

1 **Title: A Genome-Wide Association Analysis Reveals Epistatic Cancellation of Additive**
2 **Genetic Variance for Root Length in *Arabidopsis thaliana***

3

4 **Short Title: Epistatic Cancellation of Additive Genetic Variance**

5

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22

1 **Abstract**

2 Efforts to identify loci underlying complex traits generally assume that most genetic variance is
3 additive. Here, we examined the genetics of *Arabidopsis thaliana* root length and found that the
4 narrow-sense heritability for this trait was statistically zero. This low additive genetic variance
5 likely explains why no associations to root length could be found using standard additive-model-
6 based genome-wide association (GWA) approaches. However, the broad-sense heritability for
7 root length was significantly larger, and we therefore also performed an epistatic GWA analysis
8 to map loci contributing to the epistatic genetic variance. This analysis revealed four interacting
9 pairs involving seven chromosomal loci that passed a standard multiple-testing corrected
10 significance threshold. Explorations of the genotype-phenotype maps for these pairs revealed
11 that the detected epistasis cancelled out the additive genetic variance, explaining why these loci
12 were not detected in the additive GWA analysis. Small population sizes, such as in our
13 experiment, increase the risk of identifying false epistatic interactions due to testing for
14 associations with very large numbers of multi-marker genotypes in few phenotyped individuals.
15 Therefore, we estimated the false-positive risk using a new statistical approach that suggested
16 half of the associated pairs to be true positive associations. Our experimental evaluation of
17 candidate genes within the seven associated loci suggests that this estimate is conservative; we
18 identified functional candidate genes that affected root development in four loci that were part of
19 three of the pairs. In summary, statistical epistatic analyses were found to be indispensable for
20 confirming known, and identifying several new, functional candidate genes for root length using
21 a population of wild-collected *A. thaliana* accessions. We also illustrated how epistatic
22 cancellation of the additive genetic variance resulted in an insignificant narrow-sense, but
23 significant broad-sense heritability that could be dissected into the contributions of several

1 individual loci using a combination of careful statistical epistatic analyses and functional genetic
2 experiments.

3

4 **Author summary**

5 Complex traits, such as many human diseases or climate adaptation and production traits in
6 crops, arise through the action and interaction of many genes and environmental factors. Classic
7 approaches to identify contributing genes generally assume that these factors contribute mainly
8 additive genetic variance. Recent methods, such as genome-wide association studies, often
9 adhere to this additive genetics paradigm. However, additive models of complex traits do not
10 reflect that genes can also contribute with non-additive genetic variance. In this study, we use
11 *Arabidopsis thaliana* to determine the additive and non-additive genetic contributions to the
12 phenotypic variation in root length. Surprisingly, much of the observed phenotypic variation in
13 root length across genetically divergent strains was explained by epistasis. We mapped seven
14 loci contributing to the epistatic genetic variance and validated four genes in these loci with
15 mutant analysis. For three of these genes, this is their first implication in root development.
16 Together, our results emphasize the importance of considering both non-additive and additive
17 genetic variance when dissecting complex trait variation, in order not to lose sensitivity in
18 genetic analyses.

19

20 **Introduction**

21 Identifying the loci underlying quantitative phenotypes is among the central challenges in
22 genetics. The current quantitative genetics paradigm is based on the assumption of additive gene
23 action, despite an increasing body of evidence showing that non-additive effects are crucial to

1 many, if not most, biological systems [1–4]. The key arguments for remaining within the
2 additive paradigm are that many genetic architectures with non-additive gene action display
3 considerable additive genetic variance in populations [5] and that additive model-based
4 approaches have facilitated detection of thousands of loci associated with many complex traits
5 [6]. However, the amount of additive variance contributed by epistasis for example, will vary
6 with the allele-frequencies in populations and therefore be a dynamic property of studied
7 populations [7–13 and references therein]. Therefore, the focus on additive variation alone can
8 leave a considerable amount of genetic variance unexplored [3,14], and for a full dissection of a
9 complex trait, it is often necessary to explore alternative approaches that capture non-additive
10 variation. Statistical epistasis has been shown to be pervasive in both humans and model
11 organisms [2,14]. Epistatic contributions to complex traits have primarily been identified in
12 experimental populations using candidate gene approaches [1,2,4,15–19] and QTL mapping, for
13 example [1,20]. Genome-wide association (GWA) analyses have more recently emerged as an
14 effective method for mapping loci underlying phenotypic variation for complex traits also in
15 natural populations. However, GWA analyses to detect epistatic interactions are still under-
16 utilized in efforts to identify loci contributing to the genetic architecture of complex traits. To
17 obtain a more complete compilation of the loci contributing to the variation of a complex trait, it
18 would therefore be valuable to also account for epistatic interactions in addition to additive
19 effects in existing and future datasets.

20

21 The drawbacks of studying epistasis in natural populations of higher organisms include the cost
22 and time to collect large datasets, the lack of power due to noisy phenotypes, and the difficulties
23 in performing the necessary follow-up studies to replicate the epistatic associations, identify the

1 polymorphic genes, and dissect the underlying mechanisms leading to statistical epistatic
2 associations. Efforts to utilize epistatic interactions for detection of novel loci, however, are of
3 importance for enhancing sensitivity in statistical analyses [21], for selecting the most promising
4 analytical approaches and experimental designs for future work, and for providing insights as to
5 what type of genetic and biological mechanisms can lead to non-additive genetic variance and
6 genotype-phenotype relationships. Public resources have recently been developed for studying
7 genotype-phenotype associations with GWA analyses in model organisms such as *Arabidopsis*
8 *thaliana* [22] and *Drosophila melanogaster* [23]. Many powerful experimental approaches are
9 available in these species, making them attractive for studies aiming to functionally dissect more
10 complex genetic mechanisms such as those underlying statistical epistatic associations. For these
11 reasons, these resources have the potential to contribute to the fundamental work of clarifying
12 how statistical epistatic analyses should more generally be interpreted and used in efforts to
13 dissect the genetic architectures of complex traits.

14 Two major challenges have often been discussed in relation with GWA analyses based on
15 statistical epistatic models: the lack of sufficiently large experimental datasets to obtain
16 reasonable statistical power and the computational complexity when screening the genome for
17 multi-locus epistasis. One of the causes of low power is the need to use stringent, multiple-
18 testing corrected significance-thresholds to account for the large number of statistical tests
19 performed when exploring epistasis between all possible combinations of loci [24]. As allele-
20 frequencies are often skewed in the populations used for GWA analyses, power is decreased
21 further due to the small number of observations in the multi-locus, minor-allele genotype classes.
22 To some extent these challenges can be overcome by increasing the sample size and making
23 efficient use of high-performance computing [25]. Leveraging genotyped populations of model

1 organisms can also overcome statistical limitations. The genomes of model organisms are often
2 smaller, which reduces the search-space during genome-wide scans. The existence of inbred
3 populations, such the *A. thaliana* RegMap [26] or 1001 Genomes Project [14] or the *D.*
4 *melanogaster* Drosophila Genetic Reference Panel [15] collections, provide advantages in power
5 compared to analyses in general outbred populations with heterozygotes. This is particularly true
6 in analyses of epistasis due to the simpler genotype-phenotype maps. Several standard additive
7 GWA analyses have been previously performed for many traits using the publically available *A.*
8 *thaliana* collections and data sets [27,28], including several studies of *A. thaliana* root
9 phenotypes [29–31]. *A. thaliana* has a small genome (~125 Mb), large numbers of readily
10 available inbred accessions, a large knowledge base on how to efficiently score quantitative
11 phenotypes in large numbers of individuals, and powerful experimental resources, such as
12 collections of knock-out lines. This makes such populations into particularly well-suited models
13 for performing exhaustive scans for epistasis and for evaluating whether non-additive genetic
14 analyses are able to uncover novel contributions of genes to the studied phenotypes. Further,
15 these advantages lead to a smaller theoretical decrease in power than expected due to sample size
16 alone.

