

Density-independent processes decouple component and ensemble density feedbacks

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Abstract

Analysis of long-term trends in abundance provide insights into population dynamics. Population growth rates are the emergent interplay of fertility, survival, and dispersal, but the density feedbacks on some vital rates (component) can be decoupled from density feedback on population growth rates (ensemble). However, the mechanisms responsible for this decoupling are poorly understood. We simulated component density feedbacks on survival in age-structured populations of long-living vertebrates and quantified how imposed nonstationarity (density-independent mortality and variation in carrying-capacity) modified the ensemble feedback signal estimated from logistic-growth models to the simulated abundance time series. The statistical detection of ensemble density feedback was largely unaffected by density-independent processes, but catastrophic and proportional mortality eroded the effect of density-dependent survival on ensemble-feedback strength more strongly than variation in carrying capacity. Thus, phenomenological models offer a robust approach to capture density feedbacks from nonstationary census data when density-independent mortality is low.

KEY WORDS

Australia, compensation, density dependence, carrying capacity, logistic growth, stationarity

INTRODUCTION

Compensatory density feedback describes a population's ability to return to the environment's carrying capacity in response to an increase in population size¹. This phenomenon is driven by adjustments to individual fitness imposed by variation in per-capita resource availability, and associated processes of predation, competition, parasitism, and dispersal²⁻⁵. As survival and fertility rates ebb and flow in response to variation in population density, it is theoretically possible to detect the density-feedback signal in time series of abundance monitored at regular intervals over a sufficient period^{2,6}. There is now considerable evidence that survival and fertility track population trends in many vertebrate⁷⁻¹⁵ and invertebrate¹⁶⁻²² species. Therefore, given the irreplaceable importance of long-term monitoring of population size in applied ecology and conservation², assessing the presence of

44 compensatory signals in censuses of population abundance remains an essential tool in the
45 ecologist's toolbox²³.

46 The family of self-limiting population-growth models including logistic growth curves
47 ('phenomenological models' hereafter) are convenient for describing density-feedback
48 signals in abundance time series³. These models use census data to quantify the net effect of
49 population size N on the rate of instantaneous growth (r)²⁴. Expressed as a proportional
50 change in N between two time (t) steps (e.g., years or generations), the assumption is that $r =$
51 $\log_e(N_{t+1}/N_t)$ summarises the combination or 'ensemble'² of all feedback mechanisms
52 operating on individual 'component' demographic rates²⁵. The problem is that population
53 growth rates can be insensitive to variation in particular demographic rates²⁶⁻²⁸. Thus, across
54 109 observed censuses of bird and mammal populations, the strength of 'component density
55 feedback' (on demographic rates) explained only $< 10\%$ of the strength of 'ensemble density
56 feedback' (on population growth rate) using phenomenological models and after controlling for
57 time-series length and body size². The reasons for such decoupling are not well understood.

58 Determining the partial effects of different underlying mechanisms responsible for the
59 decoupling of component and ensemble density feedbacks is most often impossible for real
60 abundance time series. This analytical limitation occurs because the multiple, density-
61 dependent and -independent mechanisms generating population fluctuations change
62 themselves through time — a condition known as 'nonstationarity'²⁹. We therefore
63 constructed stochastic, age-structured populations with known, component density feedback
64 on survival and imposed nonstationarity to population size via multiple demographic
65 scenarios emulating density-independent mortality and variation in carrying capacity through
66 time. We then simulated multiannual time series of abundance from those populations and
67 estimated the strength of ensemble density feedbacks from these. Our prediction was that
68 ensemble density feedbacks should track component feedbacks if survival has a demographic
69 impact, mediated by population size, on the population growth rate of long-lived vertebrates,
70 while our demographic framework allowed the quantification of true and false detection of
71 ensemble density feedbacks.

72

73 **METHODS**

74 Our overarching aim was to simulate populations of long-living species and their time series
75 of abundance with various sources of nonstationarity. We describe below the set of test
76 species, the simulation of the base population model, component density feedbacks on
77 survival and time series of population abundance, the demographic scenarios considered, and
78 the phenomenological models used to quantify ensemble density feedbacks.

79

80 ***Test species***

81 As the variability in population growth rates is driven primarily by survival rates for slower
82 life-history species of mammals^{30,31} and birds³², we parameterised the simulated populations
83 to characterise the plausible dynamics of 21 long-lived species of extant ($n = 8$) and extinct (n
84 $= 13$) Australian vertebrates from five taxonomic/functional groups (herbivore
85 vombatiformes and macropodiformes, large omnivore birds, carnivores, and invertivore
86 monotremes), spanning mean adult body masses of 1.7–2786 kg and generation lengths of

87 2.3–21 years³³ (Table 1). These species differ in their resilience to environmental change, and
 88 represent the slow end of the slow-fast continuum of life histories³⁴ where high survival rates
 89 make it possible that reproductive efforts are dispersed over the lifetime of individuals³⁵. A
 90 full justification of the selection of our test species can be found in reference³³.

91

92 **TABLE 1** Taxonomy and life-history characteristics of the 21 test species (all native to Australia) used to
 93 simulate age-structured populations and time series of population abundance. *abb* = abbreviation of scientific
 94 name, *M* = body mass (kg), *GL* = generation length (years), *q* = projection length (years)³³.
 95

taxonomic/functional group	species	abb	<i>M</i>	<i>GL</i>	<i>q</i>	status
herbivore vombatiformes	<i>Diprotodon optatum</i>	DP	2786	18.1	724	extinct
	<i>Palorchestes azael</i>	PA	1000	15.1	604	extinct
	<i>Zygomaturus trilobus</i>	ZT	500	13.2	528	extinct
	<i>Phascolonus gigas</i>	PH	200	10.7	428	extinct
	<i>Vombatus ursinus</i>	VU	25	10.0	400	extant
herbivore macropodiformes	<i>Procoptodon goliath</i>	PG	250	8.3	332	extinct
	<i>Sthenurus stirlingi</i>	SS	150	8.1	324	extinct
	<i>Protomnodon anak</i>	PT	130	7.8	312	extinct
	<i>Simosthenurus occidentalis</i>	SO	120	7.8	312	extinct
	<i>Metasthenurus newtonae</i>	MN	55	6.0	240	extinct
	<i>Osphranter rufus</i>	OR	25	5.5	220	extant
	<i>Notamacropus rufogriseus</i>	NR	14	6.3	252	extant
large omnivore birds	<i>Genyornis newtoni</i>	GN	200	20.0	800	extinct
	<i>Dromaius novaehollandiae</i>	DN	55	5.9	236	extant
	<i>Alectura lathamii</i>	AL	2.2	6.8	272	extant
carnivores	<i>Thylacoleo carnifex</i>	TC	110	9.1	364	extinct
	<i>Thylacinus cynocephalus</i>	TH	20	5.2	208	extinct
	<i>Sarcophilus harrisii</i>	SH	6.1	3.1	124	extant*
	<i>Dasyurus maculatus</i>	DM	2	2.3	92	extant
invertivore monotremes	<i>Megalibgwilia ramsayi</i>	MR	11	16.4	656	extant
	<i>Tachyglossus aculeatus</i>	TA	4	14.1	564	extant

