, population genetic analyses have been developed to identify parts of a species range
441 that are particularly vulnerable to climate change (i.e. genomic offset; Rellstab et al., 2021). Such
442 methods analyse present-day relationships between allele frequency and the environment to predict
443 how species will fare given predicted patterns of environmental change. If the agents of local adaptation
444 are highly spatially autocorrelated, neutral population structure may be partially aligned with gradients
445 of selection, which could explain why the use of “adaptive” genetic markers and randomly chosen
446 markers seem to perform equally well in some offset analyses (Fitzpatrick et al., 2021; Láruson et al.,
447 2022; Lind et al., 2023). Violating the assumption of homogeneous local adaptation in offset analyses,
448 for example, would likely introduce noise into predictions but could potentially lead to spurious results.

449
450 [Thinking about environmental structure when building hypotheses about local adaptation](#)
451 Detailed prediction of the environmental variation that is relevant to patterns of local adaptation
452 requires an understanding of a species’ life history and physiology. The fundamental factor underlying
453 the evolution of local adaptation is the relative balance of selection and dispersal (see Introduction).
454 However, the specific pattern of environmental variation across a landscape influences whether
455 dispersing individuals are likely to encounter environments similar to those of their parents. Unlike
456 dispersal or the strength of selection, which are hard to quantify, spatial autocorrelation in the
457 environment is readily measurable. When seeking to characterise the genetic basis of local adaptation,
458 studies comparing different aspects of environmental variation should consider the pattern of such
459 variation when forming their hypotheses. For example, before comparing GEA results for the different
460 *bioclimatic* variables from WorldClim, researchers could examine how these variables are distributed
461 over space to form *a priori* hypotheses about factors underlying local adaptation. Of course, an aspect of
462 the environment may be highly autocorrelated in space, but if the variation it exhibits does not
463 correspond to varying selection pressures, then it is unlikely to be directly related to local adaptation.
464 That local adaptation is predicted to be stronger with increasing autocorrelation in the environment
465 does not imply that strong local adaptation cannot arise in highly heterogeneous environments or with
466 little spatial autocorrelation. There are numerous examples of local adaptation to environmental
467 heterogeneity that is not smoothly distributed in space. For example, heavy metal concentrations in
468 mine tailings impose selection that is so strong it overwhelms the effects of gene flow (Jain & Bradshaw,
469 1966). It must be kept in mind that the patchy distributions of environmental variation across a
470 landscape, as may be the case for heavy-metal rich soils, may be more or less autocorrelated from the
471 perspective of a given species depending on its dispersal behaviour. In my simulations, I matched the
472 granularity of dispersal with that of the environment (Figure 1A). For real species, considering
473 environmental variation at a scale relevant to how species disperse is critical and recent population
474 genetic advances may make estimating dispersal in natural populations much less time-intensive than
475 previously (Bradburd & Ralph, 2019; Smith et al., 2023). Comparing patterns of species dispersal with
476 patterns of variation in environmental variation that is plausibly relevant to selection may help identify
477 the drivers of local adaptation in natural populations.

478

479 Closing remarks

480 While it has been a long-standing expectation that the pattern of environmental variation (and
481 particularly spatial autocorrelation) will influence the evolution of local adaptation (e.g. Hadfield, 2016;
482 Levins, 1966), the simulation results and analysis of the lodgepole pine data should serve to emphasise
483 how important the spatial pattern of climatic/environmental variation can be. The spatial pattern of
484 environmental variation that a natural population has experienced will have likely shaped the evolution,
485 current patterns and genetic underpinnings of local adaptation. Thus, it likely also influences how
486 populations will respond to changes in the future.

487

488 Materials and Methods

489 Simulating spatially varying selection

490 To explore the effects of landscape structure on the outcomes of spatially varying selection, I
491 constructed a set of maps that exhibited varying degrees of spatial autocorrelation. Maps of normally
492 distributed environmental heterogeneity were constructed using the midpoint displacement algorithm
493 as implemented in the *NLMpy* package (Etherington et al., 2015). I simulated a 14x14 cell grid (i.e.
494 landscape), specifying the desired level of autocorrelation to achieve a set of 200 maps, spanning the
495 range of Moran's I values from 0.05 to 0.95 (i.e. Moran's I varied in increments of 0.0045). Simulated
496 maps were rejected if the mean value across the landscape was less than 0.4 or greater than 0.6. This
497 ensured that the approximately normal distributions of environmental values across the landscape
498 were roughly equivalent across maps.

499
500 Using SLiM v4.1 (Haller & Messer, 2023), I modelled a 2-dimensional stepping-stone meta-populations
501 with 196 demes (i.e. a 14x14 grid). Each deme contained 100 diploid individuals for total meta-
502 population size of 19,600. Migration occurred between adjacent demes in the four cardinal directions
503 except for populations at the range edge where migrants only moved back into demes they were
504 connected to. Migration rates were set at 0.07, 0.035 or 0.0175, leading to population-wide neutral F_{ST}
505 values of 0.02, 0.05 and 0.10, respectively (Figure S1A). Each diploid individual had a 10Mbp long
506 genome that recombined at a constant rate of $r=1 \times 10^{-7}$. When modelling a single trait, mutational
507 effects were distributed as $N(0,1)$ and occurred at random along the sequence at a rate of $\mu = 10^{-10}$,
508 corresponding to a mutational variance of 0.001 for the trait subject to stabilising selection. When
509 modelling two traits, the mutation rate was the same, but effects were modelled as multivariate normal
510 with means of 0, variances of 1 and covariances of 0 (i.e. mutational effects for the two traits were
511 independent). A diploid individual's phenotype for a given trait was the additive combination of the
512 effects on that trait for the alleles the individual possessed (i.e. mutations were semidominant).
513 Spatially varying stabilising selection was modelled using the maps of environmental heterogeneity to
514 specify the distribution of phenotypic optima across the landscape. An individual's relative fitness W_i
515 was calculated using the standard expression for Gaussian stabilising selection (Walsh & Lynch, 2018):

$$516 W_i = \exp \left[- \left(\frac{(\alpha_{i,d} - \theta_d)^2}{2V_s} \right) \right] \text{Equation 1}$$

517 where V_s is the variance of the Gaussian fitness function, α_i is the phenotype of the i^{th} individual in deme
518 d , and θ_d is the phenotypic optimum of deme d . When modelling stabilising selection in cases with two
519 traits, an individual's fitness was calculated as follows:

$$520 W_i = \exp \left[- \left(\frac{(\alpha_{i,1,d} - \theta_{d,1})^2 + (\alpha_{i,2,d} - \theta_{d,2})^2}{2V_s} \right) / 2 \right], \text{Equation 2}$$

521 where $\alpha_{i,1,d}$ and $\alpha_{i,2,d}$ are the values for traits 1 and 2 for individual i in deme d , respectively, and $\theta_{d,1}$ and
522 $\theta_{d,2}$ are the phenotypic optima for traits 1 and 2, respectively. In effect, an individual's relative fitness in
523 this 2-trait model is the average of the marginal fitnesses for each trait.