17 Model organism GWA studies, however, are often based on a hundred or fewer individuals, and
18 valid concerns have been raised regarding the challenges of testing for associations with
19 hundreds of thousands of markers, or millions of combinations of epistatic markers in such small
20 samples. In particular, the chance of detecting a random association will increase especially
21 when the number of individuals in the minor two-locus genotype class is small. This situation is
22 not accounted for by common statistical multiple-testing correction. Further, there is also a risk
23 of statistical epistatic associations arising due to “apparent epistasis” [32] in which the multi-

1 locus genotype tags an unobserved, single polymorphism in the genome. Here, the genetic
2 variance used to detect the association originates from a true genetic effect in the genome, but
3 this effect is hidden in a standard GWA due to its low correlation with the individual genotyped
4 markers. Instead the high-order correlation between the hidden genetic variant and a multi-
5 marker genotype leads to an epistatic variance that can be detectable in the epistatic GWA
6 analysis. Although this scenario has only previously been discussed for linked variants [32], it
7 could occur for unlinked loci if the analyzed population is small and the number of evaluated
8 combinations of loci is large. Also, this scenario goes beyond standard statistics used to account
9 for multiple testing in GWA analyses, as the increased risk for false positives is not due to the
10 strength of the statistical association. Instead, it results from the inability to properly compare
11 alternative explanatory genetic models for the association using the available genotype data.
12 Unless particular care is taken, both of these scenarios could lead to an increased false-positive
13 rate among the reported loci, as well as incorrect interpretations of the presence of biological
14 interactions if significant pairwise statistical epistasis is interpreted as evidence of an underlying
15 genetic architecture with genetic interactions. These potential risks associated with epistatic
16 GWA analyses in small experimental populations should therefore be considered when such
17 analyses are used.

18 Here, we conduct a GWA study to dissect the genetics underlying *A. thaliana* root length in a
19 population of wild-collected accessions. Root length had a statistically zero narrow-sense
20 heritability in this population, but a larger significant broad-sense heritability. The genetic
21 contribution to this trait was therefore unlikely to be detected using standard additive genetic
22 models, and it was not surprising that no locus was associated with root-length in the standard
23 additive-model based GWA analyses. An exhaustive GWA scan for two-locus interactions,

1 however, revealed four genome-wide significant interacting pairs. All of these displayed
2 genotype-phenotype relationships in which epistasis cancelled the marginal additive genetic
3 effects of the interacting loci, explaining the low levels of additive genetic variance in the
4 population. A new statistical method was used to evaluate the robustness of the GWA
5 associations in the small population, and it was estimated that half of these pairs are at risk of
6 being false positives. We also found low risk that the statistical epistatic interactions were caused
7 by “apparent epistasis” [32]. Our experimental evaluations of genes in the associated regions
8 using T-DNA insertion mutants identified functional candidates that affected root length in four
9 regions that were part of three pairs, from which three genes were newly implicated in root
10 development [33]. Together, our results illustrate the value of epistatic GWA analyses for
11 revealing novel loci contributing to complex traits in small, inbred populations. Further, we
12 noticed an increased risk for false-positive epistatic associations and “apparent epistasis” due to
13 the small sample-size but the top-ranked associations proved, despite this, to be valuable in
14 further experimental validation, as several functional candidate genes were identified in loci that
15 would have been discarded based on stringent statistics alone. We also report several promising
16 functional candidate genes for future studies to dissect the molecular mechanisms underlying
17 root development in *natural A. thaliana* populations that would have remained undetected
18 without an epistatic GWA analysis.

19

20 **Results**

21

22 *Epistasis is the primary source of genetic variance for root length*

1 To study the genetic regulation of mean root length, we generated a comprehensive data set
2 containing the *A. thaliana* root mean lengths across 93 wild-collected accessions (Fig S1, Table
3 S1). The 93 accessions used here were previously genotyped at ~215,000 single nucleotide
4 polymorphisms (SNPs) and used in a standard GWA analysis for many plant phenotypes [27].
5 We grew seedlings in a randomized design on standard plant medium in temperature-controlled
6 chambers to reduce environmental variation among individuals [34]. This experimental design
7 allows for parsing of genetic and environmental factors underlying root length [35,36].

8

9 In our dataset, mean root length varied significantly across accessions ($p < 1 \times 10^{-10}$, ANOVA;
10 Fig1). To evaluate whether the observed variation in mean root length was under genetic control,
11 we estimated both the narrow- and broad-sense heritability. This analysis used a linear mixed
12 model that simultaneously estimated the polygenic additive and epistatic variance components
13 using the R/hglm package [37]. We estimated the total broad sense heritability to be 0.24 (95%
14 CI 0.10 – 0.38) for mean root length (Table 1), showing that the trait is under genetic control. We
15 also partitioned the genetic variance into the individual additive and epistatic variance
16 components and found that almost all the variance was epistatic (0.22; $p = 1.32 \times 10^{-6}$) and only
17 a small, insignificant fraction was additive (0.01; $p = 0.63$). These heritability patterns suggest
18 that an epistatic GWA analysis might be needed to identify the loci explaining the heritable
19 phenotypic variation in root length mean.

20

1 **Table 1. Estimated narrow (h^2) and broad sense (H^2) heritabilities of root length mean and**
2 **variance in a population of 93 natural *A. thaliana* accessions.**

3 The broad-sense heritability was intermediate for root length mean, with a negligible additive
4 contribution.

| Trait | H^2 | | h^2 |
|-----------------------|--------------------|---|--|
| | ANOVA ¹ | hglm ² | hglm ² |
| Mean root length (mm) | 0.25 | 0.24 | 0.015 |
| | ($p = 0$) | (s.e. = 0.09, $p = 7.2 \times 10^{-3}$) ³ | (s.e. = 0.056, $p = 0.79$) ³ |

5 ¹Estimated using ANOVA with accession as a fixed effect, ²Estimated using the R/hglm package (hglm) by fitting a
6 linear mixed model including both additive and epistatic kinship matrices as random effects, ³p-values from Wald
7 tests for the heritability being larger than zero, and s.e. are the standard errors estimated via jackknife resampling.

8

9 ***No additive genome-wide associations were detected for mean root length***

10 Two GWA analysis methods were used to screen for additive loci contributing to the phenotypic
11 variation in mean root length. The first method was a standard single-locus additive mixed-
12 model based approach as implemented in the R-package GenABEL [38] that accounted for
13 population-structure by modelling the genomic kinships between the accessions as a random
14 effect. Significance testing was done using a Bonferroni-corrected significance threshold. The
15 second method was based on a whole-genome generalized ridge regression heteroscedastic
16 effects model (HEM), in which all SNPs were included simultaneously and their contributions
17 estimated as random effects using the R/bigRR package [39]. Here, a 5% genome-wide
18 significance threshold was determined by permutation testing. None of the tested SNPs were

1 significantly associated with mean root length in either of these analyses (Fig S1a, b). Given the
2 low estimate for the narrow-sense heritability for mean root length in this population ($h^2 = 0.01$),
3 this outcome was not surprising.

4

5 ***An epistatic genome-wide association analysis uncovered multiple interacting loci that***
6 ***contributed to mean root length***

7 To identify loci that contributed to the epistatic genetic variance for mean root length, we
8 performed an exhaustive two-locus SNP-by-SNP epistasis analysis (PLINKv1.07) [40]. Despite
9 the limited population size, several factors increased our chances of finding novel loci using this
10 approach: the small additive and large epistatic genetic variance (Table 1), the high precision in
11 estimating the phenotypic values due to extensive replication, the presence of only four two-
12 locus genotypes in the population of inbred accessions, and a reduced number of pairwise
13 combinations to test in the small genome. To further enhance power and decrease the risk of
14 false-positives, we excluded pairs of SNPs where the minor two-locus genotype class contained
15 few observations. This was achieved by applying two data filters: first, only SNPs with a minor
16 allele frequency greater than 0.25 were included, and second all SNP-by-SNP combinations with
17 fewer than four accessions in the minor genotype class were removed. In the epistatic analyses,
18 we accounted for population structure by performing the GWA analysis using a linear mixed
19 model including the genomic kinship matrix. Significance for the pairs were determined using a
20 standard Bonferroni-correction for the number of independent pair-wise tests across the
21 independent linkage blocks across the *A. thaliana* genome [41]. In total, the epistatic analysis
22 included approximately 78 million such independent tests. Six SNP-pairs, representing seven
23 genomic locations and four unique combinations of loci, passed this threshold (Table 2). Two

1 pairs were detected twice by associations to tightly linked SNPs located 1.5 kb and 505 bp apart
 2 (Chromosome 3 at 9,272,294 and 9,273,674 bp - SNPs 3_9272294/ 3_9273674 - and
 3 Chromosome at 15,862,026 and 15,862,525 bp - SNPs 5_15862026/5_15862525, respectively),
 4 and one SNP (Chromosome 3 at 66,596 bp - SNP 3_66596) was part of two unique interacting
 5 pairs (Table 2). The other two SNP-pairs contained independent SNPs ($R^2 < 0.8$).