96 * extant in Tasmania, currently extinct in mainland Australia

97

98 **Base (age-structured) population model**

99 The population model for each test species was a stochastic (parameters resampled within
 100 their uncertainty bounds) Leslie transition matrix (**M**) following a pre-breeding design, with
 101 $\omega+1$ (*i*) \times $\omega+1$ (*j*) elements (representing ages from 0 to ω years) for females only, where ω
 102 represents maximum longevity. Fertility (m_x) occupied the first row of the matrix, survival
 103 probabilities (S_x) occupied the sub-diagonal, and the final diagonal transition probability
 104 ($M_{i,j}$) was S_ω for all species — except *Vombatus ursinus* (VU; common wombat), *Thylacinus*
 105 *cynocephalus* (TC; thylacine) and *Sarcophilus harrisii* (SH; devil) for which we set $S_\omega = 0$ to
 106 limit unrealistically high proportions of old individuals in the population given the evidence
 107 for catastrophic mortality at ω for the latter two species³⁶⁻³⁸. Multiplying **M** by a population
 108 vector **n** estimates total population size (Σn) at each forecasted time step³⁹. The base model
 109 was parameterised with $n_0 = ADMw$, where **w** is the right eigenvector of **M** (stable stage
 110 distribution), and *A* is the surface area of the study zone ($A = 250,000$ km²) so that the species

111 with the lowest \mathbf{n}_0 would have an initial population of at least several thousand individuals at
112 the start of the simulations. Based on theoretical equilibrium densities (D , km⁻²) calculated
113 for each taxon³³, the species-specific carrying capacity $K = DA$.

114

115 **Density feedback on survival**

116 We simulated a compensatory density-feedback function by forcing a reduction modifier
117 (S_{red}) of the S_x vector in each model according to $\Sigma \mathbf{n}$:

$$118 \quad S_{\text{red}} = \frac{a}{1 + \left(\frac{\Sigma \mathbf{n}}{b}\right)^c} \quad [\text{eq 1}]$$

119 where the a , b , and c constants for each species are adjusted to produce a stable population on
120 average over 40 generations ($40[G]$; see below)^{40,41}. This formulation avoided exponentially
121 increasing populations, optimised transition matrices to produce parameter values as close as
122 possible to the maximum potential rates of instantaneous increase (r_m)³³, and so ensured that
123 long-term population dynamics were approximately stable at the species-specific K (see
124 previous section). Here,

$$125 \quad G = \frac{\log((\mathbf{v}^T \mathbf{M})_1)}{\lambda_1} \quad [\text{eq 2}]$$

126 $(\mathbf{v}^T \mathbf{M})_1$ is the dominant eigenvalue of the reproductive matrix \mathbf{R} derived from \mathbf{M} , and \mathbf{v} is the
127 left eigenvector³⁹ of \mathbf{M} . Thus, the total projection length in years (q) varied across the 21 test
128 species, from 92 (*Dasyurus maculatus*; DM; spot-tailed quoll) to 800 (*Genyornis newtoni*;
129 GN; mihirung) years (median = 324 years with 95 % interquartiles of [108, 762] years; Table
130 1), with one value of abundance being simulated per year. In each iteration and annual time
131 step, the S_x vector was β -resampled assuming a 5% standard deviation of each S_x and a
132 Gaussian-resampled m_x vector. We deliberately avoided applying density-feedback functions
133 to fertility to isolate the component feedback to a single demographic rate.

134

135 **Nonstationarity**

136 We added nonstationarity to our base population model through a catastrophic (density-
137 independent) mortality function to account for the probability of a catastrophic event (C)
138 scaling to generation length among vertebrates⁴²:

$$139 \quad C = \frac{pC}{G} \quad [\text{eq 3}]$$

140 where pC = probability of catastrophe set at 0.14 given this is the mean probability per
141 generation observed across vertebrates⁴². Once invoked at probability C , a β -resampled
142 proportion centred on 0.5 to the β -resampled S_x vector induced a ~ 50% mortality event for
143 that year⁴³. A catastrophic event is defined as "... any 1-yr peak-to-trough decline in
144 estimated numbers of 50% or greater"⁴². The catastrophic function recreates the demographic
145 effects of a density-independent process such as extreme weather events, fires, disease
146 outbreaks, or human harvest. However, we considered the process here as a standard
147 perturbation in all models, and then added specific types of additional perturbations per
148 scenario (see demographic scenarios below).

149

150 **Abundance time series**

151 From the base models (parameterised to incorporate age structure, density feedbacks on
152 survival, and catastrophic events in the Leslie matrices as described above), we generated

153 multiannual abundance time series up to $40[G]$ for each species. We standardised projection
154 length to $40[G]$ because there is strong evidence that the length of a time series (q) dictates
155 the statistical power to detect an ensemble density-feedback signal in logistic growth curves⁶.
156 Here, we summed the \mathbf{n} vector over all age classes to produce a total population size $N_{t,i}$ for
157 each year t of each iteration i . We rejected the first $[G]$ -equivalent years of each projection as
158 a burn-in to allow the initial (deterministic) age distribution to calibrate to the stochastic
159 expression of stability under compensatory density feedback.

160 To ascertain the degree of nonstationary in the simulated abundance time series, we used
161 Turchin's²⁹ definition of nonstationarity as temporally variant mechanisms generating
162 population fluctuations. In that conceptual context, we calculated the mean and variance of
163 return time (T_R) — defined as the time required to return to equilibrium following a
164 disturbance⁴⁴ — for each abundance time series as:

$$165 \quad \bar{T}_R = \frac{\sum_{m=1}^M T_{Rm}}{M} \quad [\text{eq 4}]$$

166 where \bar{T}_R is the mean T_R across M steps of the time series. For each m^{th} time step,

$$167 \quad T_{Rm} = S_{Cm} + S_{Fm} \quad [\text{eq 5}]$$

168 where: S_{Cm} is the number of complete time steps taken before reaching T_{Rm} , and S_{Fm} is the
169 fraction of time required to reach T_{Rm} in the M^{th} (final) step:

$$170 \quad S_{Fm} = \frac{N_p - \bar{N}}{N_p - N_a} \quad [\text{eq 6}]$$

171 where \bar{N} is the mean of the abundance time series (a proxy for K), N_p is the population size
172 prior to crossing \bar{N} , and N_a is the population size after crossing \bar{N} .

