fcfdr: an R package to leverage continuous and binary functional genomic data in GWAS

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¹ Abstract

² Summary: GWAS discovery is limited in power to detect associations that exceed the stringent

³ genome-wide significance threshold, but this limitation can be alleviated by leveraging relevant

⁴ auxiliary data. Frameworks utilising the conditional false discovery rate (cFDR) can be used to

⁵ leverage continuous auxiliary data (including GWAS and functional genomic data) with GWAS

6 test statistics and have been shown to increase power for GWAS discovery whilst controlling the

7 FDR. Here, we describe an extension to the cFDR framework for binary auxiliary data (such

⁸ as whether SNPs reside in regions of the genome with specific activity states) and introduce an

⁹ all-encompassing R package to implement the cFDR approach, fcfdr, demonstrating its utility in

¹⁰ an application to type 1 diabetes.

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Availability and implementation: The fcfdr R package is freely available at: https://github.

13 com/annahutch/fcfdr. Scripts and data to reproduce the analysis in this paper are freely available

14 at: https://annahutch.github.io/fcfdr/articles/t1d_app.html

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16 **1** Introduction

A stringent significance threshold is required to identify robust genetic associations in GWAS due to multiple testing constraints. Leveraging relevant auxiliary data has the potential to boost statistical power to exceed the significance threshold. The conditional FDR (cFDR) is a Bayesian FDR measure that additionally conditions on auxiliary data to call significant associations. The cFDR approach was originally developed to leverage GWAS *p*-values from related traits, thereby exploiting genetic pleiotropy to increase GWAS discovery^{1,2,3}, and has been shown to increase power for GWAS discovery whilst controlling the frequentist FDR¹¹.

Motivated by the enrichment of GWAS SNPs in particular functional genomic annotations¹⁴, Flexible cFDR was developed to extend the usage of the cFDR approach to the accelerating field of functional genomics⁹. However, at-present no cFDR methodology exists that permits binary auxiliary data, meaning that the approach cannot currently be used to leverage auxiliary data with a binary representation, such as whether SNPs are synonymous or non-synonymous or whether they reside in regions of the genome with specific activity states.

Here we present an extension to the cFDR approach that supports binary auxiliary data and we thus introduce a cFDR toolbox in the form of an R package (https://github.com/annahutch/ fcfdr) that supports various auxiliary data types. We demonstrate the utility of our methods and software by iteratively leveraging three distinct types of relevant auxiliary data with GWAS *p*-values for type 1 diabetes (T1D)¹² to uncover new genetic associations.

35 2 The cFDR framework

Let $p_1, ..., p_m \in (0, 1]$ be a set of *p*-values corresponding to the null hypotheses of no association between the SNPs and a trait of interest (denoted by H_0). Let $q_1, ..., q_m$ be auxiliary data values corresponding to the same *m* SNPs. Assume that *p* and *q* are realisations of random variables *P*, *Q* satisfying:

$$(P|H_0) \sim U(0,1)$$

$$P \perp Q|H_0.$$
(1)

The cFDR is defined as the probability that a random SNP is null for the trait given that the observed *p*-values and auxiliary data values at that SNP are less than or equal to values *p* and *q* respectively^{1,2}. Bayes theorem and standard probability rules are used to derive:

$$cFDR(p,q) = Pr(H_0|P \le p, Q \le q)$$

$$= \frac{Pr(P \le p|H_0, Q \le q) \times Pr(H_0|Q \le q)}{Pr(P \le p|Q \le q)}$$

$$= \frac{Pr(P \le p|H_0, Q \le q) \times Pr(Q \le q|H_0)Pr(H_0)}{Pr(P \le p, Q \le q)}.$$
(2)

To construct a conservative estimator of the cFDR, approximate $Pr(P \le p|H_0, Q \le q) \approx p$ (from property 1; note that if property 1 holds and *P* is correctly calibrated then this approximation is an equality) and $Pr(H_0) \approx 1$ (since associations are rare in GWAS):

$$\widehat{cFDR}(p,q) = \frac{p \times \overline{Pr}(Q \le q | H_0)}{\overline{Pr}(P \le p, Q \le q)},$$
(3)

where $\hat{}$ is used to denote that these are estimates under the assumption $H_0 \perp Q | P$. The methods used to estimate the cumulative densities in equation (3) vary across approaches. In the original cFDR approach they are estimated using empirical cumulative density functions^{1,10,11} whilst in Flexible cFDR they are estimated using kernel density estimation⁹.

⁵⁰ However, the \widehat{cFDR} values do not directly control the FDR¹⁰. Instead, a method proposed by ⁵¹ Liley and Wallace¹¹ can be used to generate *v*-values, which are essentially the probability of ⁵² a newly-sampled realisation (p,q) of P,Q attaining an as extreme or more extreme \widehat{cFDR} value ⁵³ than that observed, given H_0 . The *v*-values are therefore analogous to *p*-values and can be used in ⁵⁴ any conventional error-controlling multiple testing procedure that allows for slightly dependent ⁵⁵ *p*-values (e.g. the Benjamini-Hochberg procedure). The derivation of *v*-values also allows for the ⁵⁶ method to be applied iteratively to incorporate additional layers of auxiliary data.

Since binary auxiliary data can only take two values, we introduce an alternative methodology called "Binary cFDR" which is based on finding optimal rejection regions to derive *v*-values (see Supplementary Methods for full details on the Binary cFDR methodology). We show in a simulation-based analysis that applying Binary cFDR iteratively over informative auxiliary data

⁶¹ increases power whilst controlling the frequentist FDR (Supplementary Results, Supplementary
 ⁶² Fig. 2).