6
7

8 **Table 2. Exhaustive two-dimensional GWAS-scan for epistasis identifies six significant**
 9 **interactions, representing four unique pairs, associated with the root length mean.**

| Pair | SNP1 ¹ | SNP2 ¹ | Interaction (mm) | p-value ² (GWA) | p-value ³ (SS) |
|------|-----------------------------|--|---------------------|-------------------------------|------------------------------|
| 1 | rs346846662 (3_66596) | rs347034954/rs347048583 (3_9273674/3_9272294) | 0.55 | 1.5×10^{-10} | 0.003 |
| 2 | rs347005428 (1_17257526) | rs346595296/rs346511785 (5_15862026/5_15862525) | 0.63 | 1.7×10^{-10} | 0.28 |
| 3 | rs346846662 (3_66596) | rs347177728 (5_18241640) | 0.63 | 3.0×10^{-10} | 0.37 |
| 4 | rs347374065 (3_10891195) | rs347377866 (5_1027939) | 0.63 | 2.3×10^{-10} | 0.95 |

10 ¹Name of associated SNP(s) at the epistatic locus as on the Affymetrix single nucleotide polymorphism (SNP) chip
 11 (our abbreviation as Chromosome_PositionInBp); ²p-value for the interaction effect of the two loci in the epistatic
 12 GWA analysis [40]; ³Probability that the pair is a false-positive when accounting for the small sample-size (SS)
 13 [42].

14

1 ***Explorations of epistatic genotype-phenotype maps revealed an epistatic cancellation of the***
2 ***additive genetic variance***

3 The four SNP pairs that were significant in the epistatic GWA (Table 2; Fig 2a) explained no
4 additive variance, but together 51% of the phenotypic variance of mean root length (95% CI:
5 31.2% - 72.4% from bootstrap). To explore the origin of the low additive variance in the
6 presence of the revealed epistasis, we estimated the multi-locus genotype-phenotype (G-P) maps
7 and their respective genotype frequencies in the analyzed population. The G-P maps were highly
8 related for the four pairs: the minor allele double-homozygote was always associated with the
9 longest mean root length and the shortest mean root lengths were associated with the genotypes
10 combining one major- and one minor-allele homozygote (Fig 2b; Fig S2). Although this type of
11 epistasis will lead to marginal additive genetic effects for the two loci for many allele-
12 frequencies, there are also combinations of allele-frequencies where the epistasis will cancel the
13 additive genetic variance completely. We illustrate this epistatic cancellation of the additive
14 genetic variance for the G-P map and allele-frequencies of the most significant pair in Fig 2c.
15 For this pair, there are minimal differences in the average (marginal) mean root lengths for the
16 two genotypes at the individual loci, and hence a minor contribution to the additive genetic
17 variance. Their joint contribution to the phenotypic variance, however, is large due to the large
18 differences in root lengths among the four two-locus genotype classes. The epistatic GWA
19 analysis is able to reveal these pairs in this population as the differences between these genotypes
20 contribute to the epistatic genetic variance instead.

21

22 ***Evaluating the robustness of statistical epistatic associations to a small population size***

23 We next applied a conservative genome-wide test to estimate the risk that the epistatic GWA
24 associations were false positives due to the small population size. A pseudo-marker was created

1 for each pair of SNPs in the epistatic scan, resulting in an epistatic pseudo-genome with 78
2 million markers. Given the epistatic effect estimated for the four pairs that reached genome-wide
3 significance in the epistatic GWA, we computed the genome-wide probability of these pairs
4 being false-positives using the R/p.exact package [42], assuming conservatively that all the
5 pseudo-markers are independent [42] (Table 3). The pair 3_10891195/5_1027939 was found to
6 be sensitive to population size ($p = 0.94$). The pairs 1_17257526/5_15862026 and
7 3_66596/5_18241640 were less sensitive ($p = 0.28$ and $p = 0.37$, respectively), and hence there is
8 a 90% probability that at least one of these two signals is not false. The last pair
9 3_66596/3_9273674 was not sensitive to population-size ($p = 0.003$). We conclude that it is
10 essential to consider the effect of population-size when interpreting the results from GWA
11 analyses in small populations. In our particular case, half of the original associations were at risk
12 of being false-positives due to small populations size.

13

14

1 **Table 3. Estimation of the risk that epistatic pairs identified in the GWA analysis are due**
 2 **to high-order LD to unobserved functional variants (“apparent epistasis”).**

3 Using the whole-genome re-sequencing data from the reference 1001-genomes *A. thaliana*
 4 collection, we estimated the high-order LD (r^2) between the four pseudo-markers representing
 5 our epistatic pairs and all sequencing variants that were not genotyped in our GWA analysis. A
 6 bootstrap approach was used to estimate the mean and max r^2 between the epistatic pseudo-
 7 markers and all the genome-wide sequencing-variants. The risk that an association might be due
 8 to “apparent epistasis” was calculated as the risk of observing an $r^2_{\max} > 0.8$ in this analysis.

| Significant pairs in two-dimensional epistatic GWA analysis | | | | |
|---|--------------|--------------|--------------|--------------|
| Locus 1 [†] | 3_66596 | | 1_17257526 | 3_10891195 |
| Locus 2 [†] | 3_9273674 | 5_18241640 | 5_15862026 | 5_1027939 |
| r^2_{mean} | 0.01 ± 0.001 | 0.01 ± 0.001 | 0.01 ± 0.003 | 0.01 ± 0.001 |
| r^2_{max} | 0.52 ± 0.10 | 0.66 ± 0.15 | 0.78 ± 0.11 | 0.78 ± 0.09 |
| P($r^2_{\text{max}} > 0.8$) | 0.002 | 0.18 | 0.42 | 0.42 |
| MAF _{PM} | 0.065 | 0.043 | 0.043 | 0.043 |

9 [†]Locus detected as part of epistatic pair named as Chromosome_PositionInBp ; r^2_{mean} : mean high-order LD between
 10 epistatic pseudo-marker and genome-wide sequence variants ± standard error; r^2_{max} : maximum high-order LD
 11 between epistatic pseudo-marker and genome-wide sequence variants ± standard error; P($r^2_{\text{max}} > 0.8$): probability of
 12 observing a high-order LD larger than 0.8 in a random sample of 93 accessions with MAF_{PM} for the epistatic
 13 pseudo-marker; MAF_{PM}: minor-allele frequency for the epistatic pseudo-marker, i.e. the frequency of the minor-
 14 allele double-homozygote for the epistatic pair.

15

16 ***Quantifying the risk of “apparent epistasis” for significant statistical epistatic associations***

17 For “apparent epistasis” to be present, the hidden functional variant needs to be in high linkage
 18 disequilibrium (measured as r^2) with one of the “epistatic pseudo-markers” described above.