173 The variance of T_R is:

$$174 \quad \text{Var}(T_R) = \frac{\sum_{m=1}^M (T_{Rm} - \bar{T}_R)^2}{M-1} \quad [\text{eq 7}]$$

175 Thus, when $\bar{T}_R \ll \text{Var}(T_R)$ (i.e., $\bar{T}_R/\text{Var}(T_R) \ll 1$), the time series is considered to be highly
176 nonstationary⁴⁴.

177

178 **Demographic scenarios**

179 We generated 10,000 abundance time series over $40[G]$ for each test species in each of nine
180 demographic scenarios that combined different types and magnitudes of nonstationarity in the
181 form of density-independent (catastrophic and proportional) mortality and variation in
182 carrying capacity (K) through time. Each times series represented the idiosyncratic
183 demography of a unique population occupying an area of 250,000 km² with zero dispersal
184 (see above).

185 We split the nine scenarios into two main groups: **(1)** eight testing the probability of a
186 false negative (reduced detection of ensemble density feedback when a component feedback
187 on survival existed), and **(2)** one testing the probability of a false positive (evidence of
188 ensemble density feedback when a component feedback on survival was absent) (see details
189 in Table 2). The false-negative scenarios included three subcategories: **(1.1)** *i.* fixed K with
190 no perturbations other than the stochasticity imposed by resampling demographic rates in the
191 Leslie matrices; **(1.2)** fixed K with generationally scaled catastrophes centred on 50%
192 mortality *ii.* leading to $\bar{r} \cong 0$, *iii.* as in *ii.*, but with an additional, single ‘pulse’ perturbation of
193 90% mortality applied across the entire age structure at 20 generations, *iv.* a ‘harvest’-like

194 process where a consistent proportion of individuals is removed from the \mathbf{n} vector at each
195 time step to produce $\bar{r} \cong -0.001$ (i.e., weak, monotonic decline in average population size), or
196 v . as in *iv*, but where the resultant $\bar{r} \cong -0.01$ (i.e., strong, monotonic decline in average
197 population size); and **(1.3)** K fluctuations with v_i . stochastically resampled K with a constant
198 \bar{K} and a constant variance (via resampling the b parameter in equation [1]), **vii.** as in *vi*, but
199 where the resampling variance doubles over the projection interval (via a linear increase in
200 the standard error used to resample the b parameter in equation [1]), and **viii.** as in *vi*, but
201 where K declines at a rate of 0.001 over the projection interval (via decreasing the b
202 parameter in equation [1]). **2.** The false-positive scenario **2ix.** tested for false positives in the
203 ensemble signal by imposing a density-independent mortality via an increase in the
204 probability of catastrophe p_c in equation [3] to produce $\bar{r} \cong 0$ over $40[G]$. In that scenario,
205 we removed the component density-feedback on survival (i.e., setting $S_{\text{red}} = 1$) —
206 theoretically, populations lack a carrying capacity in the absence of density feedbacks.
207

208 **TABLE 2** Demographic scenarios to quantify the detection of ensemble density-feedback signals in time series
 209 of abundance using phenomenological models (logistic growth curves) if a component density feedback on
 210 survival is present (1. H_0 : false negatives), or absent (2. H_0 : false positives). All scenarios were simulated over
 211 40 generations across 21 vertebrate species. Time series obtained from simulated age-structured populations
 212 (Leslie matrices) occupying 250,000 km² with no dispersal. G = generation, N = population abundance, K =
 213 carrying capacity; \bar{r} = long-term mean instantaneous rate of population change, SD = standard deviation. See
 214 test species in Table 1.
 215

scenario	catastrophe type	description
1. H_0: false negatives (component feedback)		
<i>1.1 no catastrophic mortality or fluctuation in K</i>		
<i>i. $K_{\text{fixed}}, \bar{r} \cong 0$</i>	none	stochastically resampled survival rates in age-structured population
<i>1.2 catastrophic mortality (50%) and stable K</i>		
<i>ii. $K_{\text{fixed}}, \bar{r} \cong 0$; sustained catastrophic mortality</i>	generationally scaled	as in <i>i</i> , but with catastrophes
<i>iii. $K_{\text{fixed}}, \bar{r} \cong 0$; pulsed catastrophic mortality</i>	generationally scaled	as in <i>ii</i> , but with a single 90% mortality pulse implemented at $20G$
<i>iv. $K_{\text{fixed}}, \bar{r} \cong -0.001$; sustained proportional mortality</i>	generationally scaled	as in <i>ii</i> , but with proportional removal of individuals from the \mathbf{n} vector such that $\bar{r} = -0.001$ (slowly declining population)
<i>v. $K_{\text{fixed}}, \bar{r} \cong -0.01$; sustained proportional mortality</i>	generationally scaled	as in <i>iv</i> , but where $\bar{r} = -0.01$ (rapidly declining population)
<i>1.3 catastrophic mortality (50%) and fluctuation in K</i>		
<i>vi. $K_{\text{stochastic}}, \bar{r} \cong 0$</i>	generationally scaled	as in <i>ii</i> , but normally distributed K varying randomly at each time step (SD = 5%)
<i>vii. $K_{\text{stochastic}}$ with increasing variance; $\bar{r} \cong 0$</i>	generationally scaled	as in <i>vi</i> , but variance in K increased linearly from 5% to 10%
<i>viii. $K_{\text{stochastic}}$ declining, forcing $\bar{r} < 0$</i>	generationally scaled	as in <i>vi</i> , but K also decreases on average at a rate of -0.001
2. H_0: false positives (no component feedback)		
<i>ix. no K; $\bar{r} \cong 0$</i>	temporally scaled	probability of catastrophe increased over time such that $\bar{r} \cong 0$ (~ average stability)