524
525 To achieve an equilibrium of migration, selection and drift, meta-populations evolved for 100,401
526 generations. Initially, meta-populations evolved under stabilising selection with an optimum of 0 in all
527 demes. After 400 generations, the landscape was altered to one of the 200 maps of environmental
528 heterogeneity and kept in that state for a further 100,000 generations. At the end of the simulation,
529 phenotypes of each individual in each deme were recorded as well as the genealogical history of the
530 meta-population stored as a tree-sequence. PySlim, tskit and msprime packages (Baumdicker et al.,
531 2022; Haller et al., 2019) were used to work with the output tree-sequence files. To calculate Weir and
532 Cockerham's F_{ST} , neutral mutations were added to the simulated population using PySlim at a rate of 10^{-8}
533 /bp.

534 Analyzing simulated data

535 Local adaptation was quantified for each deme using the “home-versus-away” (HA) method outlined by
536 Blanquart et al. (2013). Specifically, each individual’s fitness was quantified in its home deme and every
537 other possible location on the landscape. The mean local adaptation was calculated in each deme as the
538 mean difference in fitness between home and away conditions across all individuals. For each deme d
539 not on the edge of the simulated landscape, I quantified local heterogeneity in the landscape as the
540 mean sum of squares between the focal deme’s environment and that of the four adjacent demes (in
541 the cardinal directions). For each deme, across the n polymorphisms that affected the phenotype
542 additive genetic variance for the trait was calculated as $V_{A,d} = \sum_{i=1}^n p_{i,d}(1 - p_{i,d})\gamma_i^2$, where $p_{i,d}$ is the
543 allele frequency of SNP i in deme d and γ_i is the phenotypic effect of SNP i .

544
545 The contribution of individual SNPs to local adaptation was quantified as follows. For each
546 polymorphism that affected the trait(s) under selection, the presence/absence of the allele in different
547 haplotypes in different demes can be represented as a vector of 1s and 0s. By shuffling this vector, the
548 contribution of this polymorphism to local adaptation is effectively erased, while keeping its
549 contribution to additive genetic variance across the species’ range constant. For polymorphism l , I
550 recomputed all phenotypes for all individuals after shuffling allele frequencies and re-quantified local
551 adaptation as \overline{LA}_l . The relative contribution of the focal polymorphism to local adaptation is calculated
552 as:

$$553 \quad LA_{Rel,l} = (1 - \overline{LA}_l / \overline{LA}). \quad \text{Equation 3}$$

554 Note that LA_{Rel} is not strictly a proportion, as epistasis for fitness that arises in models of stabilising
555 selection means that the $\sum_{i=1}^n LA_{Rel,i} \neq 1$ for the n SNPs that affect phenotypes. Furthermore, alleles
556 that have a net negative effect on local adaptation (i.e. they are locally maladaptive) will have negative
557 LA_{Rel} values. Indeed, the total amount of local maladaptation in a meta-population was calculated as
558 the additive combination of all polymorphisms with negative LA_{Rel} .

559
560 Provenance trials were conducted on simulated data by sampling a set of 50 “planting sites” and a set of
561 100 “provenances”. The relative fitness of each provenance was computed in each of the 50 planting
562 sites. The absolute difference in phenotypic optimum for each provenance and each planting site was
563 used as environmental distance. Using *lme4* in R, I fitted a linear mixed model regressing relative fitness
564 on environmental distance with provenance as a random effect, with slopes and intercepts varying
565 across provenances.

566
567 I combined results across the 50 simulations with the lowest levels of spatial autocorrelation (weak
568 autocorrelation) and the 50 simulations with the highest autocorrelation (high autocorrelation) and
569 examined the relationship between allele frequency and phenotypic effect sizes for the alleles
570 underlying local adaptation.

571
572 For each simulation, I randomly sampled 1,000 individuals from the landscape and I recording their trait
573 values as well as neutral and phenotype affecting polymorphisms. I applied LD-pruning (with a threshold
574 of $r^2 < 0.2$) to the neutral SNPs and used these data to infer the kinship matrix using PLINK (v2; Chang et
575 al., 2015). I performed association studies on the phenotype affecting SNPs from individual simulations
576 with GEMMA (v0.98.5; Zhou & Stephens, 2012), using the inferred kinship matrix as a random effect to
577 account for population structure.

578

579 Analysis of data from the Illingworth trial

580 ClimateNA (Wang et al., 2016) was used to extract climatic data for each location in the Illingworth Trial.
581 Across the locations in the Illingworth trial, many aspects of climatic/environmental variation are highly
582 inter-correlated (Figure S11B). Because such inter-correlation would make it difficult to tease apart the
583 effects of individual aspects of climatic/environmental variation on local adaptation, I conducted a
584 principal components analysis (PCA) to separate the variation onto independent axes. I restricted the
585 analysis to the first 6 principal components as these explained 95% of climatic variation. Diameter at
586 breast height (DBM) and tree height exhibit a strong positive correlation (Pearson's $r = 0.9$), so analyses
587 were restricted solely to DBM. Trees that were dead or dying after 20 years were given a survival score
588 of 0, living trees were scored a 1.

589
590 Phenotype and survival data after 20 years for individual trees from the Illingworth trial were analyzed
591 using mixed models. Mean normalised DBH was modelled as a normally distributed variable using the
592 *lme4* package and survival using a generalised linear mixed model with a "logit" link function using the
593 *glmer* package. The normalised Euclidean distance between each individual's provenance and planting
594 site (i.e. transfer distance) in PC-space was used as a predictor in the model. Provenance, planting site
595 and planting block within sites were included as having random effects on the slope and intercept of the
596 relationship between phenotype and transfer distance.

597
598 Moran's I was calculated for each principal component of climatic variation across provenances (using
599 the *ape* package) incorporating a pairwise Haversine distance matrix as weights in the calculation.

600

601 Data accessibility

602 All the code used to perform, analyse, and plot the results of simulations is available at
603 <https://github.com/TBooker/LocalAdaptationArchitecture>. R scripts to analyse and plot the results of
604 the Illingworth trial data are available at <https://github.com/TBooker/LocalAdaptationArchitecture>, but
605 the raw data files were used by permission of the BC Ministry of Forestry.