19 Therefore, “apparent epistasis” involving physically unlinked markers is only expected for very

1 strong mutations. However, if any such hidden variant exists, by definition neither its genetic
2 effect nor its linkage disequilibrium to the “epistatic pseudo-markers” can be estimated from the
3 GWA datasets. However, if a GWA study is performed in a species with sequenced reference
4 populations, the genomic data from that population can be used to estimate the risk of observing
5 a high r^2 between the tested markers and other polymorphisms in the genome. Here, we used the
6 1001 Genomes Project *A. thaliana* genomes [22] to perform this analysis for the four significant
7 pairs in the epistatic GWA analysis. We defined bi-allelic “epistatic pseudo-markers” by
8 assigning individuals with the two-locus minor-allele homozygote genotype the minor pseudo-
9 marker genotype and all others the major pseudo-marker genotype. By using a bootstrap analysis
10 based on samples of 93 accessions from the 728 whole-genome re-sequenced accessions from
11 the 1001 Genomes Project data, we estimated the probability of observing a high r^2 ($r^2 > 0.8$)
12 between any individual SNP in the re-sequencing data and the “epistatic pseudo-markers” for the
13 four pairs. This analysis revealed that the “epistatic pseudo-markers” corresponding to the
14 identified pairs have very low r^2 to most of the hidden variants (mean $r^2 = 0.01$). There is a very
15 low risk to observe a high r^2 ($r^2 > 0.8$) to any hidden variant in our population for the most
16 strongly associated pair ($p < 0.005$). For the other three pairs, this risk is still low ($p = 0.18 -$
17 0.42) but sufficiently large to indicate that the results should be interpreted with caution (Table
18 3).

19

20 ***Functional exploration of candidate genes in the associated regions using T-DNA insertion*** 21 ***lines***

22 Our statistical analyses suggest that half of the pairs with genome-wide significant statistical
23 epistasis might be false-positives due to the small population-size. As only four pairs were found,

1 we decided to also perform an initial experimental and bioinformatics analysis that included all
2 these pairs to evaluate whether the statistical risk of false-positives was reflected also by the
3 presence of known, or novel, functional candidate genes for root-length in the associated regions.

4
5 We therefore first extracted the genotypes for 13 million genome-wide SNPs from the whole-
6 genome re-sequencing data from the 1001 Genomes Project [43]. SNPs that were in LD ($r^2 >$
7 0.8) and located within 5kb upstream or downstream of the leading SNP in the epistatic GWA
8 were considered further as candidate functional mutations. In Table 4, we list the predicted
9 functional effects of the SNPs that fulfill these criteria, together with their locations, for each
10 pair. In total, thirteen genes and two transposons were found in LD with the leading SNPs in the
11 seven regions detected in the epistatic GWA analysis to root length.

12

13

1 **Table 4. Summary of identified polymorphisms in LD with the leading SNPs from the**
 2 **epistatic GWA analysis.**

3 Four unique interacting pairs of loci were significantly associated with root length (Table 2).
 4 Here, the genes located within 1000bp of the leading SNPs, together with the mutational effects
 5 of the identified polymorphisms, are listed.

| Chr | SNP position (bp) | In pair | Locus | # Mutations (# in LD) ¹ | Mutation Effects |
|-----|--------------------------|---------|-----------|------------------------------------|--|
| 3 | 66596 | 1, 4 | At3g01180 | 1 (0) | 1 upstream |
| | | | At3g01185 | 1 (0) | 1 syn |
| 3 | 9,272,294 9,273,674 | 1 | At3g25520 | 4 (21) | 1 intron, 3 upstream |
| | | | At3g25530 | 14 (21) | 1 mis, 1 splice, 1 syn, 2 intron, 9 upstream |
| | | | At3g25540 | 17 (21) | 1 mis, 4 syn, 2 intron, 5 5'UTR, 5 upstream |
| | | | At3g25545 | 3 (21) | 3 upstream |
| 5 | 18,241,640 | 4 | At5g45120 | 6 (5) | 2 mis, 4 syn |
| 1 | 17,257,526 | 2 | At1g46624 | 1 (0) | 1 non-coding exon |
| 5 | 15,862,026 15,862,525 | 2 | At5g39610 | 2 (4) | 6 upstream ³ |
| | | | At5g39620 | 2 (4) | 6 downstream ³ |
| 3 | 10,891,195 | 3 | At3g28865 | 4 (7) | 4 non-coding exon |
| | | | At3g28870 | 2 (7) | 1 mis, 1 syn |
| | | | At3g28880 | 3 (7) | 3 upstream |
| 5 | 1,027,939 | 3 | At5g03840 | 1 (1) | 1 upstream |
| | | | At5g03850 | 1 (1) | 1 downstream |

6 ¹The number of polymorphisms associated with each listed gene, with the number of these polymorphisms that are
 7 in LD with the leading SNP in parenthesis. Some mutations are associated with more than one gene. ²Abbreviations
 8 for mutational effects: syn = synonymous variant; mis = missense variant; splice = splice site variant. ³SNPs more
 9 than 1000bp away from the nearest gene

1 Further, we identified T-DNA insertion lines for 12/15 genes and transposons comprising the
2 seven statistically epistatic loci and experimentally evaluated them for effects on root length. In
3 Table 5, we present the genes for which the T-DNA insertions significantly altered the root
4 length. Results for all evaluated T-DNA lines are presented in the Supplementary Text. For three
5 of these loci (*NAC6*, *TFL1*, *At2g28880*), we are the first to show that these underlying genes
6 contribute to root growth, and for one locus (*ATL5*), our results confirm previous findings.

7
8 ***T-DNA mutant analysis reveals four functional candidate genes for root development in***
9 ***regions detected using an epistatic genome-wide association analysis***

10 ***Decreased root length in ATL5/OLI5/PGY3 T-DNA insertion mutants***

11 Two linked SNPs on chromosome 3 (3_9272294 and 3_9273674; Table 2) were associated with
12 *A. thaliana* root length through an epistatic interaction with a distal SNP on the same
13 chromosome (3_66596). Twenty-one SNPs located in four genes were in high LD with the two
14 linked SNPs (Table 4; Table 5). For three of these (At3g25520, At3g25530 and At3g25540), we
15 were able to obtain insertion mutants and scored their mean root length compared to wild-type
16 (*Col-0*) controls. The insertion mutant for At3g25520 had significantly shorter roots than wild
17 type ($p = 1.2 \times 10^{-4}$; Tukey's post-hoc test); the insertion mutants for At3g25530 and At3g25540
18 did not differ significantly ($p = 1.00$ and $p = 0.23$, respectively) from wild type (Fig 3).

19

20

- 1 **Table 5. Function and expression patterns of genes in LD with leading SNP from epistatic**
 2 **GWAS analysis.** The gene names or proposed function is listed for all genes harboring
 3 polymorphisms in LD with the leading epistatic SNPs in the whole genome interaction analysis.

| Chr | Position (bp) | In pair | Locus | Gene/Putative function | Highest expression tissue ¹ |
|-----|---------------------------|---------|-----------|---|---|
| 3 | 66596 | 1, 4 | At3g01180 | <i>SS2</i> | cauline leaves |
| | | | At3g01185 | unknown | NA |
| 3 | 9,272,294/ 9,273,674 | 1 | At3g25520 | <i>ATL5</i> | apical meristem, seedling root |
| | | | At3g25530 | <i>LOH1</i> | rosette leaves |
| | | | At3g25540 | <i>GR1</i> | seedling root, imbibed seed |
| | | | At3g25545 | unknown | seedling root (vasculature), imbibed seed |
| 5 | 18,241,640 | 4 | At5g45120 | eukaryotic aspartyl protease family protein | adult root (meristem), imbibed seed |
| 1 | 17,257,526 | 2 | At1g46624 | gypsy-like retrotransposon | NA |
| 5 | 15,862,026/ 15,862,525 | 2 | At5g39610 | <i>NAC6</i> | seedling root (vasculature) |
| | | | At5g39620 | RAB GTPase homolog G1 | adult root vasculature, mature seed |
| 3 | 10,891,195 | 3 | At3g28865 | LINE retrotransposon | NA |
| | | | At3g28870 | histone deacetylase | mature seed |
| | | | At3g28880 | Ankyrin repeat family protein | NA |
| 5 | 1,027,939 | 3 | At5g03840 | <i>TFL1</i> | adult root vasculature (mature zone) |
| | | | At5g03850 | nucleic acid-binding, OB-fold-like protein | seedling root, apical meristem |

- 4 ¹ Locations of high expression were obtained from the BAR eFP *Arabidopsis* Browser (Winter et
 5 al. 2007).