216
 217

218 **Ensemble density feedbacks**

219 After generating 10,000 time series for each of the 21 species following the nine
 220 demographic scenarios (totalling 189,000 individual time series), we applied
 221 phenomenological models to each time series to test the statistical *evidence* for an ensemble
 222 compensatory density feedback, as well as quantify the *strength* of that feedback. Our
 223 phenomenological models included four variants of the general logistic growth curve⁴⁵
 224 following reference⁶:

$$225 \quad r = \log_e \left(\frac{N_{t+1}}{N_t} \right) = \alpha + \beta N_t + \varepsilon_t \quad [\text{eq 8}]$$

226 where N_t = population size at time t , α = intercept, β = strength of ensemble density feedback,
227 and ε_t = Gaussian random variable with a mean of zero and a variance σ^2 reflecting
228 uncorrelated stochastic variability in the instantaneous rate of population change r . Our first
229 two models are simple density-independent models (DI): (1) random walk, where $\alpha = \beta = 0$,
230 and (2) exponential growth, where $\beta = 0$. The second two variants are density-dependent or
231 density-feedback models (DF): (3) Ricker-logistic⁴⁶, and (4) Gompertz-logistic⁴⁷, where N_t
232 on the right side of equation [8] is replaced with $\log_e(N_t)$. The latter two models represent
233 alternative situations where population growth rate varies in response to unit (Ricker) or
234 order-of-magnitude (Gompertz) changes in population size¹.

235 After fitting each of the four phenomenological models to each time series, we calculated
236 their relative likelihood by means of the Akaike's information criterion (AIC) corrected for
237 finite number of samples (AIC_c). We then expressed the *evidence* for an ensemble density-
238 feedback signal Pr(DF) as the sum of AIC_c weights (w_{AIC_c} = model probability)⁴⁸ for the
239 Ricker- and Gompertz-logistic models (i.e., Σw_{AIC_c} -DF), and the *evidence* for a lack of such
240 signal as the sum of AIC_c weights for random walk and exponential growth (i.e., Σw_{AIC_c} -
241 DI). This follows the logic that if $\beta \neq 0$ between r and N_t (Ricker) or $\log_e(N_t)$ (Gompertz) is
242 more likely than $\beta = 0$ (random walk and exponential growth), then there is stronger
243 statistical support for an ensemble density feedback in the time series than not (i.e., Σw_{AIC_c} -
244 DF > Σw_{AIC_c} -DI implies Pr(DF) > 0.5).

245 We estimated the *strength* of the ensemble density-feedback signal as the negative value
246 of $\hat{\beta}$ estimated from the Gompertz-logistic model. We used the Gompertz-logistic $\hat{\beta}$, instead
247 of the Ricker-logistic $\hat{\beta}$, to estimate this strength because only the former characterises the
248 multiplicative nature of demographic rates^{2,49}. To compare the component density feedback
249 applied to survival in the stochastic age-structured models to the ensemble density feedback
250 estimated from the abundance time series under the nine demographic scenarios, we plotted
251 the negative value of Gompertz $\hat{\beta}$ relative to $1 - S_{red}$ across all 21 species modelled.

252 We tested the correlation between ensemble and component density-feedback strength,
253 and between ensemble strength and the degree of stationarity, across species by calculating a
254 bootstrapped estimate of Spearman's correlation ρ (treating relative differences in the metrics
255 as ranks). We uniformly resampled 10,000 times from the 95% confidence interval of each
256 metric for each species and demographic scenario, calculating ρ in turn, and then calculating
257 the median and 95% confidence interval of ρ . The relationships between ensemble and
258 component density-feedback strength (as well as between ensemble strength and stationarity)
259 showed some non-linearity, so we also fit simple exponential plateau models of the form $y =$
260 $y_{max} - (y_{max} - y_0)e^{-kx}$ to these relationships. Here, y_0 is the starting value of component strength,
261 y_{max} is the maximum component strength (- Gompertz $\hat{\beta}$), k = rate constant (in units of x^{-1}),
262 and x is the component strength ($1 - S_{red}$).

263

264 RESULTS

265 *Statistical evidence for density feedback*

266 For each test species, when the simulated populations were subjected to a compensatory
267 density feedback on survival (age-structured Leslie matrices), the median probability for a
268 statistical signal of ensemble compensatory density-feedback (Pr(DF) = Σw_{AIC_c} -DF; see

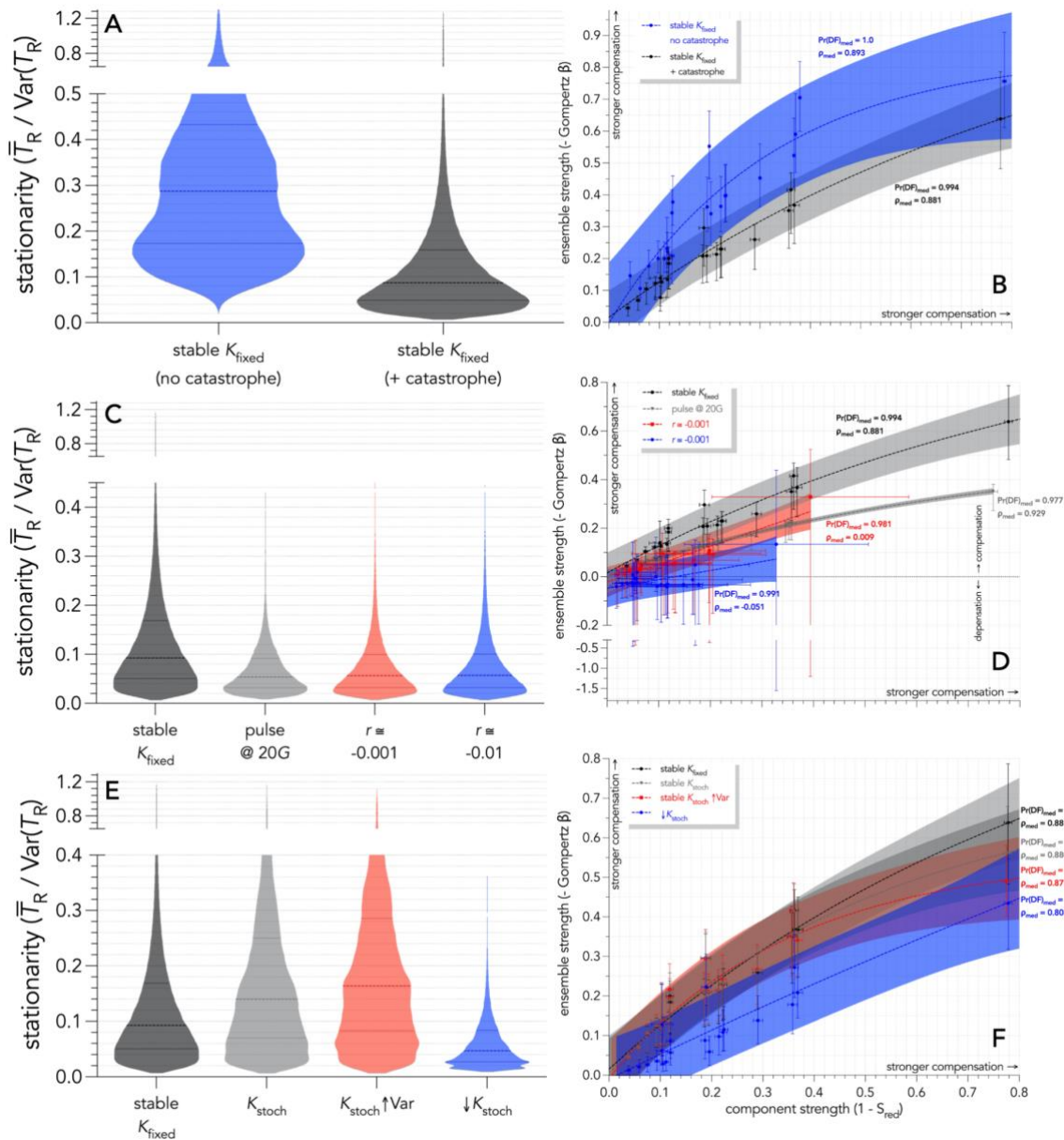
269 Materials and methods) across 10,000 times series of abundance was near unity (> 0.99) for
270 the stable ($\bar{r} \cong 0$) trajectories and most demographic scenarios (Fig. S1–S2 and S3 for
271 probability density plots of Pr(DF) across scenarios and the bootstrapped mean Pr(DF) per
272 species and scenario, respectively). Only the declining stochastic K scenario (1.3viii) had a
273 slightly smaller median Pr(DF) at 0.95. For the false-positive scenario (2ix), the median
274 Pr(DF) was 0.322. Generally, the extant dasyurid *S. harrissii* (SH; devil) and the flightless
275 bird *Dromaius novaehollandiae* (DN; emu) had the weakest evidence for density feedback
276 across the different scenarios (Fig. S3).