606

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614

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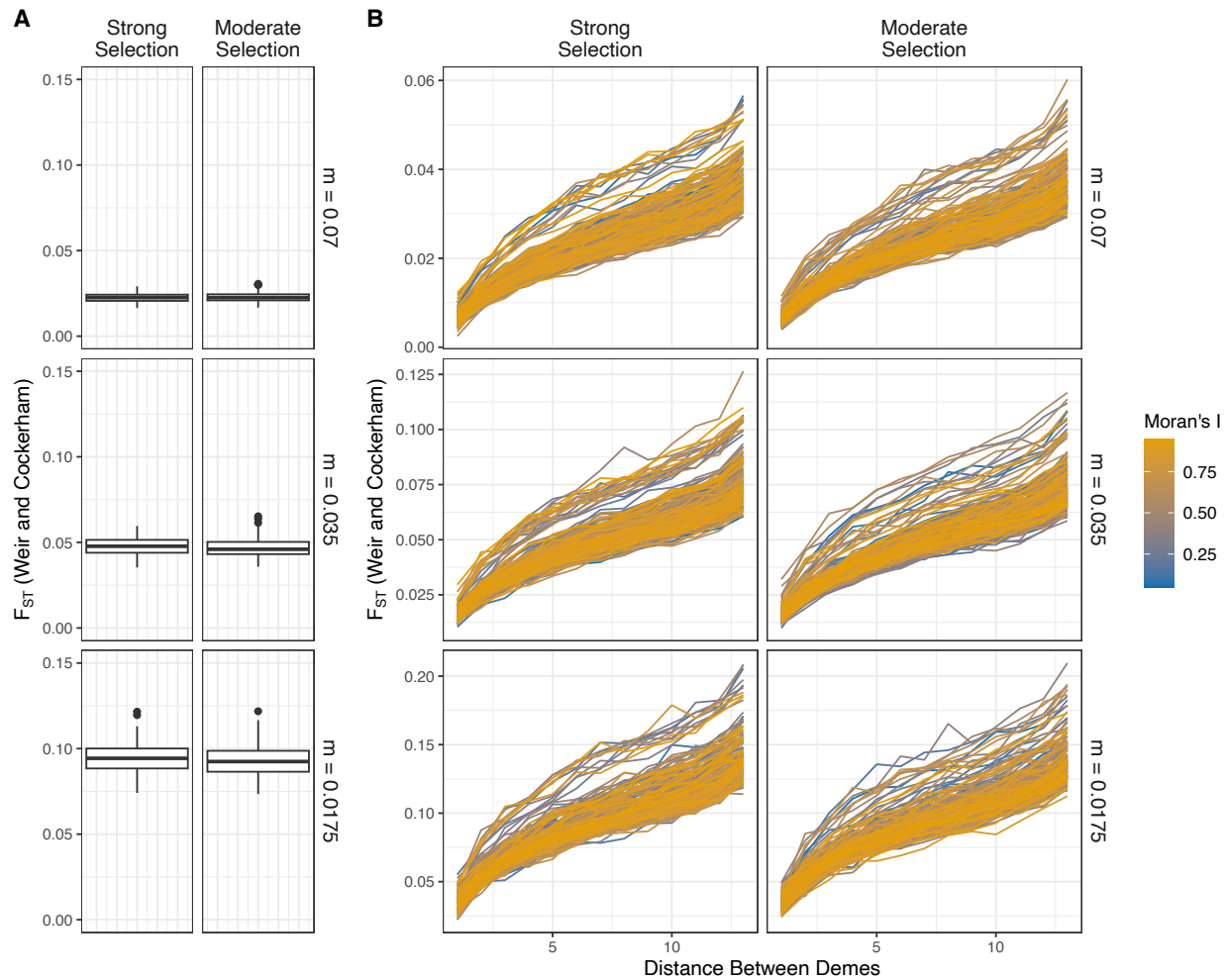
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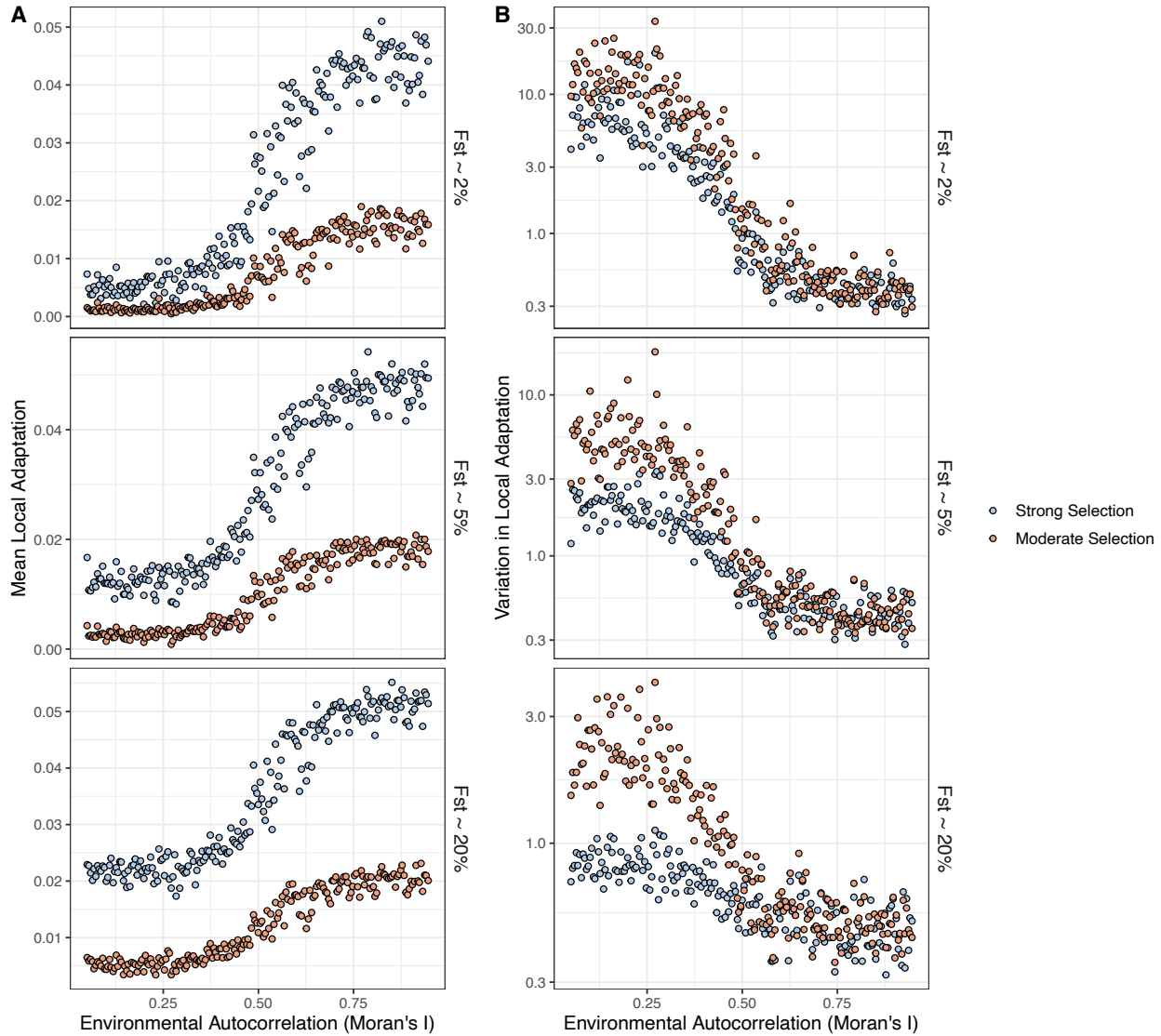
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779 Supplementary Material