6

7

1 At3g25520 encodes ATL5/OLI5/PGY3, a 5S rRNA binding protein whose promoter is highly
2 accessible in seedling roots [44], corresponding to expression in roots [45]. Our finding agrees
3 with earlier studies, in which the loss-of-function mutant *oli5-1* exhibits shorter primary roots
4 than wild-type [33]. In the tested accessions, three SNPs were found in the *ATL5* promoter and a
5 fourth in an intron, suggesting variable transcriptional regulation across accessions.

6

7 *A novel role for NAC6 in the regulation of root length*

8 Two linked SNPs on chromosome 5 (5_15862026 and 5_15862525; Table 2) were detected by
9 their statistical epistatic interaction with an SNP on chromosome 1 (17,257,526 bp - SNP
10 1_17257526; Table 2). Both of these linked SNPs, as well as the other SNPs in high LD, were
11 intergenic with the nearest gene, *NAC6* (At5g39610; Table 5), which is located 1.9 kb away.
12 *NAC6* is a transcription factor regulating leaf senescence [46] and is highly expressed in
13 senescing leaves and in maturing seeds [45]. We phenotyped the available At5g39610 insertion
14 mutant for *NAC6* and found a significant decrease in mean root length ($p = 5.82 \times 10^{-10}$; Tukey's
15 post-hoc test; Fig 4), strongly implicating *NAC6* in the regulation of root length.

16

17 *The previously uncharacterized gene At3g28880 contributes to root length*

18 A locus on chromosome 3 (10,891,195 - SNP 3_10891195; Table 2) was identified due to its
19 statistical epistatic interaction with a locus on chromosome 5 (1,027,939 bp - SNP 5_1027939;
20 Table 2). SNPs in LD with 3_10891195 were found in three genes: At3g28865, At3g28870, and
21 At3g28880 (Table 5). None of these genes previously had been associated with root length. By
22 examining available insertion mutants, we established a novel role for the gene At3g28880 in
23 root development; its mutant showed significantly decreased mean root length compared to wild-
24 type controls ($p = 4.5 \times 10^{-10}$; Fig 4), whereas At3g28865 did not. At3g28880 encodes a

1 previously uncharacterized ankyrin family protein. There is no information on developmental
2 expression patterns for this gene; although the gene's putative regulatory region is accessible in
3 seeds and seedling tissue [44]. Across the tested accessions, a number of promoter-proximal
4 SNPs were found in At3g28880, possibly altering gene regulation.

6 *Mutant TFL1 increases mean root length*

7 The SNP 5_1027939 was detected via its epistatic interaction with the At3g28880 locus on
8 chromosome 3 (Table 2). This polymorphism and the only other polymorphism in LD are both
9 intergenic between the genes At5g03840 and At5g03850 (Table 5). Phenotyping the insertion
10 mutants for these genes revealed that the mutant for At5g03840, which encodes *TFL1*, showed
11 significantly longer roots than wild-type controls ($p = 5.7 \times 10^{-10}$; Tukey post-hoc test; Fig 4),
12 whereas no significant phenotypic effect was found for the At5g03850 mutant. Interestingly,
13 these results implicate *TFL1* as a repressor of root growth.

14
15 *TFL1*'s role in floral initiation and morphology are well-established in many plant species [47].
16 It also plays a role in regulating protein storage in vegetative tissues, such as roots or seeds [48].
17 *TFL1* is highly expressed in both adult and seedling roots [45]. Our finding that *TFL1* controls
18 root length is particularly poignant as a previous study reported no root length defects for a
19 different *tfl1* mutant with multiple other pleiotropic phenotypes [49]. Our finding extends the
20 functional reach of this multi-functional regulator.

22 **Discussion**

23

1 This study was designed to identify genes regulating seedling root length. The narrow-sense
2 heritability for this trait was not significantly different from zero in our population, whereas there
3 was significant broad-sense heritability (Table 1). Notably, we observed a very large contribution
4 of epistatic variance to the phenotypic variation in mean root length. With these findings in hand,
5 we completed a comprehensive set of GWA analyses to identify additive and epistatic loci
6 affecting this trait.

7
8 Consistent with the low narrow-sense heritability for mean root length, we were unable to
9 identify any additive loci that reached the Bonferroni-corrected significance threshold in our
10 analyses. By accounting for the contribution of pairwise epistatic interactions to the genetic
11 variance of root length, however, we identified seven loci involved in four unique epistatically
12 interacting pairs of loci that together explained a large fraction of the phenotypic variance in the
13 population (Table 2). One of these pairs is highly significant also when accounting for the small
14 population-size, whereas there is a risk that two of the three remaining pairs are false-positives
15 due to this. We also showed that it is unlikely that the statistical epistatic interaction for the most
16 significant pair would result from “apparent epistasis” due to a major, hidden variant in the
17 genome. This risk is, however, not negligible for the three less significant pairs. Based on this,
18 we conclude that despite the fact that much of the genetic variance present in this population can
19 be explained by the genome-wide significant pairwise epistasis detected amongst these loci,
20 additional evidence is needed before any firm conclusions can be drawn based on the other three
21 epistatic pairs detected in the GWA analysis.

22

1 All four pairs detected in the initial epistatic GWA display an epistatic cancellation of additive
2 genetic variance (Fig 2c). The presence of this type of genetic architecture for root-length in this
3 population explains the low narrow-sense heritability in the population and why no loci could be
4 found in the standard GWA analyses. Although further studies are needed to explore how
5 general this phenomenon is, our result contrasts the common view in quantitative genetics that
6 most genetic variance for complex traits is expected to be additive [5]. It further illustrates the
7 importance of taking a richer modeling approach in GWA analyses to improve the sensitivity as
8 a too strict focus on additive effects in the initial statistical analyses might lead to important
9 findings being missed and lead to an underestimation of the full adaptive potential due to the
10 epistatic suppression of selectable additive genetic variance [10–12].

11
12 Our finding of extensive statistical epistasis in this study should not be interpreted as evidence
13 for root development being a biological trait with exceptionally high levels of biological
14 interactions. Rather it is important to notice that, although the levels of statistical epistasis
15 detected here originate in the epistatic genotype-phenotype maps for these pairs, the
16 exceptionally high levels of epistatic variance in the population is even more due to the allele-
17 frequencies for the involved loci in the population (Fig 2c). Hence if the loci displaying the G-P
18 map in Fig 2c instead, for example, would be present at intermediary and equal frequencies, they
19 would contribute high levels of additive genetic variance instead. Consequently, our results is an
20 important reinforcement of the fact that the levels of epistatic variance for a trait in a population
21 will be a dynamic outcome of changes in genotype-frequencies over time even though the
22 underlying genotype-phenotype map might be constant [7–13].

23

1 In practice, however, neither allele frequencies nor genotype-phenotype maps will be known for
2 the loci contributing to a particular trait. As illustrated by our findings for root length in this
3 population, it is important to utilize GWA analysis methods that account for both additive and
4 non-additive genetic variance to detect the contributing loci. Once this is done, further theoretical
5 and experimental work is needed to identify the underlying genotype-phenotype maps and
6 dissect the molecular underpinnings of the statistical associations regardless of whether these
7 were detected mainly via their contributions to the additive and/or epistatic genetic variance. We
8 here propose how to statistically quantify the risks of each inferred epistatic GWA association to
9 guide further experimental work based on these findings. We find that there were statistical
10 reasons to interpret three of the detected epistatic pairs with caution, both due to an increased
11 risk of false-positives due to the small population-size and an inflated risk for “apparent
12 epistasis”. Three novel functional candidate genes were, however, identified in loci detected as
13 part of these associations. This illustrates clearly that the ultimate contribution from a study will
14 rely on the combined use of appropriate statistical analyses to quantify the confidence in a
15 particular locus and an in-depth experimental dissection of the detected associations before a
16 conclusion can be reached about which loci contribute to a particular trait.
17
18 Functional dissection of epistasis, however, is a daunting task in natural populations and
19 therefore model organisms are an indispensable resource in this work. A first step in taking on
20 this challenge is to reduce the list of potential candidate genes and mutations for further, in-depth
21 functional explorations. Here, we use mutational analysis of the genes that harbor SNPs in LD
22 with the associated SNPs in the epistatic GWA analysis to identify the most likely candidate
23 genes for the inferred associations. We tested six of the seven epistatic regions with available T-

1 DNA insertion mutants. For four of the loci, we succeeded in identifying candidate genes
2 affecting root length, three of which were newly implicated in root development. The majority of
3 single-gene mutations affects do not affect root length [4,35,50]. Our results demonstrate the
4 benefits of using statistical epistatic analyses in combination with experimental analysis of T-
5 DNA insertion mutants to identify novel functional candidate genes involved in root
6 development and provide a basis for future detailed studies of the complex gene networks
7 involved in this trait (Fig 4).