277 In summary, if a component density feedback was present, the phenomenological models
278 mostly detected the ensuing ensemble feedback (true positive) — regardless of whether a
279 simulated population was perturbed via density-independent removal of individuals, or
280 altered K dynamics — in > 9 of every 10 time series; while false positives (component
281 feedback absent, ensemble feedback detected) occurred in < 4 of every 10 times series.
282

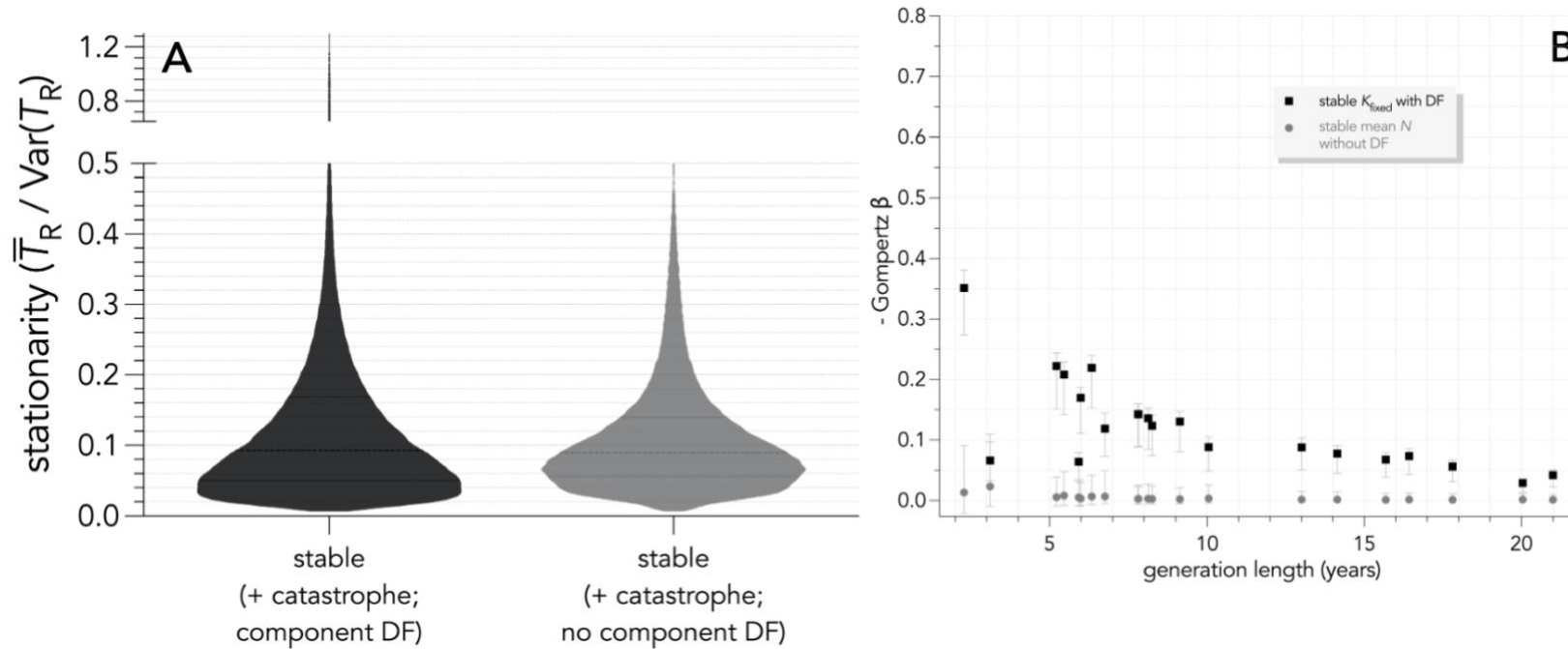
283 ***Degree of simulated stationarity***

284 The addition of the generationally scaled 50% catastrophic (density-independent) mortality
285 reduced stationarity from a median of $\bar{T}_R/\text{Var}(T_R) \sim 0.28$ (scenario 1.1i) to ~ 0.08 (scenario
286 1.2ii) (Fig. 1A). The scenarios imposing a catastrophic 90% mortality as a pulse at 20
287 generations (1.2iii), or additional proportional mortality driving a moderately (1.2iv; \bar{r}
288 $= -0.01$) or rapidly (1.2v; $\bar{r} = -0.001$) declining population over 40 generations, all reduced
289 stationarity by approximately the same amount relative to the scenario without catastrophic
290 mortality (1.1i) (Fig. 1C). For the scenarios emulating fluctuations in K (1.3vi–viii), adding
291 stochasticity to K slightly increased stationarity relative to a fixed- K scenario (Fig. 1E). Only
292 when the stochastic K was forced to decline (scenario 1.3viii), the abundance time series
293 became highly nonstationary (Fig. 1E). The false-positive scenario (2.ix) resulted in
294 negligible change to stationarity when comparing populations experiencing (Fig. 2A), or not
295 experiencing (Fig. 2B), a component density feedback on survival.