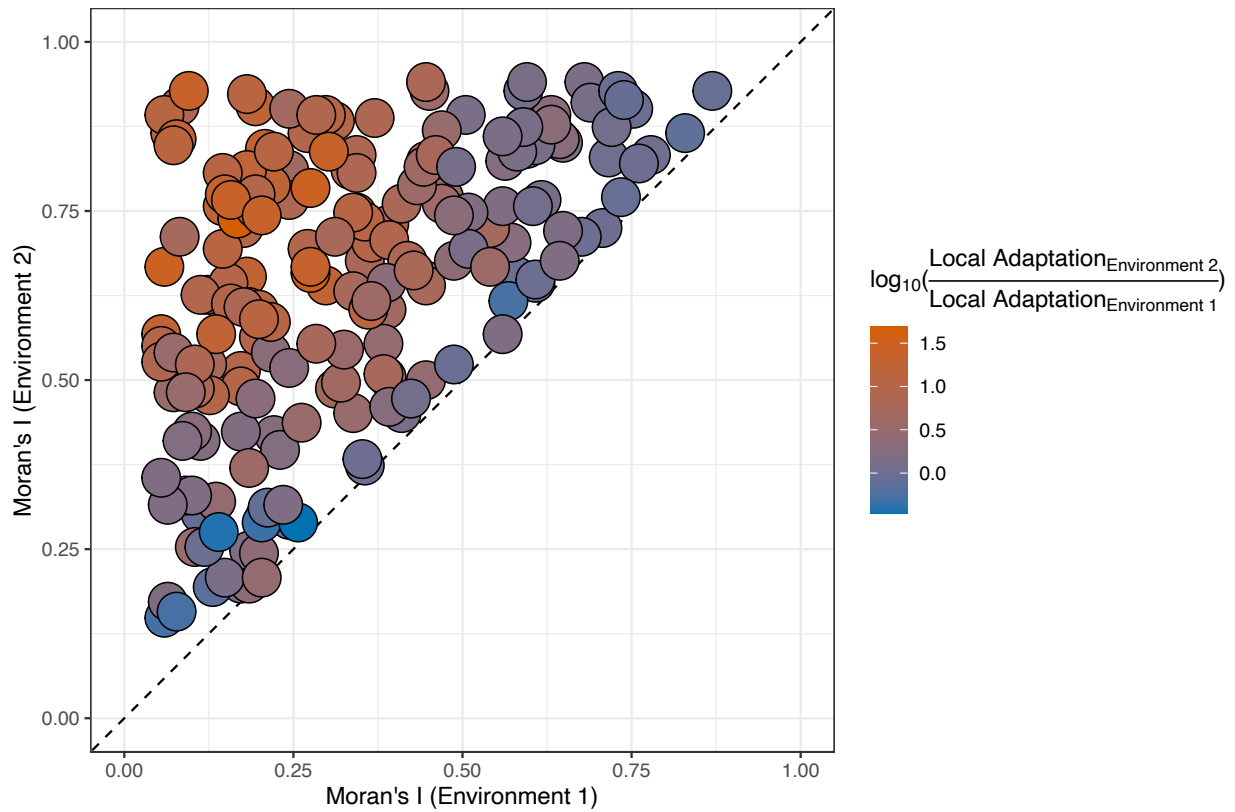
780

781 **Supplementary Figure 1** Overall F_{ST} (panel A) and isolation by distance (panel B) in simulated
782 populations. Note the varying y-axes in panel B. In the main text, F_{ST} is used to refer to the panels of
783 individual graphs. In panel A values from 200 independent simulations were used to construct the
784 boxplot and in panel B individual simulations are shown as lines. Weir and Cockerham's method for
785 calculating F_{ST} , as implemented in the *sci-kit-allele* Python package, was used.



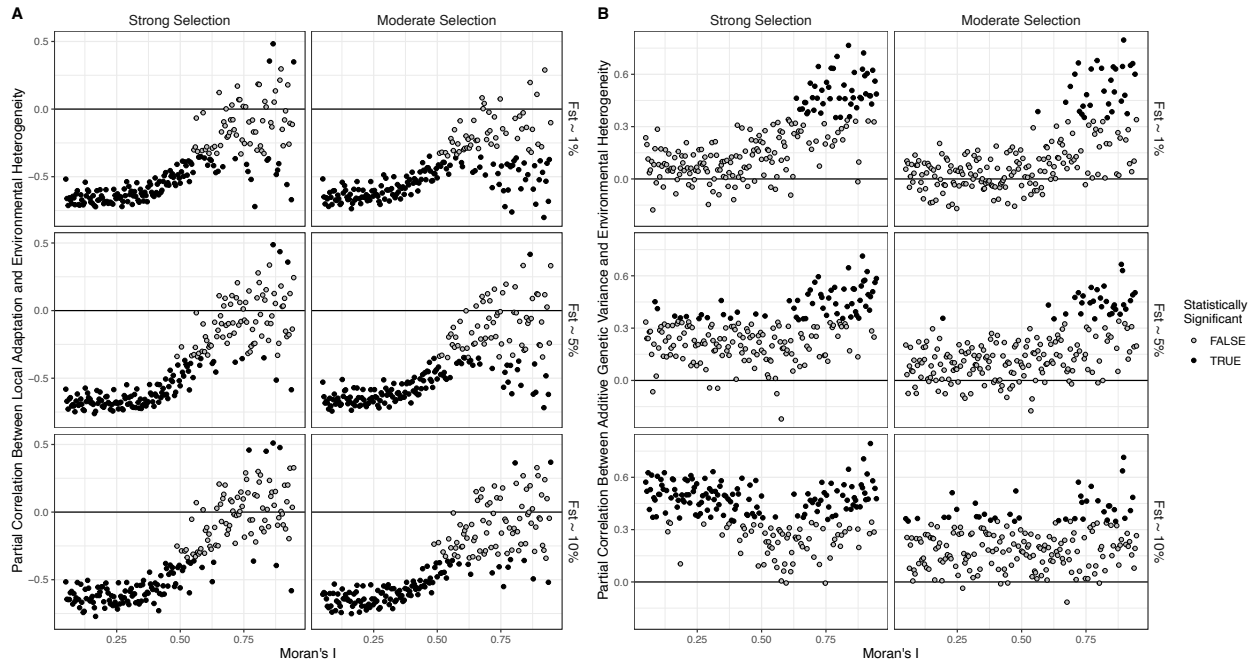
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787 **Supplementary Figure 2** The average extent of local adaptation (panel A) and coefficient of variation in
788 local adaptation (panel B) as a function of spatial autocorrelation in the environment from simulated
789 datasets. The upper cell of each column is included in Figure 1 of the main text.