8
9 It will be a major endeavor to functionally explore the biological mechanisms contributing to the
10 statistical epistatic associations. None of the genes affecting root development contain non-
11 synonymous mutations with likely phenotypic effects. Also, none of these genes have previously
12 been described to interact with their respective epistatic locus. Future functional studies will have
13 to explore the effects of the observed regulatory SNPs in transgenic plants in several
14 backgrounds to study the whether these statistical interactions are the result of an underlying
15 biological interaction or not. This endeavor will require significant time and resources, yet we
16 believe that *A. thaliana* root length represents an excellent model trait for future in-depth studies
17 of this type of complex signals arising from GWA studies.

18

19 **Materials and Methods**

20 *Phenotyping*

21 In order to reduce environmental variation among accessions, eighteen individuals from each of
22 the 93 accessions studied in Atwell et al (2010) (Table S1) were vernalized at 4° as seedlings for
23 six weeks to synchronize growth and flowering. The five most developmentally advanced

1 seedlings from each accession were then transferred to soil in a randomized design. Plants were
2 grown in long-day conditions at 22°. Flats were rotated three times per week to reduce position
3 effects on plant development. Seeds were collected over a period of three months as the plants
4 dried. Equal numbers of seeds were pooled from 3-5 parent plants to reduce the environmental
5 contribution of plant placement across flats of plants. Ethanol sterilized seeds were planted on 1×
6 Murashige and Skoog (MS) basal salt medium supplemented with 1× MS vitamins, 0.05% MES
7 (wt/vol), and 0.3% (wt/vol) phytigel in a semi-randomized design with n = 70 per accession.
8 Four sets of twenty-three accessions plus a control accession (Col-0) made up a set. Each set was
9 replicated three times providing the standard error for variance, with a total n = 210 planted for
10 each accession. The seeds were stratified at 4° for three days and then grown for seven days in
11 darkness with the plates in a vertical position. A photograph of each plate was taken, and root
12 length was measured manually using the ‘freehand’ function in ImageJ1.46r [51]. Non-
13 germination, missing organs, and delayed development were noted.

14

15 *Data standardization*

16 Thirty-one seedlings with hypocotyls less than 5mm were removed because it is likely that
17 germination was severely delayed [34]. Between-set differences were corrected by subtracting
18 the difference between Col-0 in each replicate and set and the global mean for Col-0. Systematic
19 differences between replicates were still present and corrected within each set by using the
20 residuals from a model in which root length was explained by the within-set replicates.

21

22 *Heritability estimation*

1 Heritability of root length variance and mean were estimated using the repeated measures with
2 genotype as a fixed effect using ANOVA. To parse the contribution of additive and epistatic
3 effects, a linear mixed model including both additive and epistatic effects as random effects was
4 fitted using the R/hglm package [37], i.e.

$$5 \quad y = \mu + a + b + e,$$

6 where

$$a \sim N(0, G\sigma_a^2)$$

7 are the individual-level additive genetic effects, and G is the genomic kinship matrix;

$$b \sim N(0, G \circ G\sigma_b^2)$$

8 are the individual-level epistatic effects, and G is the genomic kinship matrix; e are independent
9 and normally distributed residuals. The narrow sense heritability was estimated as the ratio of the
10 additive genetic variance to the total phenotypic variance, and the broad sense heritability was
11 estimated as the ratio of the sum of the additive and epistatic variance to the total phenotypic
12 variance.

13

14 *Genome-wide association analysis*

15 Using the transformed mean and the genotypes for the 93 accessions from Atwell et al. (2010),
16 we performed a series of analyses to detect marker-trait associations. For the standard additive
17 GWA analysis with population structure controlled, we used the function `egscore()` from
18 R/GenABEL [38]. The functions `bigRR()` and `bigRR_update()` in the R/bigRR package [39]
19 were used to perform a GWA analysis, in which all genome-wide SNP effects are modeled
20 simultaneously as random effects, and the effects were estimated via the generalized ridge

1 regression method HEM. The maximum absolute effect sizes in 1000 permutations were
2 extracted to determine an empirical 5% genome-wide significance threshold ($p = 0.0020$).

3
4 To screen the genome for pairs of epistatic SNPs, we used the PLINK –epistasis procedure [40]
5 that is based on the model:

$$6 \quad Y \sim b_0 + b_{1.A} + b_{2.B} + b_{3.AB} + e,$$

7 which considers allele by allele epistasis in $b_{3.AB}$. In the analysis, all possible pairs of SNPs with
8 a minor allele frequency $> 25\%$ were tested.

9 We filtered out the pairs where there were fewer than four accessions in the minor two-locus
10 genotype-class. Further, all combinations in which the p-value was $> 1 \times 10^{-8}$ were also removed
11 as it was considered unlikely that they would be significant after correction for population
12 structure and applying a multiple-testing corrected significance threshold. For the remaining
13 pairs, a linear mixed model including fixed, additive and epistatic effects, as in the PLINK based
14 initial scan, were fitted together with a kinship correction for population stratification, as in the
15 single locus analyses, using the package R/hglm. We derived a multiple-testing corrected
16 significance threshold for this epistatic analysis ($p = 3.2 \times 10^{-10}$) by estimating the number of
17 independent tests based on the number of estimated LD blocks in the genome. For this, we used
18 the fact that the *A. thaliana* genome is approximately 125Mb with average LD-blocks of about
19 10kb (12,500; [41]), and then applied a Bonferroni correction for the 78 million independent
20 tests performed. Seventy-eight million double-recessive pseudo markers were created, and the
21 R/p.exact package [42] was used to double-assess the exact false discovery rate of each detected
22 epistatic pair based on its estimated epistatic effect size and allele frequency. Using data on the
23 sixty-three accessions from the GWAs that were unambiguously identified in the 1001 Genomes

1 data [22], we then identified additional SNPs in LD with the leading SNP using the function *LD*
2 in the *genetics* package in R across a 10kb region around the marker.

3 *High-order LD between epistatic pseudo markers and whole-genome sequencing variants*

4 From 728 whole-genome sequences in *A. thaliana* 1001 Genomes Project
5 (<http://www.1001genomes.org>), we extracted all the sequencing variants. All four significant
6 pseudo-markers had at least four accessions with the double-recessive genotype (i.e. the minor
7 allele for the epistatic pseudo-marker) in the 728 sequenced accessions. For each of the pseudo-
8 markers, we therefore sampled 93 accessions from the 728 sequenced ones, keeping the same
9 allele frequency as for the pseudo-marker representing the epistatic pair detected in the GWA
10 dataset, and obtained the genome-wide, high-order LD r^2 distribution across all the sequenced
11 variants. This procedure was repeated 50 times for each pseudo-marker, and we calculated the
12 estimates and standard errors of the r^2_{mean} , r^2_{max} and estimated the risk for “apparent epistasis” as
13 the probability of observing an $r^2_{\text{max}} > 0.8$ in a particular random sample.

14 *Examination of SNP-associated genes*

15 The Ensembl Variant Effect Predictor based on the TAIR10 release of the *A. thaliana* genome
16 was used to determine the effects of the leading SNPs and the SNPs in high LD with them.
17 Genes were considered as candidates if they were within 1kb of a variant. Expression of the
18 candidate genes was determined using the BAR eFP Arabidopsis browser [45].

19 *Validation of interacting loci*

20 T-DNA lines were obtained for the candidate genes (Table S2). Root length was ascertained as
21 described above (n = 20). Tukey’s HSD post hoc test was used to compare root lengths between
22 the T-DNA lines and the wild-type accession (*Col-0*).