3 R package and T1D application

We present an R package that implements both Flexible cFDR and Binary cFDR, named fcfdr (https://github.com/annahutch/fcfdr), and demonstrate its utility in an application to T1D which is fully reproducible (see https://annahutch.github.io/fcfdr/articles/t1d_app. html).

We used *p*-values from an Immunochip study of T1D¹² as our primary data set. In the first iteration we used Flexible cFDR to leverage Immunochip *p*-values for a genetically related trait, rheumatoid arthritis (RA)⁶ (Fig. 1A). In the second iteration we used Binary cFDR to leverage data measuring SNP overlap with regulatory factor binding sites^{5,8,7} (Fig. 1B) and in the third iteration we used Flexible cFDR to leverage average enhancer-associated H3K27ac fold change values derived from ChIP-seq experiments conducted in T1D-relevant cell types⁴ (Fig. 1C) (see Supplementary Methods for full details on the data).

Our implementation of cFDR identified 101 SNPs as newly genome-wide significant (FDR <75 3.3e - 06 which corresponds to p < 5e - 08; Supplementary Methods). These SNPs had relatively 76 small p-values for RA (median p = 0.007 compared with median p = 0.422 in full data set), were 77 more likely to be found in regulatory factor binding sites (mean binary value was 0.406 compared 78 to 0.234 in full data set) and had larger H3K27ac fold change values in T1D-relevant cell types 79 (median fold change value was 1.44 compared with 0.576 in full data set). Similarly, 45 SNPs 80 were identified as newly not significant (i.e. they were significant in the original GWAS data set 81 but became not significant after applying cFDR). These SNPs had relatively high p-values for RA 82 (median p = 0.620), were less likely to be found in regulatory factor binding sites (mean binary 83 value was 0.044) and had smaller H3K27ac fold change values in T1D-relevant cell types (median 84 fold change value was 0.431). 85

The original GWAS identified 38 significant genomic regions (based on our definition of genomic regions, see Supplementary Methods). All of these were found to be significant in the cFDR analysis, which additionally identified 4 genomic regions that were newly significant (with lead variants:

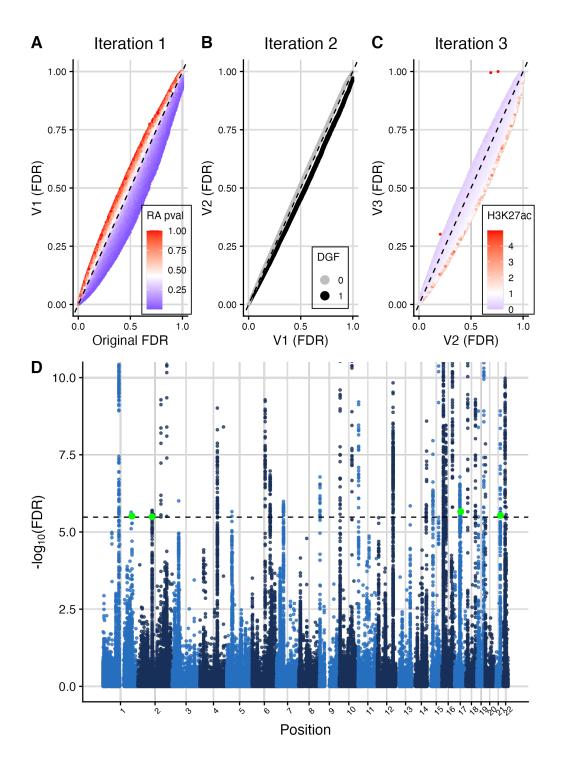


Figure 1: Summary of cFDR results for T1D application. "FDR values" were obtained from the raw *p*-values and *v*-values from each iteration of the cFDR approach using the Benjamini-Hochberg procedure. Top panel shows FDR values before and after (A) iteration 1 (B) iteration 2 and (C) iteration 3 of the method, coloured by the value of the auxiliary data (*p*-values for RA in iteration 1, DGF annotation values in iteration 2 and average H3K27ac fold change values relative to expected background counts in iteration 3). (D) Manhattan plot of ($-log_{10}$ transformed) FDR values. Green points indicate the four lead variants that were newly FDR significant after cFDR. Black dashed line at FDR significance threshold (FDR = 3.3e - 06; which was the maximum FDR value amongst SNPs with raw *p*-values $\leq 5e - 08$ - see Supplementary Methods). *y*-axis has been truncated in panel (D) to aid visualisation.

rs1052553, rs3024505, rs6518350 and rs13415583). Three of these SNPs had small p-values for 89 RA (rs1052553: RA p = 0.007; rs6518350: RA p = 0.06161 and rs13415583: RA p = 1.913e - 0690 whereas rs3024505 had RA p = 0.6008) and two of these SNPs had high H3K27ac fold change 91 values (rs3024505 had 87.4th percentile and rs6518350 had 72.7th percentile of H3K27ac fold 92 change values). Two of the lead variants overlapped regulatory factor binding sites (rs1052553 93 and rs3024505). When using a larger Immunochip study of T1D for validation (16, 159 T1D cases 94 compared to $(6,670)^{13}$, we found that three out of the four lead variants were present and that 95 these had smaller *p*-values in the validation GWAS data set than the discovery GWAS data set: 96 rs1052553 had p = 1.649e - 15, rs3024505 had p = 9.127e - 14, rs13415583 had p = 4.764e - 0997 in the validation data set ¹³ compared to p = 8.156e - 08, p = 6.394e - 08 and p = 1.062e - 0798 respectively in the discovery data set¹². 99

100 4 Conclusion

We have described a novel implementation of the cFDR approach that supports binary auxiliary data and have introduced an all-encompassing R package, fcfdr, that can be used to implement the cFDR approach for a wide variety of auxiliary data types. We have demonstrated the versatility of this tool in an application to T1D where we uncovered new genetic associations.

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