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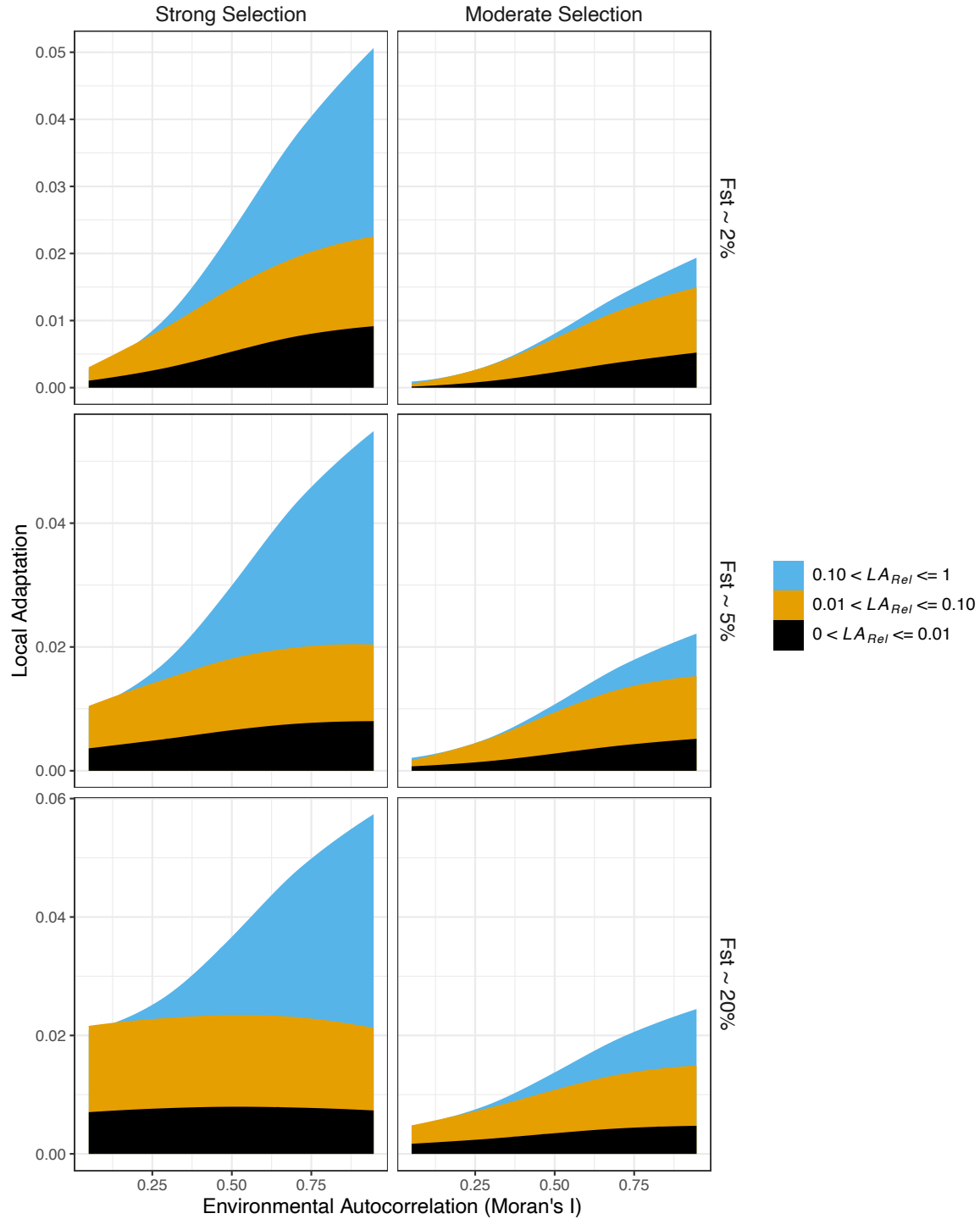
791 **Supplementary Figure 3** Comparison of local adaptation that evolves for two traits subject to spatially
792 varying selection. Selection on each trait was dictated by distinct maps of environmental
793 variation/phenotypic optima. The environment that exhibited the greater degree of spatial
794 autocorrelation (as measured by Moran's I) was designated "Environment 2". The 1:1 line is shown for
795 reference.



796

797 **Supplementary Figure 4** The effect of spatial autocorrelation in the environment on the correlations of
798 local environmental heterogeneity with local adaptation and additive genetic variance. A) The partial
799 correlation between local adaptation and local environmental heterogeneity, controlling for additive
800 genetic variance. B) The partial correlation between additive genetic variance and environmental
801 heterogeneity, controlling for local adaptation. Statistical significance was assessed after correcting for
802 multiple comparisons. The solid black line indicates the statistical null expectation of 0.

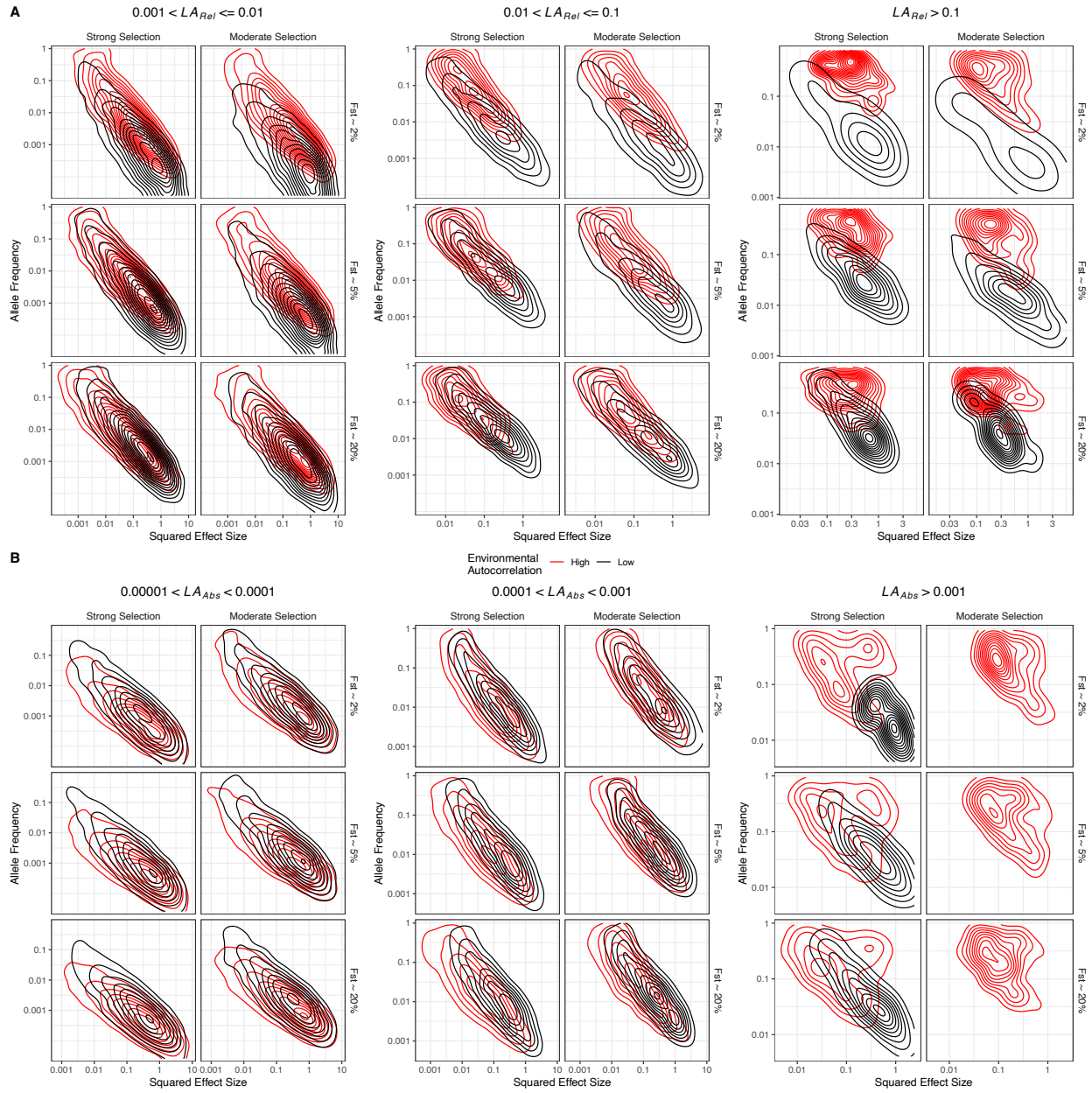
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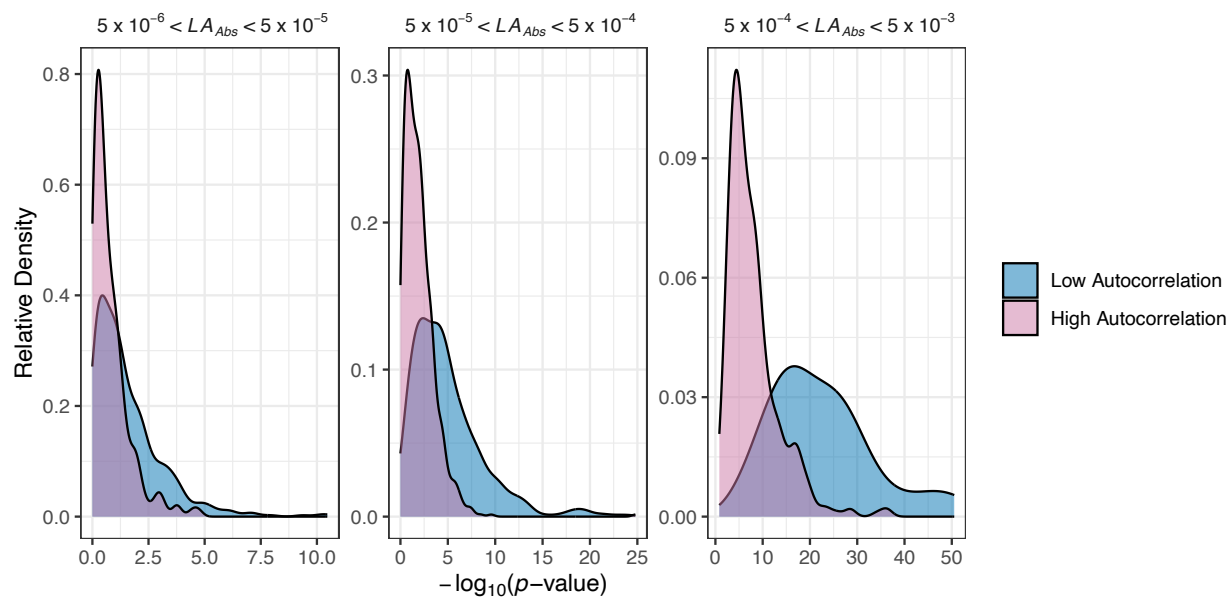
805 **Supplementary Figure 5** The distribution of locally adaptive effects as a function of spatial
806 autocorrelation in the environment. The area shown was calculated across 200 independent simulations
807 and smoothed using a LOESS regression with span 1.5. The upper right cell is included in Figure 2 of the
808 main text.