1

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5

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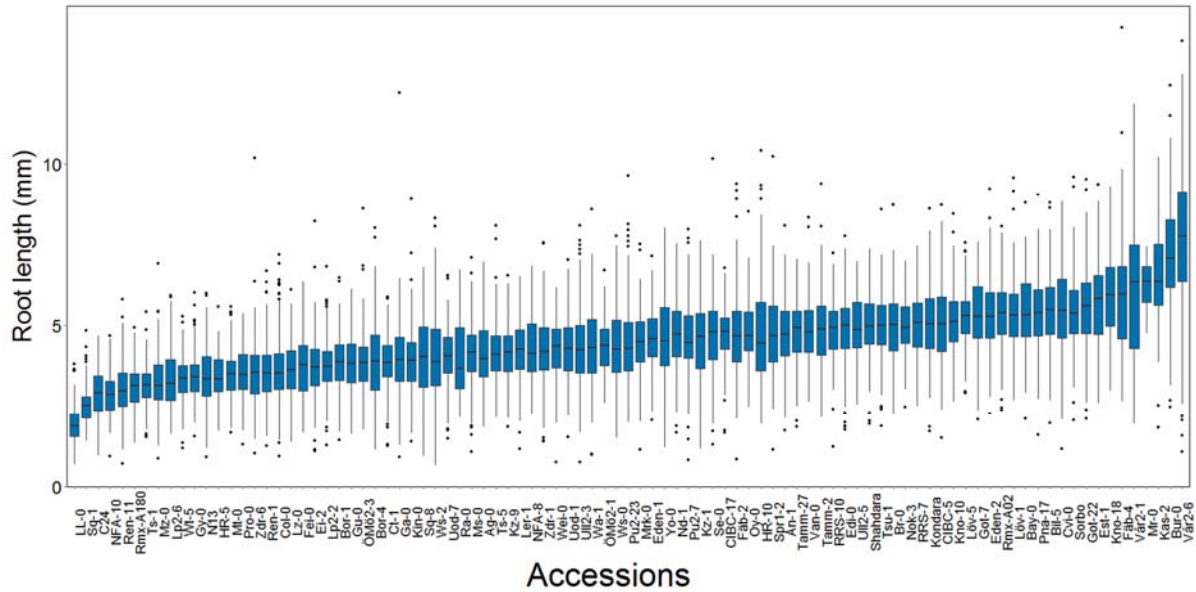
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18

19

1 **Fig 1. The distributions of within-accession root length mean and variance in a sample of**
2 **93 natural *A. thaliana* accessions.**

3 There is a large dispersion of the root length mean at day seven among the accessions ($n = 3 \times 70$
4 dark grown seedlings / genotype).

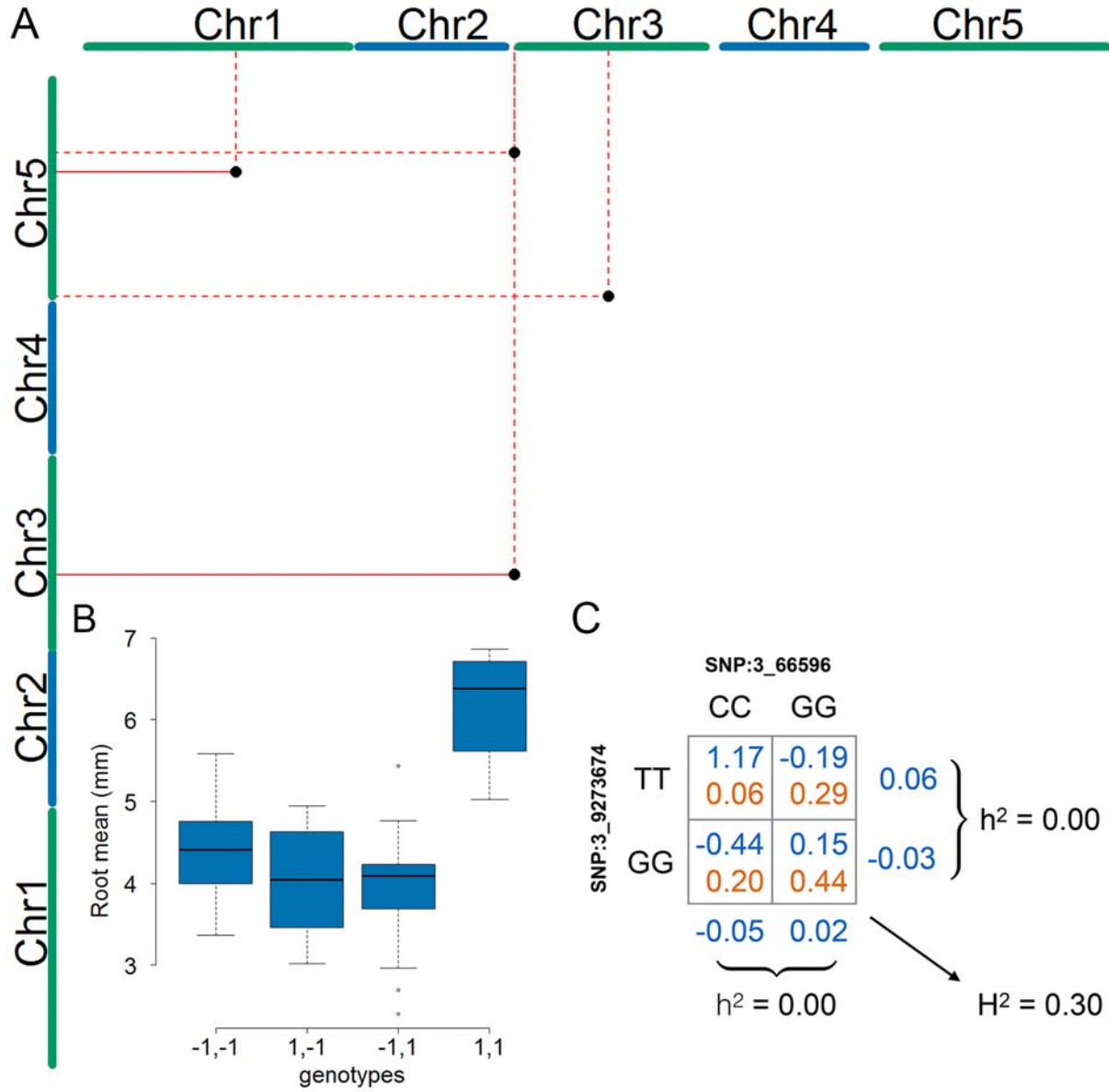


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6

1 **Fig 2. Four statistical epistatic interactions associated with mean root length.**