296 **FIGURE 1 (A, C, E)** Truncated violin plots showing the distribution of the stationarity index $\bar{T}_R / \text{Var}(T_R)$
 297 across 10,000 times series of population abundance per species and all 21 test species (see list in Table 1)
 298 obtained from age-structured populations subjected to a compensatory component density feedback on survival
 299 over 40 generations, according to nine demographic scenarios (detailed in Table 2). **(B, D, F)** Relationship
 300 between strength of ensemble (slope coefficient β of the Gompertz-logistic model $\times [-1]$) and component (1 -
 301 the modifier S_{red} on survival) density feedback. **(A-B)** Scenarios without (blue: scenario 1.1*i*) and with (grey:
 302 scenario 1.2*ii*) generationally scaled 50% catastrophic (density-independent) mortality. **(C-D)** Stable projections
 303 with carrying capacity (K) fixed (darker grey; scenario 1.2*ii*), a pulse disturbance of 90% mortality at the first 20
 304 generations (20G; lighter grey; scenario 1.2*iii*), weakly declining ($r \cong -0.001$; red; scenario 1.2*iv*), and strongly
 305 declining ($r \cong 0.01$; blue; scenario 1.2*v*). **(E-F)** Stable projections with K fixed (darker grey; scenario 1.2*ii*),
 306 varying stochastically (K_{stoch}) around a constant mean with a constant variance (lighter grey; scenario 1.3*vi*),
 307 varying stochastically with a constant mean and an increasing variance ($K_{\text{stoch}} \uparrow \text{Var}$; red; scenario 1.3*vii*), and
 308 varying stochastically with a declining mean and a constant variance ($\downarrow K_{\text{stoch}}$; blue; scenario 1.3*viii*). The fitted
 309 curves across species are exponential plateau models of the form $y = y_{\text{max}} - (y_{\text{max}} - y_0)e^{-kx}$. Shaded regions
 310 represent the 95% prediction intervals for each type. Also shown are the mean probabilities of median density
 311 feedback (Pr(DF): sum of the Akaike's information criterion weights for the Ricker- and Gompertz-logistic
 312 models across time series ($\sum w\text{AIC}_c\text{-DF}$). Compensation implies that survival and population growth wane as
 313 population abundance rises, and $\bar{T}_R \gg \text{Var}(T_R)$ implies high stationarity.



315 **FIGURE 2** (A) Truncated violin plots showing the distribution of the stationarity index $\bar{T}_R/\text{Var}(T_R)$ across 10,000 times series of population abundance per species and all
 316 21 species (see species list in Table 1) obtained from age-structured populations subjected to a compensatory component density feedback on survival over 40 generations,
 317 according to two demographic scenarios (detailed in Table 2). Demographic scenarios include carrying capacity (K) fixed with (darker grey, scenario 1.2*ii*) and without
 318 (lighter grey, scenario 2*ix*) component compensatory density-feedback on survival, the latter including an increase in the probability of 50% catastrophic (density-
 319 independent) mortality to produce stable population growth rates around 0 (see scenarios in Table 2). (B) Relationship between strength of ensemble (slope coefficient $\beta \times [-$
 320 1] of the Gompertz-logistic model) and generation length (years) across the 21 species. Probabilities of density feedback (Pr(DF) = sum of the Akaike's information criterion
 321 weights for the Ricker and Gompertz models) calculated across simulations gave median Pr(DF) = 0.994 and 0.322 for the two stable scenarios without and with component
 322 feedback on survival, respectively.