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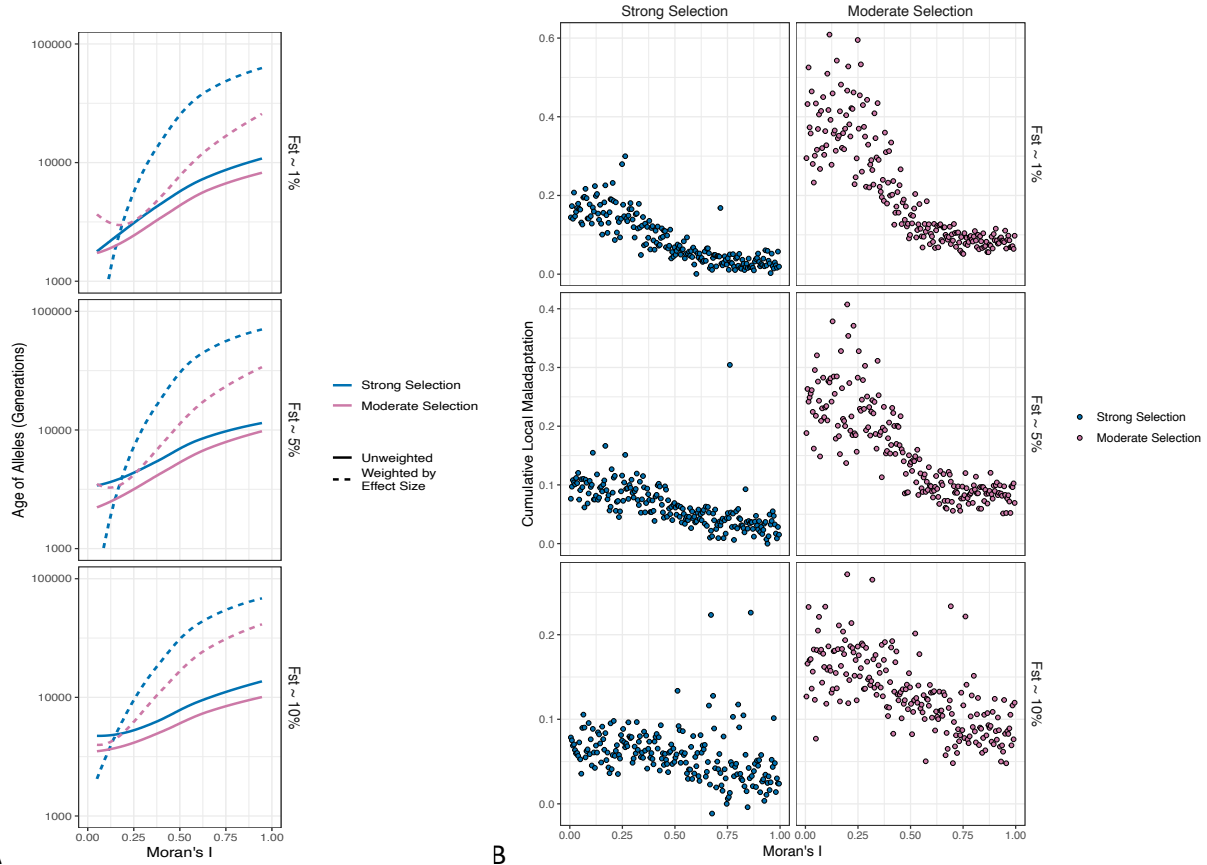


810

811 **Supplementary Figure 6** The relationship between allele frequency and the squared phenotypic effect
 812 size for polymorphisms that contribute varying degrees of local adaptation in either relative (panel A) or
 813 absolute terms (panel B).