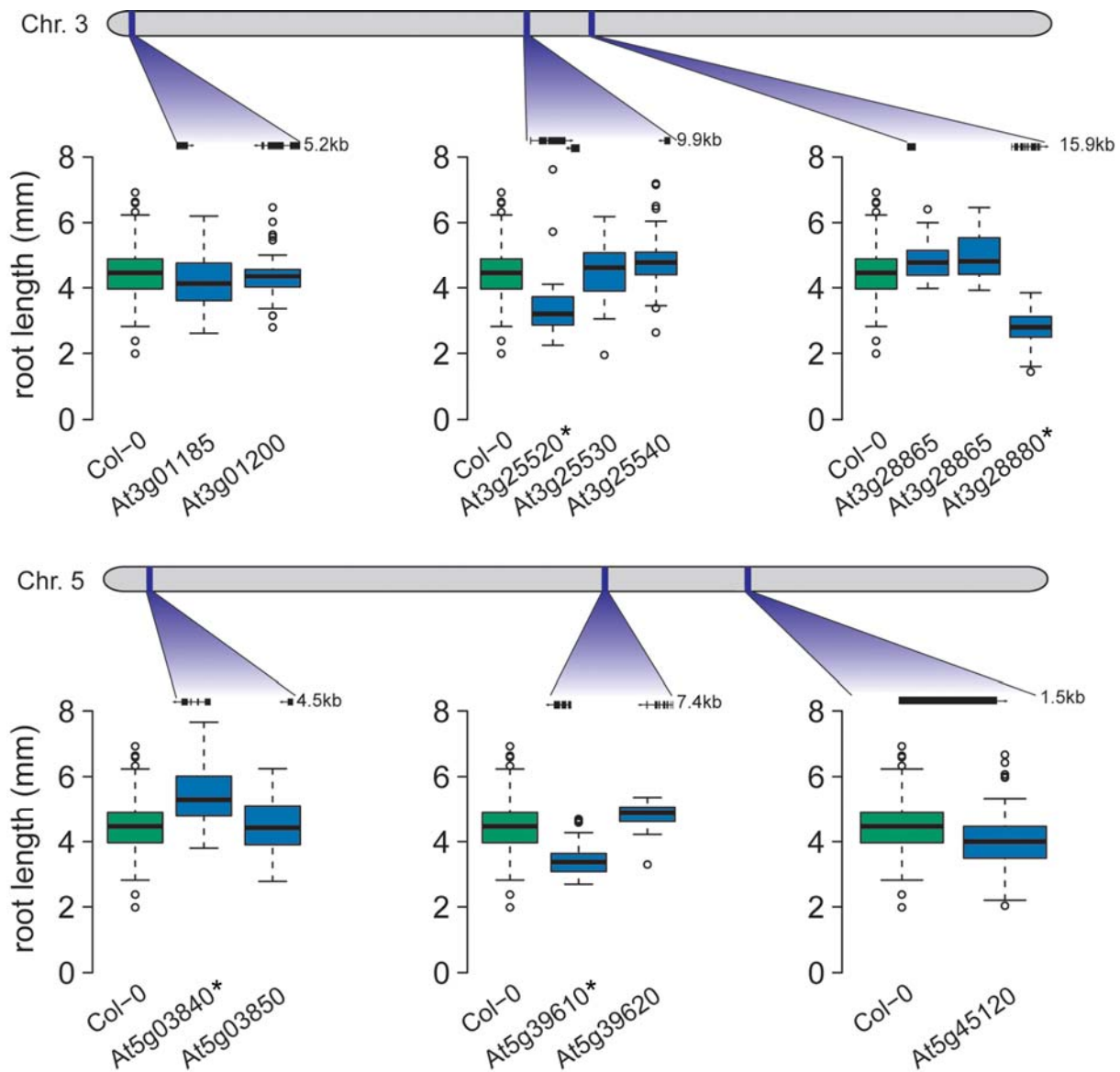
2 (A) The x- and y-axes represent the five *A. thaliana* chromosomes. The positions of the seven
3 SNPs that are part of the four significant interacting pairs (Table 2) are indicated by a black dot.
4 Solid lines indicate support for an interaction by more than one linked SNP and dotted lines
5 indicate support by a single SNP. (B) Genotype-phenotype (G-P) map of the root means for the
6 four genotype combinations constituting the interaction between the SNPs on chromosome 1
7 (17,257,526 bp) and chromosome 5 (15,862,026 bp). The major allele is indicated by -1, and the
8 minor allele is indicated by 1. This G-P-map is a representative example for the other pairs in
9 which the accessions with a minor and major allele have lower phenotypic values than those with
10 either both major or both minor alleles (Fig S2). (C) G-P map for the most significant epistatic
11 pair illustrates how this type of epistasis cancels the additive genetic variance at the allele-
12 frequencies observed for the two loci in the analyzed population of wild-collected *A. thaliana*
13 accessions.



1

2

- 1 **Fig 3. Four candidate genes exhibit altered root growth in T-DNA insertion lines.** T-DNA
- 2 lines for candidate genes in six of the seven epistatic regions were tested for root length
- 3 differences to wild-type (*Col-0*). The gene models for the tested T-DNA insertion lines are
- 4 pictured. For four of the six loci, a T-DNA line exhibited significantly different root length from
- 5 *Col-0* (significantly different lines are marked with a * in the figure).

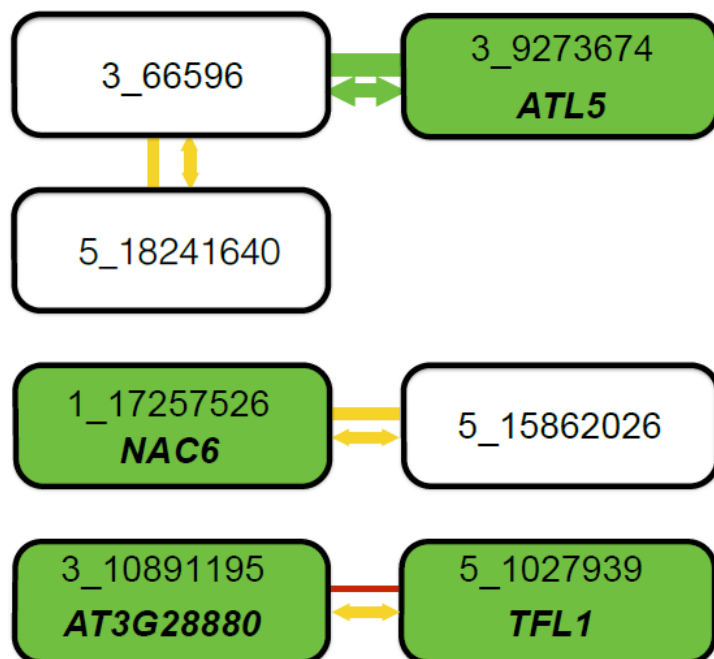


6

7

1 **Fig 4. Summary of combined statistical and functional support for loci underlying root**
2 **length.**

3 Four epistatic pairs, involving seven unique loci (connected boxes), were identified at a genome-
4 wide significance-threshold in a two-dimensional genome-wide GWA scan. Green, yellow, and
5 red lines connect pairs of loci with very low ($p = 0.003$), intermediate ($0.28 < p < 0.37$), and high
6 ($p = 0.95$) risk for the interaction being a false-positive when accounting for population size. The
7 risk of the statistical epistatic association resulting from high-order LD to an unobserved
8 functional variant in the genome (i.e. “apparent epistasis”) is illustrated by arrow color, in which
9 yellow indicates an intermediate risk ($0.18 < p < 0.42$) and green a very low risk ($p = 0.0021$).
10 Green boxes indicate loci that the T-DNA insertion line analyses suggest the named genes to be
11 involved in root development. When considering the joint statistical and functional results, two
12 pairs emerge as highly likely true positive two-locus associations: 3_66596/3_9273674 due to
13 very strong statistical support and one identified functional candidate gene, and
14 3_10891195/5_1027939 where the identification of functional candidate genes at both loci
15 suggest that the two-locus association in the original genome-wide scan is true despite the lower
16 statistical support in after the conservative statistical correction for sample-size. For the other
17 two pairs, the results are inconclusive. There is strong support for one of the two associated loci
18 (3_66596 from its statistical interaction with 3_9273674 and 1_17257526 by the detection of the
19 a functional candidate gene in the T-DNA analysis), but weaker support for the second locus.
20 Further work is thus needed to conclude whether these pairs represent true positive two-locus
21 associations, or whether they are false-positives due to the small population-size or high-order
22 LD (“apparent epistasis”) to unknown functional variants.



1

2

1 **Supporting information captions**

2 **Table S1. *A. thaliana* accessions phenotyped for GWA.**

3

4 **Table S2. T-DNA lines tested for root phenotypes.**

5

6 **Fig S1. No significant associations are found in GWAS for root length using an additive**
7 **model.**

8 (A) In the GWAS based on an additive model no SNPs reached the Bonferroni threshold. (B)
9 Using a whole-genome generalized ridge regression model, in which SNPs were modeled as
10 random effects, no SNPs reached the Bonferroni threshold.

11

12 **Fig S2. Similar genotype-phenotype maps for all significantly interacting pairs of loci.**

13 (A-C) Mean root length is displayed for all four two-locus genotype-classes for the four pairs of
14 significantly interacting loci. The major allele is indicated by -1 and the minor allele is indicated
15 by 1. The combination of two minor alleles (1,1) always has the longest root length compared to
16 the other three-allele combinations, which are more similar to one another. Pictured are mean
17 root lengths for the epistatic pairs for SNPs (A) 3_66596 and 3_9272294/3_9273674, (B)
18 3_66596 and 5_18241640, and (C) 3_10891195 and 5_1027939.