323

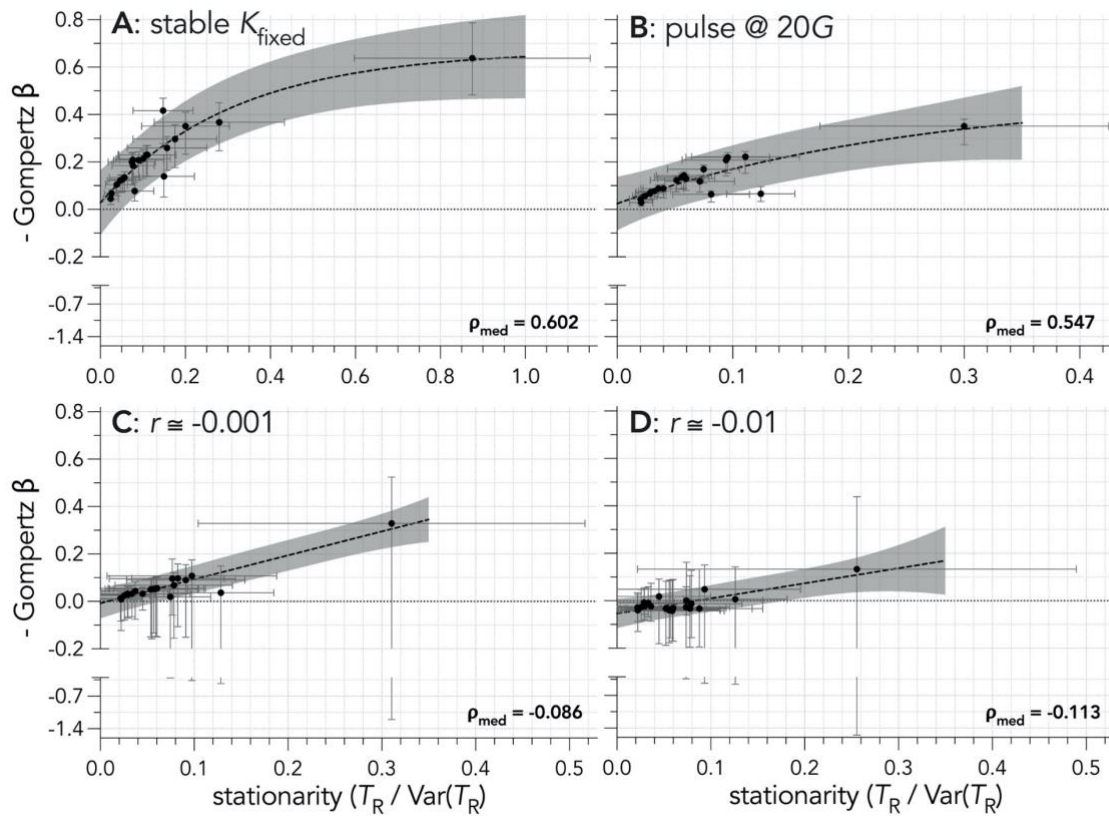
324 ***Strength of density feedback***

325 While the magnitude of statistical evidence for density feedback was largely invariant across
326 all demographic scenarios including a component density feedback on survival (Fig. S1 and
327 S2; see above), the estimated strength of the ensemble density feedback (-Gompertz β , see
328 Materials and methods) was highly sensitive to the type of perturbation the population
329 experienced. The addition of the generationally scaled 50% catastrophic (density-
330 independent) mortality under a fixed K (scenarios 1.1*i* vs. 1.2*ii*) reduced the correlation
331 (median $\rho = 0.893$ and 0.881 , respectively) and slope between ensemble feedback strength
332 and component feedback strength ($1 - S_{\text{red}}$) across the 21 test species (Fig. 1B). The
333 catastrophic pulse scenario (1.2*iii*) returned the closest correlation (median $\rho = 0.929$)
334 between ensemble and component feedback strengths, although it also depressed the slope of
335 the relationship relative to the K_{fixed} scenario (Fig. 1D). These correlations were weakest for
336 the $\bar{r} = -0.001$ and $\bar{r} = -0.01$ scenarios (1.2*v-vi*; median $\rho = 0.009$ and -0.051 , respectively),
337 which also captured a signal of depensation (population growth rate increases with
338 population size) in some abundance time series (Fig. 1D). For the demographic scenarios
339 emulating fluctuations in K (1.3), the correlation between unit change in ensemble and
340 component density feedback strength was generally higher than those where $\bar{r} < 0$ (Fig. 1F;
341 median ρ ranging from 0.803 to 0.881), with the strongest mismatch occurring when K
342 declined by a rate of 0.001 (scenario 1.3*viii*) (Fig. 1F; see also Fig. S4). For the false-positive
343 scenario (2*ix*), all estimated ensemble feedback strengths enveloped 0 (Fig. 2B), meaning that
344 the estimated slopes of the $r \sim \log_e(N_t)$ relationships could not be differentiated from zero.

345 Overall, when an ensemble density feedback was detected from time series of abundance,
346 density-independent mortality eroded the extent by which true compensatory density
347 feedbacks on survival translated into an ensemble compensatory density feedback in
348 population trends more than fluctuations in K , with the most faulty outcome in fact inferring
349 depensatory population growth rates from some populations only experiencing density
350 compensation on survival.

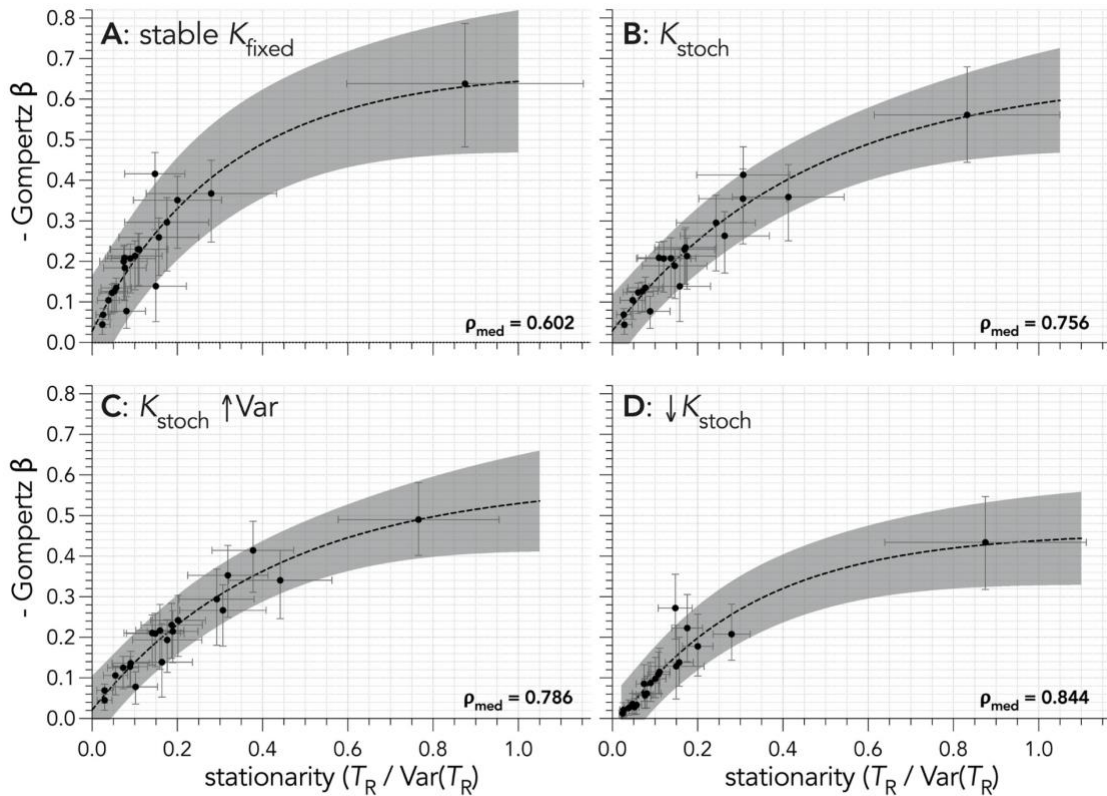
351 On the other hand, the stationarity metric $\bar{T}_R/\text{Var}(T_R)$ was a weak (median $\rho =$
352 0.547 , -0.086 , and -0.113 for the pulse, $\bar{r} = -0.001$, and $\bar{r} = -0.01$ scenarios, respectively)
353 predictor of the estimated strength of ensemble feedback when density-independent mortality
354 was imposed (Fig. 3). However, stationarity was a reasonable (median $\rho = 0.756$, 0.786 , and
355 0.844 for the $K_{\text{stochastic}}$, $K_{\text{stochastic}}$ with increasing variance, and declining $K_{\text{stochastic}}$ scenarios,
356 respectively) predictor of the ensemble signal for the fluctuating K scenarios (Fig. 4; see also
357 Fig. S4).