814
815 **Supplementary Figure 7** Results from a GWAS on 1,000 randomly chosen individuals from either high or
816 low autocorrelation environments. Each panel compares the relative density of $-\log_{10}(p\text{-values})$ from a
817 GWAS conducted on data from the 50 maps with the highest or lowest levels of spatial autocorrelation.
818



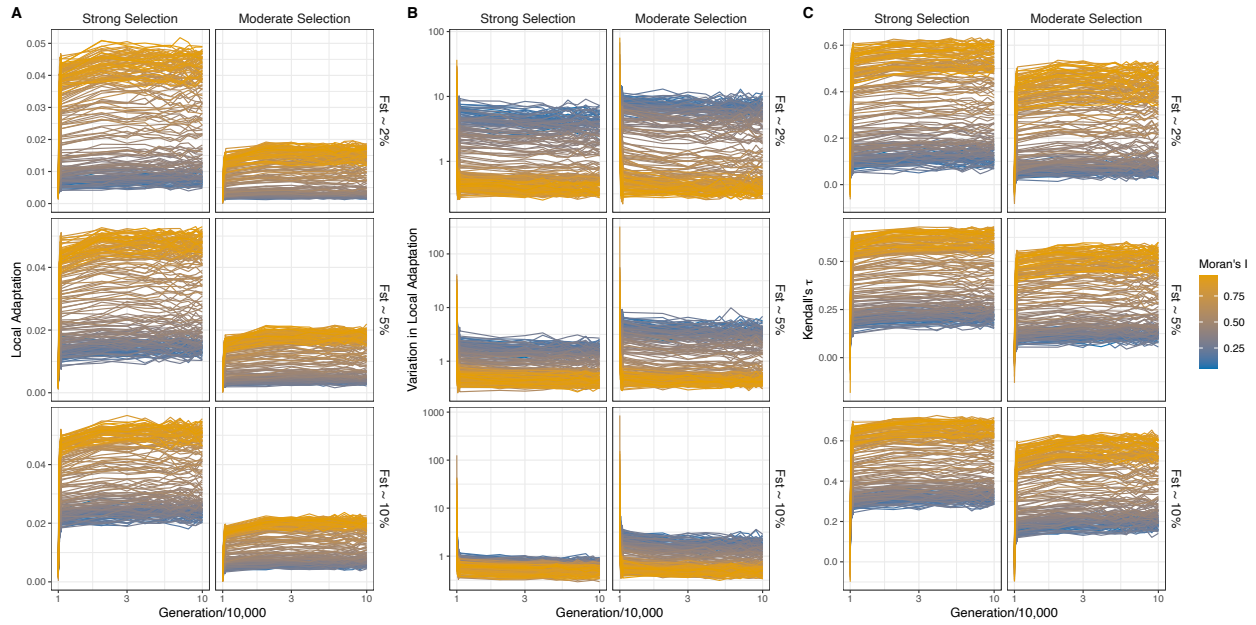
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A

B

820 **Supplementary Figure 8** A) Cumulative local maladaptation as a function of spatial autocorrelation in
821 the environment across all parameter combinations. B) The average age of locally adaptive alleles in
822 meta-populations subject to spatially varying selection. The lines represent LOESS regression curves
823 calculated with span parameters of 1.5.

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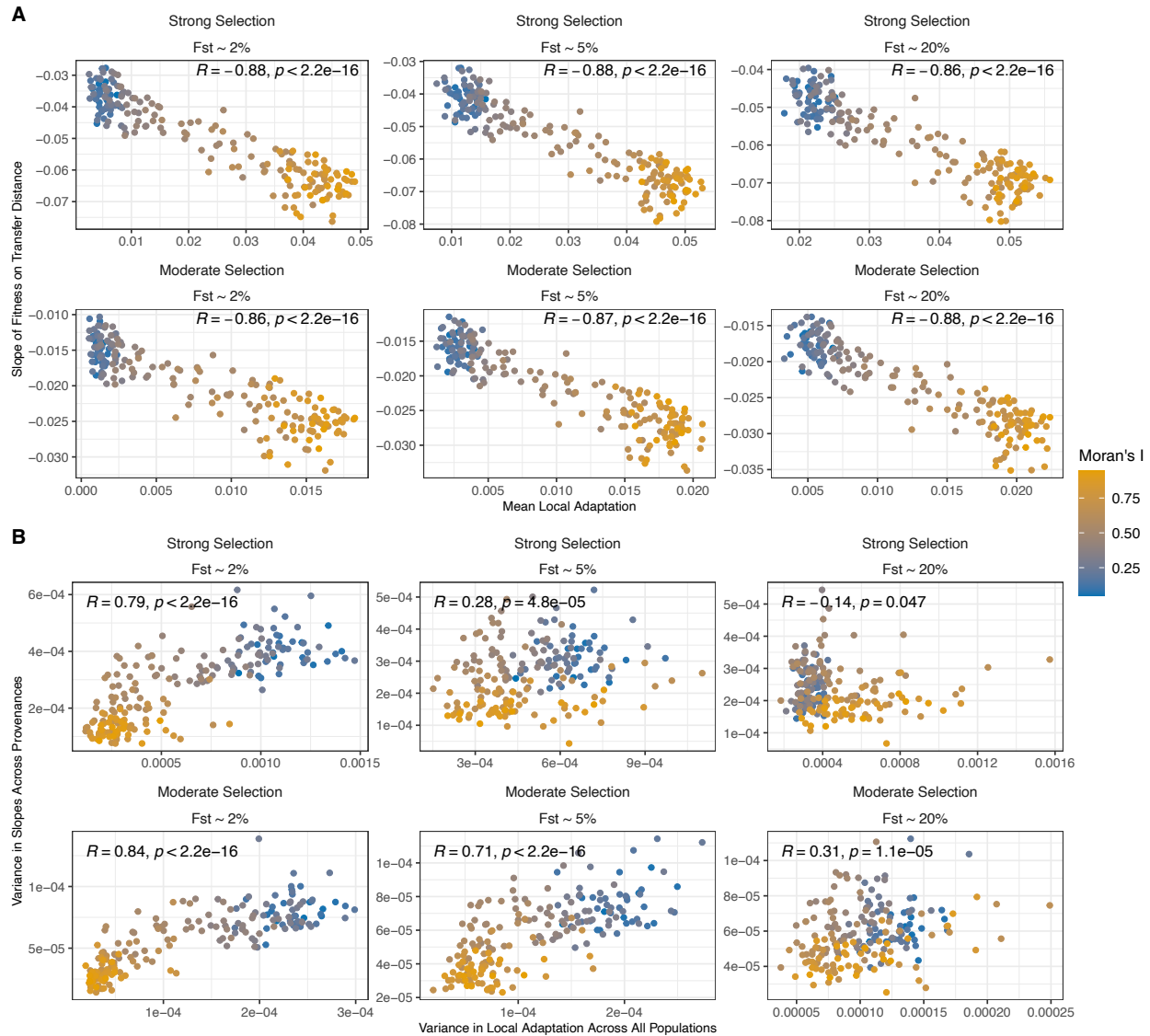


825

826 **Supplementary Figure 9** Establishment of local adaptation in the simulations. Panel A) shows the
827 average level of local adaptation across all demes. Panel B) shows the coefficient of variation in local
828 adaptation across demes. Panel C) shows the Kendall's tau rank correlation between phenotypes and
829 local optima.