358 **FIGURE 3** Relationships between the stationarity index $\bar{T}_R/\text{Var}(T_R)$ and the strength of ensemble density
359 feedback (slope coefficient $\beta \times [-1]$ of the Gompertz-logistic model) for four scenarios with 50% catastrophic
360 (density-indepedent) mortality across 21 test species (see Table 1) over 40 generations, including (A) carrying
361 capacity (K) fixed (scenario 1.2ii), (B) a pulse disturbance of 90% mortality at 20 generations (20G; scenario
362 1.2iii), (C) weakly declining ($r \cong -0.001$, scenario 1.2iv), and (D) strongly declining ($r \cong 0.01$, scenario 1.2v)
363 populations (scenarios detailed in Table 2). The fitted curves across species exponential plateau models of the
364 form $y = y_{\max} - (y_{\max} - y_0)e^{-\lambda x}$. Shaded regions represent the 95% prediction intervals for each type. ρ_{med} are the
365 median Spearman's ρ correlation coefficients for the relationship between the ensemble strength and stationarity
index across species (resampled 10,000 times; see Fig. S4 for full uncertainty range of ρ in each scenario).



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369 **FIGURE 4** Relationships between the stationarity index $\bar{T}_R/\text{Var}(T_R)$ and the strength of ensemble density
370 feedback (slope coefficient $\beta \times [-1]$ of the Gompertz-logistic model) across 21 test species (see list in Table 1)
371 over 40 generations for four scenarios (scenarios detailed in Table 2) with 50% catastrophic (density-
372 independent) mortality, including (A) carrying capacity (K) fixed (scenario 1.2ii), (B) K varying stochastically
373 (K_{stoch}) around a constant mean with a constant variance (scenario 1.3vi), (C) K varying stochastically with a
374 constant mean and increasing variance ($K_{\text{stoch}} \uparrow \text{Var}$, scenario 1.3vii), and (D) K varying stochastically with a
375 declining mean and a constant variance ($\downarrow K_{\text{stoch}}$, scenario 1.3viii). The fitted curves across species exponential
376 plateau models of the form $y = y_{\text{max}} - (y_{\text{max}} - y_0)e^{-kx}$. Shaded regions represent the 95% prediction intervals for
377 each type. ρ_{med} are the median Spearman's ρ correlation coefficients for the relationship between the ensemble
378 strength and stationarity index across species (resampled 10,000 times; see Fig. S4 for full uncertainty range
379 under each scenario).



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DISCUSSION

Our simulations reveal several new insights into how ensemble (population growth rates) and component (vital rates) density feedbacks can be decoupled. First, the statistical detection of true ensemble feedback strength through phenomenological models is little affected by nonstationarity *per se*. Second, the estimation of ensemble feedback strength through phenomenological models (logistic growth curves; see Introduction) are particularly sensitive to density-independent mortality leading to population decline, but they are less sensitive to moderate fluctuations in carrying capacity. Third, the concern that density-independent processes can invoke false evidence of ensemble signals of compensation are not borne out by our simulations, at least with respect to density-independent mortality.

The mechanisms underlying those trends are nuanced by species' life histories. For instance, in long-living terrestrial vertebrates (our focus), density feedbacks might operate on fertility to compensate for pathogen-induced adult mortality⁵⁰, those feedbacks might be stronger on survival *versus* fertility when populations are near or far from carrying capacity, respectively⁵¹, and survival can be entirely driven by climatic conditions and density-independent predation⁵². In one of the best-studied systems in this regard, Soay sheep from St. Kilda Archipelago (United Kingdom) demonstrate that the demographic role of density and weather varies across sexes and age classes in mild winters, but survival is reduced consistently in all individuals in years of bad weather and high population abundance⁵³. Much less-studied than herbivores, inter-pack aggression in carnivores with strong social hierarchies like wolves might shape survival at high densities, but be demographically irrelevant at low densities resulting from prey shortages and/or hunting or culling⁵⁴. Our study lends credence to the application of phenomenological models to the former types of studies addressing the long-term effect of vital rates on population abundance, provided there is enough information available for describing population trends.

Our approach and results do not, of course, explain all possible scenarios leading to the decoupling of ensemble and component feedback signals. For example, many other density-independent factors that we did not consider can dampen the demographic role of social and trophic interactions mediated by population size². Along with the confounding effects of sampling error^{55,56}, some of those factors include immigration⁵⁷, spatial heterogeneity in population growth rates^{58,59}, fluctuating age structure⁶⁰, and environmental state shifts^{29,61,62}. Furthermore, our choice to limit the component mechanisms to feedback on a single demographic rate (albeit, applied to all age classes) for the sake of simpler interpretation could limit the application of our conclusions. For example, additional density-feedback mechanisms operating independently on other demographic rates, such as fertility and dispersal, could potentially complicate the interpretation derived from phenomenological models.

Simulating closed populations potentially inflated the phenomenological model's capacity to detect the component signal, because permanent dispersal could alleviate per capita reductions in fitness as a population approaches carrying capacity. We also limited our projections to a standardised 40 generations, but even expanding these to 120 generations resulted in little change in the stationarity metric (Fig. S5). Complementary studies focussing on the faster end of the life-history continuum could provide further insights, even though our

range of test species still precipitated a life-history signal in terms of component (Fig. S6) and ensemble density-feedback strengths and stationarity (Fig. S7, S8) declining with increasing generation length. However, this relationship faded when the trajectories simulated declines through proportional removal. Indeed, both evidence for⁶³ and strength³⁴ of ensemble density feedback generally increase along the continuum of slow to fast life histories, because species with slow life histories are assumed to be more demographically stable when density compensation is operating⁶⁴.

While quantifying the true extent of all component density feedback mechanisms operating in real populations will remain challenging in most circumstances, phenomenological models can normally capture the evidence for and strength of the component density feedback mechanism at play. Appreciating the degree of nonstationarity and other types of perturbations affecting abundance time series can contextualise interpretations of ensemble density-feedback signals, especially where substantial density-independent mortality leads to long-term population declines. Importantly, failing to capture density feedback in applied ecological models can lead to suboptimal conservation and management recommendations and outcomes^{2,65}.

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AUTHOR CONTRIBUTIONS

CJAB conceived the idea, ran the simulations, and wrote the first draft. SHP reviewed the literature. Both authors contributed to revisions.

DATA AVAILABILITY STATEMENT

All data files and R code are openly available at <https://github.com/cjabradshaw/DensityFeedbackSims>.

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