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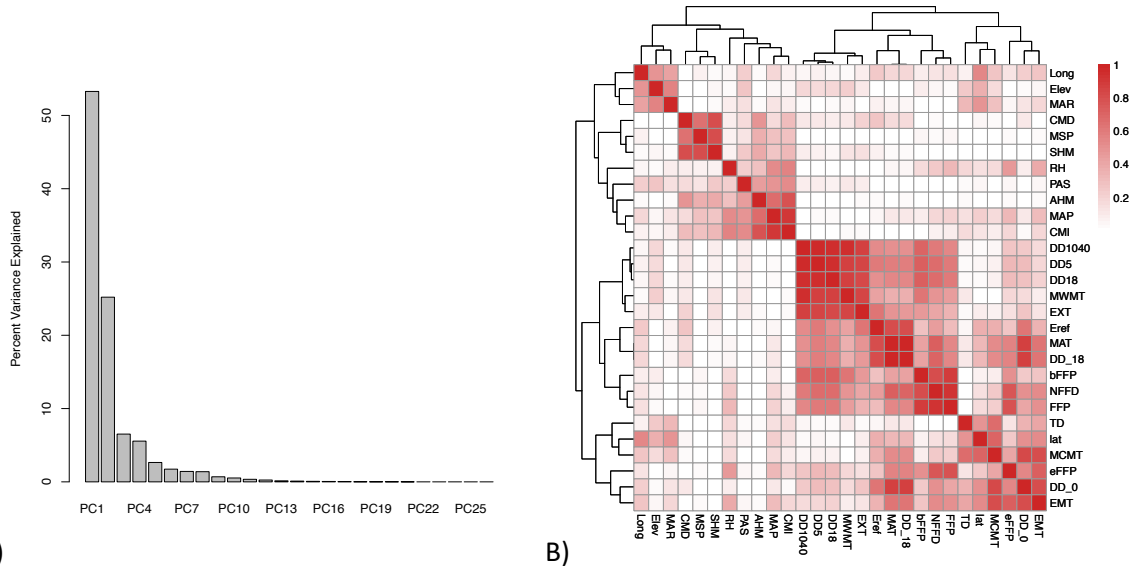


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833 **Supplementary Figure 10** Comparing local adaptation summary statistics to results of linear models
 834 applied to simulated provenance trials. Panel A compared the slopes of the relationship between
 835 relative fitness and transfer distance in simulated provenance trials to home-versus-away measure of
 836 local adaptation described by Blanquart et al., (2013). Panel B compares the variance in slopes across
 837 provenances to the variance in local adaptation across all populations. In both panels, each point
 838 summarises analyses from a single simulation. The Spearman correlation coefficient and the associated
 839 p -value are shown within each cell.

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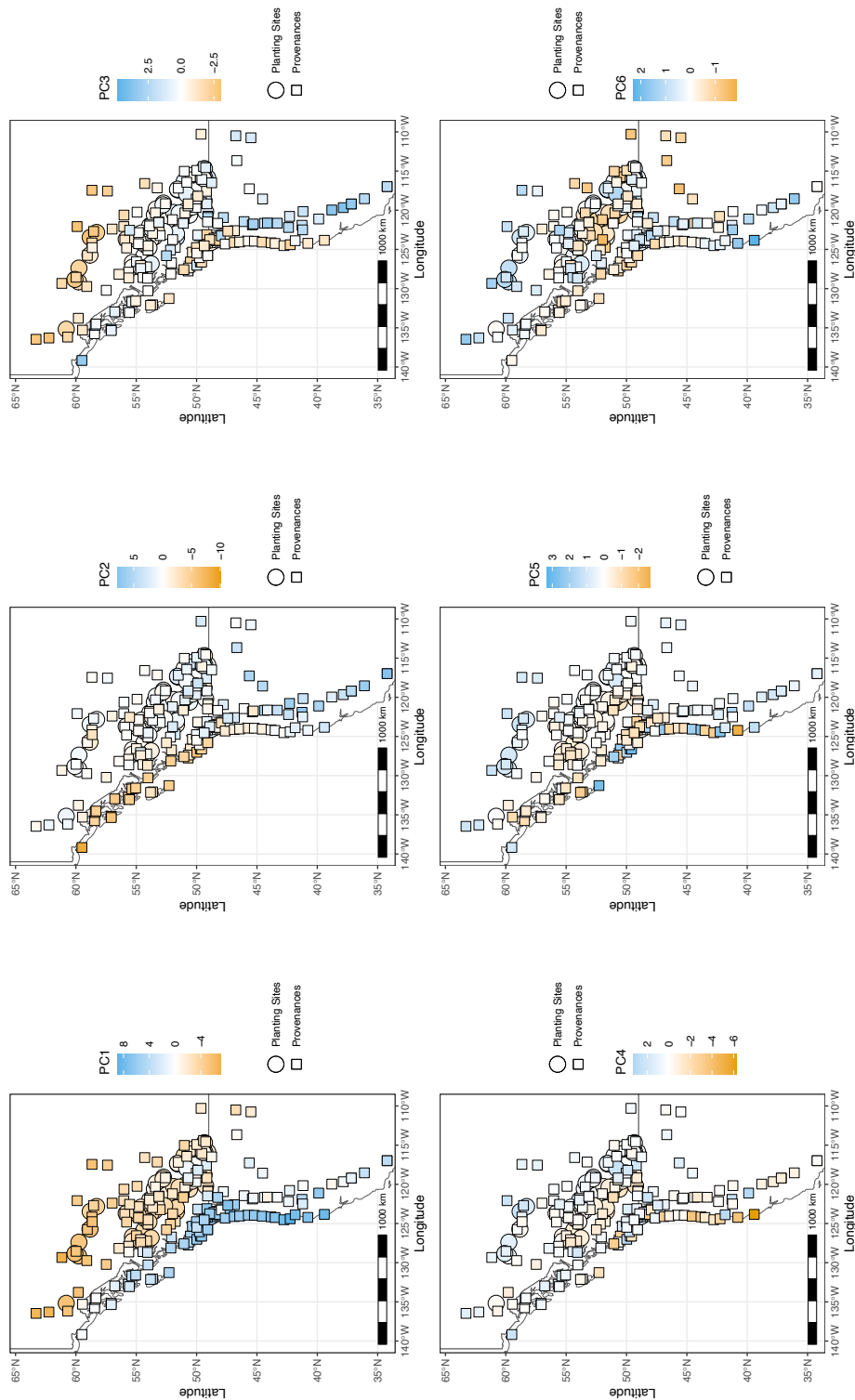
842 A)

B)

843 **Supplementary Figure 11** A) Percent variance explained by the principal component analysis conducted
844 on climatic/environmental variation in the Illingworth trial data. B) The correlation matrix for the 28
845 climatic/environmental variables for planting sites and provenances in the Illingworth trial. A key to the
846 abbreviations for the 25 annual climatic variables from ClimateBC along can be obtained from
847 <https://climatebc.ca/Help2>. Additionally, latitude (lat), longitude (Long) and elevation (Elev) are
848 included.

849

850



851

852 **Supplementary Figure 12** The spatial pattern of loadings onto the first 6 principal components of
853 environmental/climatic variation across provenances and planting sites in the Illingworth Trial. The first
854 6 principal components explained a total of 95% of the variation in the data.